Advanced Time Series Analysis: Computer Exercise 2

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Part 1

There are simulated n = 3000 where $\epsilon_t \sim \mathcal{N}(0, 1)$. ϵ_t is used as noise input for all simulations in part one.

The equation below shows the used parameters in the SETAR(2,1,1). Let us call eq. 1 and eq. 2 parameter set one (par_1) .

$$a_0 = [0.125, -0.125] \tag{1}$$

$$a_1 = [0.6, -0.4] \tag{2}$$

Simulation of the SETAR(2,1,1)

The Self-Exciting Threshold AR (SETAR) model is given by eq. 3.

$$X_{t} = a_{0}^{(J_{t})} + \sum_{i=1}^{k_{(J_{t})}} a_{i}^{(J_{t})} X_{t-i} + \epsilon^{(J_{t})}$$

$$\tag{3}$$

where J_t are regime processes. The complete model are defined in eq. 4.

$$X_{t} = \begin{cases} a_{0,1} + a_{1,1}X_{t-1} + \epsilon_{t} & for \quad X_{t-1} \leq 0 \\ a_{0,2} + a_{1,2}X_{t-1} + \epsilon_{t} & for \quad X_{t-1} > 0 \end{cases}$$

$$\tag{4}$$

The model X_t (eq. 4) has been simulated with par_1 . Its simulation is plotted in fig. ??.

Estimate the parameters using conditional least squares

```
Setar <- function(par, model) {
    #
    e_mean <- rep(NA, length(model))
#
    for (t in 2:length(model)) {
        if (model[t - 1] <= 0) {
            e_mean[t] <- par[1] + par[2] * model[t - 1]
        } else {
            e_mean[t] <- par[3] + par[4] * model[t - 1]
        }
    }
}</pre>
```

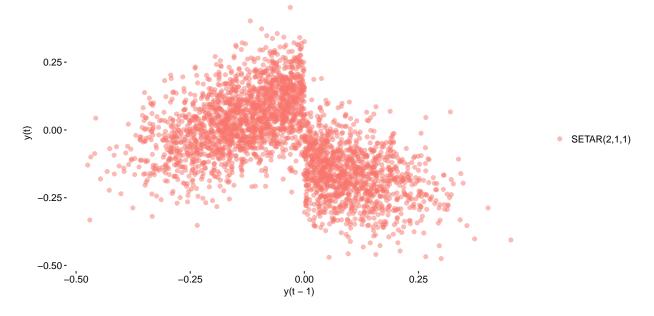


Figure 1: Two simulated SETAR(2,1,1) models using par_1 and par_2 .

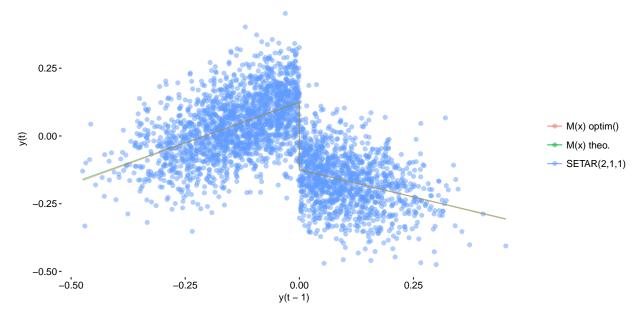
```
#
return(e_mean)
}

RSSSetar <- function(par, model) {
    # conditional mean
    e_mean <- Setar(par, model)

    ## Calculate and return the residuals
    return((model - e_mean)^2)
}

PESetar <- function(par, model) {
    # conditional mean
    e_mean <- Setar(par, model)

    ## Calculate and return the objective function value
    return(sum((model - e_mean)^2, na.rm = TRUE))
}</pre>
```



Den lodrette linje findes selfølgelig ikke!! comment !!!

Part 2

resolution 50 max_change_p 0.1 only change the slope par [2] and par [4]

N = 1:3000

2

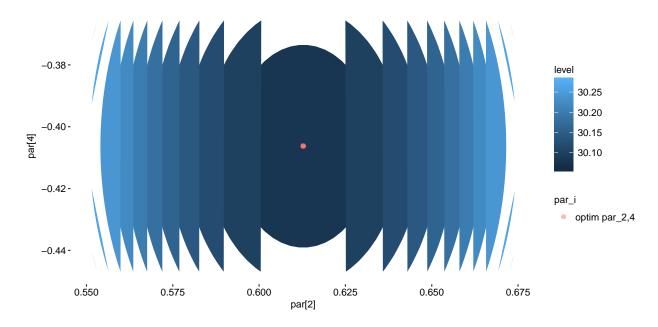
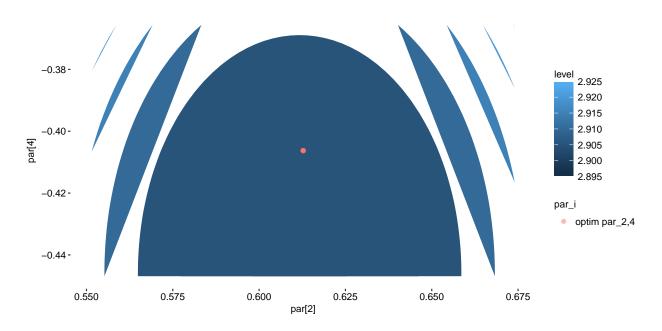
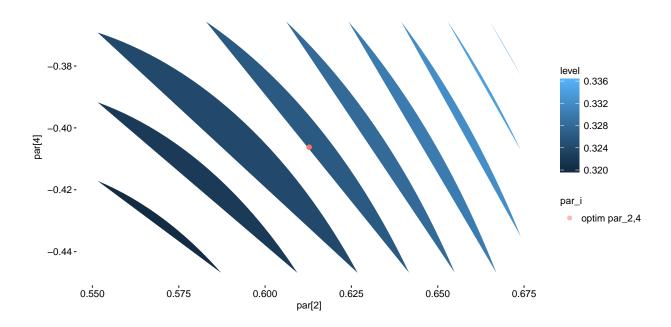


Figure 2: Contour plot of the conditional parametric model approach.

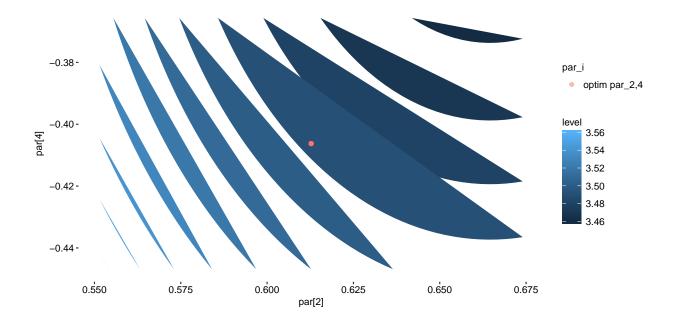
N = 1:300



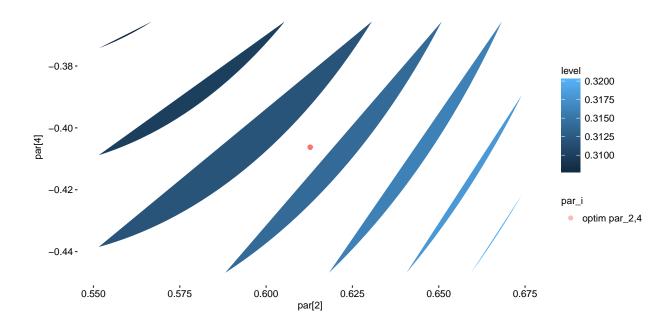
N = 1:30



N = 1001:1300



N = 1001:1030



Discuss my findings

Part 3

AUTO regression outside one -> keep growing

I will consider the following AR(2)-AR(4) non-linear doubly stochastic model, eq. 5.

$$Y_{t} = \sum_{k=1}^{2} \left(\Phi_{t-(1-k)} Y_{t-k} \right) + \epsilon_{t}$$

$$\Phi_{t} - \mu = \sum_{n=1}^{4} \left(\phi_{n} \left(\Phi_{t-n} - \mu \right) \right) + \zeta_{t}$$

$$\Phi_{t} = \sum_{n=1}^{4} \left(\phi_{n} \left(\Phi_{t-n} - \mu \right) \right) + \zeta_{t} + \underbrace{\mu \left(1 - \sum_{n=1}^{4} \left(\phi_{n} \right) \right)}_{\delta}$$
(5)

state space

$$\begin{pmatrix}
\Phi_{t} \\
\Phi_{t-1} \\
\Phi_{t-2} \\
\Phi_{t-3} \\
\delta_{t}
\end{pmatrix} = \begin{pmatrix}
\phi_{1} & \phi_{2} & \phi_{3} & \phi_{4} & 1 \\
1 & 0 & 0 & 0 & 0 \\
0 & 1 & 0 & 0 & 0 \\
0 & 0 & 1 & 0 & 0 \\
0 & 0 & 0 & 0 & 1
\end{pmatrix} \begin{pmatrix}
\Phi_{t-1} \\
\Phi_{t-2} \\
\Phi_{t-3} \\
\Phi_{t-4} \\
\delta_{t-1}
\end{pmatrix} + \begin{pmatrix}
1 \\
0 \\
0 \\
0 \\
0
\end{pmatrix} \delta_{t}$$

$$Y_{t} = (Y_{t-1} \quad Y_{t-1} \quad 0 \quad 0 \quad 0) \begin{pmatrix}
\Phi_{t} \\
\Phi_{t-1} \\
\Phi_{t-2} \\
\Phi_{t-3} \\
\delta_{t}
\end{pmatrix} + e_{t}$$

$$(6)$$

delta som state i stedet for constant \rightarrow estimate

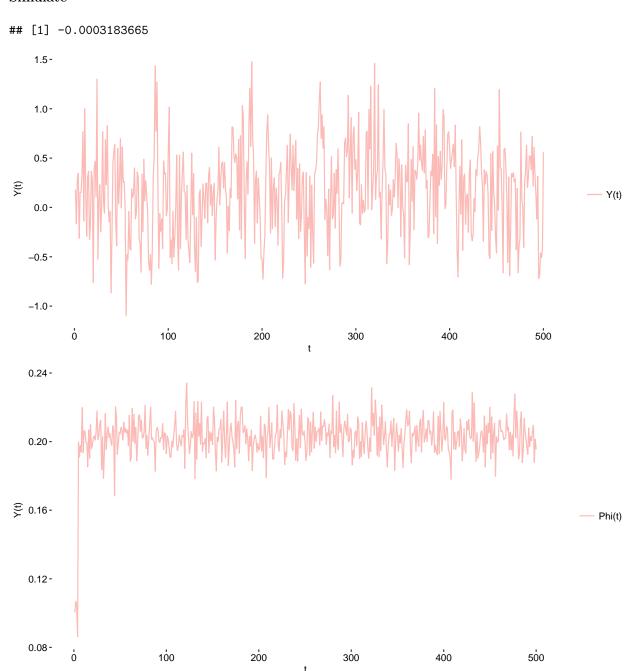
Tænk hvor havd der sker i den underlæggende proces?

hvordan er stationary conditions?

Fordele ved at se delta som et state..

Tjek for stabilitet

Simulate



Comment

Part 4

Following simple state space model is given, eq. 7.

$$x_{t+1} = ax_t + v_t y_t = x_t + e_t \tag{7}$$

where a is an unknown parameter and v_t and e_t are mutually uncorrelated white noise processes with their variences σ_v^2 and σ_e^2 .

Part 4a

The model from eq. 7 is on state space form in eq. 8

$$x_{t+1} = ax_t + v_t$$

$$y_t = x_t + e_t$$
(8)

Simulate

Simulate X time series where $a=0.4,\,\sigma_v^2=\sigma_e^2=1$ with zero mean

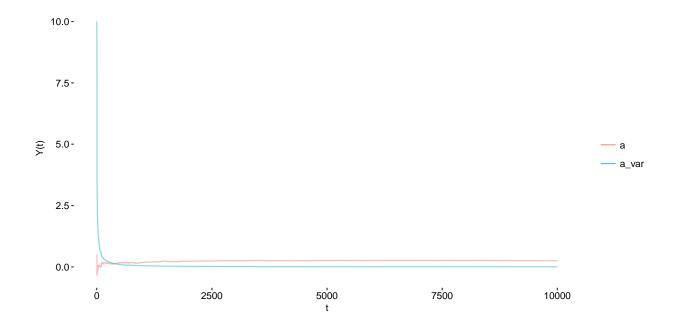
Rewrite

$$\begin{pmatrix} x_{t+1} \\ a_{t+1} \end{pmatrix} = \begin{pmatrix} a_t & 0 \\ 0 & a_t \end{pmatrix} \begin{pmatrix} x_t \\ 1 \end{pmatrix} + \begin{pmatrix} v_t \\ 0 \end{pmatrix}
y_t = \begin{pmatrix} 1 & 0 \end{pmatrix} \begin{pmatrix} x_t \\ 1 \end{pmatrix} + e_t$$
(9)

Part 4b

```
Pt <- matrix(c(Re, 0, 0, aVarInit), nrow = 2, ncol = 2)
# The state is [X a] so the differentiated observation function is
Ht <- t(c(1, 0))
# Init a vector for keeping the parameter a variance estimates
aVar <- rep(NA, length(y))
# and keeping the states
Z <- matrix(NA, nrow = length(y), ncol = 2)</pre>
Z[1, ] <- zt
## The Kalman filtering----
for (t in 1:(length(y) - 1)) {
    # Derivatives (Jacobians)
   Ft <- matrix(c(zt[2], 0, zt[1], 1), ncol = 2) # F_t-1
    # Ht does not change
   ## Prediction step
   zt = c(zt[2] * zt[1], zt[2]) #z_t/t-1 f(z_t-1/t-1)
   Pt = Ft %*% Pt %*% t(Ft) + Rv \#P_t/t-1
    ## Update step
    res = y[t] - zt[1] # the residual at time t
   St = Ht %*% Pt %*% t(Ht) + Re # innovation covariance
   Kt = Pt %*% t(Ht) %*% St^-1 # Kalman gain
    zt = zt + Kt * res # z_t/t
   Pt = (diag(2) - Kt \% + Ht) \% + Pt \#P_t/t
    ## Keep the state estimate
   Z[t + 1, ] \leftarrow zt
    ## Keep the P[2,2], which is the variance of the estimate of a
   aVar[t + 1] <- Pt[2, 2]
return(list(zt = zt, Pt = Pt, Rv = Rv, aVar = aVar, Z = Z))
```

Check for converges in worst case



a = 0.5

state	$sigma_v^2$	$sigma_a$	a mean	a sd	a_var mean	$a_var sd$
1	10	1	0.2227370	0.0204999	0.0152562	0.0004328
2	1	1	0.3955557	0.0292888	0.0008427	0.0000681
3	10	10	0.2188578	0.0208357	0.0154694	0.0004450
4	1	10	0.3955225	0.0292323	0.0008432	0.0000693

a = -0.5

state	$sigma_v^2$	$sigma_a$	a mean	a sd	a_var mean	$a_var sd$
1	10	1	0.2227370	0.0204999	0.0152562	0.0004328
2	1	1	0.3955557	0.0292888	0.0008427	0.0000681
3	10	10	0.2188578	0.0208357	0.0154694	0.0004450
4	1	10	0.3955225	0.0292323	0.0008432	0.0000693

hvordan påvirker størrelsen ad sigma_v2 og hvordan påvirkes variance of the system?

Improvements

do regulizing of the sigma vector.. add some to the diagonal $\,$