# Spatial variation in allometric growth of invasive lionfish has management implications

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# ABSTRACT

Lionfish (*Pterois volitans / miles*) are an invasive species in the Western Atlantic and the Caribbean. Improving management of invasive lionfish populations requires accurate total biomass estimates, which depend on accurate estimates of allometric growth, but sedentary species like lionfish often exhibit high levels of spatial variation in life history characteristics. We reviewed 17 published length-weight relationships for lionfish taken throughout their invasive range and found regional differences that led to significant misestimates when calculating weight from length observations. The spatial pattern we observed is consistent with findings from other studies focused on genetics or length-at-age. Here, the use of *ex situ* parameter values resulted in total biomass estimates between 76.2% and 140% of true observed biomass, and up to a threefold under- or overestimation of total weight for an individual organism. These findings can have implications for management in terms of predicting effects on local ecosystems, evaluating the effectiveness of removal programs, or estimating biomass available for harvest.

### INTRODUCTION

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Lionfish (*Pterois volitans/miles* complex) are an invasive species in the western Atlantic Western Atlantic

Ocean and Caribbean Sea, likely introduced through liberation release of aquarium-kept organisms

(Betancur-R et al., 2011). They Lionfish are the first invasive marine vertebrates established along the

North Atlantic Caribbean coasts (Schofield, 2009, 2010; Sabido-Itza et al., 2016) and their these coasts

(Schofield, 2009, 2010; Sabido-Itza et al., 2016), and they have established invasive populations in coral reefs, estuaries, mangroves, hard-bottomed areas, and mesophotic reefs (Barbour et al., 2010; Jud et al., 2011; Muñoz et al., 2016). Lionfish have established invasive populations in coral reefs, estuaries, mangroves, hard-bottomed areas, and mesophotic reefs

(Barbour et al., 2010; Jud et al., 2011; Muñoz et al., 2011; Claydon et al., 2012; Andradi-Brown et al., 2017; Gress et

A substantial amount of research describes lionfish impacts throughout its their invaded range. A meta-analysis by Peake et al. (2018) showed that invasive lionfish prey on at least 167 different species across the tropical and temperate North Western Atlantic. Their feeding behavior and high consumption rates can reduce recruitment and population sizes of native reef-fish species, and can further endanger reef fish (Green et al. (2012); Rocha et al. (2015); but see Hackerott et al. (2017) )(Green et al., 2012; Rocha et al., 2015, but see Hackerott et al. 2017 for a counterexample). For example, field experiments by Albins and Hixon (2008) showed that lionfish establishment in the Bahamas led to reduced recruitment of native fishes by nearly 80% over a five-week period in Florida. Green et al. (2012) reported that (Albins and Hixon, 2008), and prey fish biomass declined by 65% over two years as lionfish biomass increased along Bahamian coral reefs. Their trophic impacts (Green et al., 2012). However, trophic impacts of lionfish can be minimized if local lionfish their local biomass is controlled by culling (Arias-Gonzalez et al., 2011).

Governments and non-profit organizations have sought to reduce lionfish densities through removal programs and by incentivizing its consumption (Chin et al., 2016). In some cases, these have shown to significantly reduce —but not quite eliminate — but not quite eliminate — lionfish abundances at local scales (de Leon et al., 2013; Sandel et al., 2015). Complete eradication of lionfish through fishing is unlikely because of their rapid recovery rates and ongoing recruitment to shallow-water areas from persistent populations in mesophotic ecosystems (Barbour et al., 2011; Andradi-Brown et al., 2017). However, promoting lionfish consumption might create a level of demand capable of incentivizing a stable fishery while controlling shallow-water populations, thus creating alternative livelihoods and avoiding further impacts—negative effects to local biota.

The feasibility of establishing fisheries through lionfish removal programs has been extensively evaluated through field observations and empirical modeling (Barbour et al., 2011; Morris et al., 2011; de Leon et al., 2013; Johnston and Purkis, 2015; Sandel et al., 2015; Usseglio et al., 2017). Determining the feasibility of such initiatives requires modeling the change in biomass in response to changes in fishing mortality (*i.e.*, culling). A common way to model this is via length-structured population models, where fish lengths are converted to weight to calculate total biomass (Barbour et al., 2011; Côté et al., 2014; Andradi-Brown et al., 2017). The allometric length-weight relationship is thus an essential component of these models, but this relationship can vary across regions as a response to biotic and abiotic conditions (Johnson and Swenarton, 2016).

Outcomes of previous studies suggest lionfish are likely to exhibit spatial heterogeneity in the lengthweight relationship, which we summarize in two main causes. First, culling programs are effective in reducing local adult populations largely because lionfish exhibit for both behavioral and biological reasons. Important life history characteristics such as growth or natural mortality rates are often spatially variable for fish that exhibit sedentary behavior (Gunderson et al., 2008; Hutchinson, 2008; Wilson et al., 2012; Guan et al., , and in fact, high levels of site fidelity and small home ranges (Fishelson, 1997; Kochzius and Blohm, 2005; Jud and Lay . It is know that fish with sedentary behavior are likely to exhibit high levels of spatial variation in important life history characterstics such as growth or natural mortality rates (Gunderson et al., 2008; Hutchinson, 2008 . Second, genetic are two primary reasons why culling programs are effective in reducing local adult lionfish populations (Fishelson, 1997; Kochzius and Blohm, 2005; Jud and Layman, 2012; Côté et al., 2014) Genetic analysis of lionfish suggests biological differences due to the existence of also identified two genetically distinct invasive subpopulations between the northwest-Western Atlantic and the Caribbean, suggesting the existence of spatially explicit biological differences between populations as well (Betancur-R et al., 2011). Site-specific parameters are necessary to accurately estimate biomass when allometric relationships are spatially variable, and this variability is studies that calculate the length-weight relationship of lionfish report variable estimates, and these differences may be increasingly important when estimating the potential effectiveness of lionfish culling programs (Barbour et al., 2011; Morris et al., 2011; Côté et al., 2014; Johnston and Purkis, 2015). However, the region-wide differences in allometric growth parameters has remained unexplored for lionfish, despite the large number of site-specific studies reporting the influence of using ex situ parameters when estimating the length-weight relationship remains unexplored.

by using *ex situ* parameter values when estimating lionfish weight from length observations. In this study, we calculated and reported the first length-weight relationships for lionfish populations in North Carolina, Northern and Southern Gulf of Mexico, the Southern Mexican Caribbean, Bahamas, Little Cayman, Jamaica, Bonaire, Puerto Rico, and Costa Rica (Barbour et al., 2011; Darling et al., 2011; de Leon et al., 2013; Fogg et . We also collected lionfish length and weight data in the central Mexican Caribbean and report the first allometric growth equation for this region. The objective of this paper is to describe the spatial pattern of relationship for lionfish in the central Mexican Caribbean using previously collected *in situ* observations (n = 109; Villaseñor-Derbez and Herrera-Pérez (2014)). We then estimated lionfish weight in this area using previously published length-weight relationships of lionfish across the Caribbean and Western Atlanticand to discuss implications of these spatial differences for lionfish populations from ten locations across the Western Atlantic, Gulf of Mexico, and Caribbean. By comparing these weight estimates to our *in situ* length-weight observations, we showed that using *ex situ* parameter values resulted in up to a threefold under- or overestimation of lionfish weight and estimated of total biomass ranged between 76% and 140% of observed total biomass.

Here, we compare previously published Our objective was to quantify the magnitude of error caused

### METHODS

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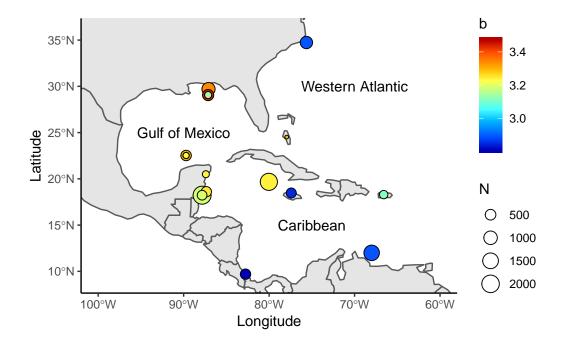
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We reviewed 12 published studies and obtained 17 length-weight relationships for the North-Western Atlantic (n =  $\frac{42}{1}$ ), Gulf of Mexico (n = 7), and Caribbean (n =  $\frac{98}{1}$ , Table 1, Fig  $\frac{3}{1}$ ). We 1). Study sites included 101 North Carolina, the Northern and Southern Gulf of Mexico, the Southern Mexican Caribbean, the Bahamas, Little Cayman, Jamaica, Bonaire, Puerto Rico, and Costa Rica (Barbour et al., 2011; Darling et al., 2011; de 103 . We have access only to the summarized information published in these studies - not the raw data authors used to make length-weight calculations. We collected information on sampling methods, sex differenti-105 ation, location, length and depth ranges, and sampling methods from each study when available. Only 106 two studies reported parameters for each gender sex specific length weight parameters (Aguilar-Perera 107 and Quijano-Puerto, 2016; Fogg et al., 2013), so we assumed both genders were included in a study 108 if gender was unspecified data were reported for both sexes combined in all other studies. Reviewed 109 studies presented information for organisms ranging from 25-475 mm in Total Length (TL) and were 110 obtained at depths between 0.5 m and 57 m. Three-Four studies explicitly stated that their organisms 111 were sampled with pole spears (Dahl and Patterson, 2014; Aguilar-Perera and Quijano-Puerto, 2016; 112 Chin et al., 2016; Sabido-Itzá et al., 2016), and five six studies mentioned that some of their organ-113 isms were obtained with pole spears (or other type of harpoon) but also hand-held nets or fish traps 114 (Barbour et al., 2011; Fogg et al., 2013; Edwards et al., 2014; Toledo-Hernández, 2014; Sandel et al., 2015; Sabido-Itz 115 , and two (Barbour et al., 2011; Fogg et al., 2013; Edwards et al., 2014; Toledo-Hernández, 2014; Sandel et al., 2015; Sandel et al., 2 116 Two studies did not specify how organisms were sampled (Darling et al., 2011; de Leon et al., 2013). Fogg et al. (2013) use spineless weight in their calculations, so their parameters likely underestimated 118 total wieght. Since no spineless to total weight conversions were available, these parameters were taken 119 as reported. 120

Locations where allometric growth parameters of lionfish (*Pterois spp*) have been reported. Circle sizes indicate sample size from each study, colors indicate the *b* coefficient from Eq. 1.

We also collected data from We also used data from 109 lionfish sampled by Villaseñor-Derbez and Herrera-Pérez (2, who used hand nets and numbered bottles to collect Total Length (TL; mm) and Total Weight (TW; g) for organisms from 10 sampling sites along the central Mexican Caribbean coast in the Summer of 2010 (Supplementary Table 1). Sampling locations included wall and carpet reefs at depths between 5.7 m and 38.1 m. All observed lionfish (n = 109) were collected using hand nets and numbered collection bottles. The use of hand nets prevented any weight loss due to bleeding and allowed better representation of small sizes by eliminating avoiding gear selectivity. Organisms were euthanized via pithingand Total Length (TL; mm) and Total Weight (TW; g) were recorded.



**Figure 1.** Locations where allometric growth parameters of lionfish (*Pterois spp*) have been reported. Circle sizes indicate sample size from each study, colors indicate the *b* coefficient from Eq. 1.

The weight-at-length relationship for lionfish in the central Mexican Caribbean was calculated with the allometric growth function:

$$TW = aTL^b (1)$$

Where a is the ponderal index and b is the scaling exponent or allometric parameter. Transforming this equation via base-10 logarithms we obtain:

$$log_{10}(TW) = b \times log_{10}(TL) + log_{10}(a)$$

This can be simplified and re-written as:

$$Y = bX + c$$

Where  $Y = log_{10}(TW)$ ,  $X = log_{10}(TL)$ , and  $c = log_{10}(a)$ . The coefficients (c and b) were estimated with We linearized the equation using  $log_{10}$ -transformation and estimated the coefficients using an Ordinary Least Squares Regression and with a heteroskedastic-robust standard error correction (Zeileis, 2004). When the b = 3, it is said that the organism exhibits a perfect isometric growth, so the b coefficient was tested against the null hypothesis of isometric growth (i.e.  $H_0:b=3$ ). Coefficients were tested with a two-tailed Student's tt-test, and the significance of the regression was corroborated with an F-test.

Some of the reviewed studies (Table 1, Fig 1) inconsistently defined a as either the ponderal index from Eq. 1 or the y-intercept (c) from Eq. ??. from the linearized log-transformed equation. Other studies incorrectly reported parameters as mm-to-g conversions when they were in fact cm-to-g conversions. We standardized each study by converting coefficients and report all parameters as TL (mm) to TW (grg) conversions. Locations where allometric studies have been performed are shown in Figure 1 and summarized in Table 1.

We obtained a total of 18 parameter pairs by combining length-weight parameters extracted from the literature and the additional pair calculated here . We used the central Mexican Caribbean as a case

study of how the use of (Fig 3). Recall that the objective of this study is not to describe differences in 150 the length-weight relationship between populations (which would require access to raw data), but rather 151 to assess how ex situ parameters influences parameter values influence the accuracy of weight estimates 152 for lionfish. We estimated TW from the TL observations we, using the central Mexican Caribbean as a 153 case study. Using each of the 18 parameter pairs, we estimated TW from the TL observations collected 154 in the central Mexican Caribbean (n = 109, with  $TL \in (34,310)$ ) using each of the 18 parameter pairs TL155 ranging from 34 mm to 310 mm) and divided predicted weights by known observed weights to obtain 156 a simple measure of over- or underestimation. Difference in mean weight ratios across the different parameter pairs were tested with a one-way analysis of variance (ANOVA) and an analysis of covariance 158 (ANCOVA):

$$R_{i,j} = \tilde{\mu} + \alpha_j + \beta T L_{ij} + e_{ij} \tag{2}$$

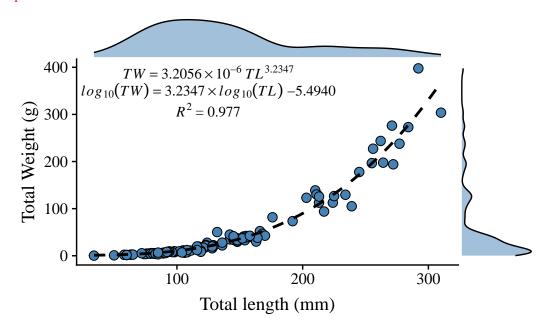
Where  $R_{ij}$  is the weight ratio for the *i*-th organism obtained with parameters from the *j*-th study,  $\tilde{\mu}$  is a constant for all individuals,  $a_i$  is the treatment effect (*i.e.*, the difference induced by each study),  $TL_{ij}$  is the covariate (*i.e.*, Total Length) for the *i*-th subject in the *j*-th group with slope  $\beta$ , and  $e_{ij}$  is the error term of the regression. Ratios were logit-transformed prior to analysis, and a *post-hoc* Tukey's test was used for *post-hoc* tests to identify groups where mean ratios did not differ. All analyses were performed in R version 3.5.1-2 (R Core Team, 2018). Raw data and code used in this work are available on github at github.com/jcyday/lionfish\_biometry.

**Table 1.** Summary of 18 allometric growth parameters available for lionfish in the invaded range from peer-reviewed literature and this study. All parameters have been adjusted to convert from millimeters to grams. n = Sample size, Sex specifies whether data was presented for Females (F), Males (M), or both sexes combined (B),  $a = scaling parameter (presented in <math>\times 10^{-5})$ , b = exponent.

Region	<u>Sex</u> _	<u>n</u>	<u>a</u>	Ď	$R^2$	Reference
Western Atlantic	<u>B</u>	<u>774</u>	2.90	2.89		Barbour et al., 2011
Western Atlantic	$\mathbf{B}$	ā	0.25	3.29	ā	Darling et al., 2011
$\widetilde{\text{GoM}}$	$\mathbf{B}$	934	0.21	3.34	0.98	Dahl & Patterson, 2014
GoM	$\mathbf{B}$	472 67 59 582	0.29	3.30	0.95	Aguilar-Perera & Quijano-Puerto, 2016
GoM	$\overset{\mathbf{F}}{\approx}$	<u>67</u>	0.12	3.47	0.95	Aguilar-Perera & Quijano-Puerto, 2016
GoM	M	59	$\underbrace{0.12}_{0.42}$ $\underbrace{0.42}_{0.14}$	3.47 3.23 3.43	0.95 0.95 0.99	Aguilar-Perera & Quijano-Puerto, 2016
<u>GoM</u>	$\mathbf{B}$	582	0.14	3.43	0.99	Fogg et al., 2013
<u>GoM</u>	$\underbrace{\mathbf{M}}_{}$	119 115 458 458	0.27	3.31	0.97	Fogg et al., 2013
GoM	$\stackrel{\mathbf{F}}{\approx}$	115	0.68	3.14	0.94	Fogg et al., 2013
Caribbean	$\mathbf{B}$	458	3.60	2.81	≅	Sandel et al., 2015
Caribbean	$\mathbf{B}$	<u>419</u>	2.80	2.85	0.87	Chin et al., 2016
Caribbean		1450	2.30	2.89	0.92	de Leon et al., 2013
Caribbean	$\bigotimes_{B}^{B}$	1887	0.30	2.89 3.24 3.18	0.92 0.97 0.99	Edwards et al., 2014
Caribbean	B	2143	0.52	3.18	0.99	Sabido-Itza et al., 2016
Caribbean	B	227	0.80	3.11	0.96	Toledo-Hernández et al., 2014
Caribbean	$\mathbf{B}$	449	$\underbrace{0.80}_{0.23}$	3.25	0.97	Sabido-Itza et al., 2016b
Caribbean	$\underbrace{\mathbf{B}}_{}$	368	0.32	3.25 3.19	0.96 0.97 0.98	Sabido-Itza et al., 2016b
Caribbean	$\mathbf{B}$	109	0.32	3.23	0.98	This study

# **RESULTS**

The length-weight relationship for organisms from the central Mexican Caribbean resulted in the eoefficient values  $a = 3.2056297 \times 10^{-6}$ , b = 3.2347391 and c = -5.4940866 (Fig 2) resulted in coefficient values of  $a = 3.205 \times 10^{-6}$  and b = 3.235 ( $R^2 = 0.977$ ,  $\frac{\text{F(df} = 1; 107)}{\text{F(df} = 1; 107)} = 6928.67$ , p < 0.001) $F_{1,107} = 6928.67$ ; p < 0.00The allometric factor (b) was significantly different from b = 3 (t(107) = 6.04; p < 0.001)indicating  $t_{107} = 6.04$ ; p < 0.001), corroborating that lionfish present allometric growth. The length-weight coef-ficients estimated in this study here were within the range identified by studies in from other regions (Table 1). Figure 2 shows the relationship between TL and TW for this region, and model fit statistics are presented in Table ??. 



**Figure 2.** Length-weight relationship for 109 lionfish sampled in the central Mexican Caribbean. Points indicate samples, dashed black line indicates curve of best fit, marginal plots represent the density distribution of each variable.

Coefficients of the linear model fit to Eq  $\ref{eq:condition}$ . Numbers in parenthesees represent heteroskedastic-robust standard errors.  $c = 5.494 (0.083)^{***} b 3.235 (0.039)^{***} F$  Statistic 6928.67\*\*\* (df = 1; 107) Observations 109 Adjusted R<sup>2</sup> 0.976 Residual Std. Error 0.096 (df = 107) *Note*:

There were ANCOVA results revealed significant differences in our predicted weights weight ratios for the central Mexican Caribbean when using each of the different pairs of parameters (F(df = 17; 1944) = 61.55; p). The lowest weight estimates resulted from using the allometric parameters  $F_{17,1943} = 24.96; p < 0.001;$  Fig 4). For example, the actual observed weights of the 109 lionfish from the Central Mexican Caribbean had a mean  $\pm$  SD of  $52.56 \pm 76.58$  g. However, if we used allometric parameter values from Banco Chinchorro in the Caribbean , with to predict weights from our observed length observations, we estimated a mean  $\pm$  SD of  $40.37 \pm 58.74$  gr (Sabido-Itzá et al., 2016), and the highest weight estimates came from the Northern Atlantic with g (Sabido-Itzá et al., 2016). If we similarly used parameter values from North Carolina in the Western Atlantic to estimate lionfish weights in the Central Mexican Caribbean, we found a mean  $\pm$  SD of  $73.76 \pm 96.11$  gr (Barbour et al., 2011). To put this in context, true observed weights were  $52.56 \pm 76.58$  gr. These correspond to g (Barbour et al., 2011). Weights predicted from these extreme parameters correspond to mean predicted-to-observed weights weight ratios of  $0.80 \pm 0.19$  and  $1.76 \pm 0.50$  (mean  $\pm$  SD), respectively.

The calculated ratio of predicted-to-observed weight ranged from Furthermore, largest errors for individual organisms collected in the central Mexican Caribbean resulted in ratios of 0.36 to and 3.51 (i.e., the tails of each violin in Fig 4). If we examined biomass (i.e., summing across all 109 organisms) instead of mean ratios, total biomass estimates were 76.2% (4, indicating that 363.53 g) and 140% (8,039.96 g) of true observed biomass (5,729.34 g). Parameters for this study estimate total biomass at 98% of

observed biomass. These misestimates come from the two most extreme sets of parameters, but results varied consistently across locations (Figs 4 and 5). Overall, the use of *ex situ* parameters can result in major under- and overestimations. led to significantly erroneous estimates of individual weight and total biomass for lionfish.

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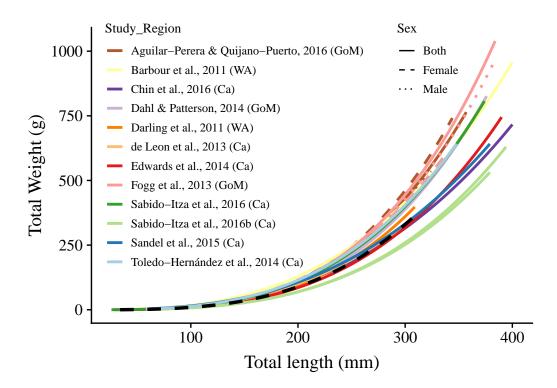
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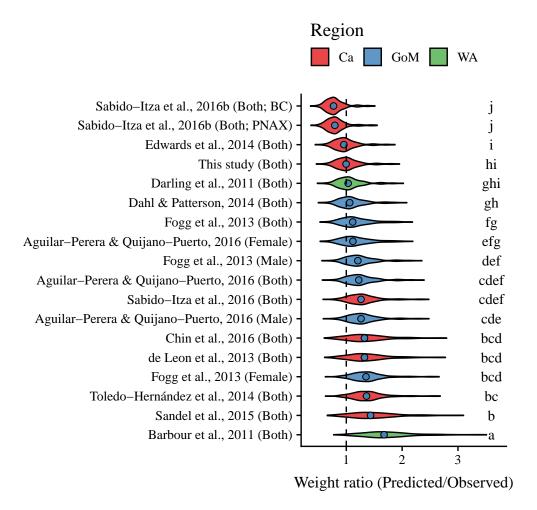
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Tukey's *post-hoc* test suggests showed that weight ratios for the central Mexican Caribbean were not different differed from those obtained with parameters from Little Cayman, the Bahamas, and some sites in the the Western Atlantic, and most sites in the Caribbean and the Gulf of Mexico (Tukey's HSD p > 0.05). The only sites where weight ratios did not differ from the central Mexican Caribbean were Little Cayman (Edwards et al., 2014), Bahamas (Darling et al., 2011), and the Northern Gulf of Mexico (Tukeys Dahl and Patterson (2014); Tukey's HSD p > 0.05). Weight All weight estimates using parameters from the Gulf of Mexico and North-Western Western Atlantic were higher on average than those than observed values, and only parameters from the Caribbean (Fig 3) produced weights smaller than observed (Fig 4). The regional average ( $\pm$  SD) of predicted-to-observed weight ratios from these three regions were  $1.24 \pm 0.309$ ,  $1.76 \pm 0.496$ , and  $1.17 \pm 0.523$ , and  $1.20 \pm 0.398$ , respectively 0.423 for the Gulf of Mexico. Western Atlantic, and Caribbean, respectively. This suggests that the smallest errors are observed when using parameters from other locations in the Caribbean. Predicted to observed weight ratios are presented in Figure 4. Spineless weight parameters from Fogg et al. (2013) still produced predicted-to-observed weight ratios > 1.

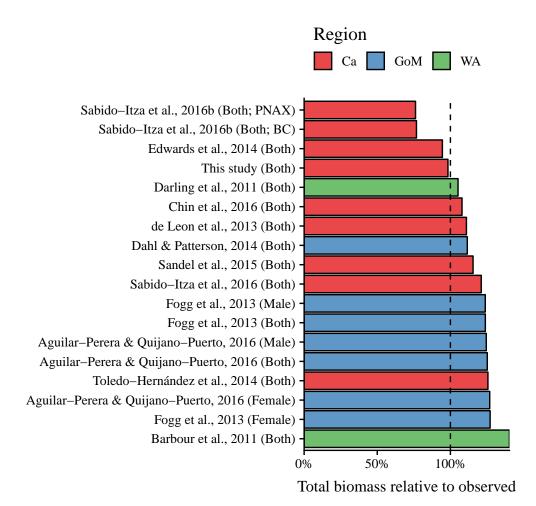
Summary of 18 allometric growth parameters available for lionfish in the invaded range from peer-reviewed literature and this study. All parameters have been adjusted to convert from millimeters to grams. n = Sample size, Sex specifies whether data was presented for Females (F), Males (M), or both genders combined (B), a = scaling parameter for Eq. 1 (presented in  $\times 10^{-5}$ ), c = y-intercept for Eq. 3, b = exponent or slope for Eq. 1 or Eq. 3, respectively. The Fit column contains the reported R<sup>2</sup> of the model fit. Region Sex n a b c Fit ReferenceCaribbeanB 458 3.6 2.81 -4.44 - Sandel et al. , 2015Caribbean B 419 2.8 2.85 -4.56 0.8715 Chin et al., 2016Caribbean B 1450 2.3 2.89 -4.64 0.96 de Leon et al., 2013Caribbean B 1887 0.3 3.24 -5.52 0.97 Edwards et al., 2014Caribbean B - 0.25 3.29 -5.60 - Darling et al., 2011Caribbean B 2143 0.52 3.18 -5.28 0.9907 Sabido-Itza et al., 2016Caribbean B 227 0.8 3.11 -5.10 0.958 Toledo-Hernández et al., 2014Caribbean B 449 0.23 3.25 -5.64 0.97 Sabido-Itza et al., 2016bCaribbean B 368 0.32 3.19 -5.50 0.98 Sabido-Itza et al., 2016bCaribbean B 109 0.32 3.23 -5.49 0.9766 This studyGoM B 934 0.21 3.34 -5.68 0.98 Dahl & Patterson, 2014GoM B 472 0.29 3.30 -5.54 0.95 Aguilar-Perera & Quijano-Puerto, 2016GoM F 67 0.12 3.47 -5.93 0.95 Aguilar-Perera & Quijano-Puerto, 2016GoM M 59 0.42 3.23 -5.38 0.95 Aguilar-Perera & Quijano-Puerto, 2016GoM B 582 0.14 3.43 -5.86 0.99 Fogg et al., 2013GoM M 119 0.27 3.31 -5.57 0.97 Fogg et al., 2013GoM F 115 0.68 3.14 - 5.17 0.94 Fogg et al., 2013North Atlantic B 774 2.9 2.89 - 4.54 - Barbour et al., 2011



**Figure 3.** Length-weight relationships (n = 18) for 12 studies and this study. The curves are shown for the range of lengths reported in each study (See Supplementary Table 2); when ranges were not present, we use the ones found in this study (34 mm - 310 mm). Colors indicate studies from which the parameters were extracted. Dotted, dashed, and solid lines show models for males, females, and combined sexes, respectively. Letters in parentheses indicate if the study comes from the Gulf of Mexico (GoM), Western Atlantic (WA), or Caribbean (Ca). The dashed black line represents the relationship estimated in this study. There are two solid green lines for Sabido-Itza et al. 2016b, one for each of the two sites for which they report parameters. A log-log version of this figure is presented in Figure S4.



**Figure 4.** Violin plot of predicted-to-observed weight ratios for when applying each of 18 different pairs of allometric parameters to the 109 lionfish collected in the central Mexican Caribbean. Sex is indicated in parentheses. Blue circles indicate median values and Like like letters indicate values that do not differ significantly. For Sabido-Itza et al, 2016b, BC and PNAX make reference to Banco Chinchorro and Parque Nacional Arrecifes de Xcalak, two sites for which they report parameters.



**Figure 5.** Estimated total biomas relative to observed biomass (5,729.34 g) for 18 pairs of allometric parameters. Sex is indicated in parentheses. For Sabido-Itza et al, 2016b, BC and PNAX make reference to Banco Chinchorro and Parque Nacional Arrecifes de Xcalak, two sites for which they report parameters.

# DISCUSSION

Our results suggest that lionfish exhibit highly variable, spatially heterogeneous allometric relationships across the their invaded range, and that this variation is related to spacerelevant for managing invasions. Moreover, we shot show that the use of ex situ parameters parameter values may lead to highly biased weight and total biomass estimates. Our comparison of observed weights to those predicted with locally-informed parameters and ex situ parameters showed that weight of an individual lionfish can be overestimated by more than a three-fold, and highlights threefold, highlighting the need to use local information. Here we discuss the implications of our findings, possible shortcommings in our analyses, and highlight potential future research directions.

Differences in length-weight relationships have traditionally been highlighted as potential pitfalls to fishery management. For example, Wilson et al. (2012) showed that small-scale variations in length-at-age and fishing mortality in other Scorpaeniformes translate to differential landings, effort, and catch per unit effort in the live fish fishery of California, and that these differences must be taken into account in management plans. The lionfish case poses the opposite scenario, where the manager desires to eradicate the species. To accurately gauge both the effectiveness of lionfish removal efforts and the resources needed to successfully manage an invasion, we must acknowledge and understand regional biological differences in important variables such as allometric growth parameters.

We detected substantial differences in weight-at-length between organisms from the Caribbean, Gulf of Mexico, and North-Western Western Atlantic. Groupings of predicted-to-observed weight ratios identified in our post hoc testing aligned with the spatial distribution of the examined studies, suggesting that these differences are may be mediated by space. These regional allometric differences mirror similar patterns in age-at-length length-at-age of lionfish across both their invaded and native regions (Pusack et al., 2016). Variation may be driven by genetics or by organisms' exposure to distinct environmental conditions. For example, Betancur-R et al. (2011) used mitochondrial DNA to demonstrate the existence of two distinct population groups, identified as the "Caribbean group" and "Northern Group", and Fogg et al. (2015) alternatively suggested that age-at-length length-at-age differences may be elimate-driven. Differences in weight-at-length could also reflect differential energy input or usage, or a combination of both. Future research is needed to determine which processes are at work here, driven by the environment.

Differences in One might be inclined to attribute all variation in the lionfish length-weight relationships have traditionally been highlighted as potential pitfalls to fishery management. For example, Wilson et al. (2012) show that small-scale variations in length-at-age and fishing mortality in other Scorpaeniformes translate to differential landings, effort, and eatch per unit effort in the live fish fishery of California, and that these differences must be taken into account in managementplans. The lionfish case poses the opposite scenario, where the manager desires to eradicate the species. To accurately gauge both the effectiveness of lionfish removal efforts and the resources needed to successfully manage an invasion, we must acknowledge and understand regional biological differences in important variables such as allometric growth parameters relationship to the spatial origin of these parameters. However, samples from the 12 studies included here were not only collected in different locations, but also at different points in time and across different depth and size ranges (See Supplementary Table 2 for an extended version of Table 1). The magnitude of the bias discovered in this study and our lack of understanding the sources driving spatial variation for lionfish highlights the need to simultaneously collect length-weight information across the invaded range to test for spatially-induced patterns and to link these findings to previously suggested environmental and genetic structures. Such an endeavor would provide insight into lionfish biology and better inform management. However, while we could not evaluate how these factors influenced length-weight estimates from previous studies without raw data, we still show that a lack of locally-calculated parameters can induce significant bias when calculating weight from length observations. We demonstrate the importance of using in situ parameters to obtain accurate weight estimates regardless of the underlying mechanisms driving variation between populations.

Applying parameter estimates to lengths outside the range of lengths originally used to estimate the parameters may also induce error. Our smallest observed organism was 34 mm in TL, and only two studies estimated parametrs with smaller organisms (Edwards et al., 2014; Sabido-Itza et al., 2016). By contrast, our largest organism had a TL of 310 mm, which is well within the range of all other studies (the next smallest maximum length was 325 mm; See Supplementary Table 2). Due to the power function describing the allometric relationship (i.e., Eq. 1), the error in weight estimates is larger when extrapolation is done for lengths that are larger than the maximum length used to estimate the

parameters. Our estimates are therefore conservative because we only used parameter pairs from other studies to estimate weights for lionfish up to 310 mm in the Central Mexican Caribbean, well within the range of lengths for which other parameters were estimated.

The results presented here have major key implications for management. For example, Edwards et al. (2014) simulated a lionfish culling program under two scenarios, one using length-at-age and length-to-weight parameters from North Carolina and one using parameters from Little Cayman. Their results show that using different parameters caused up to a four-year difference in the time required for the simulated lionfish population to recover to 90% of its initial biomass after removals ceased. Here, we show that using one set of length-weight parameters versus another for a given length can result in more than a threefold under- or overestimation of total weight. These spatially-driven for individual fish, and that total biomass estimates may range between 76% and 140% of true observed biomass. These differences become especially important when allocating resources for lionfish removal programs, incentivizing lionfish fisheries as a source of alternative livelihoods, or estimating ecosystem impacts. Research efforts focused on invasive lionfish populations need to use parameters calculated for their region to the extent possible, or at least use reasonable sets of different use different sets of parameters that provide appropriate upper and lower bounds in their results.

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