صل تحرین اور کی کی آماری حلب کی دوم

🗶 مقدمات:

$$Var(X) = E(X^{Y}) - E^{Y}(X) \Rightarrow E(X^{Y}) = Var(X) + E^{Y}(X)$$

.Theorem.

$$E(W) = E(aX + bY) = aE(X) + bE(Y)$$

Theorem.

$$\operatorname{var}(aX + bY) = a^{2} \operatorname{var} X + b^{2} \operatorname{var} Y + 2ab \operatorname{cov}(X, Y)$$

$$C \cdot v(X, Y) = 0$$

$$A \cdot V \cdot X \cdot F^{2}$$

۲X برادردگر ناارب ۱۵ اس.

$$E(\lambda \underline{X}) = \lambda E(\underline{X}) = \lambda \frac{\lambda}{\theta} = \theta$$

2. If $\hat{\Theta}_1$ and $\hat{\Theta}_2$ are unbiased estimators of the same parameter θ , what condition must be imposed on the constants k_1 and k_2 so that

$$k_1\hat{\Theta}_1 + k_2\hat{\Theta}_2$$

is also an unbiased estimator of θ ?

$$E(k, \hat{\theta}, + k, \hat{\theta},) = k, E(\hat{\theta},) + k, E(\hat{\theta},) = k, \theta + k, \theta = \theta$$

$$\Rightarrow \theta(k_1 + k_7) = \theta \Rightarrow k_1 + k_7 = 1$$

3. Based on a sample of 2 observations, consider the two estimators of μ :

$$\bar{X} = (\frac{1}{2})X_1 + (\frac{1}{2})X_2$$

and

$$W \stackrel{\Delta}{=} \left(\frac{1}{3}\right) X_1 + \left(\frac{2}{3}\right) X_2$$

(a) Prove they are unbiased.

$$E\left(\frac{x_{1}}{x_{1}} + \frac{x_{4}}{x_{4}}\right) = \frac{1}{1}\left[E\left(x_{1}\right) + E\left(x_{4}\right)\right] = \frac{1}{1}\left[E\left(x_{1}\right) + E\left(x_{4}\right)\right]$$

$$E\left(\frac{x_{1}}{x_{1}}+\frac{x_{2}}{4x_{2}}\right)=\frac{x_{1}}{1}E(x_{1})+\frac{x_{2}}{1}E(x_{2})=\frac{x_{1}}{1}V=V$$

$$Var\left(\frac{\lambda}{\lambda} + \frac{\lambda}{\lambda}\right) = \frac{1}{\xi} \left(Aar\left(\chi^{1}\right) + Aar\left(\chi^{2}\right) \right) = \frac{\xi_{1}}{\xi_{1}}$$

$$\operatorname{Aur}\left(\frac{\lambda}{X^{1}} + \frac{\lambda}{X^{1}}\right) = \frac{d}{1} \operatorname{Aur}\left(X^{1}\right) + \frac{d}{2} \operatorname{Aur}\left(X^{1}\right) = \frac{d}{2} e_{L}$$

اولی کیماترات.

مر کنی کر سر ۲ مر ۲ (X موشای تعادنی از توزیع (۴ مر) ۱ با کد. برای

التادر ما براردرع، إلا مترا، الروع را تين كند.

۲ ارامترداری ے ۲ =۲

$$w'_{i} = \frac{1}{\sum x'_{i}} = \frac{1}{\sum x'_{i}}$$

$$M_{i}^{k} = \frac{1}{\sum X_{i}^{k}} = \frac{1}{\sum X_{i}$$

$$\begin{cases} w'_1 = w'_1 \Rightarrow \hat{w} = \overline{X} \\ w'_1 = w'_2 \Rightarrow \hat{v} \Rightarrow \hat{v} = \overline{X} \end{cases} \Rightarrow \hat{v} \Rightarrow \hat{v$$

$$L(\theta) = \frac{1}{\prod_{i=1}^{n}} \theta \times_{i}^{\theta-1} = \theta^{n} \left(\prod_{i=1}^{n} \times_{i} \right)^{\theta-1}$$

$$\Rightarrow l_n(L(\theta)) = n l_n \theta + (\theta - 1) l_n(TT_{\alpha_i}) = l(\theta)$$

$$\Rightarrow \frac{l}{l\theta} l_{\alpha}(\theta) = \frac{n}{\theta} + l_{\alpha}(\pi_{\alpha_{i}}) = 0 \Rightarrow \hat{\theta} = \frac{-n}{l_{\alpha}(\tilde{\pi}_{\alpha_{i}})}$$