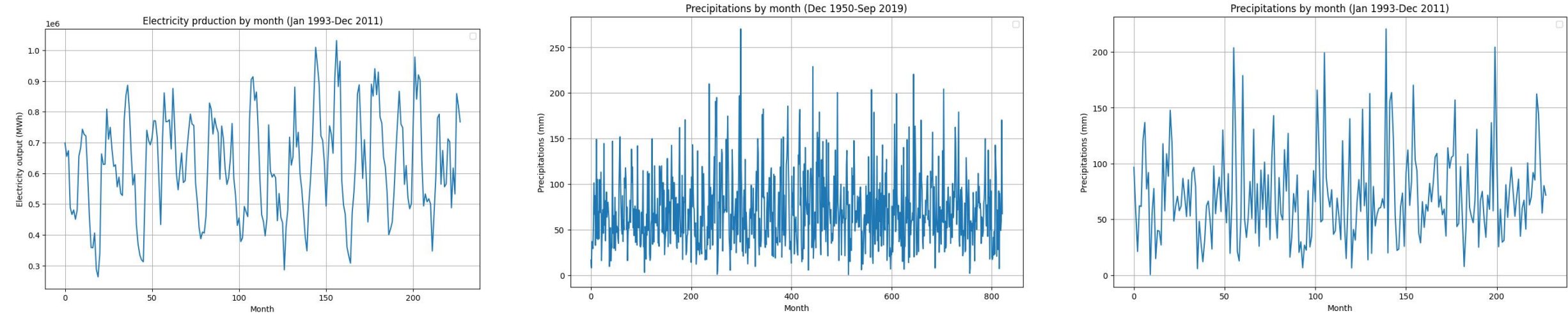


## Introduction and exploratory data analysis

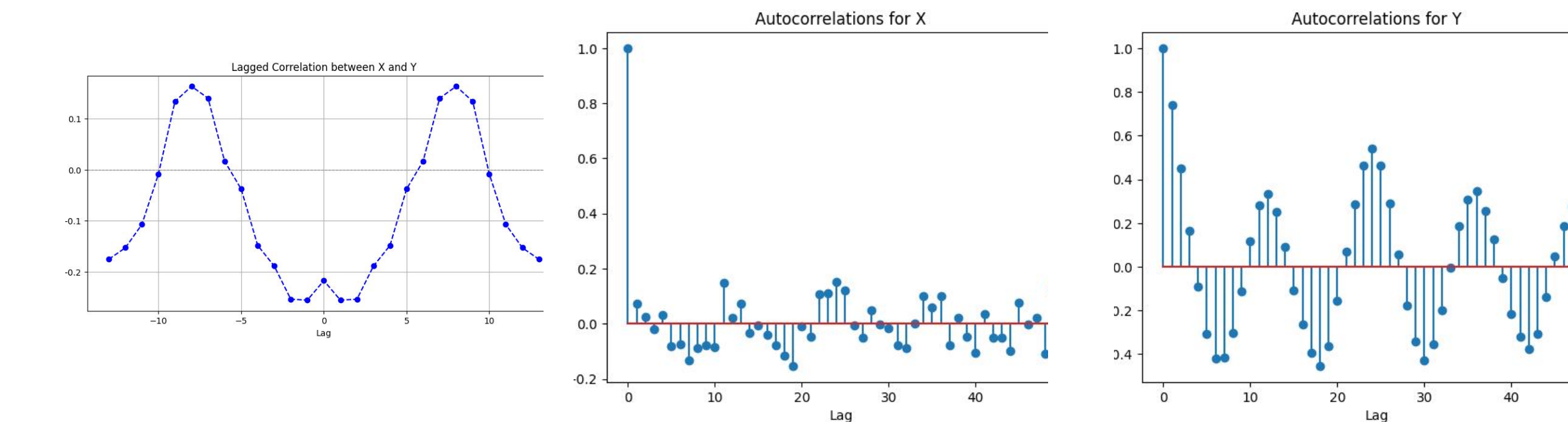
Our task consists of building a model to forecast electricity production in Trondheim using precipitation data as an explanatory variable. I will call  $X_t$  the precipitation data (exogenous variable) and  $Y_t$  the energy production (dependent variable).

Both quantities are plotted with on a monthly basis but the data for  $X_t$  spans a longer time window. (January 1993 to December 2011, 228 values VS December 1950 to September 2019, 822 values) Because of the exogenous assumption I decided to model precipitations independently, therefore employing the whole data available.



Plotting the lagged correlations between the precipitations and the energy production we see there seems to be a noticeable positive correlation around the 8 months lag mark. On the shorter lags the quantities are negatively correlated.

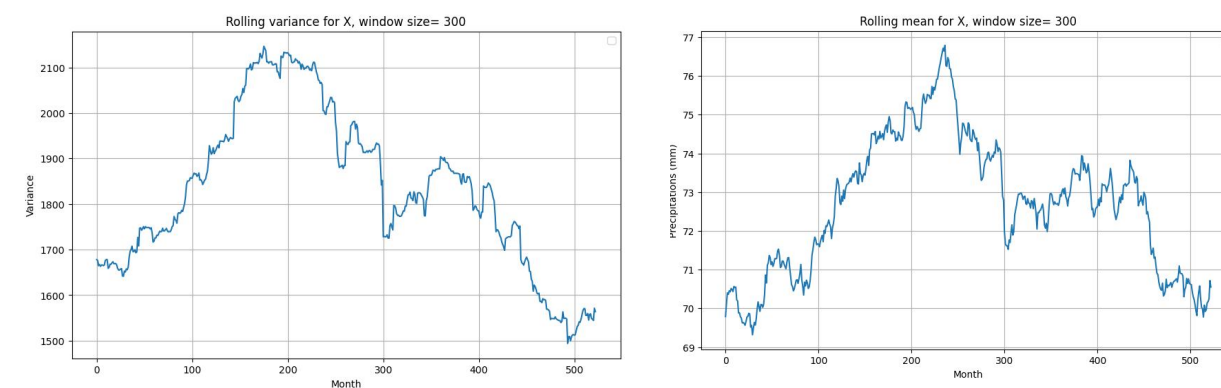
Looking at the auto-correlation functions for both X and Y we see that they exhibit seasonal behavior on annual scale. Moreover, the energy production data presents higher regularity, likely due to its relationship to energy demand.



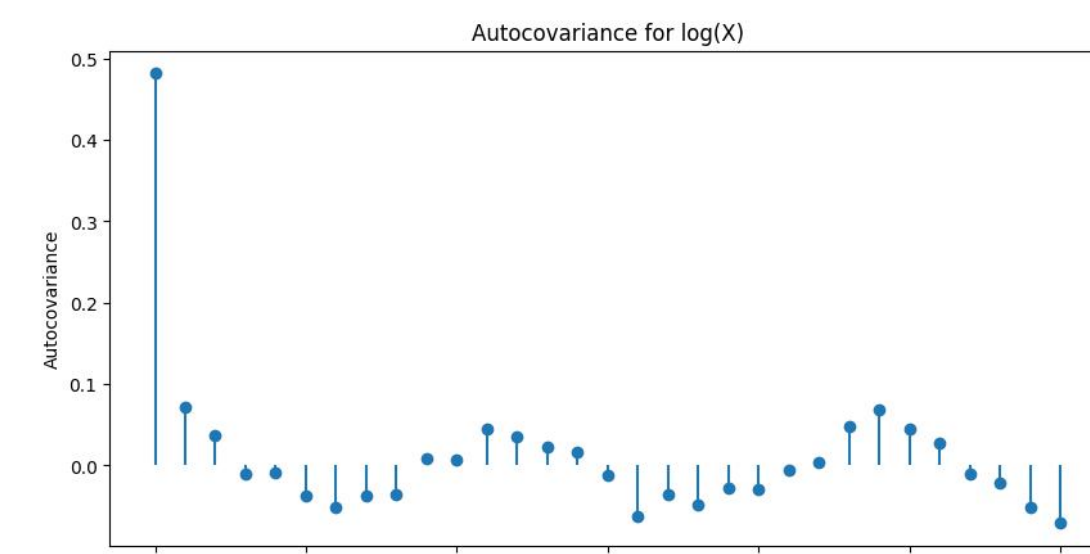
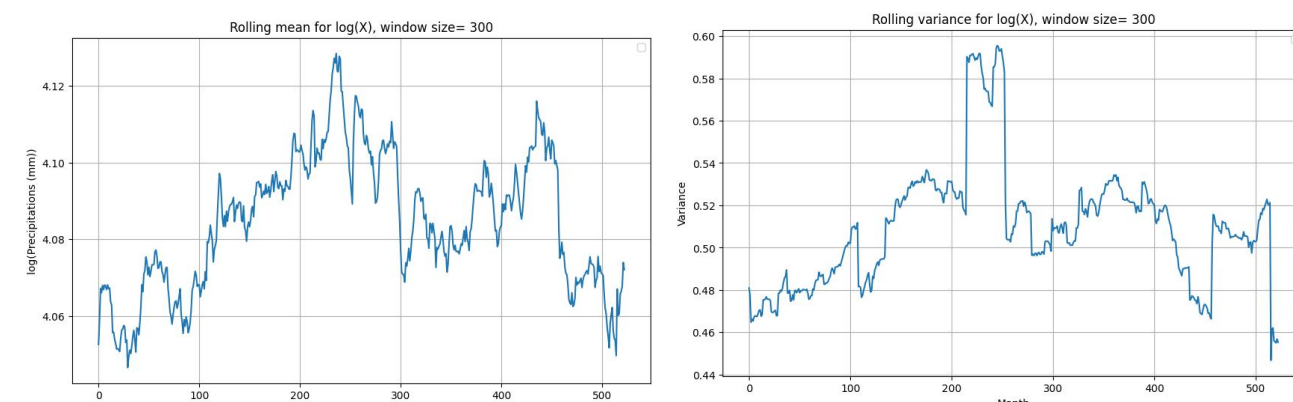
## Stationarity and Transformations

We want to explore the stationarity properties of our data in order to make a more informed modeling choice. This looks to be particularly relevant for X since its dynamics exhibits less regularity.

I plotted rolling means and rolling variances with a time window of 300 observations in order to assess potential changes in the mean function and the covariance function over time.



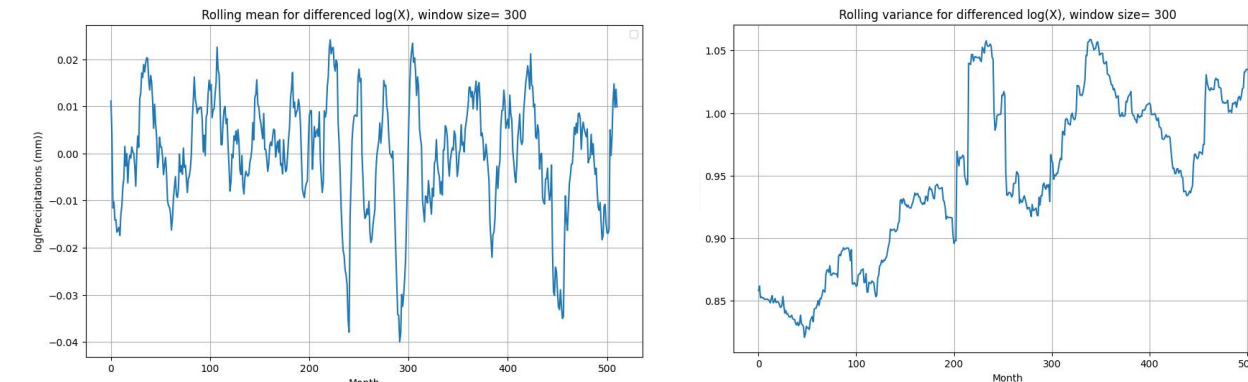
I tried to apply an bijective transformation of the data in an attempt to stabilize the mean and the variance, in particular the function:  $f(x)=\log(x)$ .



By plotting the auto-covariance for the logged version of X we see a major improvement in the regularity of its auto-regressive component.

## Seasonal differencing and SARIMA - GARCH

Before any modeling choice we need to address the seasonality of our data. By applying seasonal differencing we can get the following behavior:



Both the mean and the variance exhibit a more stable pattern, however we notice a slight upward trend in the latter. It becomes more apparent that the precipitation data is characterized by some sort of heteroskedasticity.

This fact could be addressed by a SARIMA model for the mean (that handles the seasonal differencing) supported by a GARCH model to model for the growing variance.

Analytically, our steps modeling steps can be written in this way:

**Step 1:** Log Transformation of  $X_t$

$$X_t \mapsto \log(X_t)$$

**Step 2:** Fitting of a SARIMA((p,q,d)x(PQD)s) model to handle the seasonal and the auto-regressive components

$$\Phi(B^s)\phi(B)\nabla_s^d\nabla^d\log(X_t) = \Theta(B^s)\theta(B)\epsilon_t$$

$$\phi(B) = 1 - \sum_{i=1}^p \phi_i B^i \quad \Phi(B) = 1 - \sum_{i=1}^P \Phi_i B^i \quad \nabla = 1 - B$$

$$\theta(B) = 1 + \sum_{i=1}^q \theta_i B^i \quad \Theta(B) = 1 + \sum_{i=1}^Q \Theta_i B^{is} \quad \nabla_s = 1 - B^s$$

**Step 3:** Fitting a GARCH(p,q) model on the residuals of the SARIMA to handle heteroskedasticity.

$$\epsilon_t = \sigma_t z_t \quad z_t \sim N(0,1)$$
$$\sigma_t^2 = \alpha_0 + \sum_{i=1}^p \alpha_i \epsilon_{t-i}^2 + \sum_{j=1}^q \beta_j \sigma_{t-j}^2$$

## Model Implementation and Calibration

For the implementation and fitting of the SARIMA model I used a State Space representation of the model.

Using the fact that the polynomial equation in terms of the shift operator B can be expanded (I implemented a code for handling polynomial coefficients during multiplication) I reconstructed the State equation and the Observation Equation.

Given polynomials:

$$P_F(B) = \Phi(B^s)\phi(B)\nabla_s^d\nabla^d = 1 + \sum_{i=1}^M f_i B^i$$

$$P_H(B) = \Theta(B^s)\theta(B) = 1 + \sum_{i=1}^N h_i B^i$$

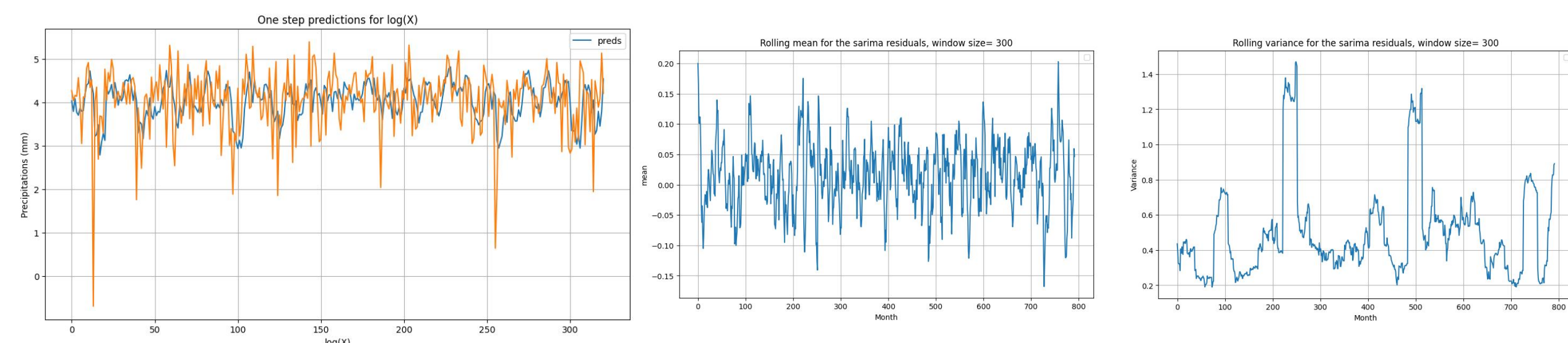
After defining  $n=\max(N+1,M)$ , we can construct the matrices:

$$F = \begin{bmatrix} -f_1 & -f_2 & \dots & -f_n \\ 1 & 0 & \dots & 0 \\ \dots & \dots & \dots & \dots \\ 0 & \dots & 1 & 0 \end{bmatrix} \quad H = [1 \quad h_1 \quad \dots \quad h_{n-1}] \quad V_1 = \begin{bmatrix} 1 & 0 & \dots & 0 \\ 0 & 0 & \dots & 0 \\ \dots & \dots & \dots & \dots \\ 0 & \dots & 0 & 0 \end{bmatrix}$$

We end up with the State Space representation:

$$\begin{cases} h_t = Fh_{t-1} + w_t & w_t \sim WN(0, V_1) \\ X_t = Hh_t \end{cases}$$

The model can then be calibrated by minimizing some loss function such as the log-likelihood or a mean square error. I decided to minimize the mean square error, in particular the error between the one step predictor obtained by the Kalman Filter and the actual values.



## SARIMAX for the Electricity Production

After fitting a GARCH model on the SARIMA residuals we get a complete description of the dynamics of X.

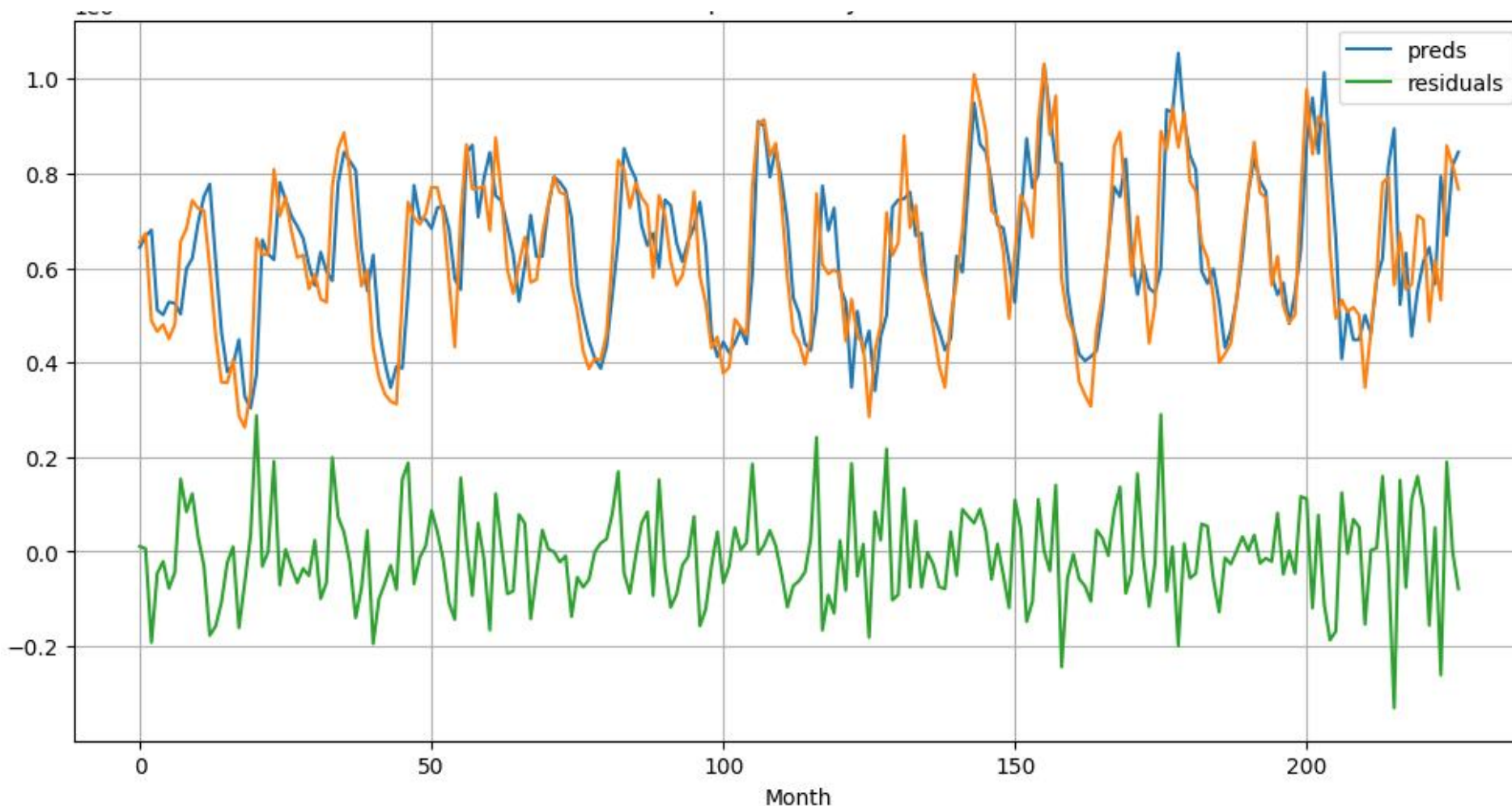
What remains to do is modeling the relationship between X and Y. To do that I extended the SARIMA model to include an exogenous input.

$$\Phi(B^s)\phi(B)\nabla_s^d\nabla^d Y_t = \Theta(B^s)\theta(B)\epsilon_t + \sum_{i=1}^k u_j X_{t-j}$$

The state space representation changes slightly as we have to take into account an additional factor dt on the observation equation.

$$\begin{cases} h_t = Fh_{t-1} + w_t & w_t \sim WN(0, V_1) \\ X_t = Hh_t + d_t \end{cases} \quad d_t = \sum_{i=1}^k u_j X_{t-j}$$

By applying again the Kalman filter given and minimizing over the the one step prediction error we can fit the model. I provided a plot of the one step predictions obtained during the calibration of the model along with the plot of the residuals..



## Forecasting

Ultimately, we can try to forecast electricity production data on longer time horizons incorporating our knowledge about the dynamics of the precipitations. My approach consisted of chosing a time horizon T and iteratively applying the following two steps:

**Algorithm:** For  $t=1, t \leq T$ , repeat:

**Step 1:** One-step prediction of the precipitations X using SARIMA-GARCH

First we apply the Kalman filter to generate the mean prediction for the SARIMA model and we simulate a random error with our calibrated GARCH model. Then we add the two contributions and take the  $\exp()$  function (to invert our initial transformation) to get a simulation for  $X(t+1)$  given the past X data.

**Step 2:** One-step prediction of the electricity production Y using SARIMAX

Given the simulated X value we can use it as an input for the SARIMAX model, adding it to the current data.

In this way we can generate an arbitrary number of trajectories for Y which can be then used to compute confidence intervals.

Below I plotted forecast performance for both X and Y along with 95% confidence intervals.

