Plotting marginal effects and computing DAME: Examples with Stata Margins

Supplement to: Zhirnov, Andrei, Mert Moral, and Evgeny Sedashov. 2023. "Taking Distributions Seriously: On the Interpretation of the Estimates of Interactive Nonlinear Models." Political Analysis 31(2): 213-234.

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About this document

This document shows how to generate plots with distribution-weighted average marginal effects (DAME) and heatmaps and contour plots for the marginal effects in Stata. We use the same examples as Zhirnov, Moral, and Sedashov (2023), but include a simpler Stata code that relies on the *margins* module.

Obtaining the datasets

We will need four datasets, which are part of the replication materials of the studies we replicate. These datasets are publicly available, and in this section, we show how you can obtain them.

Pre-electoral coalition formation (G)

The dataset was created by Sona Golder and used in her book (Golder 2006) on electoral alliances. It can be found on Matt Golder's website at http://mattgolder.com/files/interactions/interaction3.zip/ under the name "interaction3.dta".

Voter registration rules and turnout (N)

The dataset was originally used by Nagler (1991) and made public by William D. Berry, Jacqueline H. R. DeMeritt, and Justin Esarey as part of the replication materials for Berry, DeMeritt, and Esarey (2010). The file can be downloaded from https://jdemeritt.weebly.com/uploads/2/2/7/7/22771764/bde.zip; it is named as "scobit.dta".

News media and party discipline (AJLW)

The dataset is part of the published replication materials for Arceneaux et al. (2016) and can be downloaded from the Harvard Dataverse (https://dataverse.harvard.edu/dataset.xhtml?persistentId=doi:10.7910/DVN/27597). We use the original, Stata version of the "FoxNews_Master.tab" file.

Foreign direct investment and labor protest (RT)

The dataset is part of the published replication materials for Robertson and Teitelbaum (2011) and can be downloaded from Emmanuel Teitelbaum's website https://home.gwu.edu/~ejt/pages/Data.html, or directly from https://home.gwu.edu/~ejt/pages/Data_files/Robertson%20Teitelbaum%202011.dta.

Pre-electoral coalition formation (G)

Golder (2006) looks into the determinants of the pre-electoral coalition formation. One of the hypotheses is that "party system polarization increases the likelihood of pre-electoral coalitions when the electoral system is sufficiently disproportional" (p. 87). Using a dataset of party dyads in examined elections (G), the study estimates a model that predicts whether a particular party dyad enters the same alliance and includes polarization, effective electoral threshold (as a measure of the restrictiveness of the electoral system), and their interaction term as the predictors.

Load the data and estimate the main model

This analysis uses a random-effects probit model with an interaction term, which can be estimated using Stata's default xtprobit command:

```
clear all
use interaction3.dta,clear

xtprobit pec c.polarization##c.threshold c.seatshare##c.seatshare incompatibility ///
c.asymmetry##c.seatshare, re i(ident)
```

note: seatshare omitted because of collinearity.

Fitting comparison model:

```
Iteration 0: \log likelihood = -749.41197
Iteration 1: \log likelihood = -683.54241
```

Iteration 2: log likelihood = -681.29839Iteration 3: log likelihood = -681.2926log likelihood = -681.2926Iteration 4: Fitting full model: log likelihood = -681.2926rho = 0.0rho = 0.1 $log\ likelihood = -646.84249$ rho = 0.2 $log\ likelihood = -637.69168$ $log\ likelihood = -637.02544$ rho = 0.3rho = 0.4log likelihood = -640.3162Iteration 0: log likelihood = -636.85364Iteration 1: $log\ likelihood = -626.69411$ Iteration 2: log likelihood = -625.68214Iteration 3: log likelihood = -625.67261log likelihood = -625.67261 (backed up) Iteration 4: Random-effects probit regression Number of obs = 3,495Number of groups = Group variable: ident Random effects u_i ~ Gaussian Obs per group: min =1 avg = 12.6 max = 55 Integration method: mvaghermite Integration pts. = 12 Wald chi2(8) = 85.48 Log likelihood = -625.67261Prob > chi2 = 0.0000pec | Coefficient Std. err. z P>|z| [95% conf. interval] -.0138971 polarization | -.003196 .0054598 -0.59 0.558 .0075051 threshold | .0191695 .0108315 1.77 0.077 -.0020599 .0403989 c. | polarization#| c.threshold | .0005275 .0002845 1.85 0.064 -.00003 .0010851 seatshare | .052641 .0114296 4.61 0.000 .0302394 .0750426 c.seatshare#| .0001095 -5.28 -.0007928 c.seatshare | -.0005782 0.000 -.0003636 0.003 -2.99 incompatib~y | -.0075775 .0025334 -.0125428 -.0026122 -.0710249 . 2988559 -0.24 0.812 asymmetry | -.6567716 .5147218 seatshare | 0 (omitted) c.asymmetry#| c.seatshare | -.0301315 .0088792 -3.39 0.001 -.0475344 -.0127286 - 1

_cons | -2.381516 .3134815

-7.60 0.000 -2.995928 -1.767103

```
.24457
    /lnsig2u | -.5068768
                                                            -.9862252
                                                                        -.0275283
     sigma_u |
                  .7761276
                             .0949088
                                                             .6107225
                                                                          .9863301
         rho l
                   .375926
                             .0573775
                                                             .2716583
                                                                          .4931184
LR test of rho=0: chibar2(01) = 111.24
                                                         Prob >= chibar2 = 0.000
```

Before we can proceed to computing marginal effects, we need to trim the dataset to keep only the used observations and save it as a file.

```
keep if e(sample)
save temp, replace
```

Plotting marginal effects of polarization

At the end, we would like to have a canvas filled with colors that reflect the estimated marginal effect of polarization for the combinations of values of polarization and electoral threshold. To do this, we need to

- create a grid of equally spaced values of these variables,
- find the marginal effects that correspond to their combinations, and
- let Stata interpolate the values between them.

As the first step, we compute 16 equally spaced values for each of the constitutive terms:

```
use temp,clear
foreach v of varlist polarization threshold {
    qui sum `v'
    local `v'_s "`r(min)'(`=(r(max)-r(min))/15')`r(max)'"
}
```

Local macro polarization_s contains a sequence of 16 equally spaced values from the smallest value of polarization to its largest value. Local macro $threshold_s$ contains an equivalent sequence for threshold. We can feed these values into a call to a margins routine using the at() option:

```
use temp,clear
foreach v of varlist polarization threshold {
    qui sum `v'
    local `v'_s "`r(min)'(`=(r(max)-r(min))/15')`r(max)'"
}
margins, dydx(polarization) at(polarization=(`polarization_s') threshold=(`threshold_s') ///
(mean) _all) saving(temp_bg,replace)
```

We set all other covariates to their means. dydx(polarization) tells Stata to compute the marginal effects of the predicted value of the dependent variable with respect to *polarization*, and saving(temp_bg,replace) saves the estimates as an external file $temp_bg.dta$.

In addition to the canvas, we want to be able to show the markers with clusters of the values of these two variables (along with the significance of marginal effects at those points). We can use a similar approach to generate a 15x15 grid, bin observations within its cells, and compute the marginal effects at the centroids of those cells. The binning can be achieved using egen newvar=cut(oldvar), at(numrange), and the calculations of marginal effects within each cell can be done using the over() option in margins.

```
foreach v of varlist polarization threshold {
    qui sum `v'
    egen `v'_r = cut(`v'), at(`=r(min)-0.5'(`=(1+r(max)-r(min))/15')`=r(max)+0.5')
}
egen group_id=group(polarization_r threshold_r)
```

```
margins, dydx(polarization) over(group_id) at((omean) _all ///
(mean) polarization threshold) saving(temp_me,replace)
```

Here we set *polarization* and *threshold* to their means for each of the bin indicated by *group_id*; all other variables are set to their overall means.

We collapse the dataset in memory to determine the size and the coordinates for the binned observations, merge in the dataset with the estimates of the marginal effects for each of the bin (temp_me.dta), and append the dataset with the values to be used for the background color (temp_bq.dta).

```
collapse (count) obs=pec (mean) polarization threshold, by(group_id)
rename group_id _by1 /* group_id is called _by1 in the margins-generated dataset */
merge 1:1 _by1 using temp_me, nogenerate keepusing(_margin _pvalue)
gen significant=(_pvalue>=0.975)|(_pvalue<=0.025)</pre>
append using temp_bg, keep(_at? _margin)
rename _margin me_est
/* the values of the covariates are included as at1, at2, at3, etc.
  in the margins-generated dataset. The labels of these variables correspond
  to the names of the corresponding variables in the original data */
foreach v of varlist polarization threshold {
   foreach a of varlist _at? {
        local 1: variable label `a'
        if "'l'"=="'v'" {
            replace `v'=`a' if mi(`v')
            break
        }
   }
}
```

Now, use twoway contourplot to create a contour plot showing marginal effects of polarization (me_est) on two dimensions: one showing the values of polarization and another showing the values of threshold. We might want to add extra observations to anchor the scatter sizes. Additional observations 2 and 3, as defined below, will correspond to the unplotted filled and hollow markers of the same size, equal to the size of the largest marker of either type. Additional observations 1 and 4 correspond to the unplotted filled and hollow markers of the size equal to that of the smallest marker of either type.

```
gen counter=_n
qui sum counter
loc coreobs=r(max)
set obs `=`coreobs'+4'

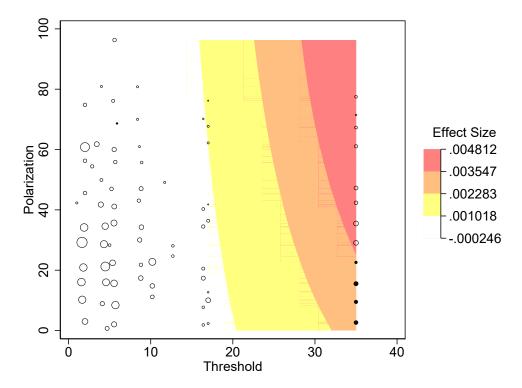
qui sum obs
replace significant=1 in `=`coreobs'+1'/`=`coreobs'+2'
replace significant=0 in `=`coreobs'+3'/`=`coreobs'+4'
replace obs=r(min) in `=`coreobs'+1'/`=`coreobs'+4'
replace obs=r(max) in `=`coreobs'+2'/`=`coreobs'+3'

/* break the ME values into steps */
qui sum me_est
matrix mimx = (r(min), r(max))*(4,3,2,1,0\0,1,2,3,4)/4
local ccuts = "`=mimx[1,2]' `=mimx[1,3]' `=mimx[1,4]'"
local zlabs = "`=mimx[1,1]' `ccuts' `=mimx[1,5]'"

local colr = "white*0.5 yellow*0.5 orange*0.5 red*0.5"
```

```
/* color ramp: from less intense to more intense colors */

twoway ///
  (contour me_est polarization threshold if !mi(me_est), ccuts(`ccuts') ccolors(`colr')) ///
  (scatter polarization threshold [fw=obs] if significant==0, ///
    msymbol(oh) mlcolor(black%95) mlwidth(vthin) msize(*.25)) ///
  (scatter polarization threshold [fw=obs] if significant==1, ///
    msymbol(o) mfcolor(black%95) mlwidth(none) msize(*.25)), ///
    xtitle(Threshold) ytitle(Polarization) ztitle("") zlabel(`zlabs') ///
    legend(off) ///
    clegend(title("Effect Size", size(medsmall) pos(12) justification(right)) ///
        width(5) height(25))
```

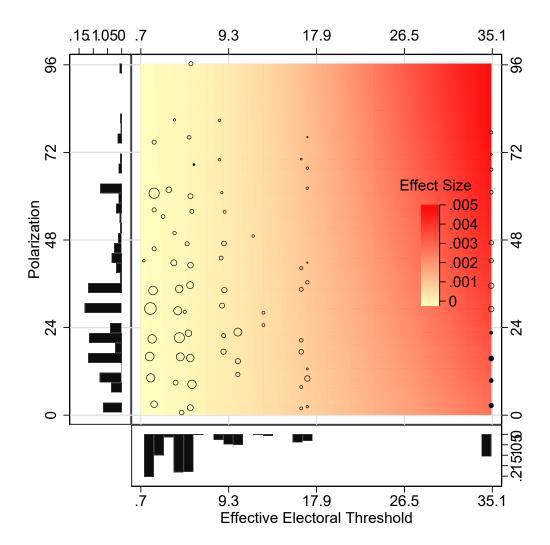


To create a heatmap instead of a contour plot, we can use the crule option instead of ccuts in the twoway contour syntax.

```
/* color ramp: from less intense to more intense colors */
loc scolr="yellow*.25" /* starting color */
loc ecolr="red*.95" /* color at the high end */

/* Specify 5 equally spaced rounded values for labels */
local nra = 5-1
foreach v of varlist polarization threshold {
    qui sum `v'
local s = (r(max)-r(min))/`nra'
local r = 10^(floor(ln(`s')/ln(10))-1)
local s = round(`s',`r')
local from = round(r(min),`r')
local to = `from' + `nra'*`s'
```

```
local `v'_r = "`from'(`s')`to'"
}
twoway (contour me_est polarization threshold if me_est!=., ///
      levels(100) crule(linear) scolor(`scolr') ecolor(`ecolr') zlab(#5, labsize(medsmall))) ///
   (scatter polarization threshold [fw=obs] if significant==0, ///
      msymbol(oh) mlcolor(black%95) mlwidth(vthin) msize(*.25)) ///
   (scatter polarization threshold [fw=obs] if significant==1, ///
      msymbol(o) mfcolor(black%95) mlwidth(none) msize(*.25)), ///
   xsca(alt) ysca(alt) xtitle("") ytitle("") ztitle("") ///
  ylab(`polarization_r', grid gmax labsize(medsmall)) ///
  xlab(`threshold_r', labsize(medsmall) grid gmax) ///
   legend(off) clegend(title("Effect Size", size(medsmall) pos(12) justification(right)) ///
   ring(0) width(5) height(25)) name(yx, replace)
/* to add histograms */
twoway histogram threshold [fw=obs], frac ysca(alt reverse) ///
   xtitle("Effective Electoral Threshold", size(medsmall)) ytitle("") ///
   xlab(`threshold_r') ylab(#4, nogrid labsize(medsmall)) ///
  fysize(20) fcolor(black%95) lwidth(vthin) lcolor(white%25) nodraw name(hy, replace)
twoway histogram polarization [fw=obs], frac xsca(alt reverse) ///
  horiz ytitle("Polarization", size(medsmall)) xtitle("") ///
   ylab(`polarization_r') xlab(#4, nogrid labsize(medsmall)) ///
   fxsize(20) fcolor(black%95) lwidth(vthin) lcolor(white%25) nodraw name(hx, replace)
gr combine hx yx hy, hole(3) imargin(zero) scale(1.1) xsize(5.5) ysize(5.5)
```



Computing and plotting DAME of polarization

To compute DAME, we first need to break the dataset into bins according to the values of the conditioning variable (threshold). We use xtile to create a variable with 10 groups of observations of approximately equal size.

```
use temp,clear
xtile group_id = threshold, nq(10)
```

Marginal effects within each bin are computed and averaged using over() option of margins:

```
margins, dydx(polarization) over(group_id) saving(temp_dame, replace)
```

To compute the marginal effects of polarization at its mean, we make a list of 21 equally spaced values spanning from the smallest to the largest value of threshold and feed it into the at() argument:

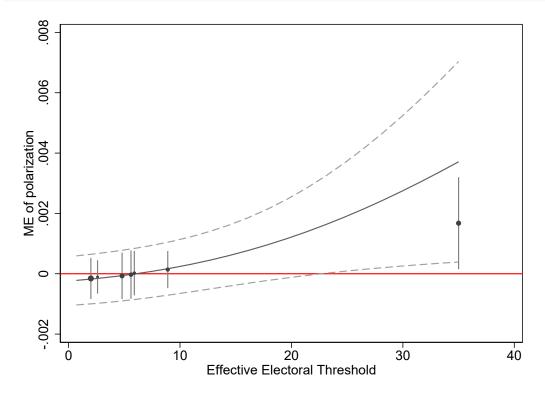
```
qui sum threshold loc cuts="\ensuremath{\text{`=r(min)'(`=(r(max)-r(min))/20')`=r(max)'"}} margins, dydx(polarization) at(threshold=(`cuts') (mean) _all) saving(temp_mem, replace)
```

Collapse the dataset in memory to find the midpoints for plotting DAME and the number of observations. Merge the dataset with DAME and append the dataset with the marginal effects at means:

```
collapse (median) threshold (count) obs=threshold, by(group_id)
rename group_id _by1
merge 1:1 _by1 using temp_dame, nogenerate keepusing(_margin _ci_lb _ci_ub)
rename (_margin _ci_lb _ci_ub)(dame lb ub)
append using temp_mem, keep(_at? _margin _ci_lb _ci_ub)
foreach v of varlist threshold {
   foreach a of varlist _at? {
      local 1: variable label `a'
      if "`l'"=="`v'" {
        replace `v'=`a' if mi(`v')
            break
      }
   }
}
rename (_margin _ci_lb _ci_ub)(mem lbm ubm)
```

Make a plot:

```
twoway (line mem threshold, lpattern(solid)) ///
(rline lbm ubm threshold, lpattern(dash)) ///
(rspike lb ub threshold) ///
(scatter dame threshold [fw=obs], msymbol(o) msize(*.25)), ///
yline(0, lcolor(red)) ytitle("ME of polarization") xtitle("Effective Electoral Threshold") ///
    legend(off)
```



Voter registration rules and turnout (N)

Nagler (1991) examines the interactive effect of education and the restrictiveness of electoral registration rules on turnout. The latter is measured using the number of days before the election when the registration closes

(closing). This variable takes on the value of 0 when the voters are allowed to register on the election day, and 30 means that the registration closes 30 days before the election day. The education is captured with an 8-category variable (neweduc), which is treated as a quantitative variable. The model also includes the squared value of neweduc. The main hypothesis is conditional: more restrictive registration rules primarily hurt less educated individuals.

Load the data and estimate the model

This expectation is captured using a probit model including the interaction terms of closing and neweduc, and closing and the squared term of education:

```
clear all
use scobit.dta,clear
drop if newvote==-1
probit newvote c.closing##c.neweduc##c.neweduc c.age##c.age ib(freq).south ib(freq).gov
```

(0 observations deleted)

```
Iteration 0: log likelihood = -63205.249
Iteration 1: log likelihood = -55865.033
Iteration 2: log likelihood = -55815.316
Iteration 3: log likelihood = -55815.275
Iteration 4: log likelihood = -55815.275
```

| Probit regression | Number of obs | = | 99,676 |
|-----------------------------|---------------|---|----------|
| | LR chi2(9) | = | 14779.95 |
| | Prob > chi2 | = | 0.0000 |
| Log likelihood = -55815.275 | Pseudo R2 | = | 0.1169 |

| newvote | Coefficient | Std. err. | z | P> z | [95% conf. | interval] |
|-----------------|-------------|-----------|--------|-------|------------|-----------|
| closing | .0006238 | .0037082 | 0.17 | 0.866 | 006644 | .0078917 |
| neweduc | . 2645073 | .041586 | 6.36 | 0.000 | .1830003 | .3460144 |
| c.closing# | | | | | | |
| c.neweduc | 0031723 | .0014993 | -2.12 | 0.034 | 0061108 | 0002337 |
| c.neweduc# | | | | | | |
| c.neweduc | .0050701 | .0041567 | 1.22 | 0.223 | 0030769 | .0132171 |
| c.closing# | | | | | | |
| c.neweduc# | | | | | | |
| c.neweduc | .0002773 | .0001504 | 1.84 | 0.065 | 0000175 | .0005722 |
| age | .0696593 | .0013051 | 53.37 | 0.000 | .0671013 | .0722172 |
| c.age#c.age | 0005061 | .0000134 | -37.71 | 0.000 | 0005324 | 0004798 |
| 1.south | 11548 | .0109736 | -10.52 | 0.000 | 1369877 | 0939722 |
| 1.gov | .0034307 | .0116343 | 0.29 | 0.768 | 0193722 | .0262335 |
| _cons | -2.743104 | .1073858 | -25.54 | 0.000 | -2.953576 | -2.532631 |

To make it easier to handle dummy variables, we explicitly declare them factors and set its mode as the

baseline category.

As earlier, we need to trim the dataset in memory to keep only the used observations and save it as a new data file.

```
keep if e(sample)
save temp, replace
```

Plotting marginal effects of voter registration rules

Find points for the background:

```
** points for the background
foreach v of varlist closing neweduc {
    qui sum `v'
    local `v'_s "(`r(min)'(`=(r(max)-r(min))/15')`r(max)')"
}
margins, dydx(closing) ///
    at(closing=`closing_s' neweduc=`neweduc_s' (mean) _continuous (base) _factor) ///
    saving(temp_bg,replace)
```

All continuous variables (except those on the axes) are set to their means; all factor variables are set to their baseline categories (their modes as specified during the estimation of model parameters).

Find points for the scatterplot:

```
foreach v of varlist closing neweduc {
    qui sum `v'
    egen `v'_r = cut(`v'), at(`=r(min)-0.5'(`=(1+r(max)-r(min))/15')`=r(max)+0.5')
}
egen group_id=group(polarization_r threshold_r)
margins, dydx(closing) over(group_id) ///
    at((omean) _continuous (base) _factor (mean) closing (median) neweduc) ///
    saving(temp_me,replace)
```

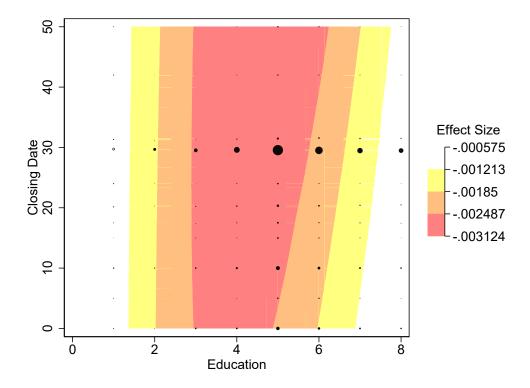
Combine information:

```
collapse (count) obs=closing (mean) closing neweduc, by(group_id)

rename group_id _by1
merge 1:1 _by1 using temp_me, nogenerate keepusing(_margin _pvalue)
gen significant=(_pvalue>=0.975)|(_pvalue<=0.025)
append using temp_bg, keep(_at? _margin)
rename _margin me_est
foreach v of varlist closing neweduc {
   foreach a of varlist _at? {
      local l: variable label `a'
      if "`l'"=="`v'" {
        replace `v'=`a' if mi(`v')
            break
      }
   }
}</pre>
```

Now, create a contour plot with the marginal effects of the registration rules. We might want to add extra observations to anchor the sizes of filled and hollow markers to the same largest and smallest values.

```
gen counter=_n
qui sum counter
loc coreobs=r(max)
set obs `=`coreobs'+4'
qui sum obs
replace significant=1 in `=`coreobs'+1'/`=`coreobs'+2'
replace significant=0 in `=`coreobs'+3'/`=`coreobs'+4'
replace obs=r(min) in `=`coreobs'+1'/`=`coreobs'+4'
replace obs=r(max) in `=`coreobs'+2'/`=`coreobs'+3'
/* Break the ME values into steps */
qui sum me_est
matrix mimx = (r(min), r(max))*(4,3,2,1,0,0,1,2,3,4)/4
local ccuts = "\"=mimx[1,2]' \"=mimx[1,3]' \"=mimx[1,4]'"
local zlabs = "`=mimx[1,1]' `ccuts' `=mimx[1,5]'"
local colr = "red*.5 orange*.5 yellow*.5 white*.5"
/* Color ramp from more intense to less intense colors */
/* Note that the replication do file uses additional graphical parameters,
which leads to different axis and legend labels from this minimal example. */
twoway (contour me_est closing neweduc if !mi(me_est), ///
 ccuts(`ccuts') ccolors(`colr')) ///
(scatter closing neweduc [fw=obs] if significant==0, ///
 msymbol(oh) mlcolor(black%95) mlwidth(vthin) msize(*.25)) ///
(scatter closing neweduc [fw=obs] if significant==1, ///
  msymbol(o) mfcolor(black%95) mlwidth(none) msize(*.25)), ///
xtitle(Education) ytitle(Closing Date) ztitle("") zlabel(`zlabs') ///
 legend(off) clegend(title("Effect Size", ///
  size(medsmall) pos(12) justification(right)) width(5) height(25))
```



Computing and plotting DAME of the restrictiveness of electoral registration rules

Find average effects by group:

```
use temp,clear
gen group_id = round(neweduc)
margins, dydx(closing) over(group_id) saving(temp_dame, replace)
```

Find marginal effects at means: set the factor variables to their baseline categories (modes) and all continuous variables (except those undere consideration) to their means.

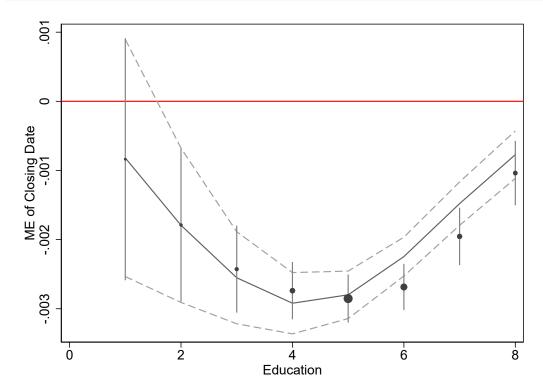
Collapse the dataset in memory, combine information:

```
collapse (median) neweduc (count) obs=neweduc, by(group_id)
rename group_id _by1
merge 1:1 _by1 using temp_dame, nogenerate keepusing(_margin _ci_lb _ci_ub)
rename (_margin _ci_lb _ci_ub)(dame lb ub)
append using temp_mem, keep(_at? _margin _ci_lb _ci_ub)
foreach v of varlist neweduc {
   foreach a of varlist _at? {
      local l: variable label `a'
      if "`l'"=="`v'" {
        replace `v'=`a' if mi(`v')
            break
   }
```

```
}
rename (_margin _ci_lb _ci_ub)(mem lbm ubm)
```

Produce a plot:

```
twoway (line mem neweduc, lpattern(solid)) ///
(rline lbm ubm neweduc, lpattern(dash)) ///
(rspike lb ub neweduc) ///
(scatter dame neweduc [fw=obs], msymbol(o) msize(*.25)), ///
yline(0, lcolor(red)) ytitle("ME of Closing Date") xtitle("Education") legend(off)
```



News media and party discipline (AJLW)

Arceneaux et al. (2016) view Congressmen as facing a choice between voting with the party and more closely following the preferences of their constituencies. Politicians are expected to vote with their party when they can afford to do so when the next election is sufficiently far, their district is safe, or they can shape the public opinion in their districts. The article presents several hypotheses about the interactive effects of these factors, but in the interest of space, we focus here only on the interaction between competitiveness and the proximity of elections.

Load the data and estimate the model

To examine the interactive effect of district competitiveness and the proximity of the next election, following Arceneaux et al. (2016), we subset the data and examine the Republicans from the districts in which Fox News was present (this is another explanatory variable in the authors' study) and build a model explaining whether a Congressman voted with their party and including a full interaction term of the polynomial of the number of days until the next election and the Democratic vote share in the preceding election in the representative's district (the larger this share, the more competitive the district). Before estimation, to make

interpretation easier, we combine dummy variables for the type of the vote into a single variable with multiple categories. We designate the modes as the baseline levels for the respective factor variables.

```
use "FoxNews_Master.dta",clear
gen dvprop=dv/100
gen vtype_char = "OtherProc" if !mi(RegPass)
foreach v of varlist RegPass Amend OtherPass ProPart Susp {
   replace vtype_char="`v'" if `v'
}
encode vtype_char,gen(votetype)
logit PartyVote c.daystoelection##c.daystoelection##c.daystoelection##c.dvprop seniorit ///
 spendgap_lag spendgap distpart_lag ib(freq).votetype ib(freq).Retirement ///
 ib(freq).qualchal_lag ib(freq).qualchal ///
 if PresencePartyUnity==1 & Republican==1 & FoxNews==1, cluster(dist2)
(2,080 missing values generated)
(68,326 real changes made)
(140,176 real changes made)
(44,820 real changes made)
(77,180 real changes made)
(60,427 real changes made)
Iteration 0:
             log pseudolikelihood = -21866.053
Iteration 1:
             log pseudolikelihood = -19635.905
Iteration 2:
             log pseudolikelihood = -19349.274
             log pseudolikelihood = -19345.854
Iteration 3:
             log pseudolikelihood = -19345.849
Iteration 4:
Iteration 5:
             log pseudolikelihood = -19345.849
Logistic regression
                                                    Number of obs = 57,086
                                                    Wald chi2(18) =
                                                    Prob > chi2 =
Log pseudolikelihood = -19345.849
                                                    Pseudo R2 = 0.1153
                               (Std. err. adjusted for 73 clusters in dist2)
                          _____
          Robust
  PartyVote | Coefficient std. err. z P>|z| [95% conf. interval]
daystoelec~n | -.0053586 .006298 -0.85 0.395 -.0177025 .0069852
         c. |
daystoelec~n#|
         c. l
               .0000157 .0000199 0.79 0.430 -.0000233
                                                                  .0000548
daystoelec~n |
         c. |
daystoelec~n#|
         c. |
```

| daystoelec~n# | | | | | | |
|----------------------------|-----------|----------|-------|-------|---------------------|----------|
| c. daystoelec~n | -1.50e-08 | 1.92e-08 | -0.78 | 0.435 | -5.27e-08 | 2.27e-08 |
| daybuddidd ii | 1.000 00 | 1.020 00 | 0.70 | 0.100 | 0.270 00 | 2.270 00 |
| dvprop | .998585 | 2.279037 | 0.44 | 0.661 | -3.468246 | 5.465416 |
| ļ. | | | | | | |
| c. daystoelec~n# | | | | | | |
| c.dvprop | | .0181061 | 0.61 | 0.542 | 0244367 | .0465377 |
| | .011000 | .010101 | 0.02 | 0.012 | | 701000 |
| c. | | | | | | |
| daystoelec~n# | | | | | | |
| c. | | | | | | |
| daystoelec~n# c.dvprop | 0000448 | .0000578 | -0.77 | 0.439 | 000158 | .0000685 |
| c.uvprop | .0000440 | .0000376 | 0.11 | 0.400 | .000150 | .0000000 |
| c. | | | | | | |
| daystoelec~n# | | | | | | |
| c. | | | | | | |
| daystoelec~n# | | | | | | |
| c. daystoelec~n# | | | | | | |
| c.dvprop | | 5.64e-08 | 1.02 | 0.307 | -5.30e-08 | 1.68e-07 |
| | | | | | | |
| seniorit | 0340132 | .0238695 | -1.42 | 0.154 | 0807966 | .0127702 |
| spendgap_lag | | .0408697 | 0.41 | 0.683 | 06341 | .0967961 |
| spendgap | | .0422487 | -1.13 | 0.259 | 1305121 3.977493 | .0350996 |
| distpart_lag | 6.657421 | 1.367335 | 4.87 | 0.000 | 3.977493 | 9.337348 |
| votetype | | | | | | |
| OtherPass | 1.261922 | .0825187 | 15.29 | 0.000 | 1.100188 | 1.423656 |
| OtherProc | 2.20396 | .0958879 | 22.98 | 0.000 | 2.016023 | 2.391897 |
| ProPart | | .0515924 | 38.47 | 0.000 | 1.883711 | 2.085949 |
| RegPass | | .0730834 | 13.35 | 0.000 | .8325338 | 1.119015 |
| Susp | 1.159763 | .1306688 | 8.88 | 0.000 | .9036568 | 1.415869 |
| 1.Retirement | 1.015441 | .1738517 | 5.84 | 0.000 | .6746975 | 1.356184 |
| 1.qualchal~g | . 2339622 | .175486 | 1.33 | 0.182 | 1099841 | .5779085 |
| 1.qualchal | | .1762658 | -0.92 | 0.356 | 5082568 | .1826924 |
| _cons | -2.277705 | 1.202385 | -1.89 | 0.058 | -4.634337 | .0789273 |

As earlier, we need to trim the dataset in Stata's memory to keep only the observations in the estimation sample and save it as a new data file.

```
keep if e(sample)
save temp, replace
```

Plotting marginal effects of election proximity

Find points for the backgound. All continuous variables (except the main ones) are set to their means, and all factor variables are set to their baseline categories.

```
foreach v of varlist daystoelection dvprop {
   qui sum `v'
```

```
local `v'_s "(`r(min)'(`=(r(max)-r(min))/15')`r(max)')"
}
margins, dydx(daystoelection) at(daystoelection=`daystoelection_s' dvprop=`dvprop_s' ///
    (mean) _continuous (base) _factor) saving(temp_bg,replace)
```

Find points for the foreground. All continuous variables (except the main ones) are set to their global means, and all factor variables are set to their baseline categories. The variables under consideration are set to their means within groups.

```
foreach v of varlist daystoelection dvprop {
    qui sum `v'
    egen `v'_r = cut(`v'), at(`=r(min)-0.5'(`=(1+r(max)-r(min))/15')`=r(max)+0.5')
}
egen group_id=group(daystoelection_r dvprop_r)
margins, dydx(daystoelection) over(group_id) at((omean) _continuous ///
    (base) _factor (mean) daystoelection dvprop) saving(temp_me,replace)
```

Combine information

```
collapse (count) obs=daystoelection (mean) daystoelection dvprop, by(group_id)

rename group_id _by1
merge 1:1 _by1 using temp_me, nogenerate keepusing(_margin _pvalue)
gen significant=(_pvalue>=0.975)|(_pvalue<=0.025)
append using temp_bg, keep(_at? _margin)
rename _margin me_est
foreach v of varlist daystoelection dvprop {
   foreach a of varlist _at? {
      local 1: variable label `a'
      if "`l'"=="`v'" {
        replace `v'=`a' if mi(`v')
        break
      }
   }
}</pre>
```

Now, create a contourplot. We might want to add extra observations to anchor the sizes of filled and hollow markers to a common scale. It is advisable to specify the color ramp in a way that more intense colors represent negative and positive values of higher magnitude.

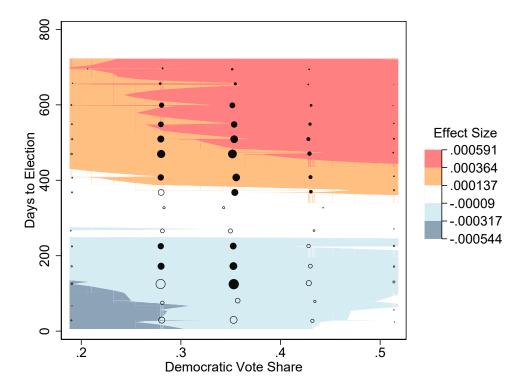
```
gen counter=_n
qui sum counter
loc coreobs=r(max)
set obs `=`coreobs'+4'

qui sum obs
replace significant=1 in `=`coreobs'+1'/`=`coreobs'+2'
replace significant=0 in `=`coreobs'+3'/`=`coreobs'+4'
replace obs=r(min) in `=`coreobs'+1'/`=`coreobs'+4'
replace obs=r(max) in `=`coreobs'+2'/`=`coreobs'+3'

* Break the ME values into steps
qui sum me_est
matrix mimx = (r(min), r(max))*(5,4,3,2,1,0\0,1,2,3,4,5)/5
local ccuts = "`=mimx[1,2]' `=mimx[1,3]' `=mimx[1,4]' `=mimx[1,5]'"
local zlabs = "`=mimx[1,1]' `ccuts' `=mimx[1,6]'"
```

```
loc colr= "navy*.5 ltblue*.5 white*.5 orange*.5 red*.5"
/* Color ramp has intense colors at both ends */

/* Note that the replication do file uses additional graphical parameters,
    which leads to different axis and legend labels from this minimal example. */
twoway (contour me_est daystoelection dvprop if me_est!=., ccuts(`ccuts') ccolors(`colr')) ///
(scatter daystoelection dvprop [fw=obs] if significant==0, ///
    msymbol(oh) mlcolor(black%95) mlwidth(vthin) msize(*.25)) ///
(scatter daystoelection dvprop [fw=obs] if significant==1, ///
    msymbol(o) mfcolor(black%95) mlwidth(none) msize(*.25)), ///
xtitle(Democratic Vote Share) ytitle(Days to Election) ///
    ztitle("") zlabel(`zlabs') ///
    legend(off) ///
    clegend(title(`"Effect Size"', size(medsmall) pos(12) justification(right)) width(5) height(25))
```



Computing and plotting DAME of election proximity

To compute the distribution-weighted average marginal effects, we first need to break the dataset into bins according to the values of the conditioning variable (the Democratic vote share, dvprop). Here, we use deciles:

```
use temp,clear
xtile group_id = dvprop, nq(10)
margins, dydx(daystoelection) over(group_id) vce(unconditional) saving(temp_dame, replace)
```

Compute the marginal effect of the proximity of the next election at its mean, 21 equally-spaced values of the Democratic vote share, and the means or modes of all other covariates:

```
qui sum dvprop
loc cuts="\[ =r(min)'(\[ =(r(max)-r(min))/20')\[ =r(max)'"\]
margins, dydx(daystoelection) at(dvprop=(\[ cuts') (mean) _continuous (base) _factor) ///
```

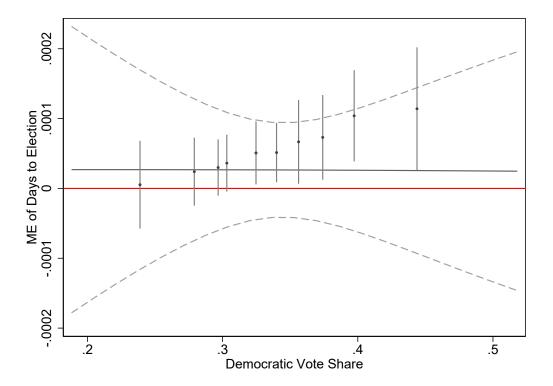
```
vce(unconditional) saving(temp_mem, replace)
```

Combine data:

```
collapse (median) dvprop (count) obs=dvprop, by(group_id)
rename group_id _by1
merge 1:1 _by1 using temp_dame, nogenerate keepusing(_margin _ci_lb _ci_ub)
rename (_margin _ci_lb _ci_ub)(dame lb ub)
append using temp_mem, keep(_at? _margin _ci_lb _ci_ub)
foreach v of varlist dvprop {
    foreach a of varlist _at? {
        local l: variable label `a'
        if "`l'"=="`v'" {
            replace `v'=`a' if mi(`v')
            break
        }
    }
}
rename (_margin _ci_lb _ci_ub)(mem lbm ubm)
```

Produce a plot:

```
twoway (line mem dvprop, lpattern(solid)) ///
(rline lbm ubm dvprop, lpattern(dash)) ///
(rspike lb ub dvprop) ///
(scatter dame dvprop [fw=obs], msymbol(o) msize(*.25)), ///
yline(0, lcolor(red)) ytitle("ME of Days to Election") xtitle("Democratic Vote Share") ///
  legend(off)
```



Foreign direct investment and labor protest (RT)

Robertson and Teitelbaum (2011) study the response of the local labor to foreign direct investment. The article argues that FDI flows lead to more labor protests, and this effect is more substantial when there are fewer democratic means for resolving such conflicts.

Load the data and estimate the model

Since the dependent variable is a count of protests, we use a negative binomial regression (with random effects, as in the original study). The right-hand side of the model equation includes an interaction of political regime (Polity 2 score) and the natural log of FDI flows: the latter's effect is expected to be conditional on the values of the former. Both variables are lagged.

```
clear all
use "Robertson Teitelbaum 2011.dta", clear
tsset country year
gen l_l_flows=L.l_flows
gen l_polity2=L.polity2
gen l_dispute=L.dispute
gen l_demflows=l_l_flows*l_polity2
xtnbreg dispute c.l_l_flows##c.l_polity2 l_dispute open_penn l_gdp_pc_penn gdp_grth ///
  inflation_1 urban xratchg l_pop time, re
keep if e(sample)
Panel variable: country (strongly balanced)
Time variable: year, 1979 to 2006
        Delta: 1 unit
(882 missing values generated)
(635 missing values generated)
(138 missing values generated)
(1,074 missing values generated)
Fitting negative binomial (constant dispersion) model:
Iteration 0:
              log likelihood = -19743.067 (not concave)
Iteration 1:
              log likelihood = -19693.709 (not concave)
Iteration 2:
              log likelihood = -19669.166 (not concave)
              log likelihood = -19419.284 (not concave)
Iteration 3:
              log likelihood = -13856.193 (not concave)
Iteration 4:
Iteration 5:
              log likelihood = -11966.287 (not concave)
Iteration 6:
              log likelihood = -11310.599
                                            (not concave)
              log likelihood = -10199.071
Iteration 7:
Iteration 8:
              log likelihood = -9186.4985
                                            (backed up)
                                            (backed up)
Iteration 9:
              log likelihood = -6216.4714
Iteration 10: log likelihood = -3115.8101
                                            (backed up)
Iteration 11: log likelihood = -2711.2173
Iteration 12: log likelihood = -1142.8628
Iteration 13: log likelihood = -1049.5449
```

```
Iteration 14: log likelihood = -994.70282
Iteration 15: log likelihood = -994.26758
Iteration 16: log likelihood = -994.26733
Iteration 17: log likelihood = -994.26733
Iteration 0:
            \log likelihood = -1138.5175
Iteration 1: log likelihood = -1048.9393
Iteration 2:
             log\ likelihood = -1043.1845
Iteration 3:
             log\ likelihood = -1043.1752
Iteration 4:
             log likelihood = -1043.1752
Iteration 0:
             log likelihood = -1043.1752
                                       (not concave)
Iteration 1:
             log likelihood = -974.09467
             log likelihood = -939.21383
                                       (not concave)
Iteration 2:
Iteration 3:
             log likelihood = -819.68488
             log likelihood = -786.2915
Iteration 4:
             log likelihood = -781.86279
Iteration 5:
Iteration 6:
             log\ likelihood = -781.78374
Iteration 7:
             log\ likelihood = -781.7837
Fitting full model:
Iteration 0:
             log likelihood = -768.40601
Iteration 1:
             log\ likelihood = -724.05204
Iteration 2:
             log\ likelihood = -718.95793
Iteration 3:
             log\ likelihood = -718.92085
Iteration 4:
             log likelihood = -718.92082
Random-effects negative binomial regression
                                               Number of obs = 2,348
Group variable: country
                                               Number of groups =
                                                                  131
Random effects u_i ~ Beta
                                               Obs per group:
                                                          min =
                                                          avg = 17.9
                                                          max =
                                                                   25
                                               Wald chi2(12) = 162.80
Log likelihood = -718.92082
                                               Prob > chi2
                                                            = 0.0000
______
   dispute | Coefficient Std. err. z P>|z|
                                                   [95% conf. interval]
______
              .3956329 .0843602
                                                    .23029
  l l flows |
                                   4.69 0.000
                                                              .5609758
               .2741761 .0597606 4.59 0.000 .1570476
  1_polity2 |
                                                              .3913047
c.l_l_flows#|
c.l_polity2 |
             -.0323768
                         .0087734
                                   -3.69
                                          0.000
                                                -.0495724
                                                             -.0151813
                         .0192369
  l_dispute |
              .0594588
                                   3.09
                                         0.002
                                                   .0217551
                                                              .0971625
                                   0.21
  open_penn |
              .0007637
                         .0036998
                                          0.836
                                                  -.0064878
                                                              .0080152
             -.1925234
                                   -0.82
                                         0.410
                                                  -.6506558
                                                              .2656089
l_gdp_pc_p~n |
                        .2337453
   gdp_grth |
                                  -1.84 0.066
                                                -.0591221
             -.028622
                        .0155615
                                                              .001878
                        .0000684
inflation_1 |
             .0001256
                                   1.84 0.066
                                                  -8.48e-06
                                                              .0002597
                                   2.36 0.018
      urban |
              .0219114
                        .0092731
                                                  .0037364
                                                              .0400863
```

```
-.0044268
   xratchg |
           .0031663 .0038741 0.82 0.414
                                                     .0107595
                              3.41 0.001 .1654766
     l_pop | .3889047 .113996
                                                   .6123328
                              1.98 0.048
                                            .000325
     time | .0322821 .0163049
                                                     .0642392
     _cons | -10.22589 2.576873 -3.97 0.000
                                         -15.27647 -5.175317
  ______
                                            .5605223
     /ln r |
             1.07164
                    .2607792
                                                   1.582758
    /ln s | -.2429911 .3254917
                                           -.8809432 .3949609
        r |
            2.920165
                    .7615182
                                            1.751587
                                                     4.868364
        s | .7842785 .2552761
                                            .4143919
                                                     1.484326
LR test vs. pooled: chibar2(01) = 125.73
                                         Prob >= chibar2 = 0.000
```

(1,516 observations deleted)

As earlier, we need to trim the dataset in memory to keep only the observations in the estimation sample and save it as a new data file.

```
keep if e(sample)
save temp, replace
```

Plotting marginal effects of logged FDI flows

Compute values for the background:

```
foreach v of varlist l_l_flows l_polity2 {
    qui sum `v'
   local `v'_s "(`r(min)'(`=(r(max)-r(min))/15')`r(max)')"
}
margins, dydx(l_l_flows) at(l_l_flows=`l_l_flows_s' l_polity2=`l_polity2_s' (mean) _all) ///
   saving(temp_bg,replace)
```

We bin the values of logged flows to avoid overplotting later on.

```
foreach v of varlist l_l_flows l_polity2 {
   qui sum `v'
    egen v' r = cut(v'), at(=r(min)-0.5'(=(1+r(max)-r(min))/10')=r(max)+0.5')
egen group_id=group(l_l_flows_r l_polity2_r)
margins, dydx(l_l_flows) over(group_id) at((omean) _all (mean) l_l_flows l_polity2) ///
  saving(temp_me,replace)
```

Combine information:

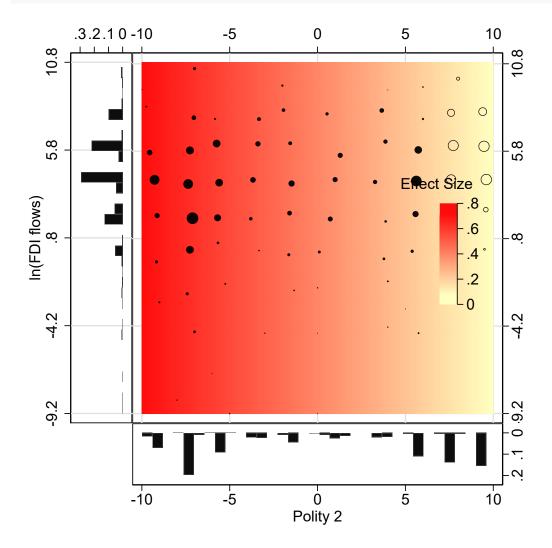
```
collapse (count) obs=1_1_flows (mean) 1_1_flows 1_polity2, by(group_id)
rename group id by1
merge 1:1 _by1 using temp_me, nogenerate keepusing(_margin _pvalue)
gen significant=(_pvalue>=0.975)|(_pvalue<=0.025)</pre>
append using temp_bg, keep(_at? _margin)
rename _margin me_est
foreach v of varlist l_l_flows l_polity2 {
   foreach a of varlist _at? {
        local 1: variable label `a'
        if "'1'"=="'v'" {
            replace `v'=`a' if mi(`v')
            break
```

```
}
}
```

Now, create a heatmap. We might want to add extra observations to anchor the marker sizes on the scatter plot.

```
gen counter=_n
qui sum counter
loc coreobs=r(max)
set obs `=`coreobs'+4'
qui sum obs
replace significant=1 in `=`coreobs'+1'/`=`coreobs'+2'
replace significant=0 in `=`coreobs'+3'/`=`coreobs'+4'
replace obs=r(min) in `=`coreobs'+1'/`=`coreobs'+4'
replace obs=r(max) in `=`coreobs'+2'/`=`coreobs'+3'
* Color ramp from less intense to more intense colors
loc scolr="yellow*.25"
loc ecolr="red*.95"
* Labs: numlist specification for 5 equally spaced values for each variable
local nra = 5-1
foreach v of varlist l_l_flows l_polity2 {
qui sum `v'
local s = (r(max)-r(min))/nra'
local r = 10^{(floor(ln(`s')/ln(10))-1)}
local s = round(`s', `r')
local from = round(r(min), r')
local to = `from' + `nra'*`s'
local `v'_r = "`from'(`s')`to'"
/* Note that the replication do file uses additional graphical parameters,
   which leads to different axis and legend labels from this minimal example. */
twoway (contour me_est 1_1_flows 1_polity2 if me_est!=., levels(100) crule(linear) ///
   scolor(`scolr') ecolor(`ecolr') zlab(#5, labsize(medsmall))) ///
(scatter l_l_flows l_polity2 [fw=obs] if significant==0, ///
   msymbol(oh) mlcolor(black%95) mlwidth(vthin) msize(*.25)) ///
(scatter l_l_flows l_polity2 [fw=obs] if significant==1, ///
   msymbol(o) mfcolor(black%95) mlwidth(none) msize(*.25)), ///
xsca(alt) ysca(alt) xtitle("") ytitle("") ztitle("") ///
ylab(`l_l_flows_r', grid gmax labsize(medsmall)) ///
xlab(`l_polity2_r', labsize(medsmall) grid gmax) ///
clegend(title("Effect Size", size(medsmall) pos(12) justification(right)) ring(0) width(5) height(25))
legend(off) nodraw name(yx, replace)
twoway histogram l_polity2 [fw=obs], frac ysca(alt reverse) xtitle("Polity 2", size(medsmall)) ytitle("
xlab(`l_polity2_r') ylab(#4, nogrid labsize(medsmall)) ///
fysize(20) fcolor(black%95) lwidth(vthin) lcolor(white%25) nodraw name(hy, replace)
twoway histogram l_l_flows [fw=obs], frac xsca(alt reverse) horiz ytitle("ln(FDI flows)", ///
```

```
size(medsmall)) xtitle("") ///
ylab(`l_1_flows_r', grid gmax labsize(medsmall)) xlab(#4, nogrid labsize(medsmall)) ///
fxsize(20) fcolor(black%95) lwidth(vthin) lcolor(white%25) nodraw name(hx, replace)
gr combine hx yx hy, hole(3) imargin(zero) scale(1.1) xsize(5.5) ysize(5.5)
```



Computing and plotting DAME of the logged FDI flows:

To compute the distribution-weighted average marginal effects, we first need to break the dataset into 4 bins based on the values of the conditioning variable:

```
use temp,clear
xtile group_id = 1_polity2, nq(4)
margins, dydx(1_1_flows) over(group_id) saving(temp_dame, replace)
```

Compute the marginal effect of 1_1_flows at its mean:

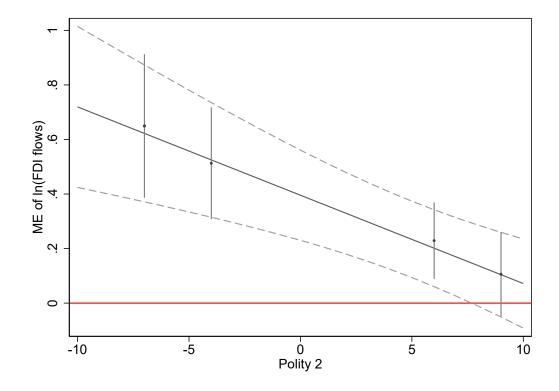
```
qui sum l_polity2
loc cuts="\[ =r(min)'(1)\[ =r(max)''' \]
margins, dydx(l_l_flows) at(l_polity2=(\[ cuts'\])) atmeans saving(temp_mem, replace)
```

Combine information:

```
collapse (median) l_polity2 (count) obs=l_polity2, by(group_id)
rename group_id _by1
merge 1:1 _by1 using temp_dame, nogenerate keepusing(_margin _ci_lb _ci_ub)
rename (_margin _ci_lb _ci_ub)(dame lb ub)
append using temp_mem, keep(_at? _margin _ci_lb _ci_ub)
foreach v of varlist l_polity2 {
    foreach a of varlist _at? {
        local l: variable label `a'
        if "`l'"=="`v'" {
            replace `v'=`a' if mi(`v')
            break
        }
    }
}
rename (_margin _ci_lb _ci_ub)(mem lbm ubm)
```

Produce a plot:

```
twoway (line mem l_polity2, lpattern(solid)) ///
(rline lbm ubm l_polity2, lpattern(dash)) ///
(rspike lb ub l_polity2) ///
(scatter dame l_polity2 [fw=obs], msymbol(o) msize(*.25)), ///
yline(0, lcolor(red)) ytitle("ME of ln(FDI flows)") xtitle("Polity 2") legend(off)
```



This document

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References

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