

# Supplementary material: The state learner – a super learner for right-censored data

Anders Munch<sup>1,\*</sup> and Thomas A. Gerds<sup>1</sup>

<sup>1</sup>Section of Biostatistics, University of Copenhagen

\*Address for correspondence: [a.munch@sund.ku.dk](mailto:a.munch@sund.ku.dk)

December 29, 2024

This document contains proofs of the results stated in the main paper. Section [S1](#) contains a proof of the consistency result from Section 5.1 of the main paper; Section [S2](#) contains proofs of the oracle inequalities from Section 5.2 of the main paper; and Section [S3](#) demonstrates transience of the second order remainder structure stated in Section 5.3 of the main paper.

We start by recalling some notation. We use  $F$  to denote a function that describes the conditional state occupation probabilities of the observed data, as defined in equation (5) of the main paper. We denote the integrated Brier score as  $\bar{B}_\tau(F, O) = \int_0^\tau B_t(F, O) dt$ , where  $B_t$  is defined

$$B_t(F, O) = \sum_{j=-1}^2 (F(t, j, X) - \mathbb{1}\{\eta(t) = j\})^2.$$

Finally, recall that we let  $\mathcal{H}_\mathcal{P}$  denote the function space consisting of all conditional state occupation probability functions for some measure  $P \in \mathcal{P}$ , and that we equipped this space

with the norm

$$\|F\|_{P_0} = \left\{ \sum_{j=-1}^2 \int_{\mathcal{X}} \int_0^\tau F(t, j, x)^2 dt H_0(dx) \right\}^{1/2}, \quad (1)$$

for some fixed measure  $P_0 \in \mathcal{P}$ . We use  $F_0$  to denote the conditional state occupation probability function associated with the measure  $P_0$ .

## S1 Consistency

Define  $\bar{B}_{\tau,0}(F, o) = \bar{B}_\tau(F, o) - \bar{B}_\tau(F_0, o)$  and  $R_0(F) = P_0[\bar{B}_{\tau,0}(F, \cdot)]$ .

**Lemma 1.**  $R_0(F) = \|F - F_0\|_{P_0}^2$ , where  $\|\cdot\|_{P_0}$  is defined in equation (1).

*Proof.* For any  $t \in [0, \tau]$  and  $j \in \{-1, 0, 1, 2\}$  we have

$$\begin{aligned} & \mathbb{E}_{P_0} [(F(t, j, X) - \mathbb{1}\{\eta(t) = j\})^2] \\ &= \mathbb{E}_{P_0} [(F(t, j, X) - F_0(t, j, X) + F_0(t, j, X) - \mathbb{1}\{\eta(t) = j\})^2] \\ &= \mathbb{E}_{P_0} [(F(t, j, X) - F_0(t, j, X))^2] + \mathbb{E}_{P_0} [(F_0(t, j, X) - \mathbb{1}\{\eta(t) = j\})^2] \\ &\quad + 2 \mathbb{E}_{P_0} [(F(t, j, X) - F_0(t, j, X))(F_0(t, j, X) - \mathbb{1}\{\eta(t) = j\})] \\ &= \mathbb{E}_{P_0} [(F(t, j, X) - F_0(t, j, X))^2] + \mathbb{E}_{P_0} [(F_0(t, j, X) - \mathbb{1}\{\eta(t) = j\})^2], \end{aligned}$$

where the last equality follows from the tower property. Hence, using Fubini, we have

$$P[\bar{B}_\tau(F, \cdot)] = \|F - F_0\|_{P_0}^2 + P_0[\bar{B}_\tau(F_0, \cdot)].$$

□

*Proof of Proposition 1 of the main paper.* The result follows from Lemma 1. □

## S2 Oracle inequalities

Recall that we use  $\mathcal{F}_n$  to denote a library of learners for the function  $F$ , and that  $\hat{\varphi}$  and  $\tilde{\varphi}$  denotes, respectively, the discrete super learner and the oracle learner for the library  $\mathcal{F}_n$ ,

c.f., Section 4 of the main paper.

*Proof of Corollary 1 of the main paper.* First note that minimising the loss  $\bar{B}_\tau$  is equivalent to minimising the loss  $\bar{B}_{\tau,0}$ , so the discrete super learner and oracle according to  $\bar{B}_\tau$  and  $\bar{B}_{\tau,0}$  are identical. By Lemma 1,  $R_0(F) \geq 0$  for any  $F \in \mathcal{H}_\mathcal{P}$ , and so using Theorem 2.3 from [van der Vaart et al., 2006] with  $p = 1$ , we have that for all  $\delta > 0$ ,

$$\begin{aligned} & \mathbb{E}_{P_0} \left[ R_0(\hat{\varphi}_n(\mathcal{D}_n^{-k})) \right] \\ & \leq (1 + 2\delta) \mathbb{E}_{P_0} \left[ R_0(\tilde{\varphi}_n(\mathcal{D}_n^{-k})) \right] \\ & \quad + (1 + \delta) \frac{16K}{n} \log(1 + |\mathcal{F}_n|) \sup_{F \in \mathcal{H}_\mathcal{P}} \left\{ M(F) + \frac{v(F)}{R_0(F)} \left( \frac{1}{\delta} + 1 \right) \right\} \end{aligned}$$

where for each  $F \in \mathcal{H}_\mathcal{P}$ ,  $(M(F), v(F))$  is some Bernstein pair for the function  $o \mapsto \bar{B}_{\tau,0}(F, o)$ .

As  $\bar{B}_{\tau,0}(F, \cdot)$  is uniformly bounded by  $\tau$  for any  $F \in \mathcal{H}_\mathcal{P}$ , it follows from section 8.1 in [van der Vaart et al., 2006] that  $(\tau, 1.5P_0[\bar{B}_{\tau,0}(F, \cdot)^2])$  is a Bernstein pair for  $\bar{B}_{\tau,0}(F, \cdot)$ . Now, for any  $a, b, c \in \mathbb{R}$  we have

$$\begin{aligned} (a - c)^2 - (b - c)^2 &= (a - b + b - c)^2 - (b - c)^2 \\ &= (a - b)^2 + (b - c)^2 + 2(b - c)(a - b) - (b - c)^2 \\ &= (a - b) \{ (a - b) + 2(b - c) \} \\ &= (a - b) \{ a + b - 2c \}, \end{aligned}$$

so using this with  $a = F(t, j, x)$ ,  $b = F_0(t, j, x)$ , and  $c = \mathbb{1}\{\eta(t) = j\}$ , we have by Jensen's

inequality

$$\begin{aligned}
& P_0[\bar{B}_{\tau,0}(F, \cdot)^2] \\
& \leq 2\tau \mathbb{E}_{P_0} \left[ \sum_{j=-1}^2 \int_0^\tau \left\{ (F(t, j, X) - \mathbb{1}\{\eta(t) = j\})^2 - (F_0(t, j, X) - \mathbb{1}\{\eta(t) = j\})^2 \right\}^2 dt \right] \\
& = 2\tau \mathbb{E}_{P_0} \left[ \sum_{j=-1}^2 \int_0^\tau (F(t, j, X) - F_0(t, j, X))^2 \right. \\
& \quad \left. \times \{F(t, j, X) + F_0(t, j, X) - 2\mathbb{1}\{\eta(t) = j\}\}^2 dt \right] \\
& \leq 8\tau \mathbb{E}_{P_0} \left[ \sum_{j=-1}^2 \int_0^\tau (F(t, j, X) - F_0(t, j, X))^2 dt \right] \\
& = 8\tau \|F - F_0\|_{P_0}^2.
\end{aligned}$$

Thus when  $v(F) = 1.5P_0[\bar{B}_{\tau,0}(F, \cdot)^2]$  we have by Lemma 1

$$\frac{v(F)}{R_0(F)} = 1.5 \frac{P_0[\bar{B}_{\tau,0}(F, \cdot)^2]}{P_0[\bar{B}_{\tau,0}(F, \cdot)]} \leq 12\tau,$$

and so using the Bernstein pairs  $(\tau, 1.5P_0[\bar{B}_{\tau,0}(F, \cdot)^2])$  we have

$$\sup_{F \in \mathcal{H}_{\mathcal{P}}} \left\{ M(F) + \frac{v(F)}{R_0(F)} \left( \frac{1}{\delta} + 1 \right) \right\} \leq \tau \left( 13 + \frac{12}{\delta} \right),$$

For all  $\delta > 0$  we thus have

$$\begin{aligned}
\mathbb{E}_{P_0} \left[ R_0(\hat{\varphi}_n(\mathcal{D}_n^{-k})) \right] & \leq (1 + 2\delta) \mathbb{E}_{P_0} \left[ R_0(\tilde{\varphi}_n(\mathcal{D}_n^{-k})) \right] \\
& \quad + (1 + \delta) \log(1 + |\mathcal{F}_n|) \tau \frac{16K}{n} \left( 13 + \frac{12}{\delta} \right),
\end{aligned}$$

and then the final result follows from Lemma 1.  $\square$

*Proof of Corollary 2 of the main paper.* By definition of the oracle and Lemma 1,

$$\mathbb{E}_{P_0} \left[ \|\tilde{\varphi}_n(\mathcal{D}_n^{-k}) - F_0\|_{P_0}^2 \right] \leq \mathbb{E}_{P_0} \left[ \|\varphi_n(\mathcal{D}_n^{-k}) - F_0\|_{P_0}^2 \right]$$

for all  $n \in \mathbb{N}$ . The results then follows from Corollary 1 of the main paper.  $\square$

### S3 Transience of the second order remainder structure

Recall that we let  $\Lambda_1$  denote the conditional cumulative hazard function for one of the event times of interest and  $\Gamma$  the conditional cumulative hazard function for censoring, c.f., Section 2 of the main paper.

*Proof of Proposition 2 of the main paper.* For notational convenience we suppress  $X$  in the following. The final result can be obtained by adding the argument  $X$  to all functions and averaging. We use the relations from equation (8) in the main paper to write

$$\begin{aligned}
& \int_0^\tau w(s) \left\{ \Gamma(s) - \hat{\Gamma}_n(s) \right\} [\Lambda_1 - \hat{\Lambda}_{1,n}](ds) \\
&= \int_0^\tau w(s) \left\{ \int_0^s \frac{F(du, -1)}{F(u-, 0)} - \int_0^s \frac{\hat{F}_n(du, -1)}{\hat{F}_n(u-, 0)} \right\} \left[ \frac{F(ds, 1)}{F(s-, 0)} - \frac{\hat{F}_n(ds, 1)}{\hat{F}_n(s-, 0)} \right] \\
&= \int_0^\tau w(s) \left\{ \int_0^s \left( \frac{1}{F(u-, 0)} - \frac{1}{\hat{F}_n(u-, 0)} \right) F(du, -1) \right. \\
&\quad \left. + \int_0^s \frac{1}{\hat{F}_n(u-, 0)} [F(du, -1) - \hat{F}_n(du, -1)] \right\} \\
&\quad \times \left[ \left( \frac{1}{F(s-, 0)} - \frac{1}{\hat{F}_n(s-, 0)} \right) F(ds, 1) + \frac{1}{\hat{F}_n(s-, 0)} (F(ds, 1) - \hat{F}_n(ds, 1)) \right] \\
&= \int_0^\tau \int_0^s w(s) \left( \frac{1}{F(u-, 0)} - \frac{1}{\hat{F}_n(u-, 0)} \right) \left( \frac{1}{F(s-, 0)} - \frac{1}{\hat{F}_n(s-, 0)} \right) F(du, -1) F(ds, 1) \\
&\quad + \int_0^\tau \int_0^s w(s) \left( \frac{1}{F(u-, 0)} - \frac{1}{\hat{F}_n(u-, 0)} \right) \frac{F(du, -1)}{\hat{F}_n(u-, 0)} (F(ds, 1) - \hat{F}_n(ds, 1)) \\
&\quad + \int_0^\tau \int_0^s \frac{w(s)}{\hat{F}_n(u-, 0)} [F(du, -1) - \hat{F}_n(du, -1)] \left( \frac{1}{F(s-, 0)} - \frac{1}{\hat{F}_n(s-, 0)} \right) F(ds, 1) \\
&\quad + \int_0^\tau \int_0^s \frac{w(s)}{\hat{F}_n(u-, 0)} [F(du, -1) - \hat{F}_n(du, -1)] \frac{1}{\hat{F}_n(s-, 0)} (F(ds, 1) - \hat{F}_n(ds, 1)) .
\end{aligned}$$

Consider the first term on the right hand side. By the mean value theorem,

$$\frac{1}{F(t-, 0)} - \frac{1}{\hat{F}_n(t-, 0)} = \frac{-1}{\tilde{r}_n(t)^2} [F(t-, 0) - \hat{F}_n(t-, 0)] ,$$

where  $\tilde{r}_n(t)$  is some value between  $\hat{F}(t-, 0)$  and  $\hat{F}_n(t-, 0)$ . Letting  $w_n^*(t) = -\tilde{r}_n(t)^{-2}$ , we

may write

$$\begin{aligned}
& \int_0^\tau \int_0^s w(s) \left( \frac{1}{F(u-, 0)} - \frac{1}{\hat{F}_n(u-, 0)} \right) \left( \frac{1}{F(s-, 0)} - \frac{1}{\hat{F}_n(s-, 0)} \right) F(du, -1) F(ds, 1) \\
&= \int_0^\tau \int_0^s w(s) w_n^*(u) \left( F(u-, 0) - \hat{F}_n(u-, 0) \right) \\
&\quad \times w_n^*(s) \left( F(s-, 0) - \hat{F}_n(s-, 0) \right) F(du, -1) F(ds, 1) \\
&= \int_0^\tau \int_0^s w_n^a(s, u) \left( F(u-, 0) - \hat{F}_n(u-, 0) \right) \left( F(s-, 0) - \hat{F}_n(s-, 0) \right) F(du, -1) F(ds, 1),
\end{aligned}$$

where we have defined  $w_n^a(s, u) = w(s)w_n^*(s)w_n^*(u)$ . By the assumption that  $F(\cdot, 0)$  and  $F_n(\cdot, 0)$  are uniformly bounded away from zero on  $[0, \tau]$ , it follows that  $w_n^*$  is uniformly bounded, and hence  $w_n^a(s, u)$  is also uniformly bounded, because  $w$  was assumed uniformly bounded. The same approach can be applied to the three remaining terms which gives the result.  $\square$

## References

A. W. van der Vaart, S. Dudoit, and M. J. van der Laan. Oracle inequalities for multi-fold cross validation. *Statistics & Decisions*, 24(3):351–371, 2006.