

Cognitive Ability, Education, and Fertility Risk*

Agustín Díaz Casanueva[†]

January 17, 2023

**Preliminary and Incomplete.
Do Not Cite.**

Abstract

The paper provides empirical evidence using NLSY79 data that cognitive ability affects pregnancy timing and intention. Then, I build and estimate a life cycle model to understand whether education groups' differences in opportunity cost and contraception efficiency explain this effect. I find that these mechanisms cannot explain why women in the lowest cognitive ability decile have nine times more pregnancies than those in the top. Finally, I use the model to show that policies that decrease contraception cost decrease early pregnancies but do not improve college attendance, as college is too costly for this group of women, even without a child.

*

[†]University of Pennsylvania, andiaz@sas.upenn.edu

1 Introduction

One of the most important economic decisions is when having children, as the timing affects parents' and children's welfare, fertility decisions have far-reaching social and economic consequences. Since the introduction of the pill, women potentially have almost complete certainty regarding pregnancies. However, in the US, women report that 41% of pregnancies are not at the right time or even wanted. Moreover, this fraction increase for younger women; for example, among teens aged 15 to 19 years old, the fraction increases to 75% ¹, which suggests that an important fraction of women have children, not at the desired moment. In this paper, I study the mechanism behind the fact that teens and out-of-wedlock pregnancies, which in most cases are unintended, are mainly concentrated in low cognitive ability women. For example, women in the bottom ability quartile have nine times more teen pregnancies than women in the top quartile.

Heckman et al. (2006) show that noncognitive and cognitive ability affect multiple social outcomes, including teen pregnancies. However, the channels through which ability affect fertility is an open question. For this reason, in this paper, I study whether the relationship between ability, formal education, and wages can fully explain the role of ability in fertility. The two most studied explanations of why attending college affect pregnancies are the allocation of time theory and differences in contraception efficiency by education groups. The allocation of time theory (Becker, 1965) proposes that highly educated women have a higher opportunity cost of having children early in terms of labor and marriage opportunity. This theory has been explored in multiple studies showing its empirical relevance (Caucutt et al., 2002; Rosenzweig, 1999; Greenwood et al., 2000). The theory of differences in contraception efficiency by education groups comes from Rosenzweig and Schultz (1989), which found that college educated individuals are more efficient at using contraception. This paper shows that the relationship between ability and education partially explains the relationship between ability and fertility timing. However, a relevant fraction remains to be explained, making future research relevant.

The economic and social consequences of unintended pregnancies are multiple. For example, women whose first child was unintended are more likely to be single mothers or enter in

¹<https://www.cdc.gov/reproductivehealth/contraception/unintendedpregnancy/index.htm>

shot-gun marriage with higher divorce probability (Kozlov, 2021). As a result, their children are more likely to grow up in single-parent houses performing worse in multiple educational and economic outcomes (Kearney and Levine, 2017). Moreover, mothers with early or unintended pregnancies also perform worse economically and socially (Amador, 2017; Foster et al., 2018; Levine and Painter, 2003). Additionally, out-of-wedlock pregnancies are more common in low income women decreasing intergenerational mobility.

The paper is related to several branches of economic literature. First, it relates to the literature on dynamic models with endogenous family formation and fertility choices. For example, Regalia et al. (2011) study how the change in relative wages affects the number of single mothers; Caucutt et al. (2002) study how child investment, marriage, and labor market outcomes shape women’s fertility timing; Choi (2017) shows that to simultaneously account for the observed heterogeneity in births and abortion across educational groups, differences in fertility risk by education are necessary; Filote et al. (2019) study how the welfare state affects teenage childbearing behavior through a parental investment model with risky sexual behavior; Kozlov (2021) study how unintended pregnancies affect marriage and divorce. Finally, Seshadri and Zhou (2022) study how heterogeneity in fertility planning affects children’s investment and shapes intergenerational mobility.

Second, it is related to the empirical literature on education and fertility. For example, Rosenzweig and Schultz (1989) shows that contraception method efficiency increases with education, and Musick et al. (2009) found that the gradient in fertility by education comes primarily from unintended childbearing, not differences in opportunity cost, which suggests a missing mechanism in the literature through which education affects fertility. Finally, it relates to the literature on the return on cognitive ability, noncognitive ability, and education on behavioral outcomes. For example, Heckman et al. (2006) found that cognitive and noncognitive abilities decrease the probability of teenage pregnancy; at the top of the cognitive and noncognitive ability distribution, the likelihood of unintended pregnancy is close to zero. Heckman et al. (2018) found that increasing cognitive endowments positively affects education, wages, smoking, and health outcomes. Meanwhile, increasing noncognitive endowments primarily affects smoking and health outcomes. Additionally, education affects the present value of health and smoking. Hai and Heckman (2022) found that an extra year

of schooling reduces smoking by 21%.

I provide two contributions. First, I document a positive relationship between cognitive ability, fertility timing, and intention, which decreases with age. Second, I build and estimate a dynamic choice model with endogenous education, marriage, and fertility, in which contraception efficiency depends on cognitive ability. Then, I estimate the model to quantify the importance of cognitive ability and education on women’s fertility timing and risk.

In order to study the relationship between ability, early and unintended pregnancies, I use the National Longitudinal Survey of Youth 1979 (NLSY79) as it measures cognitive ability and women’s fertility and labor outcomes until their late forties. As said before, women in the lowest ability quartile are nine times more likely to have their first pregnancy between 14 and 17 years old than women in the top quartile (29% vs. 3%). However, after 22 years old, this gap disappears.

Next, I estimate a static linear model with the ability as a latent variable. This method allows control by measurement error and simultaneous causation bias ([Hansen et al., 2004](#)). I found that with age, the importance of cognitive ability in the first pregnancy decreases and even changes sign. The magnitude of cognitive ability is quantitatively significant, with its interdecile range decreasing the probability of pregnancy by 98% between 14-17 years old and 65% between 18-22 years old. Interestingly, even when between 22-29 years old, women in the top ability decile are 27% more likely to have their first pregnancy than the ones in the bottom; they are 7% less likely to have an unintended pregnancy. This finding is consistent with research that argues that cognitive ability improves women’s fertility planning (see, [Commendador \(2007\)](#)).

Then, I build and estimate a dynamic life cycle model with endogenous education, fertility, and marriage, finding that differences in opportunity cost by ability and differences in contraception cost by education cannot fully account for the relationship between ability and fertility. Finally, I use the model to show that, consistent with empirical findings, improvements in contraception have a small effect on college attendance, as women who have teen pregnancies are too disadvantaged to attend college even without a child.

2 Empirical Evidence

In this section, I provide empirical evidence on the effect of cognitive ability on fertility using the National Longitudinal Survey of Youth (NLSY79) data. First, I describe the dataset. Second, I show descriptive statistics relating cognitive ability, early pregnancies, education, and marriage. Finally, I estimate a linear static model at different ages to measure the effect of cognitive ability on fertility across the lifecycle.

2.1 Data description

The National Longitudinal Survey of Youth 1979 follows a representative sample of American youth born between 1957-64 who were 14-22 years old when first interviewed in 1979. I use the observations from women whose first pregnancies end in a live birth. For this reason, in the paper, childbirth and pregnancy are synonymous. The survey has three essential characteristics for the paper: First, in 1980, participants' cognitive ability was measured through a cognitive test. Second, participants have been followed for more than 40 years, and women in the sample have already finished their reproductive life. Third, mothers in the survey were asked if the pregnancy was intended at the moment of conception.

In 1980 NLSY participants answered ten intelligence tests referred to as the Armed Services Vocational Aptitude Battery (ASVAB). Using a sub-sample of the ASVAB test, it is possible to construct an approximate score of a general cognitive ability test known as the Armed Forces Qualifications Test (AFQT). The test is the standard measure used in the literature to approximate cognitive ability. I use this measure to rank women in cognitive ability percentiles. After dropping the observations with missing ASVAB values, I have a panel of 5939 women.

Table 1. Observation Summary

Age First Pregnancy	Low Ability			High Ability			Total
	HSD	HS	Coll	HSD	HS	Coll	
14-17	233	425	19	13	126	22	838
18-21	178	818	52	12	434	54	1548
22-29	82	587	81	2	551	406	1709
No child at 30	67	528	93	11	544	601	1844
Total	560	2358	245	38	1655	1083	5939

I divide women’s fertility life into three periods. I focus on first birth before the age of 30, as 90% of firstborn children are born before this age. In the first period, women are in high school age (14 to 17 years old). In the second period, women are of college age (18-21 years old). Finally, in the third period, they are young adults between 22 and 29. I classify individuals as having high or low ability depending if they are above or below the sample mean. Table 1 shows the number of observations by age at first pregnancy, ability, and education. The relationship between ability and education outcomes is clear as almost no high ability women drop out high school and few low ability graduate from college.

2.2 Descriptive Statistics

In this subsection, I describe the data characteristic that I focus on. The purpose of the paper is to understand how cognitive ability affects fertility. Given that pregnancies affect the cost of education, labor supply, and marriage. I focus on the data characteristics that connect cognitive ability, pregnancy timing, education, wages, and marital outcomes.

2.2.1 Cognitive Ability and Age at First Childbirth

Table 2 displays the joint distribution between age groups, ability quartiles, and pregnancy outcomes conditional on not having childbirth in the previous age bin. For example, the interpretation of the first column, third row, is that half of the women in the first ability quartile who did not have childbirth before 22 years old had one between 22-29 years old.

Table 2. Joint Distribution First and Unwanted First Pregnancy by Age and Ability

	Ability Quartile			
	First Quartile	Second Quartile	Third Quartile	Four Quartile
Age First Pregnancy	Pregnancy Probability			
14-17	26%	15%	9%	3%
18-21	54%	43%	28%	16%
22-29	48%	48%	45%	45%

We observe a positive relationship between ability and age at first birth that disappears with age. Between 14 and 17 years old, 26% of women in the bottom ability quartile had their first pregnancy; meanwhile, only 3% in the top cognitive quartile. This difference means nine times more pregnancies for women in the lowest ability quartile compared to those in the top. As women grow, the ratio reduces to 3.37 between 18-21 and just 1.06 between 22-29 years old.

2.2.2 Pregnancy Timing and Education

Table 3 shows the joint distribution of women’s age at first pregnancy and the highest educational level achieved. First, we observe that most pregnancies occur at ages when women are expected to have completed their final education level. For example, 82% of high school graduates had their first child after high school age, and 73% of college graduates had their first pregnancy after college age.

Table 3. Conditional Distribution Age First Pregnancy by Education Outcomes

Age at First Pregnancy	Education Outcome		
	High School Dropout	High School Graduate	College Graduate
14-17	48%	18%	6%
18-21	40%	50%	21%
22-29	12%	32%	73%

Education outcomes and pregnancies cause mutually. On the one hand, children are costly in time and money. Pregnancies make women more likely to drop out or not pursue higher education degrees, causing a relationship between early pregnancies and low educational

outcomes. On the other hand, women with low cognitive ability attain lower education levels, which allows them to have children early, relating cognitive ability and pregnancy timing. For this reason, in the following section, I use a dynamic model to answer whether low ability women have low education achievement because they are young mothers or whether they have children young because it is the optimal time as pursuing higher education is too costly.

2.2.3 Early Pregnancies and Marriage

[Becker \(1991\)](#) argues that out-of-wedlock children decrease the likelihood of the mother’s future marriage by reducing her net resources and raising the cost of searching for a husband. [Bronars and Grogger \(1994\)](#) find that women with unplanned births were, on average, 13 percentage points less likely to be married while their children were young, the differential narrowed only to 9 percentage points by the time their children were preteenagers.

Model tractability does not allow me to develop a model with a frequency such that I can study how out-of-wedlock pregnancies affect marriage timing. So instead, I study the out-of-wedlock birth effects on the probability of marriage during the mothers’ lifetime and the husbands’ quality measured by their income. Women with out-of-wedlock birth have a similar probability of marriage during their lifetime than women without it. However, on average, their husbands have lower wages.

Table 4. Probability of ever marriage single mother vs non single mothers

	High School Dropout	High School Graduate	College Graduate
Age at Pregnancy	Single Mothers		
14-17	77%	86%	91%
18-21	73%	84%	88%
22-29	48%	73%	83%
	Non Single Mothers		
	82%	83%	83%

Table 4 compares the marriage probabilities of single mothers to non-single mothers. In the case of single mothers, I look at the marriage probability conditional on their age at first pregnancy. The marriage probability for single mothers increases with education and decreases with age at childbirth. For example, 86% of high school graduate single mothers

between 14 and 17 years old married after birth compared to 73% of single mothers between 22-29 years old with the same education level. Additionally, 73% of 18 to 21 years old high school dropout single mothers got married compared to 88% of college graduates. For non-single mothers, the fraction who married is around 83% for the three education groups. Out-of-wedlock pregnancies only decrease the probability of finding a husband for high school dropouts.

Table 5. Average Husband Wage by Education and Women Children at Marriage.

Age Fir. Preg.	High School Dropout		High School Graduate		College Graduate	
	Child	No Child	Child	No Child	Child	No Child
14-17	36873	34546				
18-21	37473	41224	44785	48923		
22-29	43760	30448	47420	58889	63248	81521

Table 5 compares husbands' wages for women with and without an out-of-wedlock birth after finishing their highest educational degree. The table shows the husband's average yearly wage by women's education, age at the first childbirth, and whether she had an out-of-wedlock child. High school and college graduates without an out-of-wedlock children have higher-income spouses. In the case of high school dropouts, husband income does not differ considerably. Surprisingly, high school dropout women with an out-of-wedlock child between 22-29 years old had, on average, a higher income husband. However, this group has few observations, and the sample size probably causes the result. In the case of high school graduates with their first pregnancy between 18 and 21 years old, the average penalty for an out-of-wedlock pregnancy is \$4000 yearly; meanwhile, for women that had their first child between 22 and 29 years old, the penalty is \$11000. For college graduates, the penalty increases to \$18000.

2.3 The effect of cognitive ability on fertility

In subsection 2.2.1, I discussed the differences in fertility timing by cognitive ability. Two important mechanisms that delay childbirth are education and wages. Both increase the opportunity cost of having children young for high ability individuals. Additionally, high cognitive ability is related to family background. For this reason, I estimate a static linear

model at three moments of women’s life cycle: teens (14-17 years old), college age (28-21 years old), and young adults (22-29 years old) to measure the direct effect of cognitive ability on fertility.

An additional concern is that the dataset provides an intelligence test. Measurement error and reverse causality generate bias as the test is affected by family background and formal education. As [Heckman et al. \(2006\)](#) argue *”we note that there is an important distinction between intelligence tests (i.e., IQ tests) and achievement tests. Although IQ is fairly well set by age 8, achievement tests have been demonstrated to be quite malleable”*. The simple least square model overpredicts the impact of ability tests and understates the contribution of formal education on different outcomes as it imputes the impact of family background and education on ability. I follow [Cawley et al. \(1996\)](#); [Heckman et al. \(2006\)](#); [Hansen et al. \(2004\)](#) to estimate ability as a latent variable correcting reverse causality bias and measurement error.

As agents are forward-looking, future education and income affect teenager decision. Because a static linear model does not capture the effect of future outcomes on current fertility decisions, I estimate the model at three points of time explained before. First, for teens, the effect of ability is upward bias, as is capturing future wages and marital opportunities. Some of the education outcomes are already realized in the case of women at college age, but still, future income affects the opportunity cost and biases the estimators. Finally, for young adults, education is completely realized, and we already observe an important fraction of the wage realizations allowing a better imputation of the effect of ability, education, and wage on fertility. Additionally, I estimate the effect of ability on unintended childbirth, which provides additional evidence on whether ability affects women’s fertility control.

The estimated linear model at each age is:

$$\begin{aligned} Y_i^{14-17} &= \beta_0 + \alpha_\theta \theta_i + \beta_X X_i + \epsilon_i \\ Y_i^{18-21} &= \beta_0 + \alpha_\theta \theta_i + \beta_{HS} HS_i + \beta_{CA} CA_i + \beta_X X_i + \epsilon_i \\ Y_i^{22-29} &= \beta_0 + \alpha_\theta \theta_i + \beta_w w_i + \beta_{HS} HS_i + \beta_C C_i + \beta_X X_i + \epsilon_i \end{aligned}$$

where Y_i^{age} is an outcome at a particular age (had a child and had an unintended child),

θ is a latent measure of cognitive ability, HS is whether graduate from high school, CA is a dummy for attending college, C is a dummy whether graduate from college, w is the mean wage during the period and X are demographic controls (race, both parent education, and broken home at 14). The measurement system used to identify the latent variable α is:

$$T_i = \beta_T X_{i,D} + \alpha_i + \epsilon_{i,T}$$

where T is a vector of cognitive ability measures (math reasoning, numerical operations, and coding speed), X_T is a vector of control for each measure (race, mother education, broken home at 14, and years of completed education at the moment of the test). Finally, the linear model is estimated, maximizing a likelihood function assuming that F_θ follows a mixture of normal distributions.

Table 6 shows the models' results. First, we can observe that cognitive ability affects pregnancy timing and intention even after controlling for education and wages. In the case of pregnancy, higher ability decrease the pregnancy probability for teens and women at college age. However, as high-ability women wait until they are young adults, ability increases pregnancy probability for this age group. Consistent with the hypothesis that high ability women are better at planning pregnancy, cognitive ability decreases the probability of unintended pregnancies for the three age groups; however, the effect is small in young adults.

Teenagers between 14-17 years old in the top cognitive decile are 98% less likely to have a pregnancy and 65% less likely to have an unintended pregnancy than those in the bottom decile. For 18-21 years old, the difference between the top and the bottom deciles is 28% for a pregnancy and 33% for an unintended pregnancy. Finally, women between 22-29 years old in the top decile are 27% more likely to have a pregnancy but 7% less likely to have an unintended pregnancy than their pairs in the bottom.

Education achievement and income are relevant determinants of fertility timing. Higher wages and college education decrease the probability of pregnancy and unintended pregnancy. For 18-21 years old, the interdecile range effect of ability is a decrease of 28% in the probability of pregnancy, and attending college reduces pregnancy probability by 80%. In the case of unintended pregnancies, the interdecile range effect of ability on pregnancy probability is

a reduction of 33%; meanwhile, college attendance reduces the pregnancy probability by 72%. Between 22-29 years old, a college education reduces the pregnancy probability by 17% and the unintended pregnancy probability by 14%. The interdecile range of wages is a 42% decrease in pregnancies and 90% in unintended pregnancies. High school graduation does not affect pregnancies or unintended pregnancies once controlled by ability. However, attending college reduces the probability of pregnancies and unintended pregnancies to almost zero. The fact that college is relevant even after controlling for wages supports the theory that college increases contraception efficiency.

Table 6. Latent Factor Model: Pregnancies and Unintended Pregnancies

	(1)	(2)	(3)	(4)	(5)	(6)
	Preg.	Unint. Preg.	Preg.	Unint. Preg.	Preg.	Unint. Preg.
	14-17 yrs old	14-17 yrs old	18-21 yrs old	18-21 yrs old	22-29 yrs old	22-29 yrs old
Cog. Ab.	-.1049*** (.0124)	-.0414*** (.0112)	-.0774*** (.0200)	-.0391** (.0163)	.1017*** (.0267)	-.0060 (.0215)
Wage					-.0826*** (.0117)	-.0430*** (.008)
HSG			.0435*** (.0167)	.003 (.0146)		
Att. Coll.			-.2706*** (.0175)	-.1038*** (.0152)		
College					-.0791*** (.0223)	-.0161 (.0170)
Change in Probability						
$\Delta(d_{10} - d_1)$ Cog. Ab.	-98%	-65%	-28%	-33%	27%	-7%
$\Delta(d_{10} - d_1)$ Wage					-42%	-90%
HS			13%	2%		
Att. Col.			-80%	-72%		
College					-17%	-14%
Controls	Yes	Yes	Yes	Yes	Yes	Yes
Observations	4390	3616	3801	3031	2412	1753

Standard errors in parentheses

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

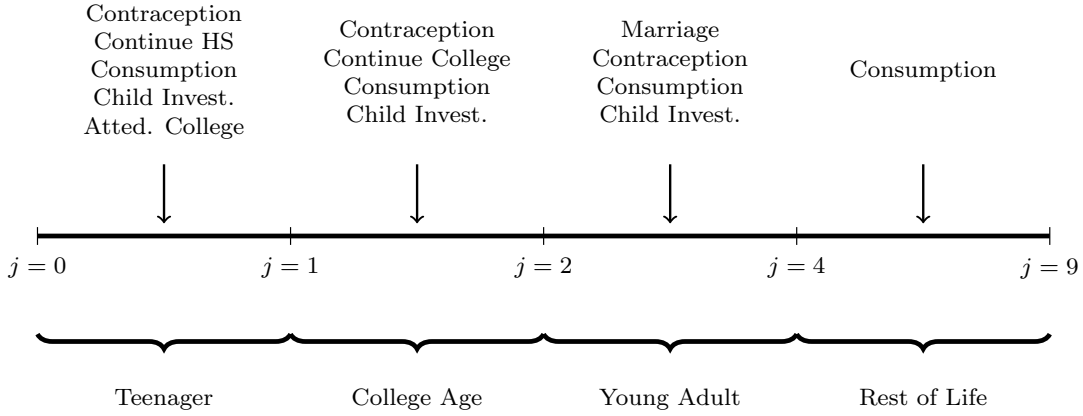
The previous results suggest that ability affects fertility beyond opportunity costs associated with education and wage. However, I can not control for the possibility that lower educational outcomes are caused by low ability women having more pregnancies which hinder their education outcomes, or that low ability women have children earlier because it is too costly to attend college. For this reason, in the following sections, I build and estimate a dynamic model to quantify the importance of cognitive ability and formal education in

fertility.

3 Model

In this section, I develop and estimate a life cycle model to measure whether cognitive ability affects fertility beyond education, wages, and marriage. Specifying a model allows dealing with the endogeneity caused by the fact that agents are forward looking, and their contraception choice depends on future opportunities related to their cognitive ability. Additionally, the model allows analyzing how fertility changes as future opportunities or contraception costs vary.

Figure 1. Women Attending College Life Cycle



The model has nine periods, each period represents four years of a woman's life. The timing of the model decision is depicted in figure 1. During the first four periods, women are fertile, deciding contraception. Then, in the period that a child is born, they decide on the children's investment. In the first period, they are teenagers (14-17 years old); in the second, they are at college age (18-21 years old); then, individuals are young adults for two periods (22-29 years old), and finally, they consume until their death (30-49 years old). Women differ by ability, marital status, motherhood, wage, and education. During each fertile period, women make different decisions: graduate from high school, attend and graduate from college, marry in the case of being single, use contraception in the case of not having a previous child, and invest in the child in the case of having one. Within each period, decisions are sequential, allowing us to divide them into sub-periods. In order to keep

tractability, there is no saving.

In the model, each woman can only have one child who lives only during the period that it was born. During that period, mothers can invest money to increase the child's human capital. Child investment is a one-in-a-lifetime investment with a positive return. The optimal age for having a child is when parents have higher resources, which is increasing in age. However, waiting is costly, given the contraception cost. Mothers are altruistic to their children, and the utility received is increasing in children's monetary investment. Child investment is a reduced form to create incentives for high income people to postpone parenthood and is also consistent with the fact that child investment is increasing in income. However, modeling child investment realistically is beyond the scope of the paper. Each period single women meet a potential husband, and they can marry. There is no divorce.

A fundamental dimension in the model is contraception. By contraception cost, I consider all monetary, physical, psychological, and social costs associated with avoiding pregnancy and not a particular contraception method. Contraception, apart from being costly, is imperfect; women choose the optimal amount, but pregnancy can occur even when having a child is not optimal.

The variables of the model and their notation are the following. The value of the problem at each moment is V_t^s , where t represents the period, and s is the sub-period. The state variables are ability, education, marital status, marital status at childbirth, and the period that a child was born. Ability a is continuous. Education has three levels $e \in \{HSD, HS, C\}$, where *HSD* is high school drop-out, *HS* is high school graduate, and *C* is college graduate. Marital status is denoted by $m \in \{0, 1\}$ where 0 is single, and 1 is married, $sm \in \{0, 1\}$ represents whether the woman was a single mother or not. $k_t \in \{0, 1, 2, 3, 4\}$ keeps track of the child's birth period, and it is zero whether she does not have a child. The decision variables are consumption c , child monetary investment i , and the contraception amount s . In the rest of this section, I introduce the model backward to simplify the exposition.

3.1 Fifth to Nine Period (30-49 years old)

The model focuses on ages 14 to 29 as around 90% of first childbirth and education decisions occur at this age interval. However, a relevant fraction of college wages return is backloaded

after 30 years old, so I include the ages between 30-49 years old to avoid underestimating the value of college. During this period, agents consume their household income and do not make any choices. I only model until 49 years old, given data limitations that I will explain in the estimation section. The women's problem in these periods is:

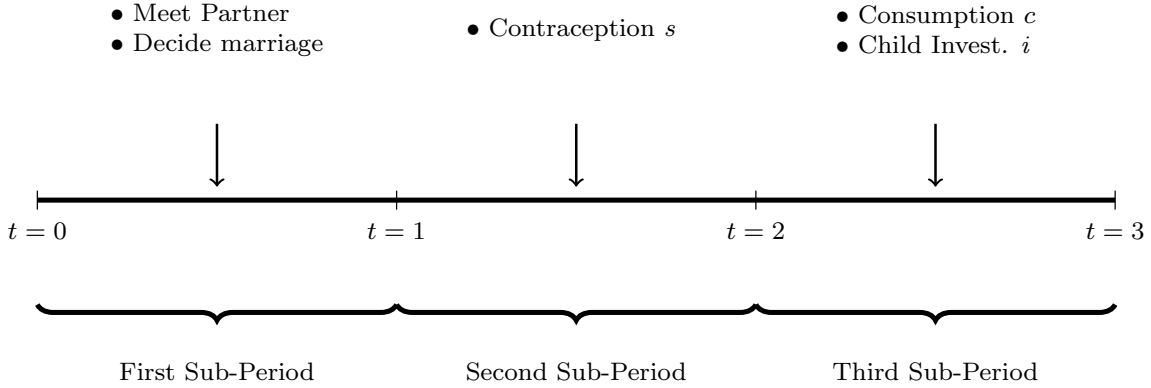
$$V_t(a, e, m, sm, k_t) = \max_c u(c) + \beta V_{t+1}(a, e, m, sm, k_t)$$

$$\phi_c(m, k_t) \cdot c = w(t, a, e, m, sm, k_t) + 1_{m=1} w^h(t, e, sm, k_t)$$

where $\phi_c(m, k_t)$ maps consumption expenditures and family composition into effective consumption. Women's wage w is a function of their age, ability, education, marital status, single motherhood, having a child, and maternal age at birth. Husband wage w^h is a function that depends on age, the woman's education, whether she was a single mother, and in the case of being a single mother, at which age she was. The estimation of the wage process is discussed in section 4.2.

3.2 Third and Fourth Period (22-29 years old)

Figure 2. No child, Single Women between 22-30 yr. old



In this period, women make three sequential decisions. Each decision is made during a sub-period. The decision timing within each period is displayed in Figure 2. In the first sub-period, single women meet a husband and decide whether they marry. In the second, women without previous children decide on contraception. Finally, in the last one, women

consume and invest in the children whether a child is born.

3.2.1 Third Sub-Period: Consumption and Child Investment

The women's problem in this sub-period depended on whether they had a child. Whether a woman has a child within the period, she has to decide on child investment and consumption, solving the following problem:

$$V_t^3(a, e, m, sm, k_t = t) = \max_{c, i} u(c) + V(i)^k + \beta V_{t+1}(a, e, m, sm, k_t = t)$$

$$\phi_c(m, k_t) \cdot c = w(t, a, e, m, sm, k_t) + 1_{m=1} w^h(t, e, sm, k_t) - i$$

where $V^k(i)$ is the utility the child provides to the mother, which depends on monetary investment i .

In the case of a woman with a child during a previous period, she only has to decide on consumption; then her problem is the following:

$$V_t^3(a, e, m, sm, k_t) = \max_c u(c) + \beta V_{t+1}(a, e, m, sm, k_t)$$

$$\phi_c(m, k_t) \cdot c = w(t, a, e, m, sm, k_t) + 1_{m=1} w^h(t, e, sm, k_t)$$

Finally, a woman without childbirth during her fertile life receives a utility $\mu_{nk}(e)$. This utility is necessary to prevent all women from having a child in the last fertile period, giving a chance to the model to replicate that in the data, 10% of women do not have a child at 30 years old. Then, this group solves the following problem:

$$V_4^3(a, e, m, sm, k_t = 0) = \max_c u(c) + \mu_{nk}(e) + \beta V_5(a, e, m, sm, k_t = 0)$$

$$\phi_c(m, k_t) \cdot c = w(4, a, e, m, sm, k_t) + 1_{m=1} w^h(4, e, sm, k_t)$$

3.2.2 Second Sub-Period: Contraception

If a woman has not had a child at the beginning of the period, she decides the amount of contraception s , solving the following problem:

$$V_t^2(a, e, m, sm, k_t) = \max_s p(t, a, e, s) \cdot V_t^3(a, e, m, sm, k_t = t) \\ + (1 - p(t, a, e, s)) \cdot V_t^3(a, e, m, sm, k_t = 0) - \phi_s s$$

where contraception has a utility cost of ϕ_s and $p(t, a, e, s)$ is the probability of having a pregnancy. The pregnancy probability is decreasing in the amount of contraception $\frac{\partial p(t, a, e, s)}{\partial s} < 0$, and at the same amount of contraception, higher ability implies lower pregnancy risk $\frac{\partial p(t, a, e, s)}{\partial a} < 0$.

In the model, women only can have one child. In the case of already had a child, they do not choose contraception, and the value at this sub-period is:

$$V_t^2(a, e, m, sm, k_t) = V_t^3(a, e, m, sm, k_t)$$

3.2.3 First Sub-Period: Marriage

Single women meet a potential husband with probability $\mu(e, t)$, which depends on women's education achievement e and they marry if the utility of marriage is higher than remaining single. Then the problem during this stage is the following:

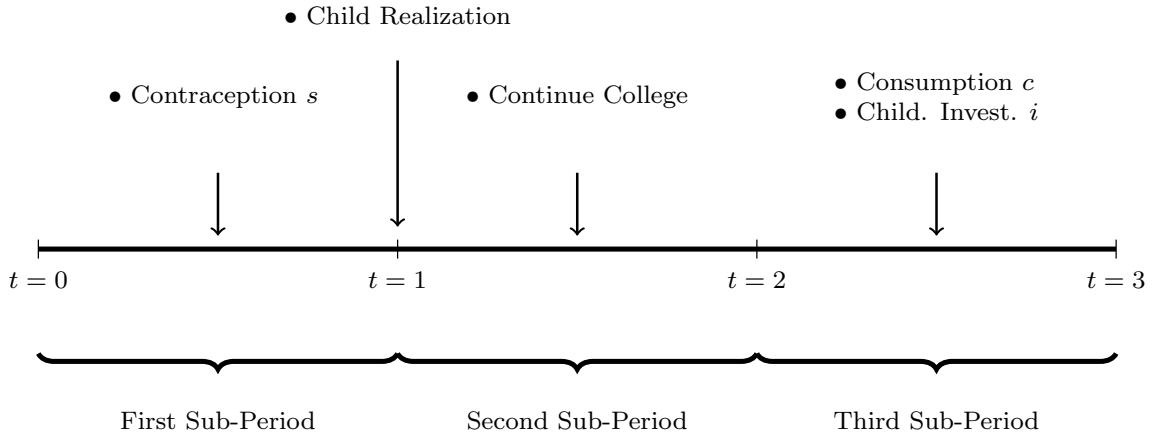
$$V_t^1(a, e, m = 0, sm, k_t) = \\ \mu(e, t) \cdot \max\{V_t^2(a, coll, m = 1, sm, k_t), V_t^2(a, e, m = 0, sm, k_t)\} \\ + (1 - \mu(e, t)) \cdot V_t^2(a, e, m = 0, sm, k_t)$$

The model does not have divorces, so married women do not make marital decisions. Then, their value function is given by:

$$V_t^1(a, e, m = 1, sm, k_t) = V_t^2(a, e, m = 1, sm, k_t)$$

3.3 Second Period (18-21 years old)

Figure 3. Attending College (18-21 years old)



In this period, women attended college or joined the labor market. Women who do not attend college have the same decision sequence as in subsection 3.2. Women attending college make three sequential decisions. First, they decide on contraception. Second, after observing their pregnancy outcome, they decide whether continue college. In the case of having a child during college, there is a double penalty. First, the cost of continuing college increases in $\kappa_{k,C}$; second, college students have few monetary resources to invest in their child. Women who continue in college receive an allowance w_c . College dropouts move to the consumption-investment sub-period with the wage of high school graduates. Finally, women decide on consumption and child investment.

3.3.1 Third Sub-Period 18-21: Consumption and Child Investment

In this subperiod, women are attending college or, in the case of dropping out in the previous subperiod, working as high school graduates. Women attending college decide on consumption and child investment in the case of having a child. Then their problem is:

$$V_2^{3,G}(a, k_t) = \max_{c,i} u(c) - \kappa_{k,C} + V^k(i) + \beta V_3^1(a, e = C, m, sm, k_t)$$

$$\phi_c(k) \cdot c = w(a, e = \text{attend-coll}, s, k_t) - i$$

where G denote the value function of continuing college and $\kappa_{k,C}$ is the cost of attending college with a child.

For women who dropped out of college, their problem is denoted by CD , and they face the consumption-investment problem of a high school graduate. So then, their value function is:

$$V_2^{3,CD}(a_t) = V_2^3(a, e = \text{HS}, m = 0, sm, k_t)$$

3.3.2 Second Sub-Period 18-21: Continue College

In this sub-period, women decide whether continue college. The decision is made after they observe their pregnancy outcomes. Then the problem is:

$$V_2^2(a, coll_a = 1) = \max_{i \in \{G, CD\}} \{V_2^{3,G} + \sigma_{Coll}\epsilon_i, V_2^{3,CD} + \sigma_{Coll}\epsilon_i\}$$

where ϵ_i is an i.i.d. extreme value shock drew from a Type I extreme value distribution, and σ_{Coll} is the scale parameter.

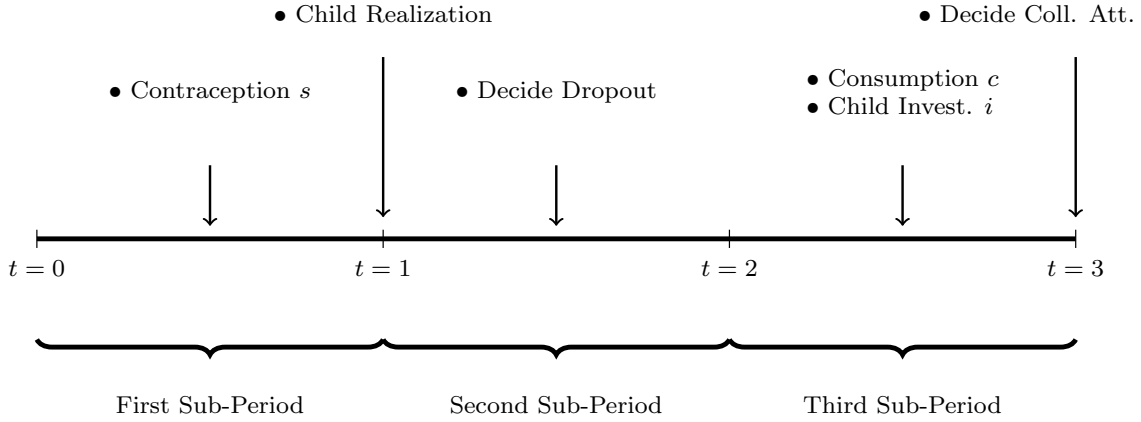
3.3.3 First Sub-Period 18-21: Contraception

In the first sub-period, college women choose the amount of contraception as in previous periods, solving the following problem:

$$V_2^1(a, coll_a = 1, k_t) = \max_s p(t, a, e = coll_a, s) \cdot V_2^1(a, coll_a = 1, k_t) \\ + (1 - p(t, a, e = coll_a, s)) \cdot V_2^1(a, coll_a = 1, k_t) - \phi(s)s$$

3.4 First Period (14-17 years)

Figure 4. Teens (14-17yr. old)



Women between 14-17 years old face the following decisions. At the beginning of the period, all attend high school and decide on contraception. After observing the fertility outcome in the second sub-period, they decide whether to continue high school. A teenager who continues high school receives an allowance in the third sub-period. Teens who drop out of high school work as high school dropouts. Finally, women who continue high school decide whether to attend college at the end of the third period.

3.4.1 Third Sub-Period 14-17: Consumption, Child Investment, and College Attendance

At the end of the period, women who graduated from high school decide whether attend college. The decision problem is:

$$V_1^{CD}(a, e = \text{hs}, sm, k_t) = \max_{i \in \{C, NC\}} \{V_2^1(a, coll_a = 1, k_t) - \kappa_c(a, k_t) + \sigma_{CD}\epsilon_i, V_2^1(a, e = HS, m = 0, sm, k_t) + \sigma_{CD}\epsilon_i\}$$

where ϵ_i is an i.i.d. extreme value shock drawn from a Type I extreme value distribution, and σ_{CD} is the scale parameter. The psych cost $\kappa_c(a, k_t)$ of attending college depends on the woman's ability and whether she has a child.

In the case of women attending high school without a child, they solve the following decision problem:

$$\begin{aligned} V_1^{3, HSG}(a, HS_a = 1, sm = 1, k_t = 1) &= \max_{c, i} u(c) + \\ &\kappa_{k, HS} + V^k(i) + \beta V_1^{CD}(a, e = \text{HS}, sm, k_t) \\ &\phi_c(k) \cdot c = w_{HS} - i \end{aligned}$$

where κ_{HS} is the utility cost of attending high school with a child, w_{HS} is the parents' allowance, and V_1^{CD} is the value of deciding college attendance.

The decision problem of women attending high school without a child is:

$$\begin{aligned} V_1^{3, HSG}(a, HS_a, k = 0, sm = 0, k_t) &= \max_c u(c) + \beta V_1^{CD}(a, e = \text{HS}, sm, k_t) \\ c &= w_{HS} \end{aligned}$$

Women who drop out of high school earn the wage of a high school dropout. In the case of having a child, they also decide on child investment. So then, their problem is:

$$\begin{aligned}
V_1^{3,HSD}(a, coll_a = 1, sm = 1, k_t = 1) &= \max_{c,i} u(c) + \kappa_{k,HS} + V^k(i) \\
&+ \beta V_2^1(a, e = HSD, m = 0, sm = 1, k_t = 1) \\
\phi_c(k) \cdot c &= w(t, a, HSD, k_t = 1) - i
\end{aligned}$$

where $w(t, a, HSD, k_t = 1)$ is the wage of a high school dropout. Women who drop out without a child solve:

$$\begin{aligned}
V_1^{3,HSD}(a, HS_a, sm = 0, k_t = 0) &= \max_c u(c) \\
&+ \beta V_2^1(a, e = HSD, m = 0, sm = 0, k_t = 0) \\
c &= w(t, a, HSD, k_t = 0)
\end{aligned}$$

3.4.2 Second Sub-Period 14-17: Continue high school

After observing their pregnancy outcome, women decide whether continue high school. Then, the problem is the following:

$$V_1^2(a, sm, k_t) = \max_{i \in HSD, HSD} \{V_1^{3,HSG} + \sigma_{HS}\epsilon_i, V_1^{3,HSD} + \sigma_{HS}\epsilon_i\}$$

where ϵ_i is an i.i.d. extreme value shock drew from a Type I extreme value distribution, and σ_{HS} is the scale parameter.

3.4.3 First Sub-Period 14-17: Contraception

In the first sub-period, all women attend high school; the only heterogeneity is the amount of ability. The decision at this stage is the amount of contraception, which is decided in the following problem:

$$V_1^1(a) = \max_s p(1, a, s) \cdot V_2^1(a, sm, k_t = 1) + (1 - p(1, a, s)) \cdot V_2^1(a, sm, k_t = 0) - \phi(s)s$$

3.5 Functional Forms and Parameters

The model is parametrized with the following functional forms to take it to the data. There are 34 unknown parameters plus the income process, which I will discuss their estimation in section 4.

3.5.1 Preferences (4)

Consumption:

Agents have standard CRRA preferences:

$$u(c) = \frac{c^{1-\gamma}}{1-\gamma} \tag{1}$$

Child Utility:

The utility that parent perceive from a child is given by:

$$V(i) = \omega_0 + \omega_1 i^{\omega_2} \tag{2}$$

where ω_0 , ω_1 , and ω_2 determine the value of having a child.

3.5.2 Fertility (9)

In the case of the pregnancy probability, I follow [Seshadri and Zhou \(2022\)](#); [Choi \(2017\)](#), who used a modified logistic function to bound the probability to one. In their papers, the only input is contraception. In the case of this paper, the pregnancy probability depends on two inputs: contraception and ability. Given its flexibility to aggregate both inputs, I use a CES production function. Finally, the probability depends on age t , education e , and

contraception s :

$$p(t, a, e, s) = \frac{2 \exp(-\lambda_{e,t}(\lambda_a \cdot a^\rho + (1 - \lambda_a) \cdot s^\rho)^{\frac{1}{\rho}})}{1 + \exp(-\lambda_{e,t}(\lambda_a \cdot a^\rho + (1 - \lambda_a) \cdot s^\rho)^{\frac{1}{\rho}})} \quad (3)$$

The parameter that determinate the fertility decision are $\lambda_{e,t}$, λ_a , ρ and $\phi(s)$. Given that $\phi(s)$ and $\lambda_{e,t}$ determinate the contraception cost, I normalize $\phi(s)$ to 1. λ_a determinate the share of ability in the contraception function and ρ is the contraception technology elasticity of substitution.

3.5.3 Marriage Market (7)

The marriage outcomes depend on the probability of meeting a potential husband $\mu(e, t)$, which depends on education and age.

3.5.4 College Attendance and Graduation (6)

College attendance is determined by the psych cost of attending college $\kappa_c(a, k_t) = \frac{\xi_c}{a^{\omega_c}} + 1_{k_t=1} \cdot \kappa_{kb}$. ξ_c is the psych cost of attending college, ω_c gives the concavity at which the psych cost decrease with ability, and κ_{kb} is the extra cost of study for teen mothers. The psych cost decrease with cognitive ability as studying is less painful for high ability kids, or at the same amount of study time; they learn more. Women who continue college having a child during that period face the extra cost $\kappa_{k,C}$. Finally, the extreme value shock scale parameters for attending and continuing college are given by σ_{Coll} and σ_{CD} .

3.5.5 High School Graduation (2)

High school graduation depends on the continuation values, the cost of attending high school with a child $\kappa_{k,HS}$, and the extreme value shock scale parameter σ_{HS} .

4 Model Estimation

The parameters in the model are divided into three sets: those that I take from the literature, the wage process that is estimated exogenously to the model, and those that are estimated

within the model using the simulated method of moments, conditional on the two previous sets.

4.1 Exogenous Parameter

Table 7 shows the exogenous parameters taken from the literature: discount factor, relative risk aversion, and the economies of scale in consumption.

Table 7. Exogenous Parameters

Parameter	Interpretation	Value	Source
Preferences			
β	Discount Factor	.96	Regalia et al. (2011)
γ	Relative Risk Aversion	.43	Regalia et al. (2011)
ϕ_s	Contraception Cost	1.00	Normalization
Household Parameters			
$\phi_c(m, k)$	Economies of Scale in Consump.	{.5, .7}	OECD

4.2 Wages Profiles

Women’s and husbands’ wage profiles are essential in the model because they determine the opportunity cost of early and out-of-wedlock pregnancies. Women’s wages vary by age, education, marital status, age at childbirth, and ability. Husbands’ wage depends on the wife’s education, the wife’s marital status at childbirth, and the wife’s age at birth. Because NLSY79 individuals are currently in their early fifties, wage profiles are estimated using OLS for individuals between 14 to 49 years old. The decision to not include wages after their fifties can understate the value of college or the child penalty. However, wages later in life are discounted heavily when fertility decisions are taken in the model. Wages are reported by NLSY79 surveyed individuals and are deflated to 2016 prices. Individuals who make less than 2.5 dollars per hour or work less than 2000 hours a year are not considered. Finally, the estimated equations for women’s wages and husbands’ wages are:

$$\begin{aligned}
w_{i,t} &= \alpha_0 + \alpha_t \cdot t + \alpha_{t^2} \cdot t^2 + \alpha_e e_i + \alpha_{HA} HA_i + \alpha_k k_i + \alpha_{et} \cdot e_i \cdot t + \\
&\alpha_{et^2} \cdot e_i \cdot t^2 + \alpha_{eHA} \cdot e_i \cdot HA_i + \alpha_{kHA} \cdot k_i \cdot HA_i + \alpha_{eHA} \cdot e_i \cdot HA_i + \\
&\alpha_{ek} \cdot e_i \cdot k_i + \alpha_{ekHA} \cdot e_i \cdot k_i \cdot HA_i + \alpha_m \cdot m_{i,t} + \epsilon_{i,t} \\
\\
w_{i,t}^h &= \gamma_0 + \gamma_t \cdot t + \gamma_{t^2} \cdot t^2 + \gamma_m \cdot m_i + \gamma_{e^f t} \cdot e_i^f \cdot t + \\
&\gamma_{e^f t^2} \cdot e_i^f \cdot t^2 + \gamma_{e^f m} \cdot e_i^f m_i + \epsilon_{i,t}
\end{aligned}$$

where w is the woman's wage, w^h is the husband's wage, t is the women's age, i is the individual, e and e^f is the women education level (HSD, HS, C), HA is a dummy whether the individual test above the mean cognitive ability test score, k is a dummy that indicates the woman age at the first childbirth, m is a dummy whether the woman is married, and m^f is a dummy whether the woman is married at the moment of the childbirth.

Early pregnancies decrease women's wages for all education groups and the penalty increases in ability. Women who married after an out-of-wedlock pregnancy had husbands with lower wages. The size of the husband's wage penalty increases with the women's age at childbirth, and college educates bear the highest penalty. As a result, the cost of an early pregnancy increases in ability, as high ability women are more likely to attend college and have a college graduate husband.

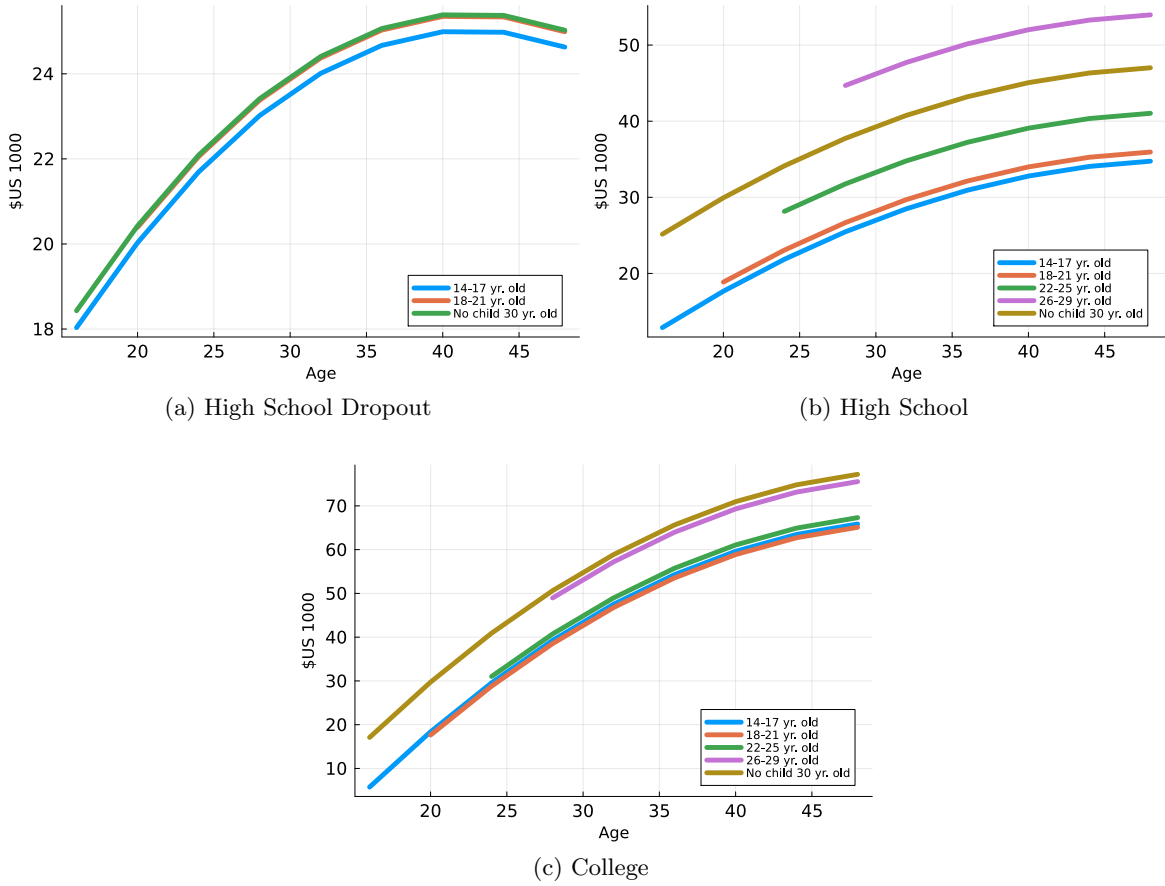


Figure 5. Women wage profile by education and age first birth

Women's wages profile by education and age at first childbirth are shown in figure 5. As expected, college women have the highest earnings, followed by high school graduates, and finally, high school dropouts have the lowest wages. High school dropout profiles are shown in figure 5a. Early childbirth has a negligible effect on this group's earnings; a birth between 14-17 years old only reduces wages by \$400. For high school dropouts, there are not enough observations to calculate the wage profiles for women with first childbirth between 22-29 years old, as for them, having the first child at these ages is infrequent. In the model, I assume they have the same profiles as high school dropout women without kids before 30. High school graduates' wage profiles are displayed in figure 5b. Women who have children between 26-29 years old have the highest wages, and early children bring a considerable wage penalty. For example, women with a child between 14-17 years old each year earn \$12200 less

than women who had the first child between 26-29 years old. Finally, college women's wage profiles are in figure 5c. We can see two wage paths, women who had a child before 26 and women who had a child after 26 or did not have one before 30. The first group earns around \$12000 less than the second. The child penalty for college and high school graduates is similar in monetary terms. However, the relative magnitude is higher for high school graduates as they have lower wages.

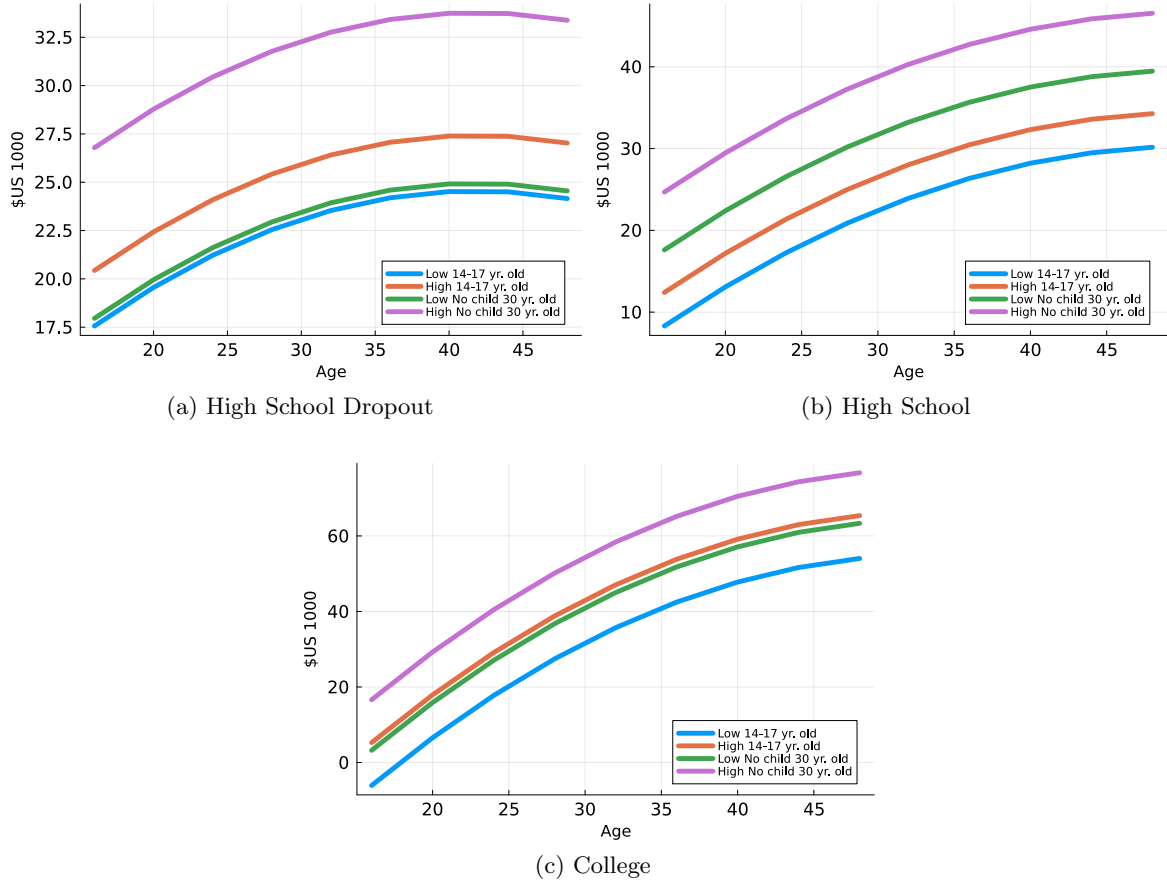


Figure 6. Women wage profile by education and ability

Figure 6 shows the relationship between ability and wage by women's education. The effect of ability on wage is increasing with education. For simplicity, I plot four wage profiles: high and low ability women with childbirth between 14-17 years old and high and low ability women without a child at 30. Cognitive ability increases wages in every education group. Figure 6a shows the case of high school dropouts. High ability women with childbirth between

14-17 years old earn \$3000 more yearly than low ability. In the case of high ability women without a child at 30, they earn \$9000 more yearly than low ability. Next, women high school graduates are depicted in figure 6b. The difference between high and low ability is \$4000 yearly for women with childbirth between 14-17 years old and \$7000 yearly for women without childbirth at 30 years old. Finally, college graduates are in figure 7, high ability women with childbirth between 14-17 make \$11000 more yearly than low ability, and this difference is \$13000 for women without a child at 30 years old. The penalty wage associated with early pregnancy is increasing in ability. In the case of low ability high school dropouts, are \$400 yearly and \$6300 yearly in the case of high ability. Low ability high school graduates face an early pregnancy penalty of \$6300 yearly and \$12200 yearly for the high ability. Finally, for college graduates, the penalty for low ability is \$9300 yearly and \$11340 yearly for high ability.

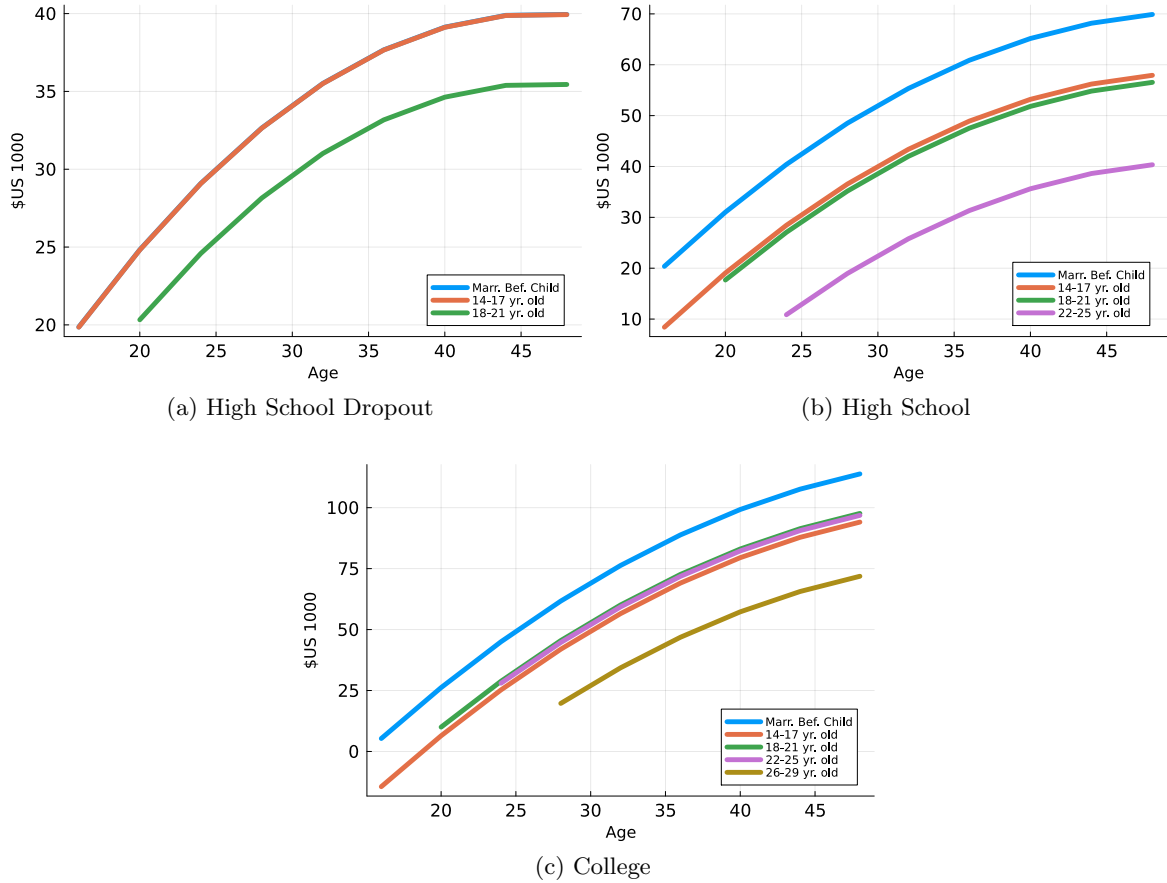


Figure 7. Husband wage profile by marital status at birth and age first birth

Finally, I analyze husbands' wages. Early pregnancies are associated with out-of-wedlock births. Single motherhood affects the husband's quality measured by his wage. Figure 7 shows husbands' wages by education, whether they marry before childbirth, and in the case of an out-of-wedlock pregnancy, the mother's age at birth. Single mothers have husbands with lower wages in all education groups. The penalty is increasing in age and education; the highest penalty is for college women who had an out-of-wedlock pregnancy between 26-29 years old. Figure 7a shows high school dropouts. For this group, an out-of-wedlock child between 18-21 years old is associated with a husband that earns \$4000 less than the husbands of women who married before having the child. High school graduates are shown in figure 7b. For them, an out-of-wedlock child reduces husbands' wages by \$12000 between 14-21 years old and by \$30000 between 23-26 years. Finally, figure 7b shows that college graduates with an out-of-wedlock child between 14-25 years old earn \$25000 less and between 26-29 years old, the penalty increases to \$42000 less.

4.3 Endogenous Parameters

The remainder 30 unknown parameters are estimated through 38 data moments using the Simulated Method of Moments. Table 8 shows the parameters' respective interpretations. The model aims to understand how ability, education, wage, and marriage shape fertility. For this reason, I estimate the model using moments that relate cognitive ability, education, and marital outcomes with fertility.

Table 8. Parameters estimated through the indirect method of moments

Contraception Technology and Child Return		
Parameter	Value	Interpretation
λ_a (1)	.08	Ability Share Contr. Tech.
ρ (1)	-.40	Elasticity of Substitution Contr. Tech.
$\lambda_{e,t}$ (7)	Table A1	Scale Parameter by Ability and Education
$\omega_0 \omega_1 \omega_2$ (3)	2.39, 0.09, .55	Child Return
Marriage Outcomes		
Parameter	Value	Interpretation
$\mu(e, t)$ (5)	Table A2	Probability Meeting a Husband
Education Outcomes		
Parameter	Value	Interpretation
$\kappa_{Coll.} \xi_c \omega_c$ (3)	-27, -201, 1.19	College Psych Cost
κ_{HS}, κ_{kb} (2)	-2, -16	Cost Child Dif. Edu. Stages
$\mu_{nk}(e)$ (3)	34, 14, 1	Terminal Value no Child
$w(HS), w(Coll)$ (2)	US\$48600, US\$10000	Teen and Coll. Allowance
$\sigma_{HS}, \sigma_{CD}, \sigma_C$ (3)	50, 68, 13	EV Shock HS, Attend Coll., Grad Coll.

Table 8 shows the estimated parameters. The share of ability in contraception is 0.08; even when it is small, it is relevant to explain the differences in pregnancies by ability, as discussed in section 5.2. The elasticity of substitution between ability and contraception is -0.4, making ability and contraception complements. The cost of teen pregnancy during high school is low, which indicates that in the model, the reason women with pregnancy are dropping high school is not the child. The estimated parent allowance during high school is \$48600, and the allowance during college is \$10000. In section 6, I discuss how contraception costs vary with education and ability.

5 Results

In this section, I discuss the estimation results. First, I discuss how the model fits the data. Second, I argue that ability is a necessary input in the fertility function, and it is not possible to explain the relationship between fertility and ability only through education and wages. Finally, I quantify the difference in contraception cost by education and ability in terms of equivalent consumption.

5.1 Model Fit

This subsection compares the model moments with the empirical counterpart used to estimate the parameters. In general, the model does a good job considering its simplicity and the fact that it uses 38 moments for 30 parameters. In order to simplify the exposition, I group the moments into three sets: moments that connect ability with pregnancies, moments related to education outcomes, and moments related to marital outcomes.

5.1.1 Cognitive Ability and Pregnancies

Table 9. Fit Model Pregnancy and Ability

Age First Pregnancy	Ability Quartile			
	First Quartile	Second Quartile	Third Quartile	Four Quartile
	Pregnancy Probability No Prev. Preg.			
14-17	30%	7%	5%	4%
	(26%)	(15%)	(9%)	(3%)
18-21	59%	32%	23%	21%
	(54%)	(43%)	(28%)	(16%)
22-29	64%	51%	33%	29%
	(48%)	(48%)	(45%)	(45%)

Values in parenthesis are model outcomes, and values without parenthesis are the data values.

Table 9 shows the moments that relate cognitive ability with pregnancy timing. These moments are the fraction of women who had their first child between 14-17, 18-21, and 22-29 years old by cognitive ability quartiles, conditionally on being childless at the beginning of each age interval. Between 14-17 and 18-21 years old, the fraction of pregnancies for the bottom and top quartiles generated by the model is similar to the data. The model can replicate the positive correlation between ability and pregnancy. However, this relationship is considerably more concave than in the data. In the case of women between 22-29 years old, the model fails as it generates a positive relationship between ability and pregnancies considerably higher than in the data.

5.1.2 Education Outcomes

Table 10. Education Outcomes and Pregnancy Moments

Moments	Data	Model
Drop out High School No Pregnancy (<18)	0.06	0.06
Drop out High School Pregnancy (<18)	0.28	0.27
Attend College No Pregnancy (<18)	0.41	0.38
Attend College Pregnancy (<18)	0.08	0.08
Attend College Ability Q1	0.12	0.10
Attend College Ability Q2	0.25	0.32
Attend College Ability Q3	0.40	0.46
Attend College Ability Q4	0.67	0.50
Graduate College No Pregnancy (<22)	0.62	0.51
Graduate College Pregnancy (<22)	0.30	0.38

The model replicates the relationship between cognitive ability, pregnancies, and education attainment. Table 10 shows that it generates the same number of high school dropouts and college attendants. Even when the model generates a smaller college attendance for the top ability quartile than in the data, it generates the right college attendance for the lowest ability quartile, which is the one that we are particularly interested in as a result of the high number of teen pregnancies within this group. Although the relationship between ability and college attendance is smaller at the top of the ability distribution than in the data, the fit is pretty good considering the lack of financial friction to attend college and the relationship between ability and parent income, which is not modeled. Finally, the model generates a gap in college graduation between women with childbirth during college compared to those without; however, it is smaller than in the data.

5.1.3 Marital Outcomes

Table 11. Fit Marital Outcomes and Pregnancies by Education Group

Moments	Data	Model	Moments	Data	Model
Married 18-21 yr. old HSD, NPP	0.39	0.34	Sing. Mom 22-29 yr. old HS, NPP	0.12	0.10
Married 18-21 yr. old HS, NPP	0.41	0.34	Sing. Mom 22-29 yr. old Coll, NPP	0.02	0.02
Married 22-29 yr. old HSD, NPP	0.61	0.50	Preg. 18-21 yr. old HSD, NPP	0.42	0.47
Married 22-29 yr. old HS, NPP	0.66	0.68	Preg. 18-21 yr. old HS, NPP	0.35	0.40
Married 22-29 yr. old Coll, NPP	0.68	0.64	Preg. 18-21 yr. old Att. Coll, NPP	0.16	0.17
Sing. Mom 18-21 yr. old HSD, NPP	0.44	0.36	Preg. 22-29 yr. old HSD, NPP	0.43	0.48
Sing. Mom 18-21 yr. old HS, NPP	0.29	0.24	Preg. 22-29 yr. old HS, NPP	0.43	0.37
Sing. Mom 22-29 yr. old HSD, NPP	0.22	0.18	Preg. 22-29 yr. old Coll, NPP	0.42	0.49

NPP = No Previous Pregnancy.

Finally, we have the moments related to contraception and marriage choices conditional on education and age. These moments are total pregnancies, single motherhood, and fraction married. Table 11 shows that the model is close to matching marriage, single mothers, and total pregnancies. However, the number of high school graduates' total pregnancies is considerably overpredicted.

5.2 Mechanism Decomposition and the Importance of Ability

The paper argues that cognitive ability affects pregnancy timing beyond education, wages, and marriage opportunities. In this section, I analyze how the model performs when contraception is independent of cognitive ability, and cognitive ability only affects fertility decisions through education and wage opportunity cost. First, I estimate the model with homogeneous contraception technology in education and ability. In this case, the only heterogeneity in contraception is by age $\lambda_{1t}^{HSD} = \lambda_{1t}^{HS} = \lambda_{1t}^{Coll}, \lambda^a = 0, \rho = 0$. I call this the “baseline model”. Then, I make contraception depend on education. Finally, contraception depends on education and ability $\lambda^a \geq 0, \rho \neq 0$. Then, each model is estimated separately, and the fit with the data is assessed using the sum squared error (SSE) normalized by the data moment:

$$SSE = \sum \left(\frac{m_i - m_i(\hat{\theta})}{m_i} \right)^2$$

where m_i is the data moments and $m_i(\hat{\theta})$ is the model moments for the parametrization

θ . As in the previous section, I group the moments into three categories: Cognitive Ability and Pregnancies, Education, and Marital Moments. The top of table 12 shows the total SSE for each different version of the model, and the bottom displays the percentual improvement in the model fit when each mechanism is added ($1 - \frac{SSE_1}{SSE_0}$). The full set of moments for each model specification is shown in table C1. Figure 8 shows the relationship between pregnancies and ability for women between 14-17 and 18-21 years old for the three model specification and the data.

Table 12. Decomposing the Model Fit

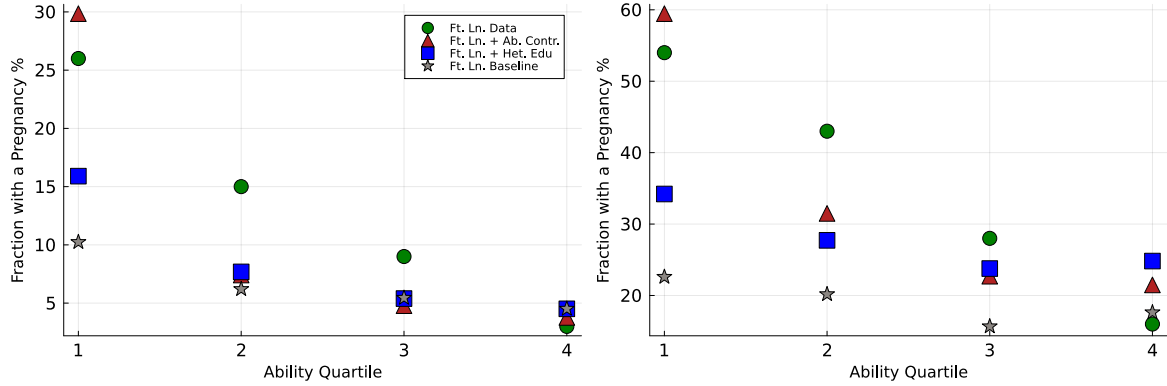
	(1)	(2)	(3)
	Baseline	Baseline + Educ. Het.	Baseline + Educ. Het. + Ab. Cont.
Total SSE	2.77	2.49	1.73
Pregnancies and Ability Moments SSE	1.61	1.73	1.12
Education Moments SSE	0.24	0.48	0.32
Marital Moments SSE	0.92	0.60	0.29
Fit Improvement ($1 - \frac{SSE_1}{SSE_0}$)		+ Educ. Het.	+ Ab. Cont.
Total Fit		10%	30%
Pregnancies and Ability Moments		13%	21%
Education Moments		-103%	33%
Marital Moments		51%	51%

As shown in figures 8a and 8b, the “baseline” model is not able to generate a positive relationship between ability and education. Adding heterogeneity by education improves a 10% the total fit of the model and a 13% the fit of the moments relating ability with pregnancies. As education strongly correlates with ability, adding heterogeneity in this dimension increases the slope at which higher ability decreases teen pregnancies. However, the model is still far from explaining the high number of pregnancies in low ability women. Furthermore, the improvement in the fit of pregnancies and ability moments is made at the expense of decreases by half the model performance in fitting education moments. Finally, adding the ability to the contraception technology improves the fit for the previous model in 30%

and 21% in the moments that relate pregnancies with ability. Additionally, now the model can explain the high pregnancy rate of the first and last ability quartiles but underpredict pregnancy rates for the second and third.

Figure 8. Fraction of Pregnancies by Age and Cognitive Ability Quartile

(a) First Period: Pregnancy between 14-17 years old (b) Second Period: Pregnancy between 18-21 years old



Three parameters mainly drive the relationship between ability and pregnancies: the share of ability on the contraception function λ_a , the cost of attending college ξ_c , and ω_h that give the concavity at which ability decreases the college psych cost. Figure 9c shows how the cost of college shape pregnancies between 14-17 years old when the contraception technology does not depend on ability ($\lambda_h = 0$). Figure 9d shows the case in which contraception technology depends on ability ($\lambda_h = -0.4$). A reduction in the cost of college cost decreases pregnancies for all ability groups. However, there is no point in the space parameter that generates a gap between groups of the magnitude we see in the data. The addition of the ability to the contraception technology creates a gap in the level of pregnancies between the low and high ability individuals, similar to the data.

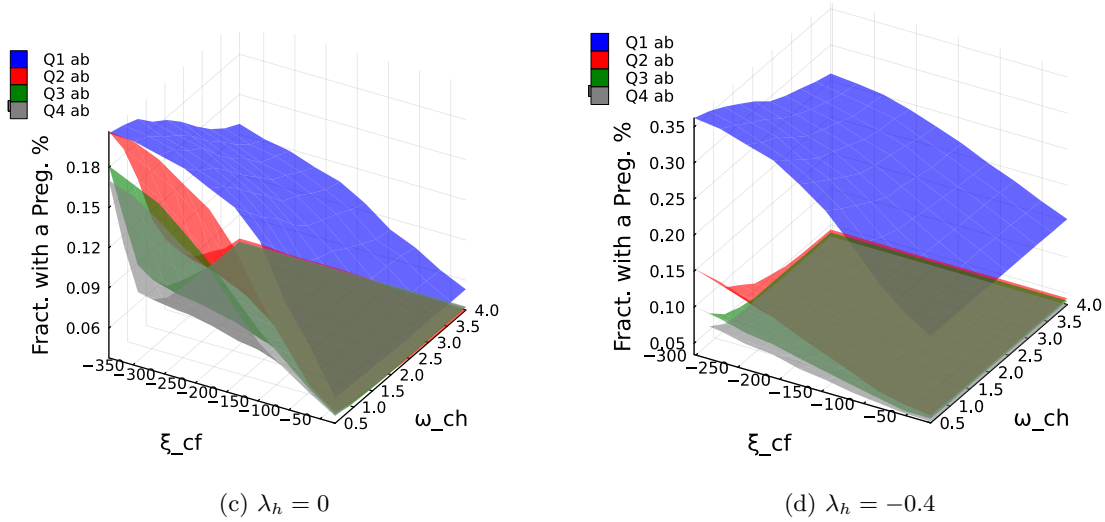
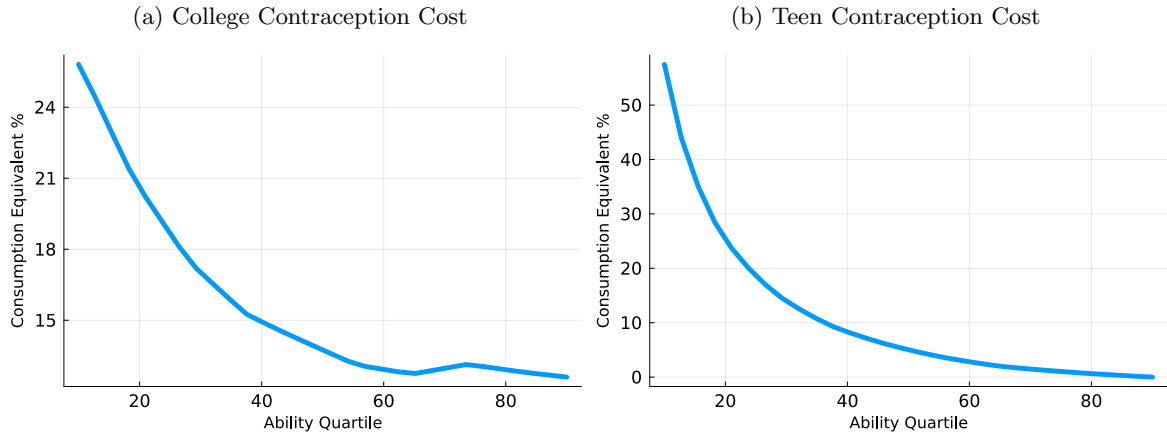


Figure 9. Relationship between ability and pregnancies.

5.3 Contraception Cost by Ability and Education

In this subsection, I quantify the value of better contraception technology in terms of consumption. In the model, formal education and ability decrease contraception costs. First, I quantify the difference in contraception cost by education. Second, I quantify the difference in contraception cost by ability.

Figure 10. Equivalent Consumption by Ability Percentiles



To measure the difference in contraception cost by education, in figure 10a, I calculate the increase in each period consumption necessary to have the same utility as in a model where

all education groups have the same contraception cost as college graduates. For women in the first ability decile, equalizing contraception cost by education groups is equivalent to increasing each period consumption by 24%; meanwhile, for women in the top decile is around 10%. Low-ability women benefit most from this contraception change as they are less likely to attend college; meanwhile, college attendance in the top decile is high, so the contraception cost remains the same for most women in the decile. In figure 10b, I quantify the difference in contraception cost by ability. Given that the only heterogeneity in the first period is in ability, I calculate the consumption increase in the first period necessary to make an individual indifferent to have the same ability as someone in the percentile 90, finding that someone in the first decile needs to increase by 50% her initial consumption.

6 Contraception Cost and Education Opportunities

Single motherhood and teen pregnancy are essential policy concerns as these women are considerably disadvantaged in multiple economic and social dimensions. Additionally, the effect of mistiming in children's outcomes is persistent over time, affecting multiple economic dimension as human capital accumulation, labor force participation, intergenerational inequality, and family formation.

An open question is the effect of the children on their mothers' outcomes. Whether early pregnancies cause disadvantages or women who have early pregnancies are disadvantaged early in life, with pregnancies not increasing their disadvantage, it is fundamental to design policies that can help them. For example, whether the cause of teen mothers' low education achievement is the pregnancy's timing, reducing contraception costs would have significant consequences on education achievement; however, if early inequalities cause the low education outcomes of these women, policies that decrease the cost of children for teens and single mothers will not improve educational achievement and can increase the number of early pregnancies. The effect of the U.S. welfare system on nonmarital and teenage pregnancies has been discussed since [Becker \(1991\)](#), who argue that at least part of the high number of nonmarital childbirth is caused by government aid. [Rosenzweig \(1999\)](#) find supportive evidence that welfare policies induce young women to choose to have a child outside of marriage. The relevance of this question is such that in 1996 the Personal Responsibility and

Work Opportunity Reconciliation Act reformed the welfare system, stating that one of its goals was reducing nonmarital and teen pregnancies ([Lichter and Jayakody, 2002](#)).

Previous research has found that the correlation between early pregnancies and poor educational outcomes highly overstates the effect of the child on the mother’s outcomes. [Hotz et al. \(2005\)](#) using miscarried as an IV found that teenage pregnancies have an insignificant effect on high school graduation at 28. [Levine and Painter \(2003\)](#) using a propensity score matching to deal with endogeneity found that only half of the correlation between teenage pregnancy is causal and teenage pregnancy decrease college attendance by 20% half of what is found on simple regression. In order to assess this question, I use the model to realize three counterfactuals:

1. High ability for contraception: All women are in the top ability percentile for contraception purposes.
2. High ability for education and wages: All women have the same college cost and income profile as women in the top ability percentile.
3. High ability for contraception, education, and wages: All women are in the top ability percentile for contraception, college cost, and wages.

Table 13. Counterfactual Results

	(1)	(2)	(3)
Moments	Same Cont. Ab.	Same Ab. Edu. + Same Ab. Wage	Same Cont. Ab. + Same Ab. Edu. + Same Ab. Wage
Attend Coll	5%	77%	83%
Preg < 18	-87%	-33%	-91%
Preg < 22	-46%	-30%	-49%

The counterfactual results are displayed in table [13](#). In the first column, we see how pregnancies and education change when low ability women have the same contraception cost as women in the top-ability percentile. Consistent with previous empirical studies, I found

that reducing contraception costs has little effect on college attendance. In the model, give to low ability women the same contraception cost as high ability women only increase college attendance by 5% as women who are teen mothers are in the lowest ability quartiles, which does that the cost of attending college is too high even without a child. However, even when education outcomes do not change, decreasing the contraception cost reduces teen pregnancies by almost 90% and before 22 by 46%. Additionally, we show in subsection that there is a considerable increase in welfare for low ability women. Column 2 shows the effect of reducing the psych cost of attending college for low ability individuals and keeping the contraception cost fixed. In this case, college attendance increase by 77%, and teen pregnancies are reduced by a third. Finally, column 3 shows the effect of reducing contraception and the psych college cost. In this case, college attendance increase by 80%, and teen pregnancies are reduced by almost half.

In conclusion, differences in contraception costs and education opportunities cause the difference in teen pregnancies by ability. However, even when policies that decrease contraception costs are welfare improving, especially for low ability individuals, as shown in figure 10, implementing these policies without improving education opportunities for low ability individuals will have little impact on educational outcomes.

7 Reduce Contraception Cost

Table 14. Reduce Contraception Cost

Moments	Data	Coll. Cost	Moments	Data	Coll. Cost
Preg. < 18 yr. old	0.11	-24%	Pregnancy 18-21 yrs. old HSD, NPP	0.47	-46%
Preg. < 22 yr. old	0.40	-15%	Pregnancy 18-21 yrs. old HS, NPP	0.40	-38%
Preg. < 29 yr. old	0.64	-42%	Pregnancy 18-21 yrs. old Attend Coll, NPP	0.17	-10%
Drop HS No Preg	0.06	130%	Single Mom 22-29 yrs. old HSD, NPP	0.18	-98%
Drop HS Preg	0.27	-6%	Single Mom 22-29 yrs. old HS, NPP	0.10	-94%
Attend Coll, NPP No Preg	0.38	-12%	Single Mom 22-29 yrs. old Coll, NPP	0.02	0%
Attend Coll, NPP Preg	0.08	12%	Married 22-29 yrs. old HSD, NPP	0.50	-16%
Grad. Coll., NPP No Preg	0.51	-5%	Married 22-29 yrs. old HS, NPP	0.68	3%
Grad. Coll., NPP Preg	0.38	6%	Married 22-29 yrs. old Coll, NPP	0.64	1%
Married 18-21 yrs. old HSD, NPP	0.35	-2%	Pregnancy 22-29 yrs. old HSD, NPP	0.48	-97%
Married 18-21 yrs. old HS, NPP	0.34	-1%	Pregnancy 22-29 yrs. old HS, NPP	0.37	-89%
Single Mom 18-21 yrs. old HSD, NPP	0.36	-52%	Pregnancy 22-29 yrs. old Coll, NPP	0.49	-2%
Single Mom 18-21 yrs. old HS, NPP	0.24	-35%			

NPP = No Previous Pregnancy.

In this section, I use the model to analyze the effect on educational outcomes, marriage, and pregnancies of a policy that reduces the cost of contraception for all women to the cost of a college graduate. In order to do it, I set the contraception cost (λ) for high school dropouts equal to the contraception cost of high school graduates between 18-21 years old and between 22-29 years old, the contraception cost of high school dropouts and high school graduates is set to the cost of a college graduate.

If all women have contraception costs of college graduates, pregnancies decrease in every age group. Teen pregnancies decrease by 24%, before 22 by 15%, and before 29 by 42%. The decrease in teenage pregnancy is driven by expectations about the future, given that the contraception cost between 14-17 years old remains the same. The decrease in the contraception cost increases high school dropouts; as discussed in the previous section, teenagers drop high school because of the cognitive cost of education and not because of teen pregnancies. In the model, for low ability women, access to better contraception technology is an important reason to acquire education. When this motive disappears, education is less attractive, then high school dropouts double, and college attendance decrease by 12%. Single moth-

ers decrease between 18-21 years old by half for high school dropouts and by 35% for high school graduates. Between 22-29 years old, noncollege single mothers almost disappear but remain constant for college graduates as their contraception costs remain unchanged. Finally, the number of marriages remains constant for high school graduates and college graduates but decline by 20% for high school dropout, as marriage work as insurance in the case of a pregnancy.

8 Conclusion

In this paper, I empirically and quantitatively explore cognitive ability's effects on fertility timing. The main finding is that education, wage, and marital opportunity cost cannot fully explain differences in fertility timing and pregnancy intention across women with different levels of cognitive ability. Women in the lowest ability quartile have nine times more teen pregnancies than women in the top. However, this difference in pregnancies disappears after 22 years old. The cognitive ability remains relevant in fertility after controlling for education and wage, decreasing the number of pregnancies at early ages.

The paper develops and estimates a dynamic life-cycle model with endogenous education, wages, and marriage to evaluate how cognitive ability affects women's fertility. First, consistent with the empirical results, cognitive ability affects teen and young adult fertility risk beyond the differences in time opportunity cost and fertility risk by education. Then, through the lens of the model, I show the welfare gains of improvement in contraception cost by ability. Finally, I use the model to analyze whether single mothers' poor educational outcomes result from the cost of early pregnancy or previous inequalities. Even when reducing teen pregnancies is welfare improving, college attendance remains constant as the cause of poor educational outcomes is that education is too costly even without a child.

Finally, this work contributes to the increasing literature showing cognitive ability's importance in noneconomic outcomes and highlights its importance in understanding fertility timing and intention. Additionally, it provides a framework to understand the effects on early pregnancies and educational outcomes of policies targeting contraception cost and cognitive ability. In future work, I expect to extend the model to allow for children's human capital accumulation and evaluate the effect of early pregnancies on children's educational outcomes

and intergenerational inequality with persistent cognitive ability between generations.

References

- Amador, Diego**, “The consequences of abortion and contraception policies on young women’s reproductive choices, schooling and labor supply,” *Documento CEDE*, 2017, (2017-43).
- Becker, Gary S**, “A Theory of the Allocation of Time,” *The economic journal*, 1965, 75 (299), 493–517.
- , *A treatise on the family: Enlarged edition*, Harvard university press, 1991.
- Bronars, Stephen G and Jeff Grogger**, “The economic consequences of unwed motherhood: Using twin births as a natural experiment,” *The American Economic Review*, 1994, pp. 1141–1156.
- Caucutt, Elizabeth M., Nezih Guner, and John Knowles**, “Why Do Women Wait? Matching, Wage Inequality, and the Incentives for Fertility Delay,” *Review of Economic Dynamics*, 2002, 5 (4), 815–855.
- Cawley, John, Karen Conneely, James Heckman, and Edward Vytlačil**, “Cognitive Ability, Wages, and Meritocracy,” Working Paper 5645, National Bureau of Economic Research July 1996.
- Choi, Sekyu**, “Fertility risk in the life cycle,” *International economic review*, 2017, 58 (1), 237–259.
- Commendador, Kathleen**, “The relationship between female adolescent self-esteem, decision making, and contraceptive behavior,” *Journal of the American Academy of Nurse Practitioners*, 2007, 19 (11), 614–623.
- Filote, Andra, Georgi Kocharkov, Jan Mellert et al.**, “Teenage childbearing and the welfare state,” Technical Report, mimeo 2019.
- Foster, Diana Greene, M Antonia Biggs, Lauren Ralph, Caitlin Gerdtts, Sarah Roberts, and M Maria Glymour**, “Socioeconomic outcomes of women who receive and women who are denied wanted abortions in the United States,” *American journal of public health*, 2018, 108 (3), 407–413.
- Greenwood, Jeremy, Nezih Guner, and John A Knowles**, “Women on welfare: A macroeconomic analysis,” *American Economic Review*, 2000, 90 (2), 383–388.
- Hai, Rong and James J Heckman**, “The Causal Effects of Youth Cigarette Addiction and Education,” Technical Report, National Bureau of Economic Research 2022.
- Hansen, Karsten T, James J Heckman, and Kathleen J Mullen**, “The effect of schooling and ability on achievement test scores,” *Journal of econometrics*, 2004, 121 (1-2), 39–98.

- Heckman, James J, John Eric Humphries, and Gregory Veramendi**, “Returns to education: The causal effects of education on earnings, health, and smoking,” *Journal of Political Economy*, 2018, *126* (S1), S197–S246.
- , **Jora Stixrud, and Sergio Urzua**, “The effects of cognitive and noncognitive abilities on labor market outcomes and social behavior,” *Journal of Labor economics*, 2006, *24* (3), 411–482.
- Hotz, V Joseph, Susan Williams McElroy, and Seth G Sanders**, “Teenage childbearing and its life cycle consequences exploiting a natural experiment,” *Journal of Human Resources*, 2005, *40* (3), 683–715.
- Kearney, Melissa S and Phillip B Levine**, “The economics of nonmarital childbearing and the marriage premium for children,” *Annual Review of Economics*, 2017, *9*, 327–352.
- Kozlov, Egor**, “Collective Household Models with Limited Commitment.” PhD dissertation, Northwestern University 2021.
- Levine, David I and Gary Painter**, “The schooling costs of teenage out-of-wedlock childbearing: Analysis with a within-school propensity-score-matching estimator,” *Review of Economics and Statistics*, 2003, *85* (4), 884–900.
- Lichter, Daniel T and Rukamalie Jayakody**, “Welfare reform: How do we measure success?,” *Annual review of sociology*, 2002, pp. 117–141.
- Musick, Kelly, Paula England, Sarah Edgington, and Nicole Kangas**, “Education differences in intended and unintended fertility,” *Social Forces*, 2009, *88* (2), 543–572.
- Regalia, Ferdinando, José-Victor Rios-Rull, and J Short**, “What accounts for the increase in the number of single households,” *Unpublished manuscript*, 2011.
- Rosenzweig, Mark R**, “Welfare, marital prospects, and nonmarital childbearing,” *Journal of Political Economy*, 1999, *107* (S6), S3–S32.
- and **T Paul Schultz**, “Schooling, information and nonmarket productivity: contraceptive use and its effectiveness,” *International Economic Review*, 1989, pp. 457–477.
- Seshadri, Ananth and Anson Zhou**, “Intergenerational mobility begins before birth,” *Journal of Monetary Economics*, 2022.

A Calibration Parameters

Appendix Table A1. Contraception Cost by Education and Age Group

Education / Age Group	13-17 yr. old	18-21 yr. old	22-29 yr. old
High School Dropout	1.18	2.70	0.38
High School Graduate	1.18	0.63	0.91
College	1.18	0.57	6.28

Appendix Table A2. Probability Meeting a Husband

Education / Age Group	18-21 yr. old	22-29 yr. old
High School Dropout	0.33	0.29
High School Graduate	0.48	0.31
College		0.33

B Perfect Contraception

Appendix Table B1. Reduce Contraception Cost

Moments	Data	Perfect Contr.	Moments	Data	Perfect Contr.
Preg. < 18 yr. old	0.11	-94%	Pregnancy 18-21 yrs. old HSD, NPP	0.47	-100%
Preg. < 22 yr. old	0.40	-93%	Pregnancy 18-21 yrs. old HS, NPP	0.40	-93%
Preg. < 29 yr. old	0.64	-82%	Pregnancy 18-21 yrs. old Attend Coll, NPP	0.17	-96%
Drop HS No Preg	0.06	82%	Single Mom 22-29 yrs. old HSD, NPP	0.18	-72%
Drop HS Preg	0.27	63%	Single Mom 22-29 yrs. old HS, NPP	0.10	-95%
Attend Coll, NPP No Preg	0.38	-3%	Single Mom 22-29 yrs. old Coll, NPP	0.02	-100%
Attend Coll, NPP Preg	0.08	-100%	Married 22-29 yrs. old HSD, NPP	0.50	-20%
Grad. Coll., NPP No Preg	0.51	-6%	Married 22-29 yrs. old HS, NPP	0.68	3%
Grad. Coll., NPP Preg	0.38	39%	Married 22-29 yrs. old Coll, NPP	0.64	0%
Married 18-21 yrs. old HSD, NPP	0.35	-5%	Pregnancy 22-29 yrs. old HSD, NPP	0.48	-69%
Married 18-21 yrs. old HS, NPP	0.34	0%	Pregnancy 22-29 yrs. old HS, NPP	0.37	-95%
Single Mom 18-21 yrs. old HSD, NPP	0.36	-100%	Pregnancy 22-29 yrs. old Coll, NPP	0.49	-30%
Single Mom 18-21 yrs. old HS, NPP	0.24	-91%			

NPP = No Previous Pregnancy.

C Model Mechanism Contribution

Appendix Table C1. Pregnancy and Wages

Moments	Data	Baseline + Educ. Het + Ab. Cont.	Baseline + Educ. Het.	Baseline
% Preg Q1 14-17 yrs. old	0.26	0.30	0.16	0.11
% Preg Q2 14-17 yrs. old	0.15	0.07	0.08	0.07
% Preg Q3 14-17 yrs. old	0.09	0.05	0.05	0.05
% Preg Q4 14-17 yrs. old	0.03	0.04	0.05	0.05
Drop HS No Preg	0.06	0.06	0.07	0.06
Drop HS Preg	0.28	0.27	0.21	0.29
Att. Coll. No Preg	0.41	0.38	0.35	0.35
Att. Coll. Preg	0.08	0.08	0.09	0.07
Att. Coll. Q1	0.12	0.10	0.07	0.11
Att. Coll. Q2	0.25	0.32	0.30	0.29
Att. Coll. Q3	0.40	0.46	0.46	0.44
Att. Coll. Q4	0.67	0.49	0.49	0.48
Graduate College No Preg	0.62	0.51	0.55	0.55
Graduate College Preg	0.30	0.38	0.33	0.36
% Preg AQ1 18-21 yrs. old NPP	0.54	0.59	0.34	0.30
% Preg AQ2 18-21 yrs. old NPP	0.43	0.31	0.28	0.25
% Preg AQ3 18-21 yrs. old NPP	0.28	0.23	0.24	0.20
% Preg AQ4 18-21 yrs. old NPP	0.16	0.21	0.25	0.21

NPP = No Previous Pregnancy.

Appendix Table C2. Pregnancy and Wages

Moments	Data	Baseline	Baseline	Baseline
		+ Educ. Het	+ Educ. Het.	
		+ Ab. Cont.		
Married 18-21 yrs. old HSD, NPP	0.39	0.35	0.29	0.44
Married 18-21 yrs. old HS, NPP	0.41	0.34	0.32	0.36
Single Mom 18-21 yrs. old HSD, NPP	0.44	0.36	0.36	0.35
Single Mom 18-21 yrs. old HS, NPP	0.29	0.24	0.19	0.17
Pregnancy 18-21 yrs. old HSD, NPP	0.42	0.47	0.48	0.57
Pregnancy 18-21 yrs. old HS, NPP	0.35	0.40	0.33	0.29
Pregnancy 18-21 yrs. old Attend Coll.	0.16	0.17	0.15	0.11
% Preg AQ1 22-29 yrs. old NPP	0.48	0.64	0.50	0.51
% Preg AQ2 22-29 yrs. old NPP	0.48	0.51	0.47	0.49
% Preg AQ3 22-29 yrs. old NPP	0.45	0.33	0.44	0.46
% Preg AQ4 22-29 yrs. old NPP	0.45	0.29	0.39	0.45
Single Mom 22-29 yrs. old HSD, NPP	0.22	0.18	0.14	0.18
Single Mom 22-29 yrs. old HS, NPP	0.12	0.10	0.10	0.13
Single Mom 22-29 yrs. old Coll., NPP	0.02	0.02	0.02	0.02
Married 22-29 yrs. old HSD, NPP	0.61	0.50	0.56	0.51
Married 22-29 yrs. old HS, NPP	0.66	0.68	0.57	0.83
Married 22-29 yrs. old Coll., NPP	0.68	0.64	0.65	0.99
Pregnancy 22-29 yrs. old HSD, NPP	0.43	0.48	0.57	0.40
Pregnancy 22-29 yrs. old HS, NPP	0.43	0.37	0.45	0.52
Pregnancy 22-29 yrs. old Coll., NPP	0.42	0.49	0.42	0.35

NPP = No Previous Pregnancy.