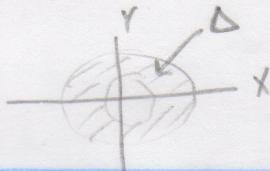


2.3



$$\Pr[\text{error}] = \Pr\left[\bigcup_{i=1}^m x \notin R_i\right]$$

$$\leq \sum_{i=1}^m \Pr[x \notin R_i] \quad \text{(Union Bound)}$$

$$= \sum_{i=1}^m (1 - \Pr[R_i])^m \quad \text{(Independence)}$$

Let us assume  $\Pr[R_i] \geq \frac{\epsilon}{2}$ , then:

$$\Pr[\text{error}] \leq 2(1 - \frac{\epsilon}{2})^m \leq 2e^{-\frac{m\epsilon}{2}} \quad (1+x \leq e^x \text{ inequality})$$

Solving for m:

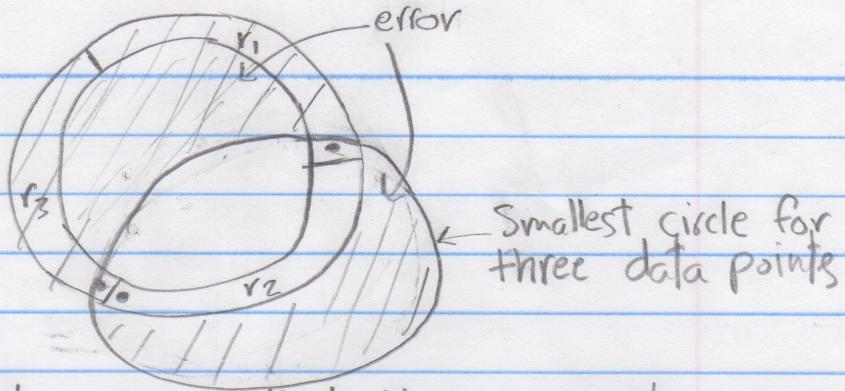
$$\Pr[R(n) \leq \epsilon] \leq 1 - \delta \quad (\text{Substituting } 2e^{-\frac{m\epsilon}{2}} \text{ for error})$$

$$2e^{-\frac{m\epsilon}{2}} \leq \delta \quad (\text{take the log})$$

$$-\frac{m\epsilon}{2} \leq \log(\delta) \quad (\text{multiply by } -\frac{2}{\epsilon})$$

$$\boxed{m \geq \frac{2}{\epsilon} \log\left(\frac{2}{\delta}\right)} \quad \text{So, Class is PAC-Learnable}$$

2.4



This example shows that this concept class is not PAC-learnable. The error here is clearly greater than  $\epsilon$  and so, there might be data points that are missed outside of the three regions.

2.6

(a) The probability of  $R'$  missing  $r_j, j \in [1, 4]$  is the product of  $P(X_i \notin r_j)$  or  $P(X_i \in r_j \wedge \text{flipped positive } X_i)$

$$\begin{aligned} &= P(X_i \notin r_j \vee (X_i \in r_j \wedge \text{flipped positive } X_i)) \\ &= P(X_i \notin r_j) + P(X_i \in r_j \wedge \text{flipped positive } X_i) \\ &= (1 - P(X_i \in r_j)) + \eta \Pr(X_i \in r_j) \\ &= (1 - \eta)(1 - P(X_i \in r_j)) + \eta \quad (\text{four regions}) \\ &\leq (1 - \eta)(1 - \epsilon_4) + \eta \\ &= (1 - \epsilon_4) + \eta \epsilon_4 \leq 1 - \frac{\epsilon}{4}(1 - \eta') \end{aligned}$$

(b) Assume  $P(R') \geq \epsilon_4$

$$\Pr[\text{error}] \leq 4(1 - \epsilon_4(1 - \eta')) \leq 4e^{-\frac{m\epsilon}{4}(1 - \eta')}$$

$$\Pr[\text{error}] \leq 1 - 8$$

$$4e^{-\frac{m\epsilon}{4}(1 - \eta')} \leq 8 \quad (\text{Take the log})$$

$$-\frac{m\epsilon}{4}(1 - \eta') \leq \log(8) \quad (\text{Multiply by } \frac{4}{\epsilon(1 - \eta')})$$

$$m \geq \frac{4}{\epsilon(1 - \eta')} \log\left(\frac{4}{8}\right)$$

3.6] The VC-Dimension of a hypothesis set  $H$  is the size of the largest set that can be fully shattered by  $H$ :

$$VCdim(H) = \max\{m : \Gamma_H(m) = 2^m\}$$

Any  $2K$  points on the real line can be shattered by  $K$  intervals. For  $2K+1$  points, assume that the points are alternatively labeled by  $+1$  or  $-1$ . So, we will have  $K+1$  positive points & need  $2K+1$  intervals to shatter since an interval will not have two consecutive points. As a result no  $2K+1$  set of points can be shattered. So, the VC dimension of subsets of the real line formed by the union of  $K$  intervals is  $2K$ .

3.5 We are going to use the Growth Function.  
If  $G$  is a function taking values  $\{-1, +1\}$ , then

$$R_m(G) \leq \sqrt{\frac{2 \ln \Pi_G(m)}{m}}$$

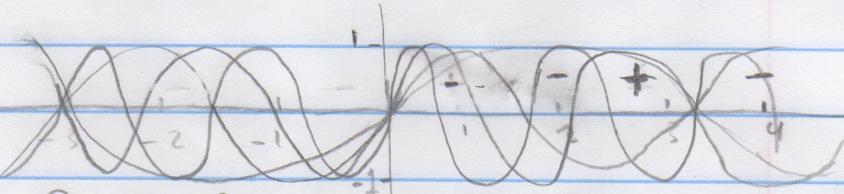
where  $\Pi_G(m)$  is Growth Function. To connect that to VC dimensions and Radamacher complexity, we use Sauer's Lemma:

$$\Pi_H(m) \leq \sum_{i=0}^d \binom{m}{i} \equiv \phi_d(m)$$

Where  $\phi_d(m) \leq E[\phi(m)] + \sqrt{\ln \frac{1}{\delta}}$ . The above contradicts Professor Jester's bound on Radamacher Complexity.

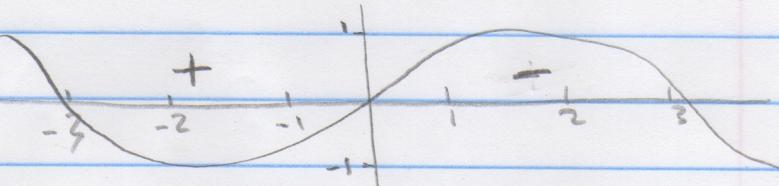
3.12

(a)



Let  $X \in \mathbb{R}$  be fixed then, the four points labeled  $- - + -$  cannot be shattered

(b) By making the frequency of the sine function greater, any few points of opposite signs can be separated by a sine cycle



3.19 (a)

$$\begin{aligned}\text{error}(f_0) &= \Pr[f_0(S) \neq X] \quad (\text{Prediction Rule}) \\ &= \Pr[f_0(S) = X_A \wedge X = X_B] + \Pr[f_0(S) = X_B \wedge X = X_A] \\ &= \Pr[N(S) < \frac{m}{2} \mid X = X_B] \Pr[X = X_B] + \\ &\quad \Pr[N(S) > \frac{m}{2} \mid X = X_A] \Pr[X = X_A] \\ &= \frac{1}{2} \Pr[N(S) < \frac{m}{2} \mid X = X_B] + \frac{1}{2} \Pr[N(S) > \frac{m}{2} \mid X = X_A] \\ &\geq \frac{1}{2} \Pr[N(S) > \frac{m}{2} \mid X = X_A]\end{aligned}$$

(b)

Using the Binomial distribution,

$$\text{error}(f_0) \geq \frac{1}{2} \Pr[N \geq \frac{m+1}{2}] = \frac{1}{2} \Pr\left[N \geq \frac{\lceil m \rceil \epsilon}{1-\epsilon^2}\right]$$

Using Tail bound,

$$\text{error}(f_0) \geq \frac{1}{4}(1 - \sqrt{1 - e^{-m\epsilon^2}}). \text{ This agrees with } 3.57$$

(c)

$$\text{Since } \Pr[N(S) > \frac{m}{2} \mid X = X_A] \geq \Pr[N(S) \geq \frac{m+1}{2} \mid X = X_A]$$

$$\text{So, } \text{error}(f_0) \geq \frac{1}{2} \Pr[N(S) \geq \frac{m+1}{2} \mid X = X_A]$$

(d)

$$\begin{aligned}\text{if error} > \delta, \text{ then } \frac{1}{4} \left[ 1 - \left[ 1 - e^{-\frac{2cm(2)\epsilon^2}{1-\epsilon^2}} \right]^{\frac{1}{2}} \right] < \delta \\ e^{-\frac{2cm(2)\epsilon^2}{1-\epsilon^2}} < 1 - (1 - 4\delta)^2 = 4\delta(2 - 4\delta) = 8\delta(1 - 2\delta)\end{aligned}$$

$$m > 2 \left[ \frac{1 - \epsilon^2}{2\epsilon^2} \log \frac{1}{8\delta(1 - 2\delta)} \right]$$

The lower Bound changes based on  $\frac{1}{\epsilon^2}$

3.19

(e)  $\text{error}(f) = \sum \Pr[S \wedge X_B] + \sum \Pr[S \wedge X_A]$

If  $N(S) \geq \frac{m}{2}$ ,  $\Pr[S|X_A]$  and if  $N(S) < \frac{m}{2}$ ,  $\Pr[S|X_A] \geq \Pr[S|X_B]$

$$\text{error}(f) \geq \sum \Pr[S \wedge X_B] + \sum \Pr[S \wedge X_A]$$

$$\text{error}(f) \geq \frac{1}{2} \sum \Pr[S|X_B] + \frac{1}{2} \sum \Pr[S|X_A] +$$

$$\frac{1}{2} \sum \Pr[S|X_B] + \frac{1}{2} \sum \Pr[S|X_A]$$

$$= \frac{1}{2} \sum \Pr[S|X_B] + \frac{1}{2} \sum \Pr[S|X_A]$$

$$= \text{error}(f_0)$$