Ministério da Ciência, Tecnologia e Inovações



Centro Brasileiro de Pesquisas Físicas



Métodos para Análise de grande volume de dados e Astroinformática

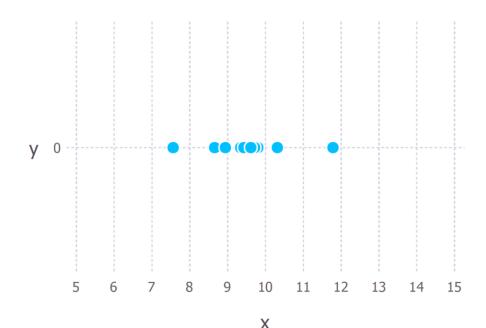
Clécio Roque De Bom - debom@cbpf.br



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Maximum likelihood estimation is a method that determines values for the parameters of a GIVEN model.



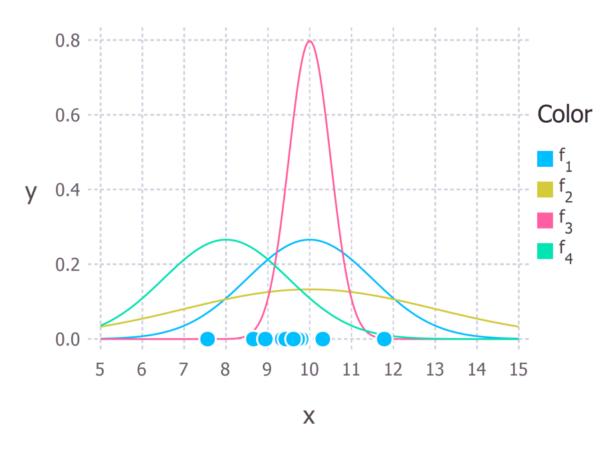
Independent and identically distributed random variables

the i.i.d. assumption requires that the observation of any given data point does not depend on the observation of any other data point (each gathered data point is an independent experiment) and that each data point is generated from same distribution family with the same parameters.

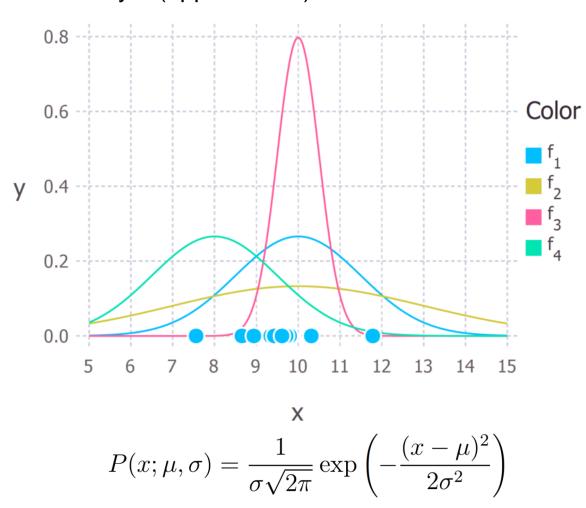
Figure credit, adapted from: https://towardsdatascience.com/probability-concepts-explained-maximum-likelihood-estimation-c7b4342fdbb1

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Assume that the data generation process can be described by a (approximate) Gaussian distribution.



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X

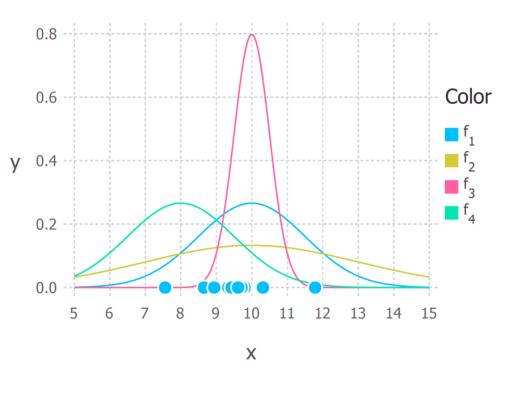
$$P(x;\mu,\sigma) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

$$\downarrow f_1 \\ f_2 \\ f_3 \\ f_4 \end{pmatrix} P(9,9.5,11;\mu,\sigma) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(9-\mu)^2}{2\sigma^2}\right) \times \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(9.5-\mu)^2}{2\sigma^2}\right) \times \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(11-\mu)^2}{2\sigma^2}\right)$$

Now, all we need to do is to maximize it for the parameters!

 $P(x_1, x_2, ..., x_n | \theta) \rightarrow f(x_1, x_2, ..., x_n | \theta)$

$$\widehat{\theta}_{MLE} = argmax_{\theta} \prod_{i}^{n} f(x_{i}|\theta)$$



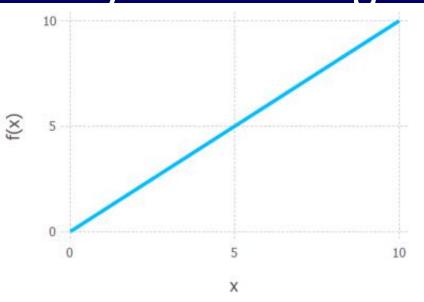
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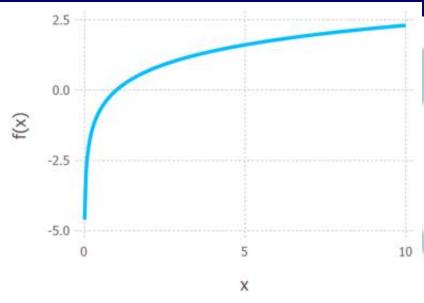
$$\widehat{\theta}_{MLE} = argmax_{\theta} \prod_{i}^{n} f(x_{i}|\theta)$$

$$argmax_{\theta} \prod_{i}^{n} f(x_{i}|\theta) \rightarrow \frac{\partial}{\partial \theta} \prod_{i}^{n} f(x_{i}|\theta) = 0$$

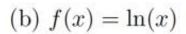
$$\frac{\partial}{\partial \theta} \prod_{i}^{n} f(x_{i}|\theta) \sim \frac{\partial}{\partial \theta} \ln \langle \prod_{i}^{n} f(x_{i}|\theta) \rangle = \frac{\partial}{\partial \theta} \sum_{i}^{n} \ln \langle f(x_{i}|\theta) \rangle$$
$$= \sum_{i}^{n} \frac{\partial}{\partial \theta} \ln \langle f(x_{i}|\theta) \rangle = 0$$

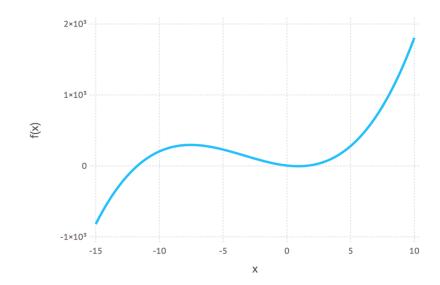
Monotonically inscreasing function



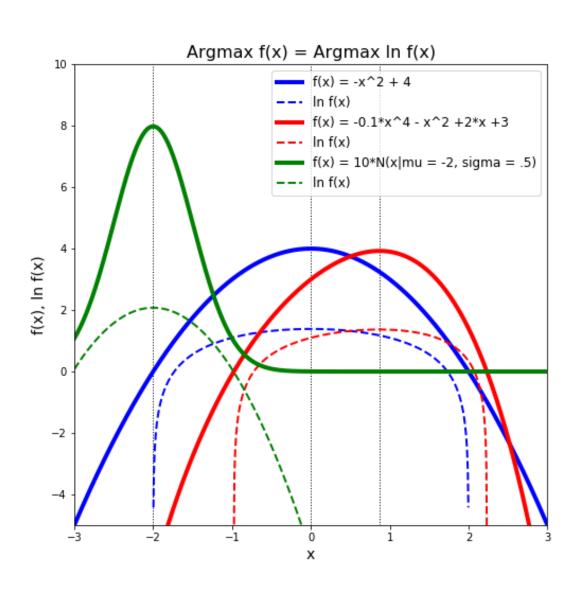


(a)
$$f(x) = x$$





Monotonically inscreasing function



$$\sum_{i}^{n} \frac{\partial}{\partial \theta} \ln \langle f(x_{i}|\theta) \rangle = \sum_{i}^{n} \nabla_{\mu,\sigma} \ln \langle f(x_{i}|\mu,\sigma) \rangle = 0$$

$$\sum_{i}^{n} \nabla_{\mu} \ln \langle \frac{1}{\sigma \sqrt{2\pi}} e^{\frac{-(x_{i} - \mu)^{2}}{2\sigma^{2}}} \rangle = \sum_{i}^{n} \nabla_{\mu} - \frac{1}{2} \ln(2\pi\sigma^{2}) - \frac{(x_{i} - \mu)^{2}}{2\sigma^{2}}$$

$$\sum_{i}^{n} \nabla_{\mu} - \frac{1}{2} \ln(2\pi\sigma^{2}) - \frac{(x_{i} - \mu)^{2}}{2\sigma^{2}} = -\frac{1}{2\sigma^{2}} \sum_{i}^{n} \nabla_{\mu} (x_{i} - \mu)^{2}$$

$$= -\frac{1}{2\sigma^2} \sum_{i=1}^{n} -2(x_i - \mu) = \frac{1}{\sigma^2} \sum_{i=1}^{n} (x_i - \mu)$$

$$\sum_{i}^{n} \frac{\partial}{\partial \theta} \ln \langle f(x_{i}|\theta) \rangle = \sum_{i}^{n} \nabla_{\mu,\sigma} \ln \langle f(x_{i}|\mu,\sigma) \rangle = 0$$

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$$= -\frac{1}{2\sigma^2} \sum_{i}^{n} -2(x_i - \mu) = \frac{1}{\sigma^2} \sum_{i}^{n} (x_i - \mu)$$

$$\frac{1}{\sigma^2} \sum_{i}^{n} (x_i - \mu) = 0 \rightarrow \hat{\mu}_{MLE} = \frac{1}{n} \sum_{i}^{n} x_i$$

$$\sum_{i}^{n} \nabla_{\sigma} \ln \langle \frac{1}{\sigma \sqrt{2\pi}} e^{\frac{-(x_{i} - \mu)^{2}}{2\sigma^{2}}} \rangle = \sum_{i}^{n} \nabla_{\sigma} - \frac{1}{2} \ln(2\pi\sigma^{2}) - \frac{(x_{i} - \mu)^{2}}{2\sigma^{2}}$$

$$\sum_{i}^{n} \nabla_{\sigma} - \frac{1}{2} \ln(2\pi\sigma^{2}) - \frac{(x_{i} - \mu)^{2}}{2\sigma^{2}}$$

$$= -\frac{n}{2} \nabla_{\sigma} \ln \sigma^{2} - \frac{1}{2} \nabla_{\sigma} \langle \frac{1}{\sigma^{2}} \sum_{i}^{n} (x_{i} - \mu)^{2} \rangle$$

$$= -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i}^{n} (x_i - \mu)^2$$

1

Maximum Likelihood Estimation - Example

$$P(x; \mu, \sigma) = \frac{1}{\sigma \sqrt{2\pi}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right)$$
0.6
$$\int_{0.6}^{6} f_1 \qquad P(9, 9.5, 11; \mu, \sigma) = \frac{1}{\sigma \sqrt{2\pi}} \exp\left(-\frac{(9 - \mu)^2}{2\sigma^2}\right) \times \frac{1}{\sigma \sqrt{2\pi}} \exp\left(-\frac{(9.5 - \mu)^2}{2\sigma^2}\right)$$

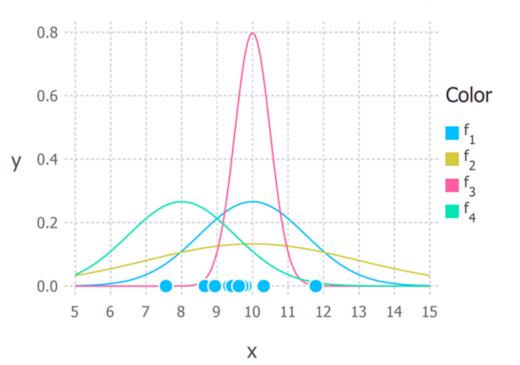
$$\times \frac{1}{\sigma \sqrt{2\pi}} \exp\left(-\frac{(11 - \mu)^2}{2\sigma^2}\right)$$
0.0
$$\times \frac{1}{\sigma \sqrt{2\pi}} \exp\left(-\frac{(11 - \mu)^2}{2\sigma^2}\right)$$

$$\ln(P(x;\mu,\sigma)) = \ln\left(\frac{1}{\sigma\sqrt{2\pi}}\right) - \frac{(9-\mu)^2}{2\sigma^2} + \ln\left(\frac{1}{\sigma\sqrt{2\pi}}\right) - \frac{(9.5-\mu)^2}{2\sigma^2} + \ln\left(\frac{1}{\sigma\sqrt{2\pi}}\right) - \frac{(11-\mu)^2}{2\sigma^2}$$

Back to our example with numbers

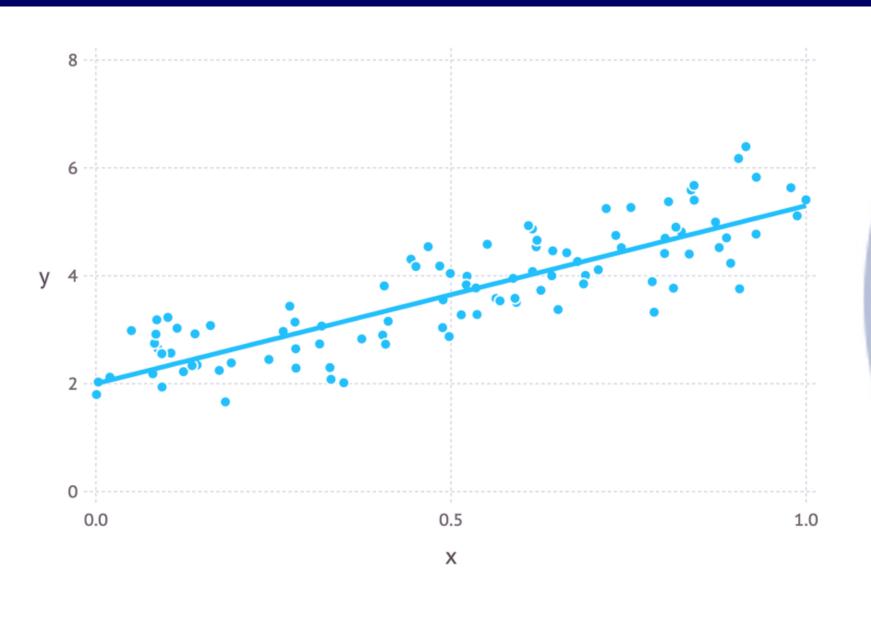
$$\ln(P(x;\mu,\sigma)) = -3\ln(\sigma) - \frac{3}{2}\ln(2\pi) - \frac{1}{2\sigma^2}\left[(9-\mu)^2 + (9.5-\mu)^2 + (11-\mu)^2\right]$$

$$\frac{\partial \ln(P(x;\mu,\sigma))}{\partial \mu} = \frac{1}{\sigma^2} \left[9 + 9.5 + 11 - 3\mu \right].$$

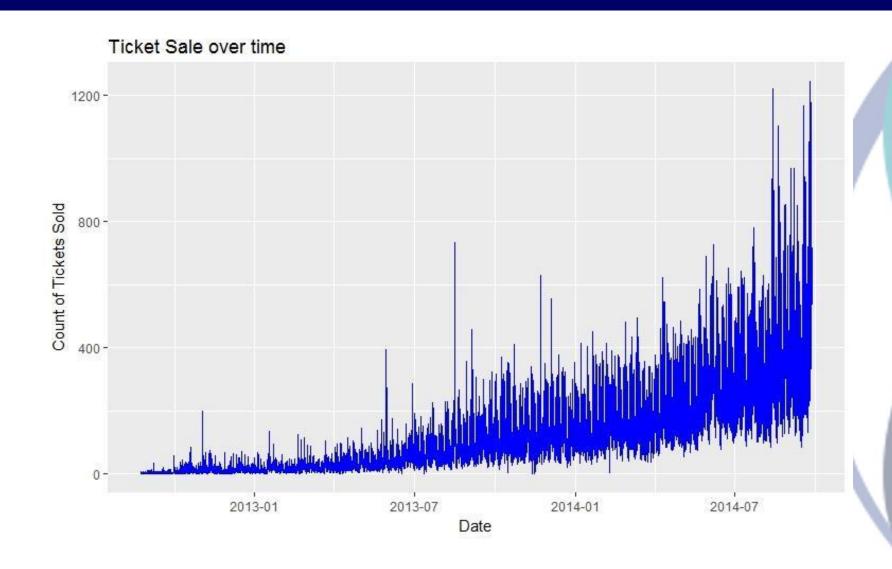


$$\mu = \frac{9 + 9.5 + 11}{3} = 9.833$$

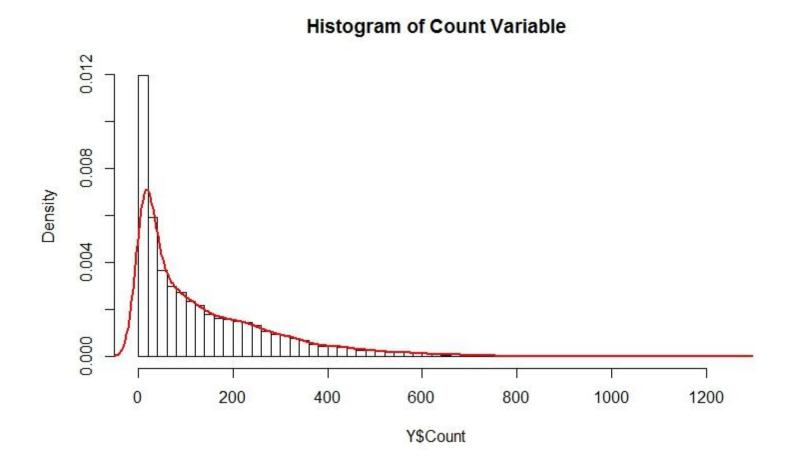
This was trivial, why bother?



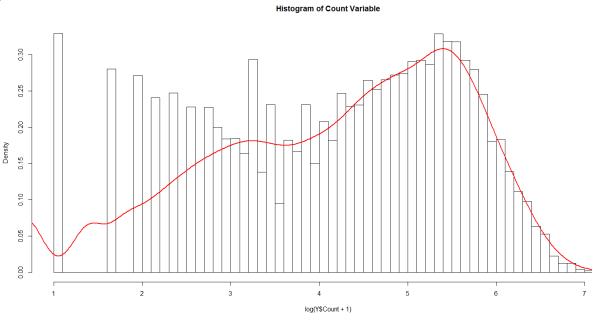
MLE



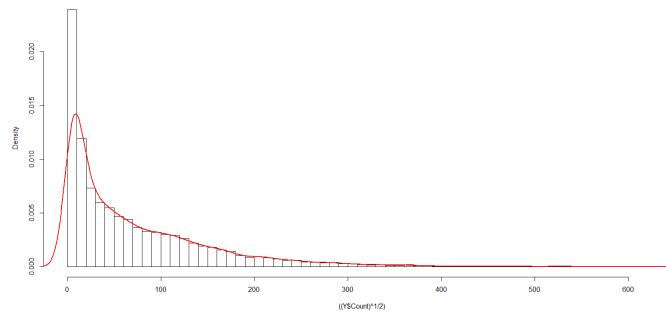
MLE

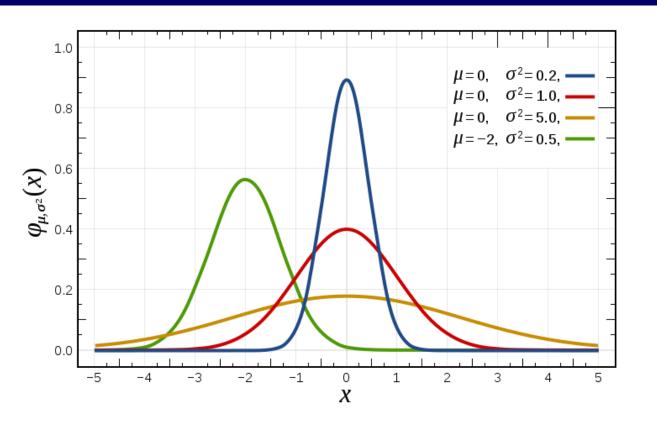


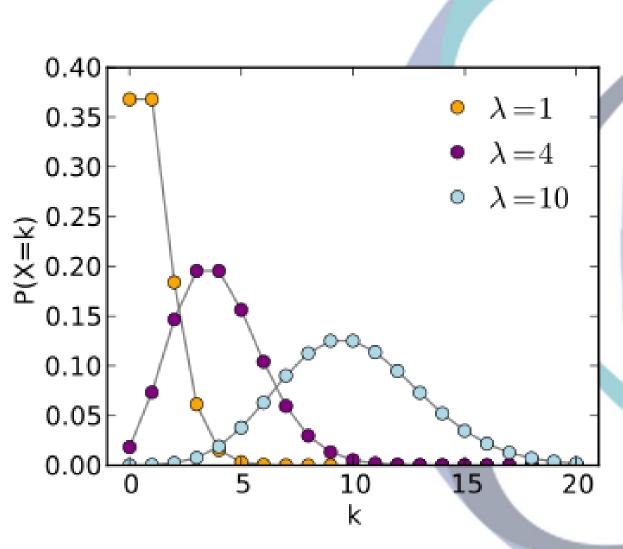
MLE

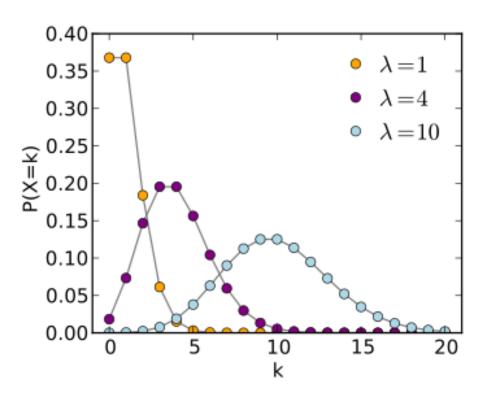












$$log(\mu_i) = x_i' \theta$$
or
$$\mu_i = exp(x_i' \theta)$$

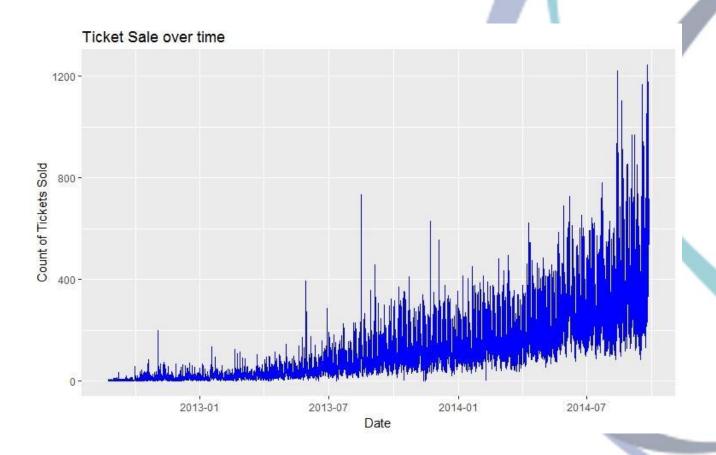
$$Pr{Y = y | \mu} = (e^{-\mu} \mu^{y})/ y!$$

LL(
$$\theta$$
) = $\sum \{y_i \log(\mu_i) - \mu_i\}$, ------ Eq. 1

$$Pr{Y = y | \mu} = (e^{-\mu} \mu^{y})/ y!$$

$$\mu = \exp(\theta_0 + age * \theta_1)$$

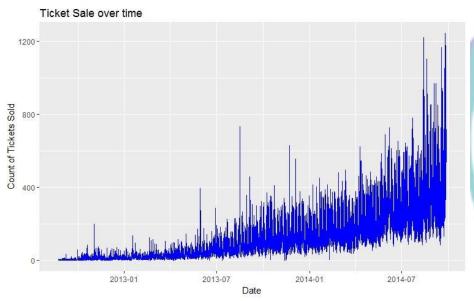
LL(
$$\theta$$
) = $\sum \{y_i \cdot (\theta_0 + age * \theta_1) - exp(\theta_0 + age * \theta_1)\}$



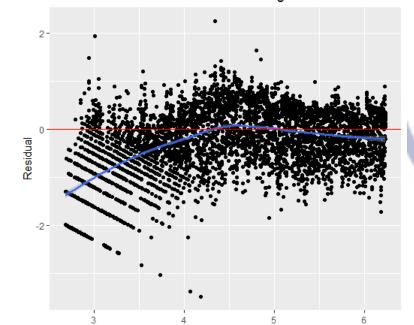
$$Pr{Y = y | \mu} = (e^{-\mu} \mu^{\nu})/ y!$$

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LL(
$$\theta$$
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Predicted Vs Residual for Poission Regression



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