ARE212- FINAL

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ARE212 Take-home Final: Exploring Measurement Error.

In this assessment we explore what happens to the least squares estimator for varying degrees of measurement error in the left hand and right hand side variables. Assume the following population model:

```
y_i^* = \beta_0 + \beta_1 \cdot x_{1i} + \beta_2 \cdot x_{2i} + \epsilon_i
Assume: E[\epsilon|x] = 0 and \epsilon is drawn from a normal distribution (0,1) x_{1i} and x_{2i} drawn from uniform normal [-200, 200] \beta_0 = 1, \beta_1 = -0.75, \beta_2 = 0.75
Set seed to 22092008.
```

1. Generate random sample

```
# Generate population data:
# Set a seed
set.seed(22092008)

# Set the population size
N <- 1e2

# Generate the data for X and E
x1 = runif(n = N, min = -200, max = 200)
x2 = runif(n = N, min = -200, max = 200)
e = rnorm(n = N, mean = 0, sd = 1)

# Generate the y variables (in anticipataion of question #2)
y = 1 -0.75*x1 + 0.75*x2 + e

# Join the data together
pop_df <- as.data.frame(cbind(y, x1, x2))</pre>
```

2. Generate y's and estimate b's with OLS

Generate y variables (done above). Estimate the three β coefficients using least squares. Calculate the difference between the true β and the estimated coefficient for each of the three β coefficients.

First, let's load our OLS functions from the last few problem sets:

```
# Function to convert tibble, data.frame, or tbl_df to matrix
to_matrix <- function(the_df, vars) {
    # Create a matrix from variables in var</pre>
```

```
new_mat <- the_df %>%
    #Select the columns given in 'vars'
    select_(.dots = vars) %>%
    # Convert to matrix
    as.matrix()
  # Return 'new mat'
 return(new_mat)
}
b_ols <- function(y, X) {</pre>
  # Calculate beta hat
 beta_hat <- solve(t(X) %*% X) %*% t(X) %*% y
  # Return beta_hat
 return(beta_hat)
ols <- function(data, y_data, X_data, intercept = T, hetsked = F, HO = O, two_tail = T, alpha = O
  # Function setup ----
    # Require the 'dplyr' package
    require(dplyr)
  # Create dependent and independent variable matrices ----
    # y matrix
    y <- to_matrix (the_df = data, vars = y_data)
    # X matrix
    X <- to_matrix (the_df = data, vars = X_data)</pre>
      # If 'intercept' is TRUE, then add a column of ones
      if (intercept == T) {
      X \leftarrow cbind(1,X)
      colnames(X) <- c("intercept", X_data)</pre>
      }
  \# Calculate b, y_hat, and residuals ----
    b <- solve(t(X) %*% X) %*% t(X) %*% y
    y hat <- X %*% b
    e <- y - y_hat
    # Inverse of X'X
    XX \leftarrow t(X) \% X
    XX_inv <- solve(t(X) %*% X)</pre>
    if (hetsked == T) {
      # For each row, calculate x_i' x_i e_i 2; then sum
     sigma_hat <- lapply(X = 1:n, FUN = function(i) {</pre>
      # Define x_i
      x_i <- matrix(as.vector(X[i,]), nrow = 1)</pre>
      # Return x_i' x_i e_i 2
      return(t(x_i) %*% x_i * e[i]^2)
      }) %>% Reduce(f = "+", x = .) }
```

```
if (hetsked == F) sigma_hat <- XX</pre>
# Useful ----
 n <- nrow(X) # number of observations
 k <- ncol(X) # number of independent variables
 dof <- n - k # degrees of freedom
 i <- rep(1,n) # column of ones for demeaning matrix
 A <- diag(i) - (1 / n) * i %*% t(i) # demeaning matrix
 y_star <- A %*% y # for SST</pre>
 X_star <- A %*% X # for SSM
 SST <- drop(t(y_star) %*% y_star)</pre>
 SSM <- drop(t(b) %*% t(X_star) %*% X_star %*% b)
 SSR \leftarrow drop(t(e) %*% e)
# Measures of fit and estimated variance ----
 R2uc <- drop((t(y_hat) %*% y_hat)/(t(y) %*% y)) # Uncentered R^2
 R2 <- 1 - SSR/SST # Uncentered R~2
 R2adj \leftarrow 1 - (n-1)/dof * (1 - R2) # Adjusted R^2
 AIC \leftarrow log(SSR/n) + 2*k/n # AIC
 SIC \leftarrow log(SSR/n) + k/n*log(n) # SIC
 s2 <- SSR/dof # s ~2
# Measures of fit table ----
 mof_table_df <- data.frame(R2uc, R2, R2adj, SIC, AIC, SSR, s2)</pre>
 mof_table_col_names \leftarrow c("$R^2_\text{uc}$", "$R^2$",
                            "$R^2_\\text{adj}$",
                            "SIC", "AIC", "SSR", "$s^2$")
 mof_table <- mof_table_df %>% knitr::kable(
   row.names = F,
    col.names = mof_table_col_names,
    format.args = list(scientific = F, digits = 4),
   booktabs = T,
    escape = F
 )
# t-test----
  # Standard error
 se <- sqrt(s2 * diag(XX_inv %*% sigma_hat %*% XX_inv)) # Vector of _t_ statistics
  # Vector of _t_ statistics
 t_stats <- (b - H0) / se
  # Calculate the p-values
 if (two_tail == T) {
 p_values \leftarrow pt(q = abs(t_stats), df = dof, lower.tail = F) * 2
 } else {
    p_values <- pt(q = abs(t_stats), df = dof, lower.tail = F)</pre>
  # Do we (fail to) reject?
 reject <- ifelse(p_values < alpha, reject <- "Reject", reject <- "Fail to Reject")
```

```
# Nice table (data.frame) of results
    ttest_df <- data.frame(</pre>
      # The rows have the coef. names
      effect = rownames(b),
      # Estimated coefficients
      coef = as.vector(b) %>% round(3),
      # Standard errors
      std_error = as.vector(se) %>% round(4),
      # t statistics
      t_stat = as.vector(t_stats) %>% round(3),
      # p-values
      p_value = as.vector(p_values) %>% round(4),
      # reject null?
      significance = as.character(reject)
      )
    ttest_table <- ttest_df %>% knitr::kable(
      col.names = c("", "Coef.", "S.E.", "t Stat", "p-Value", "Decision"),
      booktabs = T,
      format.args = list(scientific = F),
      escape = F,
      caption = "OLS Results"
    )
  # Data frame for exporting for y, y_hat, X, and e vectors ----
    export_df <- data.frame(y, y_hat, e, X) %>% tbl_df()
    colnames(export_df) <- c("y","y_hat","e",colnames(X))</pre>
  # Return ----
    return(list(n=n, dof=dof, b=b, se=se, vars=export_df, R2uc=R2uc,R2=R2,
                R2adj=R2adj, AIC=AIC, SIC=SIC, s2=s2, SST=SST, SSR=SSR,
                mof_table=mof_table, ttest=ttest_table))
}
Now ready to estimate!
X_{mat} \leftarrow c("x1", "x2")
q2 <- ols(data = pop_df, y_data = "y", X_data = X_mat, intercept = T)
q2$b %>% knitr::kable()
                                                      у
                                               0.9074893
                                    intercept
                                              -0.7498911
                                    x1
                                    x2
                                               0.7498942
diff_mat <- data_frame(</pre>
  diff0 = (q2$b[1] - 1),
  diff1 = (q2$b[2] - (-0.75)),
```

```
diff2 = (q2$b[3] - (0.75)))
diff_mat %>% knitr::kable()
```

| dif | fO | diff1 | diff2 |
|-----------|------|-----------|------------|
| -0.092510 |)7 (| 0.0001089 | -0.0001058 |

3. Introducing MEASUREMENT ERROR on y

Now assume that y_i is measured with error! It is in fact $y_i + r_i$, where r_i is drawn from a random normal distribution with mean zero and variance σ^2 . Using the y_i from the previous step, add the measurement error r_i to them and use this measured with error dependent variable as your outcome.

Estimate the three β coefficients again using least squares. Do this for $\sigma^2 = [1\ 10\ 100]$.

```
sigma2 <- c(1,10,100)
# for (s in sigma2) {
    \# Generate r_i
    r \leftarrow rnorm(N, 0, s)
   # Add new error to y
   pop_df \% <> \% mutate(y_err = y + r)
#
   # Rerun OLS
    ols(pop_df, "y_err", X_mat)$b %>% knitr::kable()
# }
\# Generate r_i
r <- rnorm(N, 0, 1)
# Add new error to y
pop_df %<>% mutate(y_err = y + r)
# Rerun OLS
ols(pop_df, "y_err", X_mat)$b %>% knitr::kable()
```

```
\begin{array}{c|c} & & & y\_{err} \\ \hline intercept & 1.0596229 \\ x1 & -0.7495957 \\ x2 & 0.7487727 \\ \end{array}
```

```
# Generate r_i
r <- rnorm(N, 0, 10)
# Add new error to y
pop_df %<>% mutate(y_err = y + r)
# Rerun OLS
ols(pop_df, "y_err", X_mat)$b %>% knitr::kable()
```

| | y_err |
|-----------|------------|
| intercept | 2.2691807 |
| x1 | -0.7457138 |
| x2 | 0.7575671 |

```
# Generate r_i
r <- rnorm(N, 0, 100)
# Add new error to y
pop_df %<>% mutate(y_err = y + r)
# Rerun OLS
ols(pop_df, "y_err", X_mat)$b %>% knitr::kable()
```

| y_err |
|------------|
| 5.6064263 |
| -0.8522662 |
| 0.7731456 |
| |

Worse coefficients!! ### 4. Measurement Error on x2 Now assume that your x_{2i} is measured with error. You observe $x_2^* = x_{2i} + r_i$, where r_i is drawn from a random normal distribution with mean zero and variance σ^2 (Just reuse the r_i from the previous step.). Use the y_i from the first step (the one measured without error) and replace x_{2i} with x_2^* in your estimation. Estimate the three β coefficients using least squares on your data. Do this for $\sigma^2 = [1 \ 10 \ 100]$.

5. Non-symmetric measurement error: POSITIVE

Repeat step 4 exactly, only now assume that your measurement error is not symmetric, but always positive. Simply take the absolute value of r_i (again using the r_i from above) before generating your x_2^* . Only do this for $\sigma^2 = 100$.

6. Non-symmetric measurement errror: NEGATIVE

Repeat step 4 exactly, only now assume that your measurement error is not symmetric, but always negative. Simply take the absolute value of r_i and multiply it times (-1) before generating your x_2^* (again using the r_i from above). Only do this for $\sigma^2 = 100$.

SCALE UP!

Repeat the above steps 10,000 times (set the seed only the first time, not each time) and calculate the bias for β_0 , β_1 and β_2 for each setting and fill in the table below. What have you learned from this exercise that you did not know before? Was there anything surprising?

simple table creation here

| Tables | Are | Are | Cool |
|----------|---------------|---------------|--------|
| col 3 is | right-aligned | right-aligned | \$1600 |
| col 2 is | centered | centered | \$12 |

| Tables | Are | Are | Cool |
|---------------|----------|----------|------|
| zebra stripes | are neat | are neat | \$1 |