


Linear probability, Probit, and Logit models

Linear probability model

- Often the dependent variable is a dummy variable.
- For example, if you would like to find the determinants of labor force participation, the dependent variable is 1 if the person is working(participate in the labor force), and 0 if the person is not working(does not participate in the labor force).

 Simplest way to predict the probability that the person would participate in the labor force is to estimate the following model using OLS.


$$\text{inlf} = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k + u$$

Where $\text{inlf} = 1$ if the person is in the labor force and $\text{inlf} = 0$ if the person is not in the labor force.

This type of the model is called **the linear probability model**.

- As long as explanatory variables are not correlated with the error term, this simple method is perfectly fine.
- However, a caution is necessary. When the dependent variable is a dummy variable, $\text{var}(u | X) = X\beta[1 - X\beta]$ where $X\beta$ is a short hand notation for $(\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k)$
- Thus, the variance depends on the values of the x-variables, which means that it is heteroskedastic. **So, you should always use the robust-standard errors.**

Exercise

 Using Mroz.dta, estimate the following linear probability model of labor force participation.

$$\begin{aligned} \text{inlf} = & \beta_0 + \beta_1 \text{nwifinc} + \beta_2 \text{educ} + \beta_3 \text{exper} \\ & + \beta_4 \text{exper} + \beta_5 \text{age} + \beta_6 \text{kidslt6} + \beta_7 \text{kidsge6} \\ & + u \end{aligned}$$

Answer

```
. reg inlf nwifeinc educ exper expersq age kidslt6 kidsge6, robust
```

Linear regression

Number of obs = 753
 F(7, 745) = 62.48
 Prob > F = 0.0000
 R-squared = 0.2642
 Root MSE = .42713


inlf	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
nwifeinc	-.0034052	.0015249	-2.23	0.026	-.0063988	-.0004115
educ	.0379953	.007266	5.23	0.000	.023731	.0522596
exper	.0394924	.00581	6.80	0.000	.0280864	.0508983
expersq	-.0005963	.00019	-3.14	0.002	-.0009693	-.0002233
age	-.0160908	.002399	-6.71	0.000	-.0208004	-.0113812
kidslt6	-.2618105	.0317832	-8.24	0.000	-.3242058	-.1994152
kidsge6	.0130122	.0135329	0.96	0.337	-.013555	.0395795
_cons	.5855192	.1522599	3.85	0.000	.2866098	.8844287


One problem of the linear probability model is that, the predicted probability can be greater than 1 or smaller than 0.

This problem can be avoided by using Probit or Logit models which are described below.

Nonetheless, you should note that, in many applications, the predicted probabilities fall mostly within the $[0,1]$ interval. Thus, it is still a reasonable model, and it has been applied in many papers.

The probit and logit models

 Consider that you would like to know the determinants of the labor force participations of married women.

 For simplicity, let us consider the model with only one explanatory variable.

☞ In probit and logit model, we assume that there is an unobserved variable (latent variable) y^* that depends on x :


$$y^* = \beta_0 + \beta_1 x + u$$

☞ You can consider y^* as the utility from participating in the labor force.

☞ We do not observe y^* . We only observe the actual labor force participation outcome.


$Y=0$ if the person is not working

$Y=1$ if the person is working.


 Then, we assume the following.


If $Y=0$ (the person is not working), then y^* must have been negative.

If $Y=1$ (the person is working), then y^* must have been positive.

 This assumption can be written as:

$$\begin{cases} \text{If } Y = 0, \text{ then } y^* < 0 \\ \text{If } Y = 1, \text{ then } y^* \geq 0 \end{cases}$$

 Intuitive interpretation is the following. If the utility from participating in the labor force is negative, the person will not work. But if the utility from participating in the labor force is positive, then the person will work.


 Given the data about x and Y , we can compute the likelihood contribution for each person in the data by making the distributional assumption about the error term u .


The model:

$$y_i^* = \beta_0 + \beta_1 x_i + u_i$$

$$\left\{ \begin{array}{l} \text{If } Y_i = 0, \text{ then } y_i^* < 0 \\ \text{If } Y_i = 1, \text{ then } y_i^* \geq 0 \end{array} \right.$$

The i -subscript denotes the observation id.

 When we assume that u_i follows the standard normal distribution (normal with mean 0 and variance 1), the model is called the **Probit model**.

 When we assume that u follows the logistic distribution, the model is called the **Logit model**.


Probit model: Likelihood function example

The model

Id	Y	X
1	0	1
2	0	4
3	1	5
4	1	6
5	1	9

$$y_i^* = \beta_0 + \beta_1 x_i + u_i$$

$$\begin{cases} \text{If } Y_i = 0, \text{ then } y_i^* < 0 \\ \text{If } Y_i = 1, \text{ then } y_i^* \geq 0 \end{cases}$$

 Suppose that you have the following data.

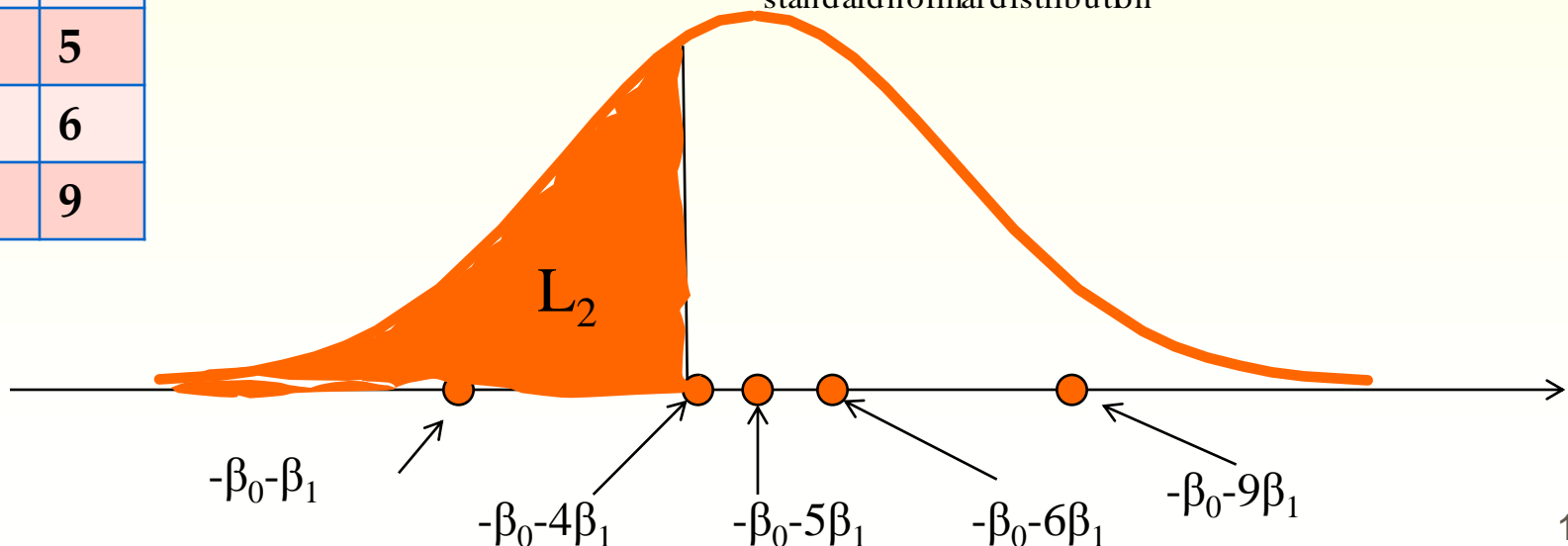
 We assume that $u_i \sim N(0,1)$

Take 2nd observation as an example. Since $Y=0$ for this observation, we know $y^* < 0$

Thus, the likelihood contribution is

$$\begin{aligned}
 L_2 &= P(y^* < 0) = P(\beta_0 + 4\beta_1 + u < 0) \\
 &= P(u < -\beta_0 - 4\beta_1) \\
 &= \int_{-\infty}^{-\beta_0 - 4\beta_1} \frac{1}{\sqrt{2\pi}} e^{-u^2} du = \underbrace{\Phi(-\beta_0 - 4\beta_1)}_{\text{Cumulative Distribution Function of standard normal distribution}}
 \end{aligned}$$

Id	Y	X
1	0	1
2	0	4
3	1	5
4	1	6
5	1	9



Now, take 3rd observation as an example.

Since $Y=0$ for this observation, we know $y^* \geq 0$

Thus, the likelihood contribution is

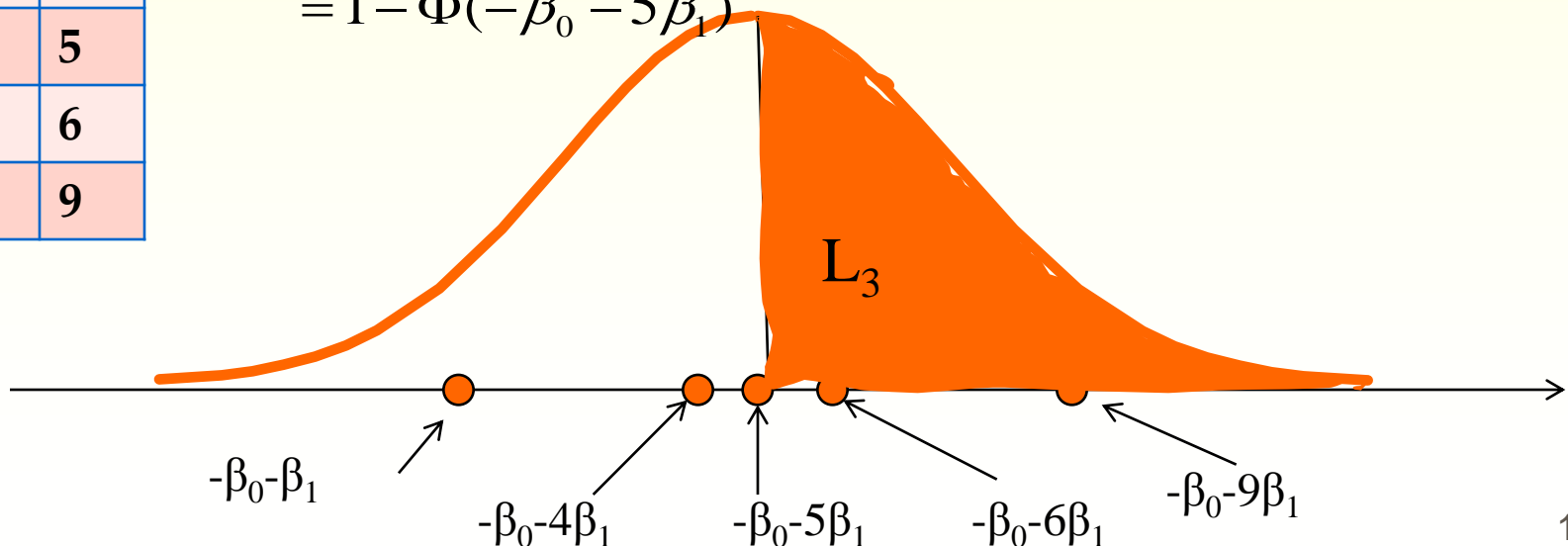
$$L_3 = P(y^* \geq 0) = P(\beta_0 + 5\beta_1 + u \geq 0)$$


$$= P(u \geq -\beta_0 - 5\beta_1)$$

$$= \int_{-\beta_0 - 5\beta_1}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-u^2} du = 1 - \int_{\infty}^{-\beta_0 - 5\beta_1} \frac{1}{\sqrt{2\pi\sigma^2}} e^{-u^2} du$$

$$= 1 - \Phi(-\beta_0 - 5\beta_1)$$

Id	Y	X
1	0	1
2	0	4
3	1	5
4	1	6
5	1	9




 Then the likelihood function is given by

Id	Y	X
1	0	1
2	0	4
3	1	5
4	1	6
5	1	9

$$\begin{aligned}
 L(\beta_0, \beta_1) &= \prod_{i=1}^5 L_i \\
 &= \left[\int_{-\infty}^{-\beta_0 - \beta} \frac{1}{\sqrt{2\pi}} e^{-u^2} du \right] \times \left[\int_{-\infty}^{-\beta_0 - 4\beta} \frac{1}{\sqrt{2\pi}} e^{-u^2} du \right] \\
 &\quad \times \left[1 - \int_{-\infty}^{-\beta_0 - 5\beta} \frac{1}{\sqrt{2\pi}} e^{-u^2} du \right] \times \left[1 - \int_{-\infty}^{-\beta_0 - 6\beta} \frac{1}{\sqrt{2\pi}} e^{-u^2} du \right] \\
 &\quad \times \left[1 - \int_{-\infty}^{-\beta_0 - 9\beta} \frac{1}{\sqrt{2\pi}} e^{-u^2} du \right] \\
 &= \Phi(-\beta_0 - \beta) \Phi(-\beta_0 - 4\beta) [1 - \Phi(-\beta_0 - 5\beta)] \\
 &\quad \times [1 - \Phi(-\beta_0 - 6\beta)] [1 - \Phi(-\beta_0 - 9\beta)]
 \end{aligned}$$

Probit model

Likelihood for more general case

 Consider the following model.

Id	Y	X
1	Y_1	x_1
2	Y_2	x_2
3	Y_3	x_3
:	:	:
n	Y_n	x_n

$$y_i^* = \beta_0 + \beta_1 x_i + u_i$$

$$\begin{cases} \text{If } Y_i = 0, \text{ then } y_i^* < 0 \\ \text{If } Y_i = 1, \text{ then } y_i^* \geq 0 \end{cases}$$

$$u_i \sim N(0,1)$$

Then, we know that

$$\begin{cases} L_i = \Phi(-\beta_0 - \beta_1 x_i) & \text{if } Y_i < 0 \\ L_i = 1 - \Phi(-\beta_0 - \beta_1 x_i) & \text{if } Y_i \geq 0 \end{cases}$$

The above can be conveniently written as:

$$L_i = [\Phi(-\beta_0 - \beta_1 x_i)]^{(1-Y_i)} [1 - \Phi(-\beta_0 - \beta_1 x_i)]^{Y_i}$$

Since normal distribution is symmetric, we have $\Phi(-\beta_0 - \beta_1 x_i) = 1 - \Phi(\beta_0 + \beta_1 x_i)$. Thus, we have

$$L_i = [1 - \Phi(\beta_0 + \beta_1 x_i)]^{(1-Y_i)} [\Phi(\beta_0 + \beta_1 x_i)]^{Y_i}$$

Thus, the likelihood function is given by

$$L = \prod_{i=1}^n L_i = \prod_{i=1}^n [1 - \Phi(\beta_0 + \beta_1 x_i)]^{(1-Y_i)} [\Phi(\beta_0 + \beta_1 x_i)]^{Y_i}$$


☞ Usually, you maximize $\text{Log}(L)$.

☞ The values of the parameters that maximize $\text{Log}(L)$ are the estimators of the probit model.

☞ The MLE is done automatically by STATA.


The Logit Model


The likelihood function example

 Consider the following model


$$y_i^* = \beta_0 + \beta_1 x_i + u_i$$

$\left\{ \begin{array}{l} \text{If } Y_i = 0, \text{ then } y_i^* < 0 \\ \text{If } Y_i = 1, \text{ then } y_i^* \geq 0 \end{array} \right.$

 In Logit model, we assume that u_i follows the logistic distribution with mean 0 and variance 1.

 The **density function** of the logistic distribution with mean 0 and variance 1 is given by the following.

$$\text{Density Function : } f(x) = \frac{e^{-x}}{1 + e^{-x}}$$

 The **cumulative distribution function** of the logistic distribution with mean 0 and variance 1 has the 'closed form'.

$$\text{Cumulative Distribution Function : } F(x) = \frac{1}{1 + e^{-x}}$$

Id	Y	X
1	0	1
2	0	4
3	1	5
4	1	6
5	1	9

Now, suppose that you have the following data.

Take the 2nd observation as an example. Since $Y_2=0$, it must have been the case that $y_2^* < 0$.

Thus, the likelihood contribution is:


$$\begin{aligned}
 L_2 &= P(y_2^* < 0) = P(\beta_0 + 4\beta_1 + u < 0) \\
 &= P(u < -\beta_0 - 4\beta_1) \\
 &= \int_{-\infty}^{-\beta_0 - 4\beta_1} \frac{e^{-u}}{1 - e^{-u}} = \frac{1}{1 + e^{-(-\beta_0 - 4\beta_1)}} = \frac{1}{1 + e^{\beta_0 + 4\beta_1}}
 \end{aligned}$$

Id	Y	X
1	0	1
2	0	4
3	1	5
4	1	6
5	1	9

Now, take the 3rd observation as an example. Since $Y_3=0$, it must have been the case that $y_3^* < 0$.

Thus, the likelihood contribution is:

$$\begin{aligned}
 L_3 &= P(y_3^* \geq 0) = P(\beta_0 + 5\beta_1 + u \geq 0) \\
 &= P(u \geq -\beta_0 - 5\beta_1) \\
 &= \int_{-\beta_0 - 5\beta_1}^{\infty} \frac{e^{-u}}{1 - e^{-u}} du = 1 - \int_{\infty}^{-\beta_0 - 5\beta_1} \frac{e^{-u}}{1 - e^{-u}} du \\
 &= 1 - \frac{1}{1 + e^{-(\beta_0 + 5\beta_1)}} = 1 - \frac{1}{1 + e^{\beta_0 + 5\beta_1}} \\
 &= \frac{e^{\beta_0 + 5\beta_1}}{1 + e^{\beta_0 + 5\beta_1}}
 \end{aligned}$$


 Thus the likelihood function for the data set is given by

Id	Y	X
1	0	1
2	0	4
3	1	5
4	1	6
5	1	9

$$L = \prod_{i=1}^5 L_i = \frac{1}{1 + e^{\beta_0 + \beta_1}} \times \frac{1}{1 + e^{\beta_0 + 4\beta_1}} \\ \times \frac{e^{\beta_0 + 5\beta_1}}{1 + e^{\beta_0 + 5\beta_1}} \times \frac{e^{\beta_0 + 6\beta_1}}{1 + e^{\beta_0 + 6\beta_1}} \times \frac{e^{\beta_0 + 9\beta_1}}{1 + e^{\beta_0 + 9\beta_1}}$$

Logit model

Likelihood for more general case


 Consider the following model.

Id	Y	X
1	Y_1	x_1
2	Y_2	x_2
3	Y_3	x_3
:	:	:
n	Y_n	x_n

$$y_i^* = \beta_0 + \beta_1 x_i + u_i$$

$$\begin{cases} \text{If } Y_i = 0, \text{ then } y_i^* < 0 \\ \text{If } Y_i = 1, \text{ then } y_i^* \geq 0 \end{cases}$$

We assume u_i follows the logistic distribution.

 Then, we know that


$$\begin{cases} L_i = \frac{1}{1 + e^{\beta_0 + \beta_1 x_i}} \text{ if } Y_i = 0 \\ L_i = \frac{e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \text{ if } Y_i = 1 \end{cases}$$


 The above can be conveniently written as:

$$L_i = \left[\frac{1}{1 + e^{\beta_0 + \beta_1 x_i}} \right]^{(1-Y_i)} \left[\frac{e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \right]^{Y_i}$$


 Thus, the likelihood function is given by

$$L = \prod_{i=1}^n L_i = \prod_{i=1}^n \left[\frac{1}{1 + e^{\beta_0 + \beta_1 x_i}} \right]^{(1-Y_i)} \left[\frac{e^{\beta_0 + \beta_1 x_i}}{1 + e^{\beta_0 + \beta_1 x_i}} \right]^{Y_i}$$

 Usually, you maximize $\log(L)$.

 The values of the parameters that maximize $\log(L)$ is the estimators of the Logit model.

Exercise

 Consider the following model.

$$y^* = \beta_0 + \beta_1 \text{nwifinc} + \beta_2 \text{educ} + \beta_3 \text{exper} \\ + \beta_4 \text{exper}^2 + \beta_5 \text{age} + \beta_6 \text{kidslt6} + \beta_7 \text{kidsge6} \\ + u$$

$\text{inlf} = 0$ if $y^* < 0$ (i.e, not work if $y^* < 0$)

$\text{inlf} = 1$ if $y^* \geq 0$ (i.e., work if $y^* \geq 0$)

 Using Mroz.dta, estimate the parameters using both Probit and Logit models.

Probit

```
. probit inlf nwifeinc educ exper expersq age kidslt6 kidsge6, robust
```

```
Iteration 0: log pseudolikelihood = -514.8732
Iteration 1: log pseudolikelihood = -405.78215
Iteration 2: log pseudolikelihood = -401.32924
Iteration 3: log pseudolikelihood = -401.30219
Iteration 4: log pseudolikelihood = -401.30219
```

```
Probit regression                                Number of obs   =       753
                                                wald chi2(7)    =      185.10
                                                Prob > chi2     =       0.0000
Log pseudolikelihood = -401.30219              Pseudo R2      =       0.2206
```

inlf	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
nwifeinc	-.0120237	.0053106	-2.26	0.024	-.0224323	-.0016152
educ	.1309047	.0258192	5.07	0.000	.0803	.1815095
exper	.1233476	.0188537	6.54	0.000	.086395	.1603002
expersq	-.0018871	.0006007	-3.14	0.002	-.0030645	-.0007097
age	-.0528527	.0083532	-6.33	0.000	-.0692246	-.0364807
kidslt6	-.8683285	.1162037	-7.47	0.000	-1.096084	-.6405735
kidsge6	.036005	.0452958	0.79	0.427	-.0527731	.124783
_cons	.2700768	.505175	0.53	0.593	-.7200481	1.260202

Logit


```
. logit inlf nwifeinc educ exper expersq age kidslt6 kidsge6, robust
```

```
Iteration 0: log pseudolikelihood = -514.8732
Iteration 1: log pseudolikelihood = -406.94123
Iteration 2: log pseudolikelihood = -401.85151
Iteration 3: log pseudolikelihood = -401.76519
Iteration 4: log pseudolikelihood = -401.76515
```

```
Logistic regression                                Number of obs   =       753
                                                    Wald chi2(7)    =      158.48
                                                    Prob > chi2     =       0.0000
Log pseudolikelihood = -401.76515                Pseudo R2      =       0.2197
```


inlf	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
nwifeinc	-.0213452	.0090781	-2.35	0.019	-.039138	-.0035523
educ	.2211704	.0444509	4.98	0.000	.1340483	.3082925
exper	.2058695	.0322914	6.38	0.000	.1425796	.2691594
expersq	-.0031541	.0010124	-3.12	0.002	-.0051384	-.0011698
age	-.0880244	.0144393	-6.10	0.000	-.1163248	-.059724
kidslt6	-1.443354	.2031615	-7.10	0.000	-1.841543	-1.045165
kidsge6	.0601122	.0798825	0.75	0.452	-.0964546	.216679
_cons	.4254524	.8597307	0.49	0.621	-1.259589	2.110494

Interpreting the parameters

 Consider the following model.

$$y_i^* = \beta_0 + \beta_1 x_i + u_i$$

$$\begin{cases} \text{If } Y_i = 0, \text{ then } y_i^* < 0 \\ \text{If } Y_i = 1, \text{ then } y_i^* \geq 0 \end{cases}$$

 In probit and Logit models, interpretation of the parameters is not straightforward.


 For the explanation purpose, consider that this model is estimating the determinants of the labor force participation.


Probit case

 Note that

$$P(y_i^* \geq 0) = 1 - \Phi(-\beta_0 - \beta_1 x_i) = \Phi(\beta_0 + \beta_1 x_i)$$

=the probability that the person participates in the labor force.

 Therefore, if β_1 is positive, an increase in x would increase the probability that the person participates in the labor force.

 If β_1 is negative, then an increase in x will decrease the probability that the person participates in the labor force.

Example

```
. probit inlf nwifeinc educ exper expersq age kidslt6 kidsge6, robust
```

```
Iteration 0: log pseudolikelihood = -514.8732
Iteration 1: log pseudolikelihood = -405.78215
Iteration 2: log pseudolikelihood = -401.32924
Iteration 3: log pseudolikelihood = -401.30219
Iteration 4: log pseudolikelihood = -401.30219
```

Probit regression

```
Number of obs   =      753
Wald chi2(7)    =     185.10
Prob > chi2     =     0.0000
Pseudo R2      =     0.2206
```

Log pseudolikelihood = -401.30219

inlf	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
nwifeinc	-.0120237	.0053106	-2.26	0.024	-.0224323	-.0016152
educ	.1309047	.0258192	5.07	0.000	.0803	.1815095
exper	.1233476	.0188537	6.54	0.000	.086395	.1603002
expersq	-.0018871	.0006007	-3.14	0.002	-.0030645	-.0007097
age	-.0528527	.0083532	-6.33	0.000	-.0692246	-.0364807
kidslt6	-.8683285	.1162037	-7.47	0.000	-1.096084	-.6405735
kidsge6	.036005	.0452958	0.79	0.427	-.0527731	.124783
_cons	.2700768	.505175	0.53	0.593	-.7200481	1.260202

Increase in non wife income will decrease the probability that the woman works.


Increase in education will increase the probability that the woman works.

Increase in the number of kids who are younger than 6 will decrease the probability that the woman works.

The Partial Effects of Probit model

(Continuous variable case)

- ☞ The sign of the parameters can tell you if the probability of the labor force participation will increase or decrease.
- ☞ We also want to know “by how much” the probability increases or decreases.
- ☞ This can be done by computing the **partial effects** (sometimes called the **marginal effects**).


 The increase in the probability due to an increase in x-variable by one unit can be computed by taking the derivative of $\Phi(\beta_0 + \beta_1 x_i)$ with respect to x_i .


Partial effect(Marginal effect)=The increase in the probability that the person works due to one unit increase in x_i

$$= \frac{\partial \Phi(\beta_0 + \beta_1 x_i)}{\partial x_i} = \underbrace{\phi(\beta_0 + \beta_1 x_i)}_{\text{Partial effect}} \beta_1$$

Partial
effect

Note, this is the **density function**, not the cumulative distribution function.

 As you can see, the partial effect depends on the value of the x -variable. Therefore, it is different for different person in the data.

 However, we want to know the overall effect of x on the probability.

 There are two ways to do so.

1. Partial effect at average
2. The average partial effect

Partial effect at average

$$PEA = \phi(\hat{\beta}_0 + \hat{\beta}_1 \bar{x}) \hat{\beta}_1$$

Stata computes this automatically.

This is called the marginal effect in Stata,

This is the partial effect evaluated at the average value of x .


Average partial effect (APE)

$$APE = \frac{1}{n} \sum_{i=1}^n \phi(\hat{\beta}_0 + \hat{\beta}_1 x_i) \hat{\beta}_1$$

This is not done automatically, but easy to do manually.

You compute the partial effect for each person, then take the average.

The partial effect of the probit model (Discrete variable case)


 Consider the following labor force participation model

$$y_i^* = \beta_0 + \beta_1 x_i + \beta_2 D_i + u_i$$

$$\begin{cases} \text{If } Y_i = 0, \text{ then } y_i^* < 0 \\ \text{If } Y_i = 1, \text{ then } y_i^* \geq 0 \end{cases}$$

where D_i is a dummy variable that is 1 if the person lives with parents, and 0 otherwise.


- ☞ You want to know the effect of living with parents on the probability that the woman participates in the labor force.
- ☞ This is a dummy variable. In such a case, the partial effect formula described before is not a good approximation.
- ☞ For a dummy variable case, a better way to compute the partial effect is given in the next slides.

 The partial effect at average for discrete case:

Partial effect at average

$$= \underbrace{\Phi(\hat{\beta}_0 + \hat{\beta}_1 \bar{x} + \hat{\beta}_2 \times 1)}_{\text{Participation probability when living with parents (at average)}} - \underbrace{\Phi(\hat{\beta}_0 + \hat{\beta}_1 \bar{x} + \hat{\beta}_2 \times 0)}_{\text{Participation probability when not living with parents (at average)}}$$

 This is computed automatically by STATA.


 The average partial effect is computed as:

$$(\text{Partial effect})_i = \underbrace{\Phi(\hat{\beta}_0 + \hat{\beta}_1 x_i + \hat{\beta}_2 \times 1)}_{\text{Participation probability when living with parents}} - \underbrace{\Phi(\hat{\beta}_0 + \hat{\beta}_1 x_i + \hat{\beta}_2 \times 0)}_{\text{Participation probability when not living with parents}}$$

Then, the average partial effect is computed as the sample average of the partial effect.

$$\text{Average partial effect} = \frac{1}{n} \sum_{i=1}^n (\text{Partial effect})_i$$

Exercise

 Using Mroz.dta, estimate the following model.

$$y^* = \beta_0 + \beta_2 \text{educ} + \beta_3 \text{exper} + u$$

$\text{inlf} = 0$ if $y^* < 0$ (i.e, not work if $y^* < 0$)

$\text{inlf} = 1$ if $y^* \geq 0$ (i.e., work if $y^* \geq 0$)

$$u \sim N(0,1)$$

Compute the effect of education on the labor force participation. Compute both the partial effect at average and the average partial effect.

```
. probit inlf educ exper
```

```
Iteration 0:  log likelihood = -514.8732
Iteration 1:  log likelihood = -456.47099
Iteration 2:  log likelihood = -455.55276
Iteration 3:  log likelihood = -455.55204
```

```
Probit regression                                Number of obs   =       753
                                                LR chi2(2)      =      118.64
                                                Prob > chi2     =       0.0000
Log likelihood = -455.55204                    Pseudo R2       =       0.1152
```

inlf	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
educ	.104287	.0219681	4.75	0.000	.0612303	.1473437
exper	.0597903	.0065816	9.08	0.000	.0468907	.07269
_cons	-1.713151	.280658	-6.10	0.000	-2.26323	-1.163071

```
. egen aveduc=mean(educ)
```

```
. egen avexper=mean(exper)
```

```
. gen avxbeta=_b[_cons ]+_b[educ ]*aveduc+_b[exper]*avexper
```

```
. gen partial_av=normalden(avxbeta)*_b[educ]
```

```
. su partial_av
```

Variable	Obs	Mean	Std. Dev.	Min	Max
partial_av	753	.0407492	0	.0407492	.0407492

Computing
partial
effect at
average,
manually.

Partial
effect

```
. probit inlf educ exper
```

```
Iteration 0: log likelihood = -514.8732
Iteration 1: log likelihood = -456.47099
Iteration 2: log likelihood = -455.55276
Iteration 3: log likelihood = -455.55204
```

Probit regression

```
Number of obs   =      753
LR chi2(2)      =     118.64
Prob > chi2     =      0.0000
Pseudo R2      =      0.1152
```

Log likelihood = -455.55204

inlf	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
educ	.104287	.0219681	4.75	0.000	.0612303	.1473437
exper	.0597903	.0065816	9.08	0.000	.0468907	.07269
_cons	-1.713151	.280658	-6.10	0.000	-2.26323	-1.163071

```
. mfx, varlist(educ)
```

```
Marginal effects after probit
y = Pr(inlf) (predict)
= .58075593
```

variable	dy/dx	Std. Err.	z	P> z	[95% C.I.]		x
educ	.0407492	.00858	4.75	0.000	.023932	.057567	12.2869

Computing
partial effect
at average
automatically.

```
. probit inlf educ exper, robust
```

```
Iteration 0: log pseudolikelihood = -514.8732
Iteration 1: log pseudolikelihood = -456.47099
Iteration 2: log pseudolikelihood = -455.55276
Iteration 3: log pseudolikelihood = -455.55204
```

```
Probit regression                                Number of obs   =       753
                                                wald chi2(2)    =       80.67
                                                Prob > chi2      =       0.0000
Log pseudolikelihood = -455.55204              Pseudo R2       =       0.1152
```


inlf	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
educ	.104287	.0217204	4.80	0.000	.0617157	.1468583
exper	.0597903	.0078708	7.60	0.000	.0443638	.0752169
_cons	-1.713151	.2802903	-6.11	0.000	-2.26251	-1.163792

```
. predict xbета, xb
. gen partial =normalden(xbета)*_b[educ]
. su partial
```

Variable	Obs	Mean	Std. Dev.	Min	Max
partial	753	.036153	.0072086	.0043549	.0416045

Computing
average
partial
effect,
“manually
”.

Exercise

 Use JPSC1.dta to estimate the following model.

$$y^* = \beta_0 + \beta_2 \text{exper} + \beta_3 (\text{livetogether}) + u$$

Work=0 if $y^* < 0$ (i.e, not work if $y^* < 0$)

Work=1 if $y^* \geq 0$ (i.e., work if $y^* \geq 0$)

$$u \sim N(0,1)$$

Livetogehter is a dummy variable for those who are living with parents.

Q1. Estimate the effect of living with parents on the labor force participation.

```
. probit work exp livetogether
```

```
Iteration 0: log likelihood = -5052.9067
Iteration 1: log likelihood = -3612.6274
Iteration 2: log likelihood = -3517.5372
Iteration 3: log likelihood = -3515.8581
Iteration 4: log likelihood = -3515.8574
```

```
Probit regression                                Number of obs   =      7537
                                                LR chi2(2)      =    3074.10
                                                Prob > chi2     =     0.0000
Log likelihood = -3515.8574                    Pseudo R2      =     0.3042
```

work	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
exp	.1967633	.0043554	45.18	0.000	.1882269	.2052996
livetogether	.2209021	.0363567	6.08	0.000	.1496442	.29216
_cons	-2.225386	.0454465	-48.97	0.000	-2.314459	-2.136313

```
. egen avexp=mean(exp)
. gen xbeta1=_b[_cons ]+_b[exp ]*avexp+_b[livetogether]*1
. gen xbeta0=_b[_cons ]+_b[exp ]*avexp+_b[livetogether]*0
. gen p1=normal(xbeta1)
. gen p0=normal(xbeta0)
. gen partial_av=p1-p0
. su partial_av
```

Variable	Obs	Mean	Std. Dev.	Min	Max
partial_av	7537	.0844965	0	.0844965	.0844965

Computing
partial effect
at average,
“manually”.

Partial
effect at
average


```
. probit work exp livetogether
```

```
Iteration 0:  log likelihood = -5052.9067
Iteration 1:  log likelihood = -3612.6274
Iteration 2:  log likelihood = -3517.5372
Iteration 3:  log likelihood = -3515.8581
Iteration 4:  log likelihood = -3515.8574
```

Probit regression

```
Number of obs   =      7537
LR chi2(2)      =    3074.10
Prob > chi2     =      0.0000
Pseudo R2      =      0.3042
```

Log likelihood = -3515.8574

work	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
exp	.1967633	.0043554	45.18	0.000	.1882269	.2052996
livetogether	.2209021	.0363567	6.08	0.000	.1496442	.29216
_cons	-2.225386	.0454465	-48.97	0.000	-2.314459	-2.136313

Computing the partial effect at average automatically.

```
. mfx, varlist(livetogether)
```

```
Marginal effects after probit
      y = Pr(work) (predict)
      = .37468279
```

variable	dy/dx	Std. Err.	z	P> z	[95% C.I.]		x
liveto~r*	.0844965	.01401	6.03	0.000	.057046	.111947	.337402

(*) dy/dx is for discrete change of dummy variable from 0 to 1

```
. probit work exp livetogether
```

```
Iteration 0:  log likelihood = -5052.9067
Iteration 1:  log likelihood = -3612.6274
Iteration 2:  log likelihood = -3517.5372
Iteration 3:  log likelihood = -3515.8581
Iteration 4:  log likelihood = -3515.8574
```

```
Probit regression               Number of obs   =       7537
                               LR chi2(2)         =      3074.10
                               Prob > chi2         =       0.0000
                               Pseudo R2          =       0.3042

Log likelihood = -3515.8574
```

work	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
exp	.1967633	.0043554	45.18	0.000	.1882269	.2052996
livetogether	.2209021	.0363567	6.08	0.000	.1496442	.29216
_cons	-2.225386	.0454465	-48.97	0.000	-2.314459	-2.136313

Computing
the average
partial effect
“manually”.

```
. gen xb1=_b[_cons]+_b[exp]*exp+_b[livetogether ]*1
. gen xb0=_b[_cons]+_b[exp]*exp+_b[livetogether ]*0
. gen pp1=normal(xb1)
. gen pp0=normal(xb0)
. gen partial2=pp1-pp0
. browse
. su partial2
```

Variable	Obs	Mean	Std. Dev.	Min	Max
partial2	7537	.0594117	.0239495	.0000651	.0879483