

MCMC Practical Session

- What can you do?
- How can you do it?
- What is important in practice?

https://github.com/aponteeduardo/cpp_lecture

`samples/metropolis_hastings.m`

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Write down your model

$$\underbrace{p(\theta|y, u, m)}_{\text{posterior}} = \frac{\overbrace{p(y|u, \theta, m)}^{\text{likelihood}} \overbrace{p(\theta|m)}^{\text{prior}}}{\underbrace{p(y|m)}_{\text{model evidence}}}$$

y : *Experimental data*

u : *Experimental design*

m : *Model*

θ : *Parameters*

Write down your model

Meaning	Variable
Experimental data	y
Experimental design	u
Model parameters	θ
Priors	$p(\theta)$
MCMC parameters	$h(\theta)$
Log likelihood	l_h
Log prior probability	l_p

Inference

- Goal samples: Draw samples $\theta_1, \dots, \theta_n$ from the distribution $p(\theta|y, u, m)$.
- Typically you are interested in:

$$E[\theta|y, u, m] \approx \frac{1}{n} \sum_{i=1}^n \theta_i$$

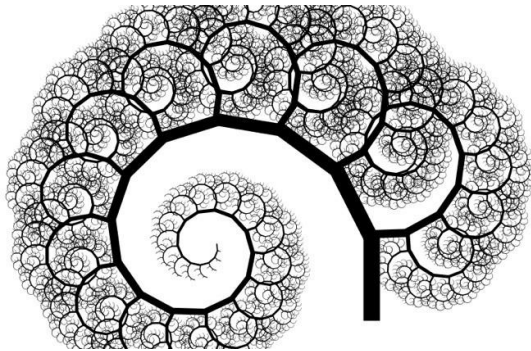
$$\text{Var}[\theta|y, u, m] \approx \frac{1}{n-1} \sum_{i=1}^n \left(\theta_i - \frac{1}{n} \sum_{j=1}^n \theta_j \right)^2$$

$$\begin{aligned} p(y^*|y, u, m) &= \int p(y^*|u, \theta, m) p(\theta|y, u, m) d\theta \\ &\approx \frac{1}{n} \sum_i^n p(y^*|u, \theta_i, m) \end{aligned}$$

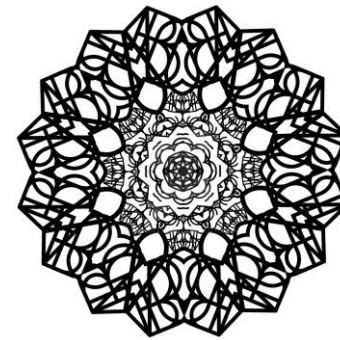
Model specification

- A simple learning task.

A



B



Decision: $y = [A, B, A, A, B, \dots]$
Monetary reward: $u = [2.1, 3.4, 0.5, \dots]$

Likelihood function

- Participants learn the expected reward of each option over trials:

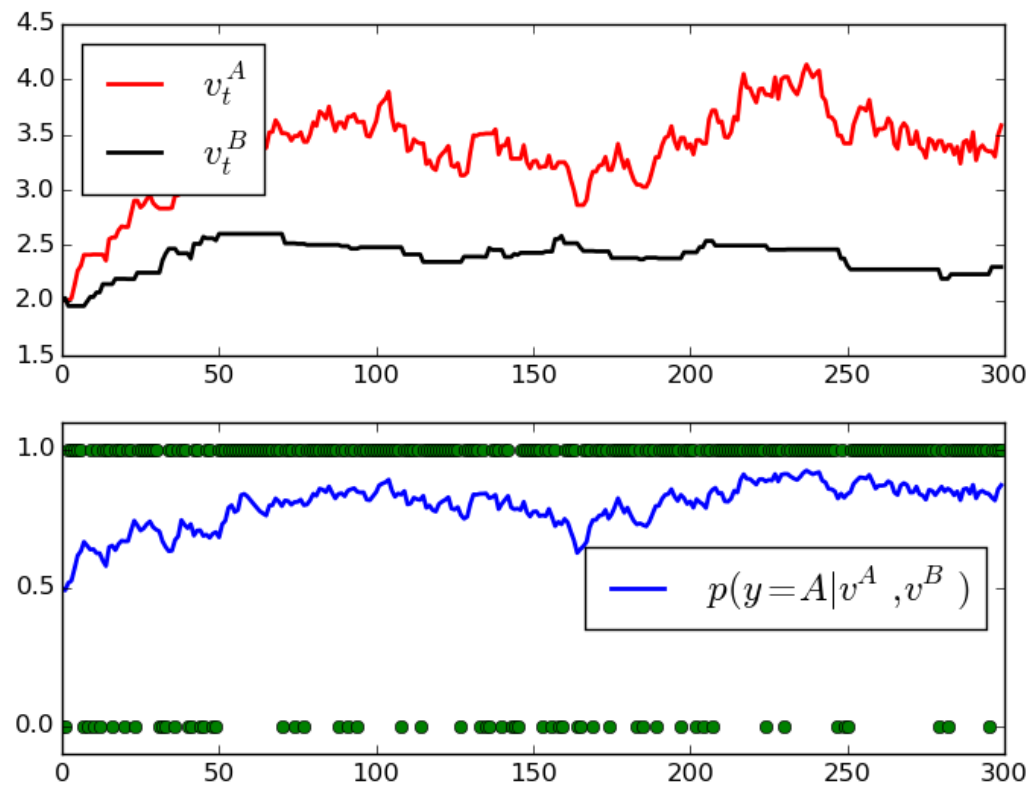
$$v_{t+1}^A = v_t^A + \alpha(u_t - v_t^A)$$

$$v_{t+1}^B = v_t^B + \alpha(u_t - v_t^B)$$

- They prefer the option with the highest expected reward:

$$p(y = A | v^A, v^B) = \frac{\exp \beta v^A}{\exp \beta v^A + \exp \beta v^B}$$

Likelihood function



Inference

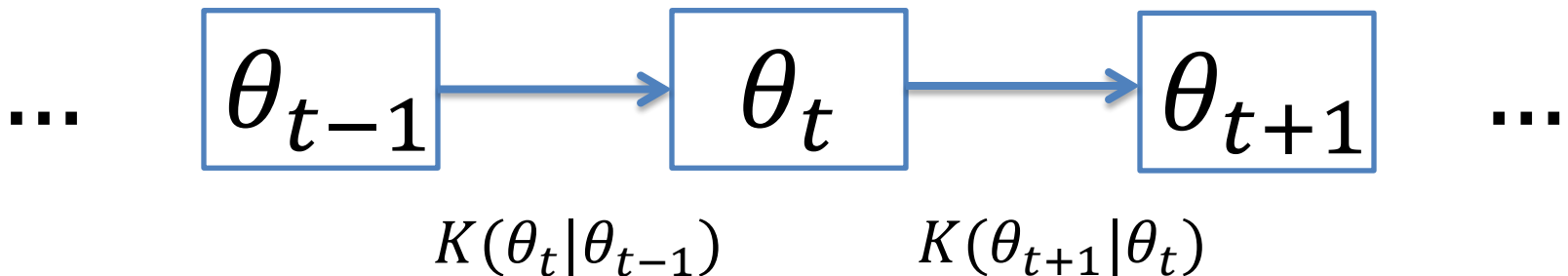
- The model has two free parameters and a few priors.
 - Parameters: α, β
 - Priors (among others): v_0^A, v_0^B .
- The algorithm has the following parameters:
 - Number of iterations.
 - Proposal distribution.

Inference

- Assumption α and β are subject specific traits.
- Given experimental data y and u , obtain samples of the posterior distribution of α and β .

Inference

- Strategy: Draw samples from a Markov chain whose equilibrium distribution is the posterior $p(\theta|y, u)$, $\theta = [\alpha, \beta]$.



Inference

- Sampling from $K(\cdot | \theta)$ is done in two steps:
 - Obtain a proposed sample from a distribution $q(\cdot | \theta)$.
 - Accept or reject the proposed sample with certain probability. You need to draw a sample from the uniform distribution in the interval $[0, 1]$.
- $K(\cdot | \theta)$ satisfies the detailed balance condition.

Inference

- Propose a random sample $\hat{\theta}$ from a simple distribution (very often a Gaussian).
- Accept with probability:

$$\min \left\{ 1, \frac{p(y|u, \hat{\theta})p(\hat{\theta})}{p(y|u, \theta_t)p(\theta_t)} \right\}$$

- If you accept $\hat{\theta}$, then $\theta_{t+1} = \hat{\theta}$.
- If you reject $\hat{\theta}$, then $\theta_{t+1} = \theta_t$.

Proposal distribution

- Gaussian!

$$\hat{\theta} = \theta_t + \varepsilon \quad \varepsilon \sim N(0, \Sigma)$$

- To sample ε you can use the fact that $\text{var}(Lx) = L^T \text{var}(x)L$. If $L^T L = \Sigma$, and $\text{var}(x) = I$, then $\text{var}(Lx) = \Sigma$.
- **chol** computes the square root of Σ , L .

Main loop

- Repeat:
 - Sample from proposal distribution
 - Evaluate the joint probability
 - Accept or reject new sample
 - Store the results

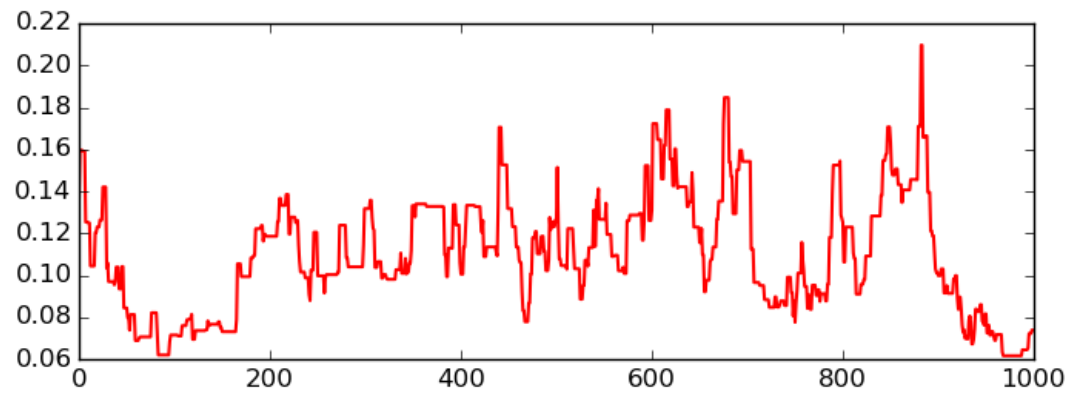
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36 % Compute the acceptance rate
37 ar = 0;
38
39 for i = 1: nburnin + niter
40     % Propose a new sample
41     ntheta = otheta + htheta.csigma * randn(size(otheta));
42
43     % Compute log likelihood conditioned on a proposed sample
44     nllh = llh(y, u, ntheta, ptheta);
45     % Compute log prior probability
46     nlpp = lpp(u, ntheta, ptheta);
47
48     % Probability of acceptance
49     paccept = min(0, (nllh + nlpp) - (ollh + olpp));
50
51     % If the samples is accepted keep it
52     if exp(paccept) > rand
53         otheta = ntheta;
54         ollh = nllh;
55         olpp = nlpp;
56
57         ar = ar + 1;
58     end
59
60     % Store your samples
61     if i > nburnin
62         post.ljp(i - nburnin) = ollh + olpp;
63         post.theta(:, i - nburnin) = otheta;
64     end
65
66 end

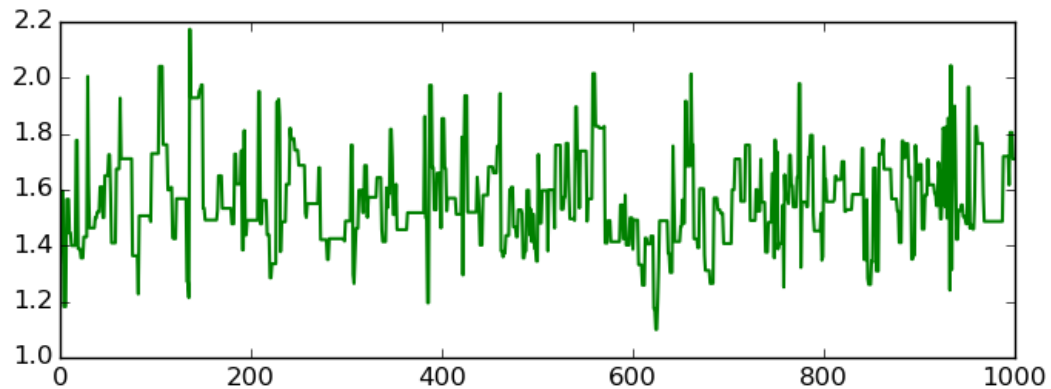
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Samples

α

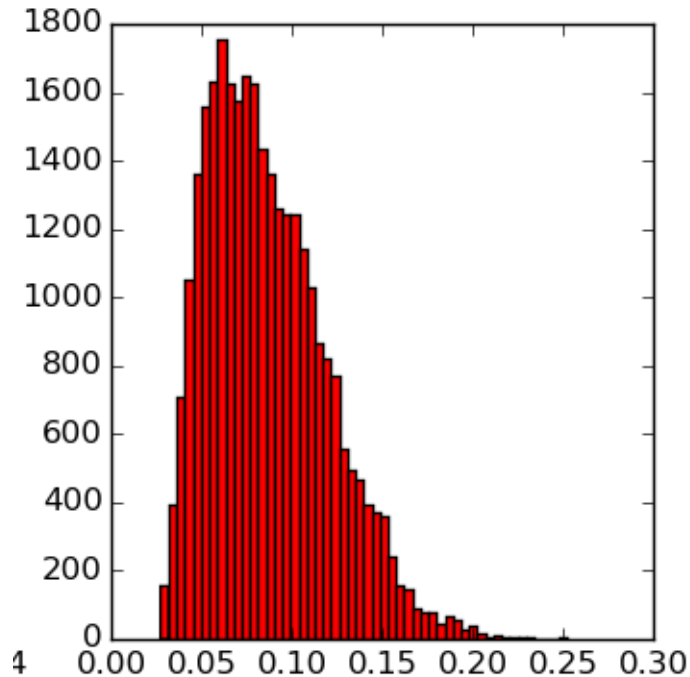


β



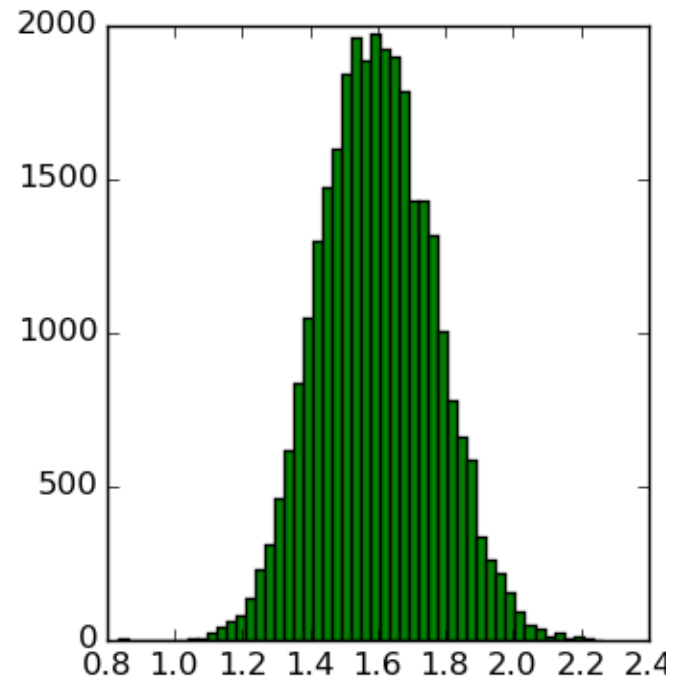
Samples

α



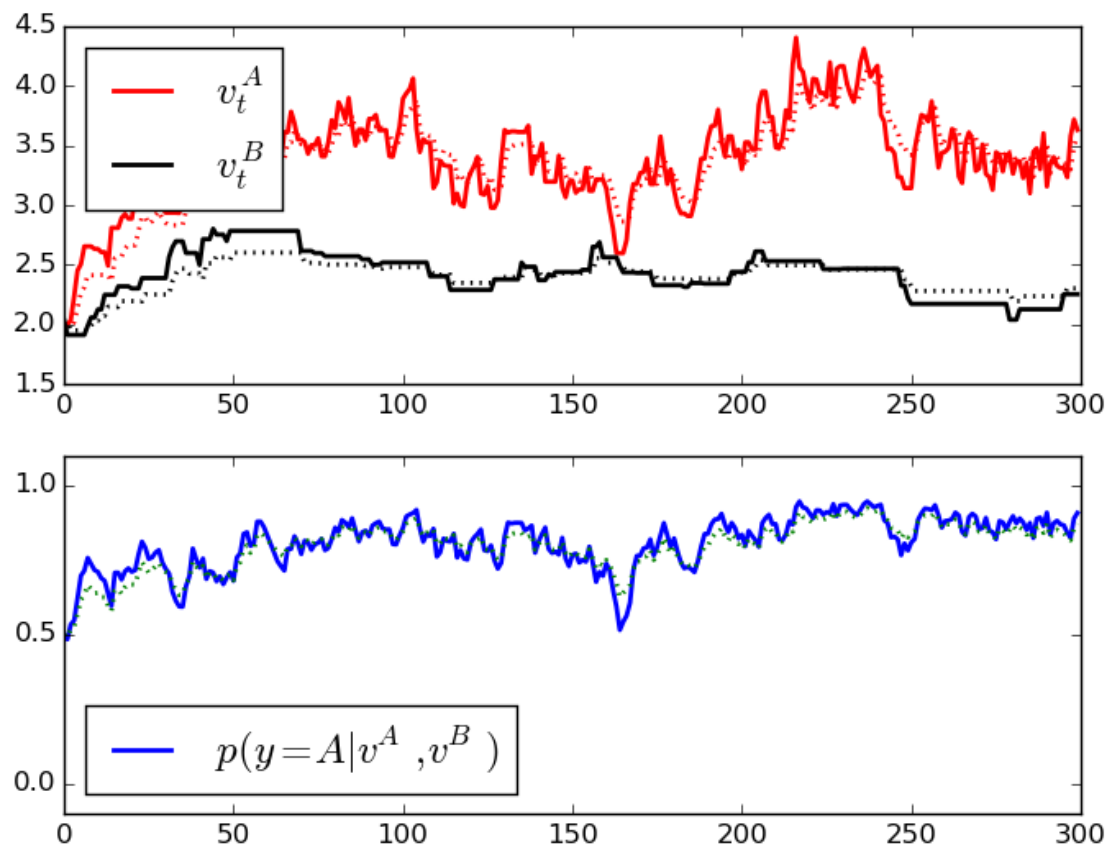
$$\alpha = 0.05 \rightarrow E[\alpha|y, u, m] = 0.08$$

β



$$\beta = 1.5 \rightarrow E[\beta|y, u, m] = 1.58$$

Samples



Some notes!

- There is no universal, principled criterion to decide whether a chain has reached equilibrium. In practice, there are very good heuristics to test this.
- Convergence is only asymptotically guaranteed.
- Efficiency: How many samples do you need to adequately sample the distribution?

Some tips

- There is a very large literature on MCMC!
- Some problems are harder than others!
- There is a trade-off between how efficient your solutions is, and how long you need to get it working. Your time is more valueable than computational time!
- Check [tapas/mpdcm/](#) for more advanced code to use in your own problem.