LECTURE 14: CATEGORIES AND INTERACTIONS

ECON 480 - ECONOMETRICS - FALL 2018

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Dummy Dependent (Y) Variables

Interaction Effects

Interactions Between a Dummy and a Continuous Variable

Interactions Between Two Dummy Variables

Interactions Between Two Continuous Variables



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 - e.g. Highest education attained (1=elementary school, 2=high school, 3=bachelor's degree, 4=graduate degree)



USING CATEGORICAL VARIABLES IN REGRESSION

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- We can easily transform a categorical variable into a set of dummy variables, one for each possible category

Example

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Can we run the following regression?

$$\widehat{Wages}_i = \hat{\beta}_0 + \hat{\beta}_1 Region_i + \epsilon_i$$



USING CATEGORICAL VARIABLES IN REGRESSION II

Example

How do wages vary by region of the country?

Let

$$Region_{i} = \begin{cases} 1 & \text{if } i \text{ is in Northeast} \\ 2 & \text{If } i \text{ is in Midwest} \\ 3 & \text{if } i \text{ is in South} \\ 4 & \text{If } i \text{ is in West} \end{cases}$$

$$\widehat{\text{Wages}}_i = \hat{eta}_0 + \hat{eta}_1 \text{Region} + \epsilon_i$$



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· Now can we run the following regression?

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USING CATEGORICAL VARIABLES IN REGRESSION III

Example

How do wages vary by region of the country?

- · Create dummy for each region:
 - · Northeast_i = 1 if i is in Northeast, else 0
 - $Midwest_i = 1$ if i is in Midwest, else 0
 - $South_i = 1$ if i is in South, else 0
 - · $West_i = 1$ if i is in West, else 0



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- · Now can we run the following regression?

$$\widehat{Wages_i} = \hat{eta_0} + \hat{eta_1} Northeast_i + \hat{eta_2} Midwest_i + \hat{eta_3} South_i + \hat{eta_4} West_i + \epsilon_i$$



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THE DUMMY VARIABLE TRAP

$$\widehat{Wages_i} = \hat{eta_0} + \hat{eta_1}$$
Northeast $+$ $\hat{eta_2}$ Midwest $+$ $\hat{eta_3}$ South $+$ $\hat{eta_4}$ West $+$ ϵ_i



THE DUMMY VARIABLE TRAP

$$\widehat{\textit{Wages}_i} = \hat{\beta_0} + \hat{\beta_1} \textit{Northeast} + \hat{\beta_2} \textit{Midwest} + \hat{\beta_3} \textit{South} + \hat{\beta_4} \textit{West} + \epsilon_i$$

• If we included *all* possible categories, they are perfectly multicollinear, an exact linear function of one another:

$$Northeast_i + Midwest_i + South_i + West_i = 1 \quad \forall i$$

• This is known as the dummy variable trap, a common source of perfect multicollinearity



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West_i omitted (arbitrarily chosen)



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- $\cdot \beta_3$:



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- β_1 : difference between West and Northeast
- β_2 : difference between West and Midwest
- β_3 : difference between West and South



THE DUMMY VARIABLE TRAP IN R

```
dtreg<-lm(wage~noreast+northcen+south+west, data=wages)</pre>
summary(dtreg)
##
## Call:
## lm(formula = wage ~ noreast + northcen + south + west, data = wages)
##
## Residuals:
     Min
          10 Median 30
##
                                  Max
## -6.083 -2.387 -1.097 1.157 18.610
##
## Coefficients: (1 not defined because of singularities)
               Estimate Std. Error t value Pr(>|t|)
##
```

(Intercept) 6.6134 0.3891 16.995 < 2e-16 ***

```
# run 4 regressions
no.noreast.reg<-lm(wage~northcen+south+west, data=wages)
no.northcen.reg<-lm(wage~noreast+south+west, data=wages)</pre>
no.south.reg<-lm(wage~noreast+northcen+west, data=wages)</pre>
no.west.reg<-lm(wage~noreast+northcen+south, data=wages)</pre>
# make output table
library("stargazer")
stargazer(no.noreast.reg, no.northcen.reg,
          no.south.reg, no.west.reg,
          type="latex", header=FALSE,
          float=FALSE. font.size="tinv")
```

| | Dependent variable: | | | |
|--------------------------------|-------------------------|----------|----------|-----------|
| | | | | |
| | (1) | (2) | (3) | (4) |
| northcen | -0.659 | | 0.324 | -0.903* |
| | (0.465) | | (0.417) | (0.504) |
| noreast | | 0.659 | 0.983** | -0.244 |
| | | (0.465) | (0.432) | (0.515) |
| south | -0.983** | -0.324 | | -1.226*** |
| | (0.432) | (0.417) | | (0.473) |
| west | 0.244 | 0.903* | 1.226*** | |
| | (0.515) | (0.504) | (0.473) | |
| Constant | 6.370*** | 5.710*** | 5.387*** | 6.613*** |
| | (0.338) | (0.320) | (0.268) | (0.389) |
| Observations | 526 | 526 | 526 | 526 |
| R^2 | 0.017 | 0.017 | 0.017 | 0.017 |
| Adjusted R ² | 0.012 | 0.012 | 0.012 | 0.012 |
| Residual Std. Error (df = 522) | 3.671 | 3.671 | 3.671 | 3.671 |
| F Statistic (df = 3; 522) | 3.099** | 3.099** | 3.099** | 3.099** |

 Constant is always mean wage for reference (omitted) category



Note: *p<0.1; **p<0.05; ***p<0.01

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- · It doesn't matter which category we omit



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Note:

USING DIFFERENT REFERENCE CATEGORIES II

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- · It doesn't matter which category we omit
- Same n, R^2 , and SER; coefficients give same results



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• In many contexts, we will want to have our dependent (Y) variable be a dummy variable



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 where $Admitted_i = \begin{cases} 1 & \text{if } i \text{ is Admitted} \\ 0 & \text{If } i \text{ is Not Admitted} \end{cases}$



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 - e.g. the probability person $\it i$ is Admitted to a program with a given GPA



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 - · requires special tools to properly interpret and extend this (logit, probit, etc)
- Feel free to write papers that have dummy Y variables (but you may have to ask me some more questions)!



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 - i.e. is there an interaction effect between sex and experience?



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 - i.e. is there an interaction effect between sex and experience?
- Do men gain more than women from an additional year of experience?
 - Note this is not the same as asking: "do men earn more than women with the same amount of experience?"



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- We look at each in turn:



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- 1. Interaction between a dummy and a continuous variable:

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$



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- 1. Interaction between a dummy and a continuous variable:

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

2. Interaction between two dummy variables:

$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + \beta_3 D_{1i} * D_{2i}$$



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2. Interaction between two dummy variables:

$$Y_{i} = \beta_{0} + \beta_{1}D_{1i} + \beta_{2}D_{2i} + \beta_{3}D_{1i} * D_{2i}$$

3. Interaction between two continuous variables:

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{1i} * X_{2i}$$



INTERACTIONS BETWEEN A DUMMY AND

A CONTINUOUS VARIABLE

• We can model this interaction by introducing a variable that is an interaction term capturing the interaction between two variables:

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- β_3 estimates the interaction term (in this case between a dummy variable and a continuous variable)
- What do the different coefficients (β 's) tell us?



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- β_3 estimates the interaction term (in this case between a dummy variable and a continuous variable)
- What do the different coefficients (β 's) tell us?
 - Again, think logically by examining each group ($D_i = 0$ or $D_i = 1$)



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

• When $D_i = 0$ (Control group):



INTERACTION EFFECTS AS TWO REGRESSIONS

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

• When $D_i = 0$ (Control group):

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i + \hat{\beta}_2(0) + \hat{\beta}_3 X_i * (0)$$



INTERACTION EFFECTS AS TWO REGRESSIONS

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$$\hat{Y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}X_{i}$$



INTERACTION EFFECTS AS TWO REGRESSIONS

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$$\hat{Y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} + \hat{\beta}_{2}(0) + \hat{\beta}_{3}X_{i} * (0)$$

$$\hat{Y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}X_{i}$$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

• When $D_i = 0$ (Control group):

$$\hat{Y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} + \hat{\beta}_{2}(0) + \hat{\beta}_{3}X_{i} * (0)$$

$$\hat{Y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}X_{i}$$

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i + \hat{\beta}_2(1) + \hat{\beta}_3 X_i * (1)$$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

• When $D_i = 0$ (Control group):

$$\begin{split} \hat{Y}_{i} &= \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} + \hat{\beta}_{2}(0) + \hat{\beta}_{3}X_{i} * (0) \\ \hat{Y}_{i} &= \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} \end{split}$$

$$\begin{split} \hat{Y}_{i} &= \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} + \hat{\beta}_{2}(1) + \hat{\beta}_{3}X_{i} * (1) \\ \hat{Y}_{i} &= \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} + \hat{\beta}_{2} + \hat{\beta}_{3}X_{i} \end{split}$$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

• When $D_i = 0$ (Control group):

$$\hat{Y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} + \hat{\beta}_{2}(0) + \hat{\beta}_{3}X_{i} * (0)$$

$$\hat{Y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}X_{i}$$

$$\hat{Y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} + \hat{\beta}_{2}(1) + \hat{\beta}_{3}X_{i} * (1)$$

$$\hat{Y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} + \hat{\beta}_{2} + \hat{\beta}_{3}X_{i}$$

$$\hat{Y}_{i} = (\hat{\beta}_{0} + \hat{\beta}_{2}) + (\hat{\beta}_{1} + \hat{\beta}_{3})X_{i}$$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

• When $D_i = 0$ (Control group):

$$\hat{Y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} + \hat{\beta}_{2}(0) + \hat{\beta}_{3}X_{i} * (0)$$

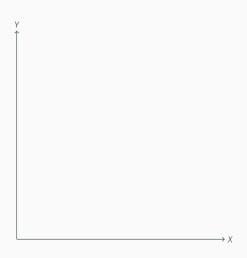
$$\hat{Y}_{i} = \hat{\beta}_{0} + \hat{\beta}_{1}X_{i}$$

• When $D_i = 1$ (Treatment group):

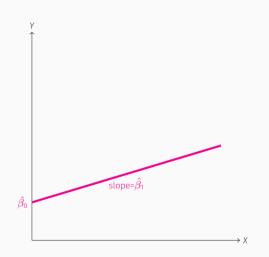
$$\begin{split} \hat{Y}_{i} &= \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} + \hat{\beta}_{2}(1) + \hat{\beta}_{3}X_{i} * (1) \\ \hat{Y}_{i} &= \hat{\beta}_{0} + \hat{\beta}_{1}X_{i} + \hat{\beta}_{2} + \hat{\beta}_{3}X_{i} \\ \hat{Y}_{i} &= (\hat{\beta}_{0} + \hat{\beta}_{2}) + (\hat{\beta}_{1} + \hat{\beta}_{3})X_{i} \end{split}$$



· So what we really have is two regression lines!

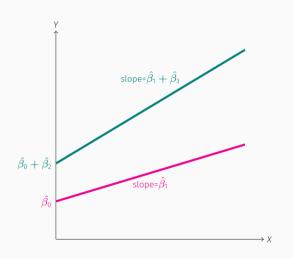






·
$$D_i = 0$$
 group:
 $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$





·
$$D_i = 1$$
 group:
 $\hat{Y}_i = (\hat{\beta}_0 + \hat{\beta}_2) + (\hat{\beta}_1 + \hat{\beta}_3)X_i$

·
$$D_i = 0$$
 group:
 $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$

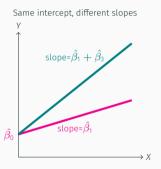


• Three distinct possibilities for the two lines $D_i = 0$ and $D_i = 1$:



INTERACTION EFFECTS AS TWO REGRESSIONS III

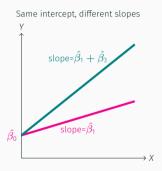
• Three distinct possibilities for the two lines $D_i = 0$ and $D_i = 1$:

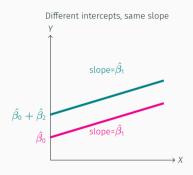




Interaction Effects as Two Regressions III

• Three distinct possibilities for the two lines $D_i = 0$ and $D_i = 1$:

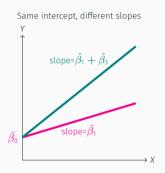


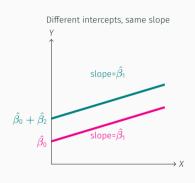


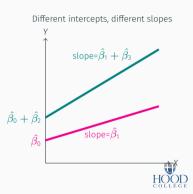


Interaction Effects as Two Regressions III

• Three distinct possibilities for the two lines $D_i = 0$ and $D_i = 1$:

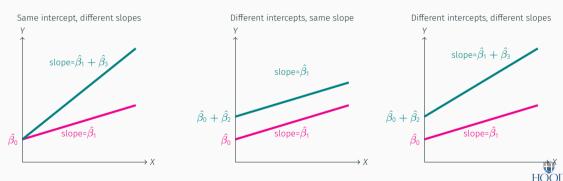






Interaction Effects as Two Regressions III

• Three distinct possibilities for the two lines $D_i = 0$ and $D_i = 1$:



· Well...four, but: what if they had the same slope and same intercept?

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

· To interpret the coefficients, compare cases after changing X by ΔX_i :



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

· To interpret the coefficients, compare cases after changing X by ΔX_i :



Interpretting the Coefficients

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- To interpret the coefficients, compare cases after changing X by ΔX_i :

$$Y_i + \Delta Y_i = \beta_0 + \beta_1 (X_i + \Delta X_i) \beta_2 D_i + \beta_3 ((X_i + \Delta X_i) D_i)$$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- To interpret the coefficients, compare cases after changing X by ΔX_i :

$$Y_i + \Delta Y_i = \beta_0 + \beta_1 (X_i + \Delta X_i) \beta_2 D_i + \beta_3 ((X_i + \Delta X_i) D_i)$$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

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$$\Delta Y_i = \beta_1 \Delta X_i + \beta_3 D_i \Delta X_i$$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

• To interpret the coefficients, compare cases after changing X by ΔX_i :

$$Y_i + \Delta Y_i = \beta_0 + \beta_1 (X_i + \Delta X_i) \beta_2 D_i + \beta_3 ((X_i + \Delta X_i) D_i)$$

$$\Delta Y_i = \beta_1 \Delta X_i + \beta_3 D_i \Delta X_i$$
$$\frac{\Delta Y_i}{\Delta X_i} = \beta_1 + \beta_3 D_i$$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

• To interpret the coefficients, compare cases after changing X by ΔX_i :

$$Y_i + \Delta Y_i = \beta_0 + \beta_1 (X_i + \Delta X_i) \beta_2 D_i + \beta_3 ((X_i + \Delta X_i) D_i)$$

$$\Delta Y_i = \beta_1 \Delta X_i + \beta_3 D_i \Delta X_i$$
$$\frac{\Delta Y_i}{\Delta X_i} = \beta_1 + \beta_3 D_i$$

- The effect of $X \to Y$ depends on the value of D_i !
- β_3 : increment to the effect of $X_i \to Y_i$ when $D_i = 1$ (vs. $D_i = 0$)



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

•
$$\beta_0$$
: Y_i for $X_i = 0$ and $D_i = 0$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- β_0 : Y_i for $X_i = 0$ and $D_i = 0$
- \cdot β_1 : Marginal effect of $X_i o Y_i$ for $D_i = 0$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- β_0 : Y_i for $X_i = 0$ and $D_i = 0$
- β_1 : Marginal effect of $X_i \to Y_i$ for $D_i = 0$
- · β_2 : Marginal effect on Y_i of difference between $D_i=0$ and $D_i=1$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- β_0 : Y_i for $X_i = 0$ and $D_i = 0$
- β_1 : Marginal effect of $X_i \to Y_i$ for $D_i = 0$
- β_2 : Marginal effect on Y_i of difference between $D_i = 0$ and $D_i = 1$
- β_3 : The **difference** of the marginal effect of $X_i \to Y_i$ between $D_i = 0$ and $D_i = 1$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- β_0 : Y_i for $X_i = 0$ and $D_i = 0$
- β_1 : Marginal effect of $X_i \to Y_i$ for $D_i = 0$
- β_2 : Marginal effect on Y_i of difference between $D_i = 0$ and $D_i = 1$
- β_3 : The difference of the marginal effect of $X_i \to Y_i$ between $D_i = 0$ and $D_i = 1$
- This is a bit awkward, easier to think about the two regression lines:



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

• For
$$D_i=0$$
 Group: $\hat{Y}_i=\hat{eta}_0+\hat{eta}_1X_i$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- For $D_i = 0$ Group: $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$
 - · Intercept (Y_i for $X_i=0$): \hat{eta}_0



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- For $D_i = 0$ Group: $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$
 - · Intercept (Y_i for $X_i=0$): \hat{eta}_0
 - · Slope (Marginal effect of X_i on Y_i for $D_i=0$ group): $\hat{\beta}_1$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- For $D_i = 0$ Group: $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$
 - · Intercept (Y_i for $X_i=0$): \hat{eta}_0
 - · Slope (Marginal effect of X_i on Y_i for $D_i=0$ group): $\hat{\beta}_1$

· For
$$D_i=1$$
 Group: $\hat{Y}_i=(\hat{eta}_0+\hat{eta}_2)+(\hat{eta}_1+\hat{eta}_3)X_i$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- For $D_i = 0$ Group: $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$
 - · Intercept (Y_i for $X_i=0$): \hat{eta}_0
 - · Slope (Marginal effect of X_i on Y_i for $D_i=0$ group): $\hat{\beta}_1$
- For $D_i=1$ Group: $\hat{Y}_i=(\hat{\beta}_0+\hat{\beta}_2)+(\hat{\beta}_1+\hat{\beta}_3)X_i$
 - · Intercept (Y_i for $X_i=0$): $\hat{eta}_0+\hat{eta}_2$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- For $D_i = 0$ Group: $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$
 - · Intercept (Y_i for $X_i=0$): \hat{eta}_0
 - · Slope (Marginal effect of X_i on Y_i for $D_i=0$ group): \hat{eta}_1
- · For $D_i=1$ Group: $\hat{Y}_i=(\hat{eta}_0+\hat{eta}_2)+(\hat{eta}_1+\hat{eta}_3)X_i$
 - · Intercept (Y_i for $X_i=0$): $\hat{eta}_0+\hat{eta}_2$
 - · Slope (Marginal effect of X_i on Y_i for $D_i=1$ group): $\hat{\beta}_1+\hat{\beta}_3$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- For $D_i = 0$ Group: $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$
 - · Intercept (Y_i for $X_i=0$): \hat{eta}_0
 - · Slope (Marginal effect of X_i on Y_i for $D_i=0$ group): $\hat{\beta}_1$
- · For $D_i=1$ Group: $\hat{Y}_i=(\hat{eta}_0+\hat{eta}_2)+(\hat{eta}_1+\hat{eta}_3)X_i$
 - · Intercept (Y_i for $X_i=0$): $\hat{eta}_0+\hat{eta}_2$
 - · Slope (Marginal effect of X_i on Y_i for $D_i=1$ group): $\hat{\beta}_1+\hat{\beta}_3$
- How can we determine if the two lines have the same slope and/or intercept (and distinguish between the 3 cases)?



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- For $D_i = 0$ Group: $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$
 - · Intercept (Y_i for $X_i = 0$): $\hat{\beta}_0$
 - · Slope (Marginal effect of X_i on Y_i for $D_i=0$ group): $\hat{\beta}_1$
- · For $D_i=1$ Group: $\hat{Y}_i=(\hat{eta}_0+\hat{eta}_2)+(\hat{eta}_1+\hat{eta}_3)X_i$
 - · Intercept (Y_i for $X_i=0$): $\hat{eta}_0+\hat{eta}_2$
 - · Slope (Marginal effect of X_i on Y_i for $D_i=1$ group): $\hat{\beta}_1+\hat{\beta}_3$
- How can we determine if the two lines have the same slope and/or intercept (and distinguish between the 3 cases)?
 - · Same intercept? t-test H_0 : $\beta_2 = 0$



$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i * D_i$$

- For $D_i = 0$ Group: $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$
 - · Intercept (Y_i for $X_i = 0$): $\hat{\beta}_0$
 - · Slope (Marginal effect of X_i on Y_i for $D_i=0$ group): $\hat{\beta}_1$
- · For $D_i=1$ Group: $\hat{Y}_i=(\hat{eta}_0+\hat{eta}_2)+(\hat{eta}_1+\hat{eta}_3)X_i$
 - · Intercept (Y_i for $X_i=0$): $\hat{eta}_0+\hat{eta}_2$
 - · Slope (Marginal effect of X_i on Y_i for $D_i=1$ group): $\hat{\beta}_1+\hat{\beta}_3$
- How can we determine if the two lines have the same slope and/or intercept (and distinguish between the 3 cases)?
 - Same intercept? t-test H_0 : $\beta_2 = 0$
 - Same slope? t-test H_0 : $\beta_3 = 0$



EXAMPLE

Example

$$\widehat{\text{wage}_i} = \hat{eta}_0 + \hat{eta}_1$$
exper $_i + \hat{eta}_2$ female $_i + \hat{eta}_3$ exper $_i * female_i$



EXAMPLE

Example

$$\widehat{wage}_i = \hat{eta}_0 + \hat{eta}_1$$
exper $_i + \hat{eta}_2$ female $_i + \hat{eta}_3$ exper $_i * female_i$

• For Males (female = 0):

$$\widehat{wage_i} = \hat{\beta_0} + \hat{\beta_1}$$
exper



Example

$$\widehat{\text{wage}_i} = \hat{eta}_0 + \hat{eta}_1 \text{exper}_i + \hat{eta}_2 \text{female}_i + \hat{eta}_3 \text{exper}_i * \text{female}_i$$

• For Males (female = 0):

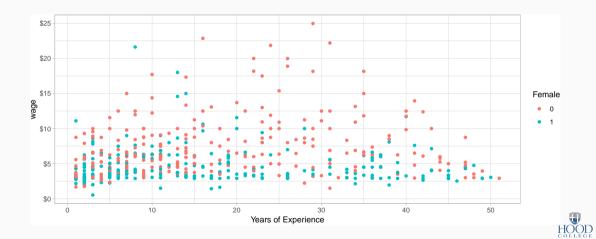
$$\widehat{wage_i} = \hat{\beta_0} + \hat{\beta_1}$$
exper

• For Females (female = 1):

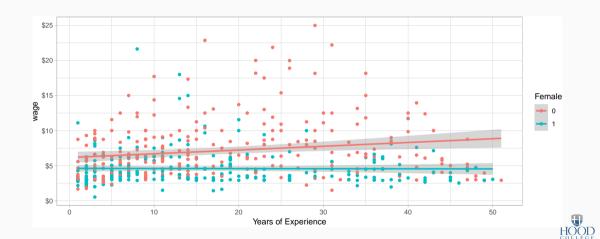
$$\widehat{wage_i} = \underbrace{(\hat{\beta}_0 + \hat{\beta}_2)}_{\text{intercept}} + \underbrace{(\hat{\beta}_1 + \hat{\beta}_3)}_{\text{slope}} exper$$



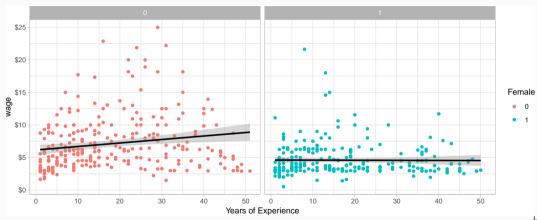
EXAMPLE II



EXAMPLE III



EXAMPLE III





EXAMPLE REGRESSION IN R

Syntax for interaction term is easy, simply add var1*var2 to the regression

```
interactionreg<-lm(wage~female+exper+female*exper, data=wages)</pre>
summary(interactionreg)
##
## Call:
## lm(formula = wage ~ female + exper + female * exper, data = wages)
##
## Residuals:
##
      Min
               1Q Median
                              30
                                     Max
## -6.3200 -1.8191 -0.9708 1.4132 17.2672
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) 6.15828 0.34167 18.024 < 2e-16 ***
## female
           -1.54655 0.48186 -3.210 0.001411 **
                         0.01544 3.472 0.000559 ***
## exper
               0.05360
## female:exper -0.05507
                         0.02217 -2.483 0.013325 *
```

EXAMPLE REGRESSION IN R II

| | Dependent variable: |
|-------------------------|-----------------------------|
| | wage |
| female | 1.547*** |
| | (0.482) |
| exper | 0.054*** |
| | (0.015) |
| female:exper | -0.055** |
| | (0.022) |
| Constant | 6.158*** |
| | (0.342) |
| Observations | 526 |
| R ² | 0.136 |
| Adjusted R ² | 0.131 |
| Residual Std. Error | 3.443 (df = 522) |
| F Statistic | 27.307*** (df = 3; 522) |
| Note: | *p<0.1; **p<0.05; ***p<0.01 |



EXAMPLE REGRESSION IN R: INTERPRETTING COEFFICIENTS

$$\widehat{\text{wage}_i} = 6.16 + 0.05 \, \text{Experience}_i - 1.55 \, \text{Female}_i - 0.06 \, \text{Experience}_i \times \text{Female}_i$$
(0.34) (0.02) (0.02)

·
$$\hat{\beta}_0$$
:



Example Regression in R: Interpretting Coefficients

$$\widehat{\text{wage}_i} = 6.16 + 0.05 \, \text{Experience}_i - 1.55 \, \text{Female}_i - 0.06 \, \text{Experience}_i \times \text{Female}_i$$
(0.34) (0.02) (0.49) (0.02)

- \cdot $\hat{eta}_{ exttt{0}}$: Males with experience of 0 years earn \$6.16
- $\cdot \hat{\beta}_1$:



Example Regression in R: Interpretting Coefficients

$$\widehat{\text{wage}_i} = 6.16 + 0.05 \, \text{Experience}_i - 1.55 \, \text{Female}_i - 0.06 \, \text{Experience}_i \times \text{Female}_i$$
(0.34) (0.02) (0.49) (0.02)

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Example Regression in R: Interpretting Coefficients

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(0.34) (0.02) (0.49) (0.02)

- \cdot \hat{eta}_0 : Males with experience of 0 years earn \$6.16
- \hat{eta}_1 : For every additional year of experience, *males* earn \$0.05
- $\cdot \hat{\beta}_2$:



Example Regression in R: Interpretting Coefficients

$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i \times \text{Female}_i$$
(0.34) (0.02) (0.49) (0.02)

- \cdot \hat{eta}_0 : Males with experience of 0 years earn \$6.16
- \cdot \hat{eta}_1 : For every additional year of experience, *males* earn \$0.05
- $\cdot \hat{\beta}_2$:



EXAMPLE REGRESSION IN R: INTERPRETTING COEFFICIENTS

$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i \times \text{Female}_i$$
(0.34) (0.02) (0.49) (0.02)

- \cdot \hat{eta}_0 : Males with experience of 0 years earn \$6.16
- \cdot \hat{eta}_1 : For every additional year of experience, *males* earn \$0.05
- \cdot \hat{eta}_2 : Women on average earn \$1.55 less than men, holding experience constant
- · $\hat{\beta}_3$:



EXAMPLE REGRESSION IN R: INTERPRETTING COEFFICIENTS

$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i \times \text{Female}_i$$
(0.34) (0.02) (0.49) (0.02)

- \cdot \hat{eta}_0 : Males with experience of 0 years earn \$6.16
- \cdot \hat{eta}_1 : For every additional year of experience, *males* earn \$0.05
- \cdot \hat{eta}_2 : Women on average earn \$1.55 less than men, holding experience constant
- · $\hat{\beta}_3$:



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- \cdot \hat{eta}_3 : Females earn \$0.06 less than men for every additional year of experience



$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i * \text{Female}_i$$
(0.34) (0.02) (0.49) (0.02)

• Regression for males (female = 0):

$$\widehat{\text{wage}}_i = 6.16 + 0.05 \text{Experience}_i$$



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• Regression for males (female = 0):

$$\widehat{\text{wage}}_i = 6.16 + 0.05 \text{Experience}_i$$

• Males with no experience earn \$6.16



$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i * \text{Female}_i$$
(0.34) (0.02) (0.49) (0.02)

• Regression for males (female = 0):

$$\widehat{\text{wage}}_i = 6.16 + 0.05 \text{Experience}_i$$

- · Males with no experience earn \$6.16
- · For every year of experience, males' wages increase by \$0.05



$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i * \text{Female}_i$$
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• Regression for males (female = 0):

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- · Males with no experience earn \$6.16
- · For every year of experience, males' wages increase by \$0.05
- Regression for females (female = 1):

$$\widehat{\text{wage}}_i = 6.16 + 0.05 \text{Experience}_i - 1.55(1) - 0.06 \text{Experience}_i * (1)$$

$$= (6.16 - 1.55) + (0.05 - 0.06) \text{Experience}$$

$$= 4.61 - 0.01 \text{Experience}$$



$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i * \text{Female}_i$$
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· Females with no experience earn \$4.61



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• Regression for males (female = 0):

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- Regression for females (female = 1):

$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{Experience}_i - 1.55(1) - 0.06 \text{Experience}_i * (1)$$

$$= (6.16 - 1.55) + (0.05 - 0.06) \text{Experience}$$

$$= 4.61 - 0.01 \text{Experience}$$



- · Females with no experience earn \$4.61
- For every year of experience, females' wages decrease by \$0.01

EXAMPLE REGRESSION IN R: HYPOTHESIS TESTING

$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i * \text{Female}_i$$
(0.34) (0.02) (0.49) (0.02)

How can we test significant differences between two regressions' slopes and intercepts?



$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i * \text{Female}_i$$
(0.34) (0.02) (0.49) (0.02)

- How can we test significant differences between two regressions' slopes and intercepts?
- Different intercepts?



$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i * \text{Female}_i$$
(0.34) (0.02) (0.49) (0.02)

- How can we test significant differences between two regressions' slopes and intercepts?
- Different intercepts?
 - · Difference between male vs. female wages for no experience?



$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i * \text{Female}_i$$
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- · How can we test significant differences between two regressions' slopes and intercepts?
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 - Is β_2 significant?



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- How can we test significant differences between two regressions' slopes and intercepts?
- Different intercepts?
 - · Difference between male vs. female wages for no experience?
 - Is β_2 significant?
 - Yes, $t = \frac{-1.55}{0.48} \approx -3.210$, p(T > t) = 0.000 (from **R** output, above)



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 - Differences between male vs. female change in wages per 1 year of experience?
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$$\widehat{\text{wage}_i} = 6.16 + 0.05 \text{ Experience}_i - 1.55 \text{ Female}_i - 0.06 \text{ Experience}_i * \text{Female}_i$$
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- · Different intercepts?
 - · Difference between male vs. female wages for no experience?
 - Is β_2 significant?
 - Yes, $t = \frac{-1.55}{0.48} \approx -3.210$, p(T > t) = 0.000 (from **R** output, above)
- Different slopes?
 - · Differences between male vs. female change in wages per 1 year of experience?
 - Is β_3 significant?
 - Yes, $t = \frac{0.06}{0.02} \approx -2.483$, p(T > t) = 0.01 (from **R** output, above)



INTERACTIONS BETWEEN TWO DUMMY

VARIABLES

INTERACTIONS BETWEEN TWO DUMMIES

$$Y_{i} = \beta_{0} + \beta_{1}D_{1i} + \beta_{2}D_{2i} + \beta_{3}D_{1i} * D_{2i}$$

• D_{1i} , D_{2i} are dummy variables



$$Y_{i} = \beta_{0} + \beta_{1}D_{1i} + \beta_{2}D_{2i} + \beta_{3}D_{1i} * D_{2i}$$

- D_{1i} , D_{2i} are dummy variables
- β_1 : effect on Y of going from $D_{1i} = 0$ to $D_{1i} = 1$



$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + \beta_3 D_{1i} * D_{2i}$$

- D_{1i} , D_{2i} are dummy variables
- β_1 : effect on Y of going from $D_{1i} = 0$ to $D_{1i} = 1$
- β_2 : effect on Y of going from $D_{1i}=0$ to $D_{2i}=1$ for $D_{1i}=0$



$$Y_{i} = \beta_{0} + \beta_{1}D_{1i} + \beta_{2}D_{2i} + \beta_{3}D_{1i} * D_{2i}$$

- D_{1i} , D_{2i} are dummy variables
- β_1 : effect on Y of going from $D_{1i}=0$ to $D_{1i}=1$
- β_2 : effect on Y of going from $D_{1i} = 0$ to $D_{2i} = 1$ for $D_{1i} = 0$
- β_3 : increment to effect on Y of going from $D_{1i}=0$ to $D_{1i}=1$ when $D_{2i}=1$ (vs. when $D_{2i}=0$)



$$Y_{i} = \beta_{0} + \beta_{1}D_{1i} + \beta_{2}D_{2i} + \beta_{3}D_{1i} * D_{2i}$$

- D_{1i} , D_{2i} are dummy variables
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- β_3 : increment to effect on Y of going from $D_{1i} = 0$ to $D_{1i} = 1$ when $D_{2i} = 1$ (vs. when $D_{2i} = 0$)
- \cdot Again, best to think logically about the possibilities (when each dummy = 0 or = 1)



$$Y_{i} = \beta_{0} + \beta_{1}D_{1i} + \beta_{2}D_{2i} + \beta_{3}D_{1i} * D_{2i}$$

 $\boldsymbol{\cdot}$ To interpret the coefficients, compare cases:



$$Y_{i} = \beta_{0} + \beta_{1}D_{1i} + \beta_{2}D_{2i} + \beta_{3}D_{1i} * D_{2i}$$

• To interpret the coefficients, compare cases:

$$E(Y_i|D_{1i} = 0, D_{2i} = d_2) = \beta_0 + \beta_2 d_2$$

• Subtracting the two, the difference between is:

$$\beta_1 + \beta_3 d_2$$



$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + \beta_3 D_{1i} * D_{2i}$$

• To interpret the coefficients, compare cases:

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$$E(Y_i|D_{1i} = 1, D_{2i} = d_2) = \beta_0 + \beta_1 + \beta_2 d_2 + \beta_3 d_2$$

· Subtracting the two, the difference between is:

$$\beta_1 + \beta_3 d_2$$

· The effect of $D_{1i} \rightarrow Y_i$ depends on d_{2i}



$$Y_{i} = \beta_{0} + \beta_{1}D_{1i} + \beta_{2}D_{2i} + \beta_{3}D_{1i} * D_{2i}$$

· To interpret the coefficients, compare cases:

$$E(Y_i|D_{1i} = 0, D_{2i} = d_2) = \beta_0 + \beta_2 d_2$$

$$E(Y_i|D_{1i} = 1, D_{2i} = d_2) = \beta_0 + \beta_1 + \beta_2 d_2 + \beta_3 d_2$$

· Subtracting the two, the difference between is:

$$\beta_1 + \beta_3 d_2$$

- The effect of $D_{1i} \rightarrow Y_i$ depends on d_{2i}
- β_3 : increment to the effect of D_1 when $D_2 = 1$



INTERACTIONS BETWEEN TWO DUMMIES: EXAMPLE

Example

$$\widehat{wage_i} = \hat{eta}_0 + \hat{eta}_1$$
femal $e_i + \hat{eta}_2$ marrie $d_i + \hat{eta}_3$ femal $e_i *$ marrie d_i



Example

Return to the gender pay gap: does it matter if person is married or single?

$$\widehat{wage_i} = \hat{eta}_0 + \hat{eta}_1$$
femal $e_i + \hat{eta}_2$ marrie $d_i + \hat{eta}_3$ femal $e_i *$ marrie d_i

· Logically, 4 possible combinations of $female_i = \{0,1\}$ and $married_i = \{0,1\}$



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- · Logically, 4 possible combinations of $female_i = \{0,1\}$ and $married_i = \{0,1\}$
 - 1. Unmarried males (female_i = 0, married_i = 0)



Example

$$\widehat{wage_i} = \hat{eta}_0 + \hat{eta}_1$$
femal $e_i + \hat{eta}_2$ marrie $d_i + \hat{eta}_3$ femal $e_i *$ marrie d_i

- · Logically, 4 possible combinations of $female_i = \{0,1\}$ and $married_i = \{0,1\}$
 - 1. Unmarried males (female $_i=0$, married $_i=0$) $\widehat{wage}_i=\hat{\beta}_0$



Example

$$\widehat{wage_i} = \hat{eta}_0 + \hat{eta}_1$$
femal $e_i + \hat{eta}_2$ marrie $d_i + \hat{eta}_3$ femal $e_i *$ marrie d_i

- · Logically, 4 possible combinations of $\textit{female}_i = \{0,1\}$ and $\textit{married}_i = \{0,1\}$
 - 1. Unmarried males (female $_i=0$, married $_i=0$) $\widehat{wage}_i=\hat{\beta}_0$
 - 2. Married males (female $_i = 0$, married $_i = 1$)



Example

$$\widehat{wage_i} = \hat{eta}_0 + \hat{eta}_1$$
femal $e_i + \hat{eta}_2$ marrie $d_i + \hat{eta}_3$ femal $e_i *$ marrie d_i

- · Logically, 4 possible combinations of $female_i = \{0,1\}$ and $married_i = \{0,1\}$
 - 1. Unmarried males (female $_i=0$, married $_i=0$) $\widehat{wage}_i=\hat{\beta}_0$
 - 2. Married males (female_i = 0, married_i = 1) $\widehat{wage}_i = \hat{\beta}_0 + \hat{\beta}_2$



Example

$$\widehat{wage_i} = \hat{eta}_0 + \hat{eta}_1$$
femal $e_i + \hat{eta}_2$ marrie $d_i + \hat{eta}_3$ femal $e_i *$ marrie d_i

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 - 1. Unmarried males (female $_i=0$, married $_i=0$) $\widehat{wage}_i=\hat{\beta}_0$
 - 2. Married males (female_i = 0, married_i = 1) $\widehat{wage}_i = \hat{\beta}_0 + \hat{\beta}_2$
 - 3. Unmarried females (female_i = 1, married_i = 0)



INTERACTIONS BETWEEN TWO DUMMIES: EXAMPLE

Example

Return to the gender pay gap: does it matter if person is married or single?

$$\widehat{wage_i} = \hat{eta}_0 + \hat{eta}_1$$
female $_i + \hat{eta}_2$ married $_i + \hat{eta}_3$ female $_i *$ married $_i$

- · Logically, 4 possible combinations of $female_i = \{0,1\}$ and $married_i = \{0,1\}$
 - 1. Unmarried males (female $_i=0$, married $_i=0$) $\widehat{wage}_i=\hat{\beta}_0$
 - 2. Married males (female_i = 0, married_i = 1) $\widehat{wage}_i = \hat{\beta}_0 + \hat{\beta}_2$
 - 3. Unmarried females (female_i = 1, married_i = 0) $\widehat{wage}_i = \hat{\beta}_0 + \hat{\beta}_1$



Interactions Between Two Dummies: Example

Example

Return to the gender pay gap: does it matter if person is married or single?

$$\widehat{wage_i} = \hat{eta}_0 + \hat{eta}_1$$
femal $e_i + \hat{eta}_2$ marrie $d_i + \hat{eta}_3$ femal $e_i *$ marrie d_i

- · Logically, 4 possible combinations of $female_i = \{0,1\}$ and $married_i = \{0,1\}$
 - 1. Unmarried males (female $_i=0$, married $_i=0$) $\widehat{wage}_i=\hat{\beta}_0$
 - 2. Married males (female_i = 0, married_i = 1) $\widehat{wage}_i = \hat{\beta}_0 + \hat{\beta}_2$
 - 3. Unmarried females (female $_i=1$, married $_i=0$) $\widehat{wage}_i=\hat{\beta}_0+\hat{\beta}_1$
 - 4. Married females (female_i = 1, married_i = 1)



Interactions Between Two Dummies: Example

Example

Return to the gender pay gap: does it matter if person is married or single?

$$\widehat{wage_i} = \hat{eta}_0 + \hat{eta}_1$$
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 - 1. Unmarried males (female_i = 0, married_i = 0) $\widehat{wage}_i = \hat{\beta}_0$
 - 2. Married males (female_i = 0, married_i = 1) $\widehat{wage}_i = \hat{\beta}_0 + \hat{\beta}_2$
 - 3. Unmarried females (female $_i=1$, married $_i=0$) $\widehat{wage}_i=\hat{\beta}_0+\hat{\beta}_1$
 - 4. Married females (female_i = 1, married_i = 1) $\widehat{wage_i} = \hat{\beta}_0 + \hat{\beta}_1 + \hat{\beta}_2 + \hat{\beta}_3$



Interactions Between Two Dummies: Conditionally Looking at the Data

```
# get average wage for unmarried men
mean(wages$wage[wages$married==0 & wages$female==0])
## [1] 5.168023
# get average wage for married men
mean(wages$wage[wages$married==1 & wages$female==0])
## [1] 7.983032
# get average wage for unmarried women
mean(wages$wage[wages$married==0 & wages$female==1])
## [1] 4.611583
# get average wage for married wommen
```



mean(wages\$wage[wages\$married==1 & wages\$female==1])

Interactions Between Two Dummies: Group Means

$$\widehat{\text{wage}_i} = \hat{eta}_0 + \hat{eta}_1$$
femal $e_i + \hat{eta}_2$ marrie $d_i + \hat{eta}_3$ femal $e_i *$ marrie d_i

| | Unmarried | Married |
|--------|-----------|---------|
| Male | \$5.17 | \$7.98 |
| Female | \$4.61 | \$4.57 |
| | | |



Interactions Between Two Dummies: Regression in R

```
reg.2dummies.interact<-lm(wage~female+married+female*married, data=wages)
summary(reg.2dummies.interact)
##
## Call:
## lm(formula = wage ~ female + married + female * married, data = wages)
##
## Residuals:
##
      Min
              10 Median 30
                                    Max
## -5.7530 -1.7327 -0.9973 1.2566 17.0184
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
                  5.1680
                             0.3614 \ 14.299 < 2e-16 ***
## (Intercept)
## female
              -0.5564
                            0.4736 -1.175 0.241
## married
          2.8150 0.4363 6.451 2.53e-10 ***
## female:married -2.8607 0.6076 -4.708 3.20e-06 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
```

Interactions Between Two Dummies: Regression in R II

| Dependent variable | | |
|-------------------------|-----------------------------|--|
| | wage | |
| female | -0.556 | |
| | (0.474) | |
| married | 2.815*** | |
| | (0.436) | |
| female:married | -2.861*** | |
| | (0.608) | |
| Constant | 5.168*** | |
| | (0.361) | |
| Observations | 526 | |
| R^2 | 0.181 | |
| Adjusted R ² | 0.176 | |
| Residual Std. Error | 3.352 (df = 522) | |
| F Statistic | 38.451*** (df = 3; 522) | |
| Note: | *n<0.1· **n<0.05· ***n<0.01 | |



$$\widehat{\text{wage}_i} = 5.17 - 0.56 \, \text{Female}_i + 2.82 \, \text{Married}_i - 2.86 \, \text{Female}_i * \text{Married}_i$$

$$(0.36) \quad (0.47) \qquad (0.44) \qquad (0.61)$$

Average Wage for each Grouping

| | Unmarried | Married |
|--------|-----------|---------|
| Male | \$5.17 | \$7.98 |
| Female | \$4.61 | \$4.57 |
| | | |

 \cdot Wage for **unmarried males**: $\hat{eta}_0 = \$$ 5.17



$$\widehat{\text{wage}_i} = 5.17 - 0.56 \, \text{Female}_i + 2.82 \, \text{Married}_i - 2.86 \, \text{Female}_i * \text{Married}_i$$

$$(0.36) \quad (0.47) \qquad (0.44) \qquad (0.61)$$

| | Unmarried | Married |
|--------|-----------|---------|
| Male | \$5.17 | \$7.98 |
| Female | \$4.61 | \$4.57 |
| | | |

- Wage for **unmarried males**: $\hat{\beta}_0 = \$5.17$
- · Wage for married males: $\hat{\beta}_0 + \hat{\beta}_2 =$ 5.17 + 2.82 = \$7.98



$$\widehat{\text{wage}_i} = 5.17 - 0.56 \, \text{Female}_i + 2.82 \, \text{Married}_i - 2.86 \, \text{Female}_i * \text{Married}_i$$
(0.36) (0.47) (0.44) (0.61)

| | Unmarried | Married |
|--------|-----------|---------|
| Male | \$5.17 | \$7.98 |
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| | | |

- \cdot Wage for **unmarried males**: $\hat{eta}_0 = \$$ 5.17
- Wage for married males: $\hat{\beta}_0 + \hat{\beta}_2 = 5.17 + 2.82 = \7.98
- Wage for unmarried females: $\hat{\beta}_0 + \hat{\beta}_1 = 5.17 0.56 = \4.61



$$\widehat{\text{wage}_i} = 5.17 - 0.56 \text{ Female}_i + 2.82 \text{ Married}_i - 2.86 \text{ Female}_i * \text{Married}_i$$

$$(0.36) \quad (0.47) \qquad (0.44) \qquad (0.61)$$

| | Unmarried | Married | |
|--------|-----------|---------|---|
| Male | \$5.17 | \$7.98 | |
| Female | \$4.61 | \$4.57 | |
| | | | _ |

- · Wage for **unmarried males**: $\hat{eta}_0 = \$5.17$
- · Wage for married males: $\hat{\beta}_0 + \hat{\beta}_2 = 5.17 + 2.82 = \7.98
- Wage for unmarried females: $\hat{\beta}_0 + \hat{\beta}_1 = 5.17 0.56 = \4.61
- Wage for married females: $\hat{\beta}_0 + \hat{\beta}_1 + \hat{\beta}_2 + \hat{\beta}_3 = 5.17 0.56 + 2.82 2.86 = 4.57





Interactions Between Two Continuous Variables

$$Y_{i} = \beta_{0} + \beta_{1}X_{1i} + \beta_{2}X_{2i} + \beta_{3}(X_{1i}X_{2i})$$

- To interpret the coefficients, compare cases after changing ΔX_1 :

$$Y + \Delta Y = \beta_0 + \beta_1(X_1 + \Delta X_1)\beta_2X_2 + \beta_3((X_1 + \Delta X_1)X_2)$$

-The difference is:

$$\Delta Y = \beta_1 \Delta X_1 + \beta_3 X_2 \Delta X_1$$
$$\frac{\Delta Y}{\Delta X_1} = \beta_1 + \beta_3 X_2$$

• The effect of X_1 depends on X_2



INTERACTIONS BETWEEN TWO CONTINUOUS VARIABLES

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- The effect of X_1 depends on X_2
- \cdot β_3 : increment to the effect of X_1 from a 1 unit change in X_2



INTERACTIONS BETWEEN TWO CONTINUOUS VARIABLES: EXAMPLE

Example

Wages on education and experience: Do education & experience interact?

$$\widehat{wage_i} = \hat{eta}_0 + \hat{eta}_1 educ_i + \hat{eta}_2 exper_i + \hat{eta}_3 educ_i \times exper_i + \epsilon_i$$



Interactions Between Two Continuous Variables: Example

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• Estimated effect of education on wages depends on the amount of experience (and vice versa)!

$$\frac{\Delta wage}{\Delta educ} = \hat{\beta}_1 + \beta_3 exper$$



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· This is a type of nonlinearity (we will examine nonlinearities next lesson)



Interactions Between Two Continuous Variables: Regression in R

```
reg.2x.interact<-lm(wage~educ+exper+educ*exper, data=wages)</pre>
summary(reg.2x.interact)
##
## Call:
## lm(formula = wage ~ educ + exper + educ * exper. data = wages)
##
## Residuals:
      Min
          10 Median 30
##
                                       Max
## -5.6747 -1.9683 -0.6991 1.2803 15.8067
##
## Coefficients:
                Estimate Std. Error t value Pr(>|t|)
##
```

Interactions Between Two Continuous Variables: Regression in R II

| | Dependent variable: | |
|-------------------------|-------------------------|--|
| | wage | |
| educ | 0.602*** | |
| | (0.090) | |
| exper | 0.046 | |
| | (0.043) | |
| educ:exper | 0.002 | |
| | (0.003) | |
| Constant | -2.860** | |
| | (1.181) | |
| Observations | 526 | |
| R^2 | 0.226 | |
| Adjusted R ² | 0.221 | |
| Residual Std. Error | 3.259 (df = 522) | |
| F Statistic | 50.713*** (df = 3; 522) | |
| Note: | *n<01·**n<005·***n<00 | |



Interactions Between Two Continuous Variables: Interpretting Coefficients

$$\widehat{\text{wage}}_i = -2.86 + 0.60 \text{ educ}_i + 0.05 \text{ exper}_i + 0.002 \text{ educ}_i \times \text{exper}_i$$
(1.181) (0.090) (0.043) (0.003)



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Changes in Education

| Exper | $rac{\Delta$ wage Δ educ | |
|-------|----------------------------------|--|
| 5 | 0.60 + 0.002(5) = \$0.61 | |
| 10 | 0.60 + 0.002(10) = \$0.62 | |
| 15 | 0.60 + 0.002(15) = \$0.63 | |



Interactions Between Two Continuous Variables: Interpretting Coefficients

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