# Homework 1

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ST562

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## 1 6.2.1

$$f_X(x) = p^2 (1-p)^{x_1} (1-p)^{x_2} = p^2 (1-p)^{x_1+x_2}$$

$$f_T(t) = \sum_{i=0}^{t} p(1-p)^i p(1-p)^{t-i} = p^2(t+1)(1-p)^t$$

$$\frac{f_X(x)}{f_T(t)} = \frac{p^2(1-p)^{x_1+x_2}}{p^2(t+1)(1-p)^{x_1+x_2}} = \frac{1}{t+1}$$

 $\frac{f_X(x)}{f_T(t)}$  is not a function of p.

## $2\quad 6.2.2$

Substituting in 1-p for q, and letting  $\hat{x} = \sum_{i=0}^{m} x_i$  and  $\hat{y} = \sum_{i=0}^{n} y_i$  gives:

$$f_{X,Y}(x,y) = \left(\prod_{1}^{m} p^{x_i} (1-p)^{1-x_i}\right) \left(\prod_{1}^{n} p^{1-y_i} (1-p)^{y_i}\right) = p^{n+\hat{x}-\hat{y}} (1-p)^{m-\hat{x}+\hat{y}}$$

 $\hat{X}$  is distributed  $Bin(p,m), \hat{Y}$  is distributed  $Bin(1-p,n), T=\hat{X}-\hat{Y}$ . Using the discrete version of convolution for subtracting random variables we have:

$$f_T(t) = \sum_{i=0}^{n} f_{\hat{Y}}(\hat{y}_i) f_{\hat{X}}(t + \hat{y}_i) =$$

$$\sum_{0}^{n} \left[ C_{\hat{y_i}}^n p^{n-\hat{y_i}} (1-p)^{\hat{y_i}} C_{t+\hat{y_i}}^m p^{t+\hat{y_i}} (1-p)^{m-(t+y_i)} \right] =$$

$$p^{n+t}(1-p)^{m-t}\sum_{i=0}^{n}C_{\hat{y_i}}^nC_{t+\hat{y_i}}^m$$

$$\frac{f_{X,Y}(x,y)}{f_T(t)} = \frac{p^{n+\hat{x}-\hat{y}}(1-p)^{m-\hat{x}+\hat{y}}}{p^{n+t}(1-p)^{m-t}\sum_{0}^{n}C_{\hat{y_i}}^{n}C_{t+\hat{y_i}}^{m}} = \frac{1}{\sum_{0}^{n}C_{\hat{y_i}}^{n}C_{t+\hat{y_i}}^{m}}$$

 $\frac{f_{X,Y}(x,y)}{f_T(t)}$  is not a function of p.

#### $3 \quad 6.2.5$

Using the resulting form of  $f_{X,Y}(x,y)$  from the previous problem:

$$f_{X,Y}(x,y) = p^{3+(x_1+x_2)-(y_1+y_2+y_3)} (1-p)^{2-(x_1+x_2)+(y_1+y_2+y_3)}$$

 $Y_1Y_2 = 1$  iff  $Y_1 = Y_2 = 1$ , so:

$$f_{Y_1Y_2}(y) = (q^2)^y (1 - q^2)^{1-y} = (1 - p)^{2y} [1 - (1 - p)^2]^{1-y}$$

Now by discrete convolution, where  $T = X_1 + Y_1Y_2$ :

$$f_T(t) = f_X(t)f_{Y_1Y_2}(0) + f_X(t-1)f_{Y_1Y_2}(1) = p^t(1-p)^{1-t}p^2 + p^{t-1}(1-p)^{1-(t-1)}(1-p^2) = p^t(1-p)^{1-t}p^2 + p^{t-1}(1-p)^{1-t}p^2 + p$$

$$p^{t+2}(1-p)^{1-t} + p^{t-1}(1-p)^{2-t}(1-p^2)$$

$$\frac{f_{X,Y}(x,y)}{f_T(t)} = \frac{p^{3+(x_1+x_2)-(y_1+y_2+y_3)}(1-p)^{2-(x_1+x_2)+(y_1+y_2+y_3)}}{p^{t+2}(1-p)^{1-t}+p^{t-1}(1-p)^{2-t}(1-p^2)}$$

Cannot cancel out p's out of the above,  $\frac{f_{X,Y}(x,y)}{f_T(t)}$  is a function of p.

## 4 6.2.6

$$f_{X,Y}(x,y) = \prod_{1}^{m} \frac{\lambda^{x_i} e^{-\lambda}}{x_i!} \prod_{1}^{n} \frac{2^{y_i} \lambda^{y_i} e^{-2\lambda}}{y_i!} = \frac{2^{(\sum_{i=1}^{n} y_i)} \lambda^{(\sum_{i=1}^{n} x_i) + (\sum_{i=1}^{n} y_i)} e^{-\lambda(m+2n)}}{(\prod_{i=1}^{m} x_i!)(\prod_{i=1}^{n} y_i!)}$$

Since  $Poisson(\lambda_1) + Poisson(\lambda_2) = Poisson(\lambda_1 + \lambda_2)$ :

$$f_T(t) = \frac{(\lambda[m+2n])^t e^{-\lambda(m+2n)}}{t!}$$

$$\frac{f_{X,Y}(x,y)}{f_T(t)} = \frac{\frac{2^{(\sum_0^n y_i)} \lambda^{(\sum_0^m x_i) + (\sum_0^n y_i)} e^{-\lambda(m+2n)}}{(\prod_1^m x_i!)(\prod_1^n y_i!)}}{\frac{(\lambda[m+2n])^t \ e^{-\lambda(m+2n)}}{t!}} = \frac{2^{(\sum_0^n y_1)} t!}{(\prod_1^m x_i!)(\prod_1^n y_i!)(m+2n)^t}$$

 $\frac{f_{X,Y}(x,y)}{f_{T}(t)}$  is not a function of  $\lambda$ .

## 5 6.2.7

$$f_{X,Y}(x,y) = \frac{2^{(\sum_{1}^{5} y_{i})} \lambda^{(\sum_{1}^{4} x_{i}) + (\sum_{1}^{5} y_{1})e^{-14\lambda}}}{(\prod_{1}^{4}) x_{i}!)(\prod_{1}^{5} y_{i}!)}$$

$$f_{T}(t) = \frac{(3\lambda)^{t} e^{-3\lambda}}{t!}$$

$$\frac{f_{X,Y}(x,y)}{f_{T}(t)} = \frac{\frac{2^{(\sum_{1}^{5} y_{i})} \lambda^{(\sum_{1}^{4} x_{i}) + (\sum_{1}^{5} y_{1})e^{-14\lambda}}}{(\prod_{1}^{4}) x_{i}!)(\prod_{1}^{5} y_{i}!)}}{\frac{(3\lambda)^{t} e^{-3\lambda}}{t!}} = \frac{2^{(\sum_{1}^{5})} \lambda^{(\sum_{1}^{2} x_{i}) + (\sum_{1}^{2} y_{i})} e^{-11\lambda}}{(\prod_{1}^{m} x_{i})(\prod_{1}^{n} y_{i})}$$

Cannot cancel p's from above,  $\frac{f_{X,Y}(x,y)}{f_T(t)}$  is a function of p.

## 6 6.2.10

(i) 
$$f_X(x) = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left\{-\frac{\sum_1^n (x_i - \mu)^2}{2\sigma^2}\right\} = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left\{-\frac{\sum_1^n (x_i^2 - 2x_i\mu + \mu^2)}{2\sigma^2}\right\} = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left\{-\frac{\sum_1^n x_i^2}{2\sigma^2}\right\} \exp\left\{n\frac{2\mu n^{-1} \sum_1^n x_i - \mu^2}{2\sigma^2}\right\}$$

(ii) 
$$f_X(x) = \frac{1}{(2\pi\sigma^2)^{n/2}} \exp\left\{-\frac{nn^{-1}\sum_1^n(x_i-\mu)^2}{2\sigma^2}\right\} 1$$

## 7 - 6.2.11

(i) 
$$f_X(x) = \sigma^{-n} \exp\left\{-\sum_{1}^{n} (x_i - \mu)/\sigma\right\} I(x_{(1)} > \mu) =$$
 
$$\sigma^{-n} \exp\left\{-\sum_{1}^{n} (x_i - \mu)/\sigma\right\} \exp\{\mu n/\sigma\} I(x_{(1)} > \mu)$$

(ii) 
$$f_X(x) = \sigma^{-n} \exp \left\{ -nn^{-1} \sum_{1}^{n} (x_i - \mu) / \sigma \right\} I(x_{(1)} > \mu)$$

(iii) 
$$f_X(x) = \sigma^{-n} \exp\left\{-\sum_{1}^{n} (x_i - \mu)/\sigma\right\} I(x_{(1)} > \mu) =$$

$$\sigma^{-n} \exp\left\{-\sum_{1}^{n} [(x_i - x_{(1)}) - (\mu - x_{(1)})]/\sigma\right\} I(x_{(1)} > \mu) =$$

$$\sigma^{-n} \exp\left\{-\left[\left(\sum_{1}^{n} x_i - x_{(1)}\right) - n\mu + nx_{(1)}\right]/\sigma\right\} I(x_{(1)} > \mu)$$

## 8 6.2.13

$$f_X(x) = I\left(\theta - \frac{1}{2} < X_{(1)}\right) I\left(X_{(n)} < \theta + \frac{1}{2}\right) 1$$

## 9 6.2.16

$$f_X(x) = 2^n \theta^{-n} \exp\left\{-\left(\sum_{i=1}^n x_i^2\right)/\theta\right\} \left(\prod_{i=1}^n x_i\right) I(x_{(1)} > 0)$$