



Tutorial 4: Generalized Linear Models
Decision Sciences & Systems (DSS)
Department of Informatics
TU München





Agenda

- 1. Generalized Linear Models
- 2. Logistic Regression
- 3. Poisson Regression
- 4. Maximum Likelihood Estimation
- 5. Evaluation and Goodness-of-Fit





Generalized Linear Models

- GLMs are a general class of linear models
- Consist of three components:
- Random: Identifies dependent variable μ and probability distribution
- Systematic: Identifies the set of explanatory variables $(X_1, ..., X_k)$
- **Link function:** Identifies function of μ that is linear

$$g(\mu) = \beta_0 + \beta_1 X_1 + \dots + \beta_k X_k$$

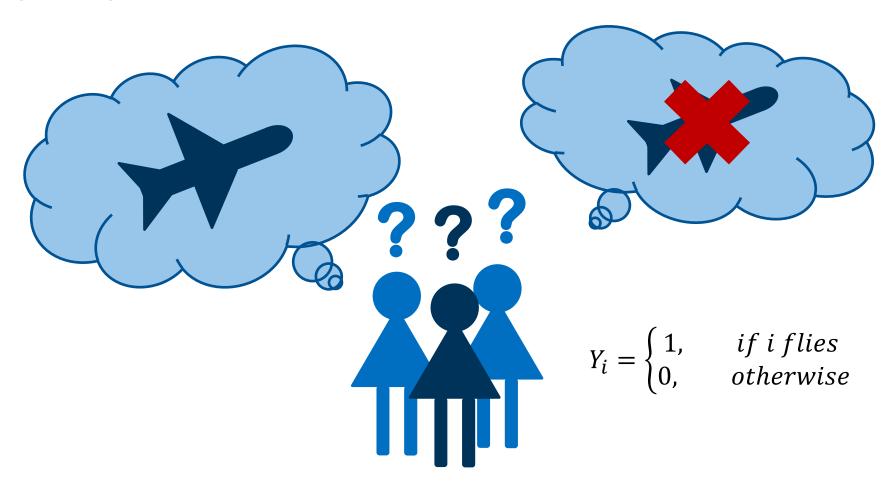
Example: Linear regression uses identity link $(g(\mu) = \mu)$

Question: Which link function could be useful for a binary dependent variable?





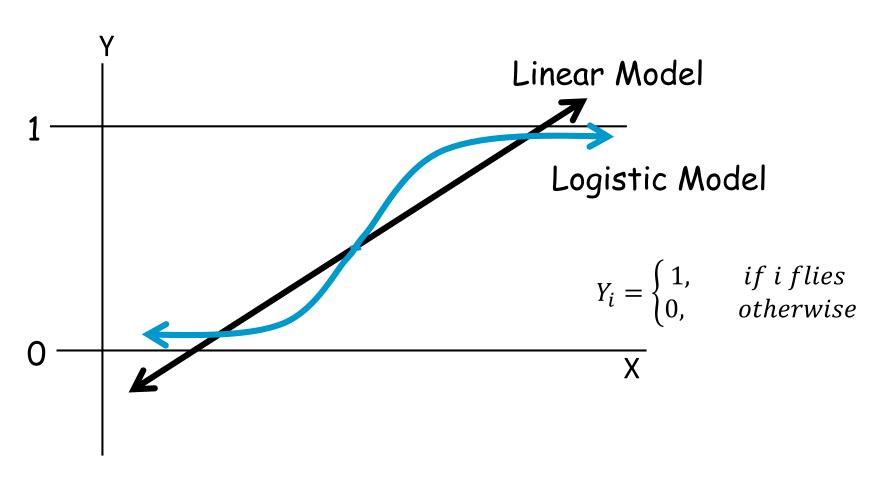
Logistic Regression - Motivation







Logistic Regression - Motivation







From Logistic Function to Logit

Logistic Function:

Logistic Regression:

$$p(x_i) = \frac{e^{x_i'\beta}}{1 + e^{x_i'\beta}}$$

transform ...

Logit:

$$\ln(\frac{p(x_i)}{1 - p(x_i)}) = x_i'\beta$$

$$\Leftrightarrow \frac{p(x_i)}{1-p(x_i)} = e^{x_i}$$

odds

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Interpreting the coefficient of logistic regression

$$x_{ij} \in x_i$$
:

$$\ln(\frac{p(x_i)}{1 - p(x_i)}) = x_i'\beta$$

$$(x_{ij}+1) \in \tilde{x}_i$$
:

$$\ln(\frac{p(\tilde{x}_i)}{1 - p(\tilde{x}_i)}) = \tilde{x}_i'\beta$$

$$\ln(\frac{p(\tilde{x}_i)}{1 - p(\tilde{x}_i)}) - \ln(\frac{p(x_i)}{1 - p(x_i)}) = \tilde{x}_i'\beta - x_i'\beta = \beta_j$$

$$\Leftrightarrow e^{\beta j} = \frac{\frac{p(x_i)}{1 - p(\tilde{x}_i)}}{\frac{p(x_i)}{1 - p(x_i)}}$$

odds ratio





Summary: Interpreting the coefficient of logistic regression

Effect of change in x_{ij} :

on log-odds (A), odds (B) and probability (C)

$$\Delta x_{ij} = 1 > 0$$

$$\Rightarrow \qquad \Delta \ln(\frac{p(x_i)}{1 - p(x_i)}) = \ln(\frac{p(\tilde{x}_i)}{1 - p(\tilde{x}_i)}) - \ln(\frac{p(x_i)}{1 - p(x_i)}) = \beta_j \tag{A}$$

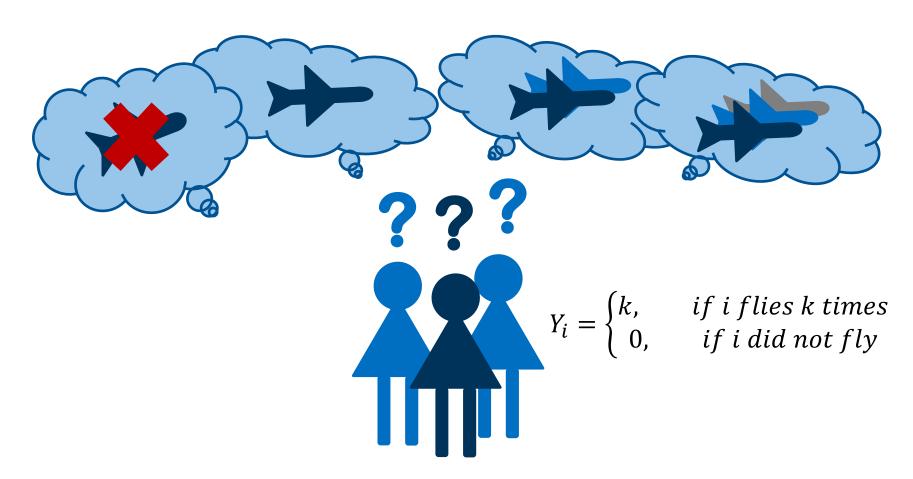
$$\Leftrightarrow \qquad e^{\beta_j} = \frac{\frac{p(\widetilde{x}_i)}{1 - p(\widetilde{x}_i)}}{\frac{p(x_i)}{1 - p(x_i)}} \tag{B), (C)}$$

β_j	$ln(\frac{p}{1-p})$ (A)	$\frac{p}{1-p}$ (B)	p (C)
$\beta_j > 0$	increases by β_j	increases by a factor of e^{eta_j}	Magnitude of increase unknown
$\beta_j < 0$	decreases by β_j	decreases by a factor of e^{eta_j}	Magnitude of decrease unknown





Poisson Regression – Motivation







From Incidence Rate to Link Function

$$\mu(x) = e^{x_i'\beta}$$

transform ...

$$\ln(\mu(x)) = x_i'\beta$$

$$\ln(\mu(x)) = x_i'\beta + \varepsilon_i$$





Interpreting the coefficient of poisson regression

$$x_{ij} \in x_i$$
:

$$\ln(\mu(x_i)) = x_i'\beta$$

$$(x_{ij}+1) \in \tilde{x}_i$$
:

$$\ln(\mu(\tilde{x}_i)) = \tilde{x}_i'\beta$$

$$\ln(\mu(\tilde{x}_i)) - \ln(\mu(x_i)) = \tilde{x}_i'\beta - x_i'\beta = \beta_j$$

$$\Leftrightarrow \qquad \beta_j = \ln(\frac{\mu(\tilde{x}_i)}{\mu(x_i)})$$

$$\Leftrightarrow \qquad e^{\beta j} = \frac{\mu(\tilde{x}_i)}{\mu(x_i)}$$

incidence rate ratio





Summary: Interpreting the coefficient of poisson regression

Effect of change in x_{ij} :

on log-incidence rate (A), incidence rate (B)

$$\Delta x_{ij} = 1 > 0$$

$$\Rightarrow \qquad \Delta \ln(\mu(x_i)) = \ln(\mu(\tilde{x}_i)) - \ln(\mu(x_i)) = \beta_j$$

$$\Leftrightarrow \qquad e^{\beta_j} = \frac{\mu(\tilde{x}_i)}{\mu(x_i)}$$
(B)

$oldsymbol{eta}_j$	$ln(\mu(x_i))$ (A)	$\mu(x_i)$ (B)
$\beta_j > 0$	increases by β_j	increases by a factor of e^{eta_j}
$\beta_j < 0$	decreases by β_j	decreases by a factor of e^{eta_j}





Maximum Likelihood Estimation

Goal: Maximize the joint probability of observing the set of dependent variables of the random sample

- Logistic regression: $L = \prod_{i=1}^n p^{y_i} (1-p)^{1-y_i}$ with $p = \frac{e^{X\beta}}{1+e^{X\beta}}$
- Poisson regression: $L = \prod_{i=1}^{n} p$ with $p = \frac{e^{X\beta y}}{y!} e^{-e^{X\beta}}$

Use numerical algorithm to find the maximum → gradient ascent

```
\begin{array}{l} k \ = \ 1, \ \text{feasible start point} \ \beta^{(1)} \in \mathbb{R}^n, \ \text{parameter} \ \varepsilon > 0 \\ \text{While} \ ( \ \left\| \nabla L(\beta^{(k)}) \right\| \geq \varepsilon \ ) \ \{ \\ \quad \bullet \ \ \text{Choose step size} \ \alpha > 0 \\ \quad \bullet \ \ \text{Set} \ \beta^{(k+1)} = \beta^{(k)} + \alpha^* \nabla L\big(\beta^{(k)}\big) \\ \quad \bullet \ k + + \\ \} \end{array}
```





Evaluation and Goodness-of-Fit

- Null deviance: -2ln(L(null))
- Residual deviance: −2 ln(L(fitted))
- McFadden R²:

$$R_{McFadden}^{2} = 1 - \frac{LL(fitted)}{LL(null)}$$

Likelihood ratio test: Does fitted model explain significantly more variance than null model?

$$D = -2\ln\left(\frac{L(null)}{L(fitted)}\right) = -2(LL(null) - LL(fitted))$$

Wald test: Is a particular coefficient significant?

$$H_0: \beta_i = 0$$





Exemplary R Output

```
> model <- glm( diabetes ~ glucose + mass + age, data = diabData, family = binomial)</pre>
> summary(model)
Call:
glm(formula = diabetes ~ glucose + mass + age, family = binomial,
    data = diabData)
Deviance Residuals:
    Min
              1Q Median
                                       Max
-2.6030 -0.6666 -0.3815 0.6765
                                   2.4804
Coefficients:
             Estimate Std. Error z value Pr(>|z|)
(Intercept) -9.677346  1.041873  -9.288  < 2e-16 ***
glucose
            0.036266 0.004906
                                 7.391 1.45e-13 ***
mass
            0.077860
                     0.020120
                                 3.870 0.000109 ***
                      0.013236
age
            0.054075
                                 4.085 4.40e-05 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
    Null deviance: 498.10 on 391 degrees of freedom
Residual deviance: 354.37 on 388 degrees of freedom
AIC: 362.37
Number of Fisher Scoring iterations: 5
```





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