Plotting marginal effects and computing DAME with R: examples

Supplement to: Zhirnov, Andrei, Mert Moral, and Evgeny Sedashov. 2022. "Taking Distributions Seriously: On the Interpretation of the Estimates of Interactive Nonlinear Models." Political Analysis (Conditional Acceptance)

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Required packages

This document illustrates how one can use the DAME package to replicate the calculations presented in Zhirnov, Moral, and Sedashov (2022). The current version of the DAME package can be installed from GitHub using devtools::install_github:

```
library(devtools)
install_github("andreizhirnov/DAME")
library(DAME)
```

The version of this package used to generate this document can be installed using its v1.0.0 release:

```
install_github("andreizhirnov/DAME@v1.0.0")
```

DAME depends on stats (R Core Team 2021), methods (R Core Team 2021), and MASS (Ripley 2021) for calculations, and ggplot2 (Wickham et al. 2021) for plotting.

In this document, we also refer to the foreign (R Core Team 2020) and dataverse (Kuriwaki et al. 2021) packages to download the below-mentioned datasets, glmmTMB (Magnusson et al. 2021) to estimate multi-level models, sandwich (Zeileis and Lumley 2021) to compute robust variance-covariance matrices, and ggExtra (Attali and Baker 2019) to add histograms to heatmaps and contour-plots.

Obtaining the datasets

We will need four datasets, which are included in the replication materials of the studies we replicated. These datasets are publicly available, and in this section, we show how you can obtain them.

Pre-electoral coalition formation (G)

The dataset was created by Sona Golder and used in her book (Golder 2006) on electoral alliances. It can be found on Matt Golder's web-page at http://mattgolder.com/files/interactions/. The following script can be used to download it using R:

Voter registration rules and turnout (N)

The dataset was originally used by Nagler (1991). It was made public by William D. Berry, Jacqueline H. R. DeMeritt, and Justin Esarey as part of the replication materials for Berry, DeMeritt, and Esarey (2010). The following script can be used to download it using R:

News media and party discipline (AJLW)

The datasets is part of the published replication materials for Arceneaux et al. (2016) and can be downloaded either directly from the Harvard Dataverse https://dataverse.harvard.edu/dataset.xhtml?persistentId=doi: 10.7910/DVN/27597 or using the dataverse package:

Foreign direct investment and labor protest (RT)

The datasets is part of the published replication materials for Robertson and Teitelbaum (2011) and can be downloaded from Emmanuel Teitelbaum's website: https://home.gwu.edu/~ejt/pages/Data.html. The following script can be used to download it using R:

```
library(foreign)
dataset <-
   read.dta("https://home.gwu.edu/~ejt/pages/Data_files/Robertson%20Teitelbaum%202011.dta")
saveRDS(dataset,"RT.rds")</pre>
```

Pre-electoral coalition formation (G)

Golder (2006) looks into the determinants of the pre-electoral coalition formation. One of the hypotheses is that "party system polarization increases the likelihood of pre-electoral coalitions when the electoral system is sufficiently disproportional" (p. 87). Using a dataset of party dyads in examined elections (G), the study estimates a model that predicts whether a particular party dyad enters the same alliance and includes polarization, effective electoral threshold (as a measure of the restrictiveness of the electoral system), and their interaction term as the predictors.

Estimate the main model

The probit model with an interaction term can be estimated using stats::glm() function:

```
dt <- readRDS("G.rds")</pre>
dt <- subset(dt, complete.cases(dt[c("pec", "polarization", "threshold", "seatshare",</pre>
                                      "incompatibility", "asymmetry")]))
m <- glm(pec ~ polarization*threshold + seatshare + I(seatshare^2) + incompatibility +
           asymmetry + asymmetry:seatshare, data=dt, family=binomial(link="probit"))
summary(m)
##
## Call:
## glm(formula = pec ~ polarization * threshold + seatshare + I(seatshare^2) +
##
       incompatibility + asymmetry + asymmetry:seatshare, family = binomial(link = "probit"),
##
       data = dt)
##
## Deviance Residuals:
##
       Min
                 1Q
                      Median
                                    30
                                            Max
## -1.1265 -0.3457 -0.2690 -0.2150
                                         2.9468
##
## Coefficients:
                             Estimate Std. Error z value Pr(>|z|)
##
## (Intercept)
                           -2.1005362  0.2040356  -10.295  < 2e-16 ***
## polarization
                           -0.0014849 0.0026606
                                                  -0.558 0.576757
## threshold
                            0.0206984 0.0060720
                                                   3.409 0.000652 ***
```

```
## seatshare
                           0.0430151 0.0085270
                                                  5.045 4.55e-07 ***
## I(seatshare^2)
                          -0.0004475
                                      0.0000825
                                                 -5.425 5.81e-08 ***
## incompatibility
                          -0.0053749
                                      0.0020063
                                                 -2.679 0.007382 **
## asymmetry
                           0.0208886
                                      0.2356696
                                                  0.089 0.929372
## polarization:threshold 0.0002361
                                      0.0001449
                                                  1.629 0.103225
## seatshare:asymmetry
                                      0.0068860
                                                 -3.554 0.000380 ***
                          -0.0244723
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
##
   (Dispersion parameter for binomial family taken to be 1)
##
       Null deviance: 1464.6 on 3494
##
                                       degrees of freedom
## Residual deviance: 1341.4 on 3486
                                       degrees of freedom
## AIC: 1359.4
##
## Number of Fisher Scoring iterations: 6
```

To replicate the original analysis, this model needs to include party-specific random effects. We can include those using glmmTMB() function from the glmmTMB package. You could achieve very similar results with lme4, which, however, has more limited functionality and is a bit slower than glmmTMB for estimating the models we show in this document.

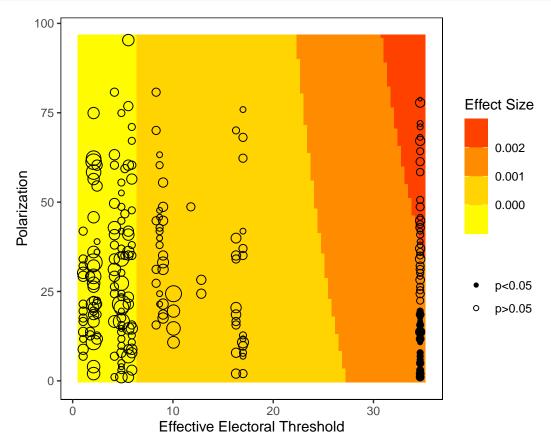
```
library(glmmTMB)
m2 <- glmmTMB(pec ~ polarization*threshold + seatshare + I(seatshare^2) + incompatibility +
                asymmetry + asymmetry:seatshare + (1|ident),
                  data=dt,
                  family=binomial(link="probit"))
summary(m2)
   Family: binomial (probit)
## Formula:
  pec ~ polarization * threshold + seatshare + I(seatshare^2) +
       incompatibility + asymmetry + asymmetry:seatshare + (1 |
##
                                                                      ident)
## Data: dt
##
##
        AIC
                 BIC
                       logLik deviance df.resid
##
     1251.3
              1312.9
                       -615.7
                                1231.3
                                           3485
##
## Random effects:
##
## Conditional model:
   Groups Name
                       Variance Std.Dev.
##
   ident (Intercept) 0.619
                                0.7868
## Number of obs: 3495, groups: ident, 278
## Conditional model:
                            Estimate Std. Error z value Pr(>|z|)
                                                -7.603 2.90e-14 ***
## (Intercept)
                          -2.4302284 0.3196440
## polarization
                          -0.0028516 0.0055633
                                                 -0.513
                                                          0.60825
                                                  1.746 0.08084 .
## threshold
                           0.0191869
                                      0.0109901
## seatshare
                           0.0511005
                                      0.0114988
                                                  4.444 8.83e-06 ***
## I(seatshare^2)
                          -0.0005592
                                      0.0001096
                                                 -5.102 3.37e-07 ***
## incompatibility
                          -0.0067494
                                                 -2.670
                                                          0.00759 **
                                      0.0025283
## asymmetry
                          -0.1017402 0.3018679
                                                 -0.337
                                                          0.73609
```

```
## polarization:threshold 0.0005055 0.0002899 1.744 0.08116 .
## seatshare:asymmetry -0.0287526 0.0089265 -3.221 0.00128 **
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

Plotting marginal effects of polarization

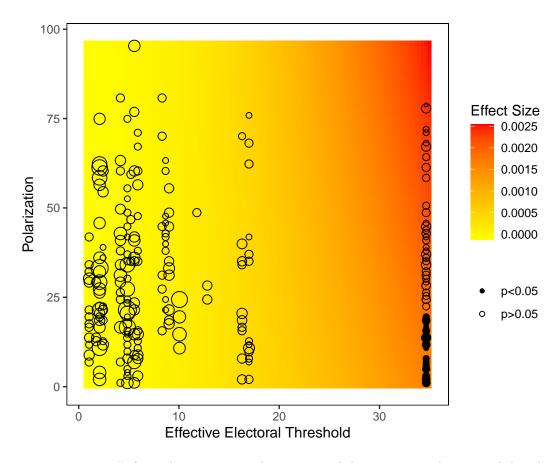
Making a contour-plot to illustrate marginal effects from a simple probit is straightforward with the plot_me() function provided that the model is estimated with a call to the glm() function. We need to specify the object with the model estimates (m), the name of the variable whose marginal effect we seek to estimate (polarization) and the conditioning variable (threshold).

```
library(ggplot2)
plot_me(model = m, x = "polarization", over = "threshold") +
   scale_fill_steps(low="yellow",high="red", n.breaks=4) +
   labs(x="Effective Electoral Threshold", y="Polarization")
```



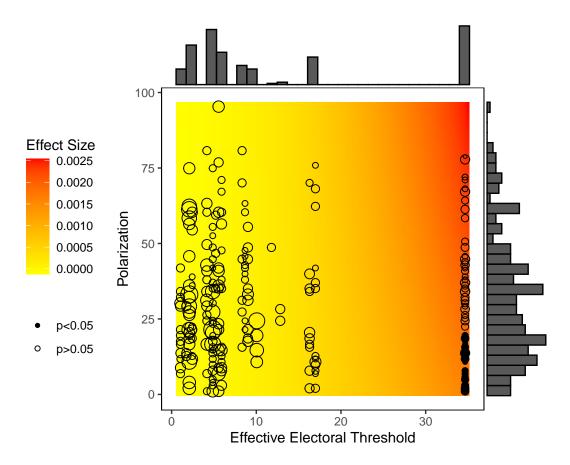
Notice that to adjust the axis labels and specify the type of the color ramp, we use the standard ggplot2 syntax. Here, scale_fill_steps() breaks down the continuous color values into the number of categories specified with the n.breaks option. To create a heatmap, we need to replace scale_fill_steps() with scale_fill_gradient():

```
(g <- plot_me(model = m, x = "polarization", over = "threshold") +
    scale_fill_gradient(low="yellow", high="red") +
    labs(x="Effective Electoral Threshold", y="Polarization"))</pre>
```



We can use ggMarginal() from the ggExtra package to attach histograms with marginal distributions of the main covariates to the heatmap:

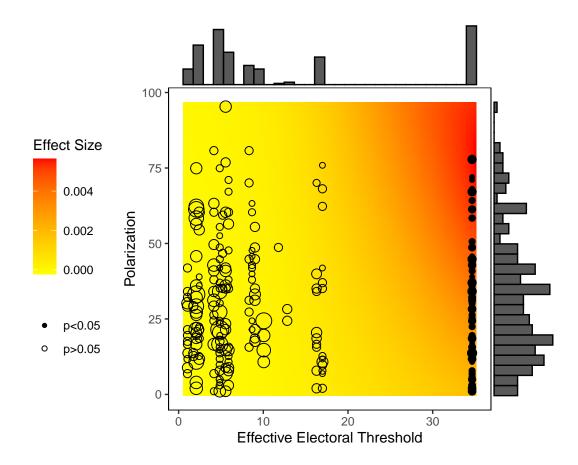
```
gt <- g + theme(legend.position="left")
ggExtra::ggMarginal(gt, type="histogram", data=dt, x=threshold, y=polarization)</pre>
```



DAME works best with model estimates from glm(). When the model estimates are produced with the lme4 or glmmTMB package, we must specify the components directly. We can extract the coefficients and variance-covariance matrix corresponding to the fixed effects using the glmmTMB::fixef() and stats::vcov() functions.

```
coef <- fixef(m2)$cond
vc <- vcov(m2)$cond</pre>
```

Since we will be setting the covariates (other than the main ones) to their central values, we can ignore the random component of the intercept, effectively setting the random effects to zero. To bring these calculations closer to what is shown in the paper, we will use simulations to produce standard errors by setting mc=TRUE:



Computing and plotting DAME of polarization

To compute the distribution-weighted average marginal effects of polarization, we can resort to the dame() function:

```
(d <- dame(model=m, x="polarization", over="threshold", nbins=10))</pre>
##
                                            1b
                                                         ub bin id
               est
## 1 -7.854570e-05 0.0001810035 -0.0004333061 0.0002762147
                                                                2.0
## 2 -5.534738e-05 0.0001512690 -0.0003518292 0.0002411345
                                                                2.6
## 3 -3.853787e-05 0.0002061198 -0.0004425253 0.0003654495
                                                                4.8
## 4 -2.078844e-05 0.0002244287 -0.0004606606 0.0004190837
                                                                5.6
     4.329523e-06 0.0002087581 -0.0004048289 0.0004134880
                                                               5.9
     6.396721e-05 0.0001744165 -0.0002778829 0.0004058173
                                                               8.9
## 7 8.938963e-04 0.0005087583 -0.0001032517 0.0018910442
                                                               35.0
```

The nbins option allows us to create 10 bins with approximately the same number of observations.

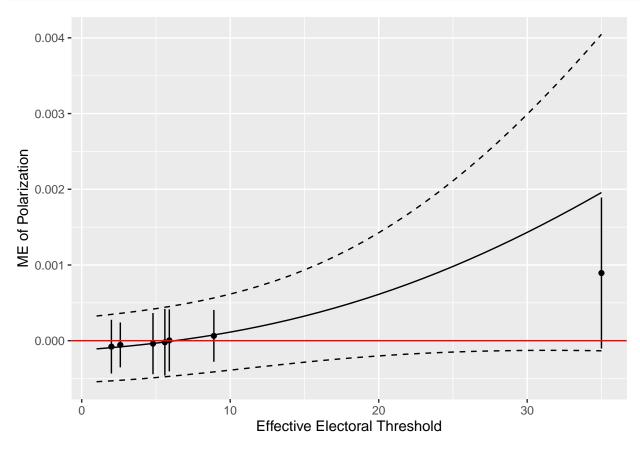
Likewise, we use mem() to compute the marginal effect of polarization at its mean:

```
mem <- mem(model=m, x="polarization", at=list("threshold"=seq(from=1,to=35, by=1)))</pre>
```

Putting both DAME and MEM on a single plot:

```
ggplot(data=d, aes(x=bin_id, y=est, ymin=lb, ymax=ub)) +
geom_point() +
geom_linerange() +
geom_line(aes(x=threshold, y=est), data=mem) +
geom_line(aes(x=threshold, y=lb), data=mem, linetype="dashed") +
```

```
geom_line(aes(x=threshold, y=ub), data=mem, linetype="dashed") +
geom_hline(yintercept=0, color="red") +
labs(x="Effective Electoral Threshold", y="ME of Polarization")
```



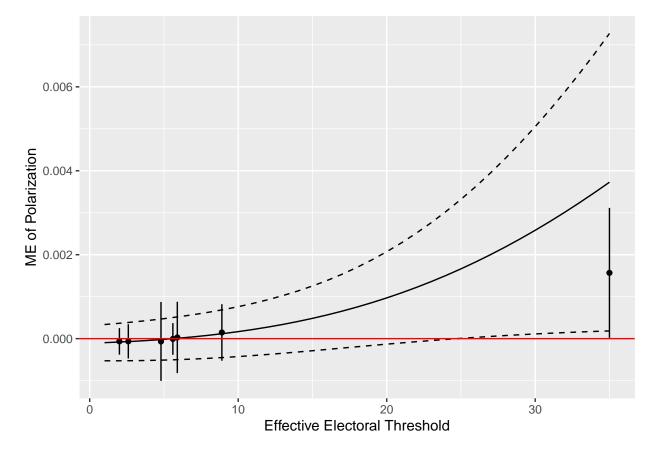
How should we deal with a model with random effects? When the estimates are produced with glmmTMB, we can easily recover the estimates for the group-specific values of the random effects (but not their sampling distribution).

```
re <- ranef(m2)$cond$ident
dt$reff <- re[as.character(dt$ident),"(Intercept)"]</pre>
```

To take into account the differential intercepts for each of the groups, we will treat random intercepts as offsets to the linear prediction and include them in the model formula:

```
x="polarization",
   at=list("threshold"=seq(from=1,to=35, by=1)))

# plot them
ggplot(data=d, aes(x=bin_id, y=est, ymin=lb, ymax=ub)) +
   geom_point() +
   geom_linerange() +
   geom_line(aes(x=threshold, y=est), data=mem) +
   geom_line(aes(x=threshold, y=lb), data=mem, linetype="dashed") +
   geom_line(aes(x=threshold, y=ub), data=mem, linetype="dashed") +
   geom_hline(yintercept=0, color="red") +
   labs(x="Effective Electoral Threshold", y="ME of Polarization")
```



Voter registration rules and turnout (N)

Nagler (1991) examines the interactive effect of education and the restrictiveness of electoral registration rules on turnout. The latter is measured using the number of days before the election when the registration closes. This variable takes on the value of 0 when the voters are allowed to register on the election day, and 30 means that the registration closes 30 days before the election day. The main hypothesis is conditional: more restrictive registration rules primarily hurt less educated individuals.

Load the data and estimate the model

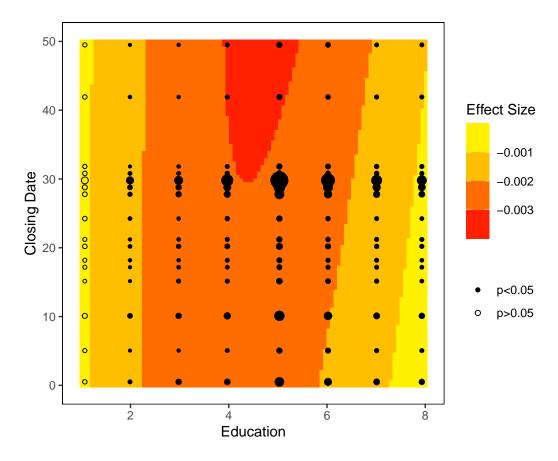
This expectation is captured using a probit model including an interaction term of closing and neweduc, as well as an intraction term of closing and neweduc squared.

```
dt <- readRDS("N.rds")</pre>
dt <- subset(dt, newvote %in% 0:1, select=c("state", "newvote", "closing", "neweduc",
                                            "age", "south", "gov"))
m <- glm(newvote ~ closing*neweduc + closing*I(neweduc^2) + age + I(age^2) + south + gov,
         data=dt,
         family=binomial(link="probit"))
summary(m)
##
## Call:
## glm(formula = newvote ~ closing * neweduc + closing * I(neweduc^2) +
       age + I(age^2) + south + gov, family = binomial(link = "probit"),
##
##
       data = dt)
##
## Deviance Residuals:
                 1Q
                      Median
                                   30
                                           Max
##
  -2.8318 -1.0797
                      0.6029
                               0.8729
                                        2.2440
##
## Coefficients:
##
                          Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                        -2.743e+00 1.077e-01 -25.479 < 2e-16 ***
## closing
                         6.243e-04 3.721e-03
                                                0.168
                                                        0.8668
## neweduc
                                                6.322 2.59e-10 ***
                         2.645e-01 4.184e-02
## I(neweduc^2)
                         5.069e-03 4.196e-03
                                                1.208
                                                        0.2270
## age
                         6.966e-02
                                    1.311e-03 53.143
                                                      < 2e-16 ***
## I(age^2)
                        -5.061e-04 1.352e-05 -37.437
                                                       < 2e-16 ***
## south
                        -1.155e-01 1.097e-02 -10.525
                                                       < 2e-16 ***
## gov
                         3.431e-03 1.162e-02
                                                0.295
                                                        0.7678
                        -3.173e-03
                                    1.509e-03
                                               -2.103
                                                        0.0355 *
## closing:neweduc
## closing:I(neweduc^2) 2.774e-04 1.518e-04
                                                1.827
                                                        0.0677 .
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
  (Dispersion parameter for binomial family taken to be 1)
##
##
       Null deviance: 126410 on 99675 degrees of freedom
## Residual deviance: 111631 on 99666 degrees of freedom
## AIC: 111651
##
## Number of Fisher Scoring iterations: 4
```

Plotting marginal effects of the registration closing date

We can use the plot_me() function to draw a contour-plot for the marginal effects of the registration closing date. Since closing can only take on integer values, it makes more sense here to use the first-difference method for calculating the marginal effect. Thus, we set discrete=TRUE and use 1-day increments.

```
plot_me(model = m, x = "closing", over = "neweduc", discrete = TRUE, discrete_step = 1) +
    scale_fill_steps(low="red", high="yellow", n.breaks=4) +
    labs(y="Closing Date", x="Education")
```

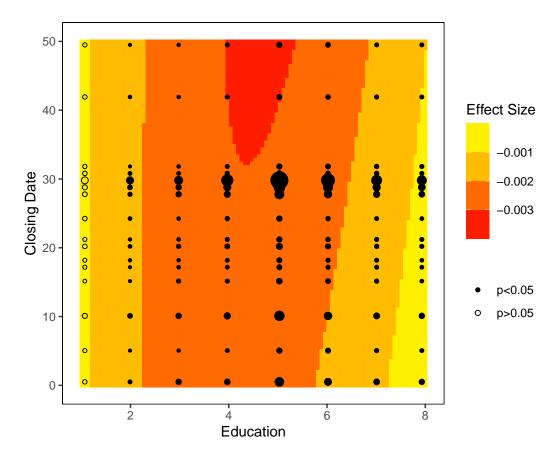


To set the dummy variables south and gov to their modes (rather than means), we can either use the at option or convert them to factors before the estimation. To use the at option, create a list with medians (which are almost always equal to modes for dummy variables):

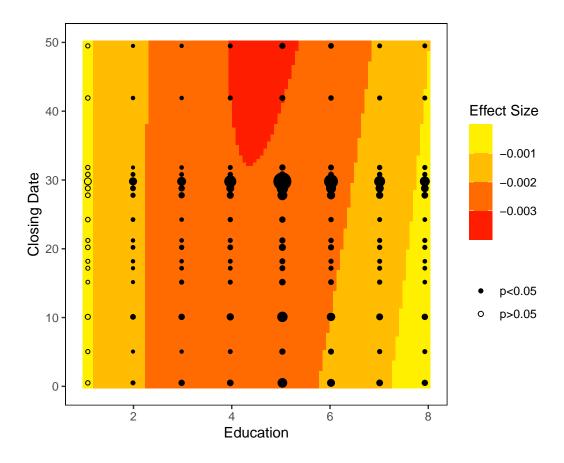
```
(medians <- lapply(dt[c("south", "gov")], median))</pre>
```

```
## $south
## [1] 0
##
## $gov
## [1] 0

plot_me(model = m, x = "closing", over = "neweduc", discrete = TRUE, discrete_step = 1, at=medians) +
    scale_fill_steps(low="red", high="yellow", n.breaks=4) +
    labs(y="Closing Date", x="Education")
```



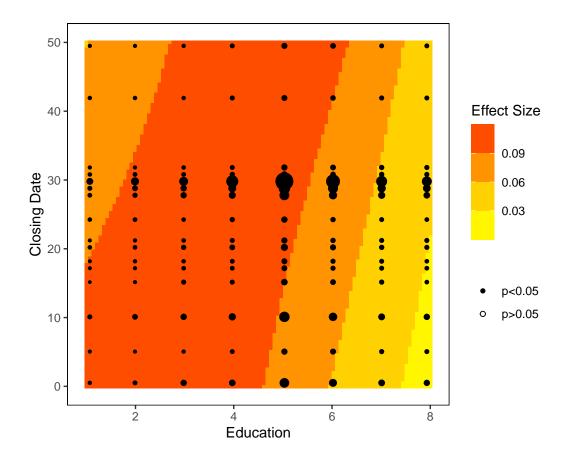
Alternatively, we could generate factor variables and use them instead of south and gov: in this case plot_me() will automatically use modes instead of means:



Plotting marginal effects of education

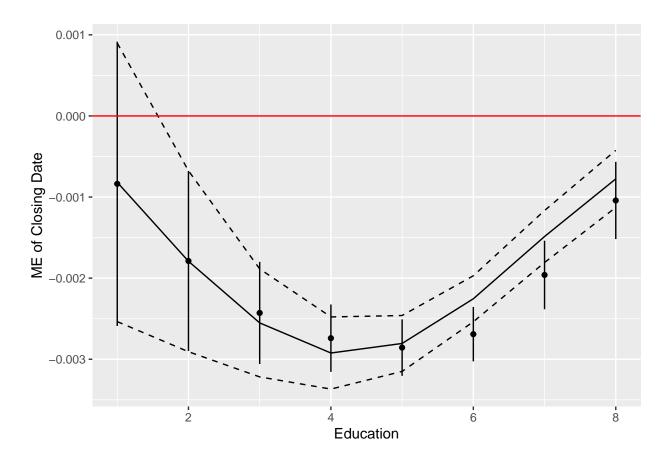
When computing the marginal effect of education, we might benefit from keeping the same orientation of axes as before. To do so, we can use the coord_flip() function from ggplot2.

```
plot_me(model = m, x ="neweduc" , over = "closing", discrete = TRUE, discrete_step = 1, at=medians) +
    scale_fill_steps(low="yellow", high="red", n.breaks=4) +
    labs(y="Education", x="Closing Date") +
    coord_flip()
```



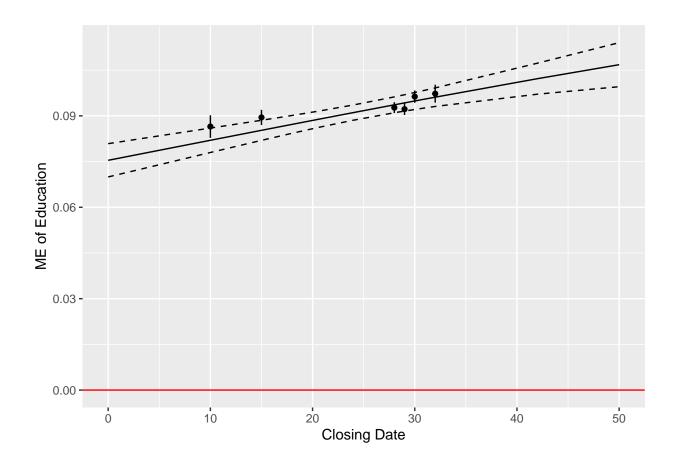
Computing and plotting DAME of the registration closing date

Turning to DAME, we can use a similar syntax as before and, since neweduc takes only 8 unique values, we can bin the observations by the unique values of this variable (use_distinct_values=TRUE).



Computing and plotting DAME of education

Change the x, over, and the values of the running variable to plot the DAME of education:



News media and party discipline (AJLW)

Arceneaux et al. (2016) view Congressmen as facing a choice between voting with the party and more closely following the preferences of their constituencies. In this theory, politicians are expected to vote with their party when they can afford to do so when the next election is sufficiently far, their district is safe, or they can shape the public opinion in their districts. The article presents several hypotheses about the interactive effects of these factors; in the interest of space, we focus here only on the interaction between the competitiveness and the proximity of elections.

Load the data and estimate the model

To examine the interactive effect of district competitiveness and the proximity of the next election, following Arceneaux et al. (2016), we subset the data and examine the Republicans from the districts in which Fox News was present (this is another explanatory variable in the authors' study) and build a model explaining whether a Congressman voted with their party and including a full interaction term of the polynomial of the number of days until the next election and the Democratic vote share in the preceding election in the representative's district (the larger this share, the more competitive the district).

Before estimating this model, we need to convert the dummy variables representing different types of votes into a single factor variable. This will make it easier to include them in the model and find their central values.

```
dt <- readRDS("AJLW.rds")
# create a factor variable with the vote type
dt$votetype <- "MiscProcedure"</pre>
```

```
for (v in c("Amend", "OtherPass", "ProPart", "RegPass", "Susp")) {
  dt$votetype[which(dt[[v]]==1)] <- v</pre>
}
dt <- subset(dt, PresencePartyUnity==1 & Republican==1 & FoxNews==1,
             select=c("id","dist2","PartyVote","daystoelection","dv","Retirement",
                      "seniorit", "qualchal_lag", "qualchal", "spendgap_lag", "spendgap",
                      "distpart_lag","votetype"))
dt <- within(dt, {
  dvprop \leftarrow dv/100
  votetype <- relevel(factor(votetype), ref="MiscProcedure")</pre>
  qualchal_lag.f <- factor(qualchal_lag, levels=0:1)</pre>
  qualchal.f <- factor(qualchal, levels=0:1)</pre>
  Retirement.f <- factor(Retirement, levels=0:1)</pre>
  })
dt <- na.omit(dt)
m <- glm(PartyVote ~ dvprop*(daystoelection + I(daystoelection^2) + I(daystoelection^3)) +
           Retirement.f + seniorit + qualchal_lag.f + qualchal.f + spendgap_lag + spendgap +
           distpart_lag + votetype,
         data=dt,
         family=binomial)
summary(m)
##
## Call:
## glm(formula = PartyVote ~ dvprop * (daystoelection + I(daystoelection^2) +
       I(daystoelection^3)) + Retirement.f + seniorit + qualchal_lag.f +
##
##
       qualchal.f + spendgap_lag + spendgap + distpart_lag + votetype,
       family = binomial, data = dt)
##
##
## Deviance Residuals:
                      Median
                                   3Q
                 10
                                           Max
                                        1.0482
## -3.0597
             0.2351
                      0.3521
                               0.6106
##
## Coefficients:
                                Estimate Std. Error z value Pr(>|z|)
##
## (Intercept)
                              -7.374e-02 4.802e-01 -0.154
                                                              0.8780
                                                               0.4246
## dvprop
                               9.986e-01 1.251e+00 0.799
## daystoelection
                              -5.359e-03 5.009e-03 -1.070
                                                              0.2847
## I(daystoelection^2)
                               1.575e-05 1.630e-05 0.966
                                                              0.3339
                              -1.502e-08 1.605e-08 -0.935
## I(daystoelection^3)
                                                              0.3496
## Retirement.f1
                              1.015e+00 2.556e-01 3.973 7.08e-05 ***
                              -3.401e-02 3.634e-03 -9.361 < 2e-16 ***
## seniorit
## qualchal_lag.f1
                               2.340e-01 3.581e-02 6.534 6.40e-11 ***
## qualchal.f1
                              -1.628e-01 3.749e-02 -4.342 1.41e-05 ***
                              1.669e-02 8.401e-03
                                                     1.987
## spendgap_lag
                                                               0.0469 *
## spendgap
                              -4.771e-02 9.783e-03 -4.877 1.08e-06 ***
                              6.657e+00 2.566e-01 25.946 < 2e-16 ***
## distpart_lag
                              -2.204e+00 9.322e-02 -23.641 < 2e-16 ***
## votetypeAmend
                              -9.420e-01 1.192e-01 -7.904 2.70e-15 ***
## votetypeOtherPass
                              -2.191e-01 1.005e-01 -2.181
## votetypeProPart
                                                               0.0292 *
## votetypeRegPass
                              -1.228e+00 1.004e-01 -12.229 < 2e-16 ***
## votetypeSusp
                              -1.044e+00 1.431e-01 -7.299 2.91e-13 ***
```

```
## dvprop:daystoelection
                               1.105e-02
                                          1.442e-02
                                                       0.767
                                                               0.4434
## dvprop:I(daystoelection^2) -4.476e-05
                                          4.710e-05
                                                      -0.950
                                                               0.3419
                                                       1.237
## dvprop:I(daystoelection^3) 5.757e-08
                                          4.655e-08
                                                               0.2161
##
## Signif. codes:
                   0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
  (Dispersion parameter for binomial family taken to be 1)
##
##
##
       Null deviance: 43732
                             on 57085
                                       degrees of freedom
## Residual deviance: 38692
                             on 57066
                                       degrees of freedom
  AIC: 38732
##
## Number of Fisher Scoring iterations: 6
```

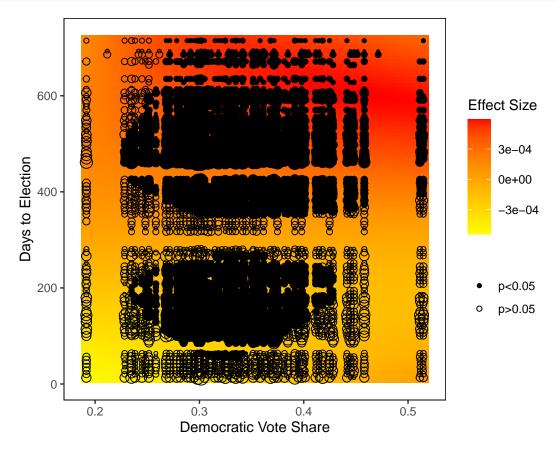
The observations are clustered by districts. Thus, we need to compute a cluster-robust variance-covariance matrix:

```
vc <- sandwich::vcovCL(m, cluster= ~ dist2 )</pre>
```

Plotting marginal effects of election proximity

We use the plot_me() function to make a heatmap for the marginal effects of election proximity.

```
plot_me(model = m, x = "daystoelection", over = "dvprop", vcov=vc) +
    scale_fill_gradient(low="yellow", high="red") +
    labs(y="Days to Election", x="Democratic Vote Share")
```

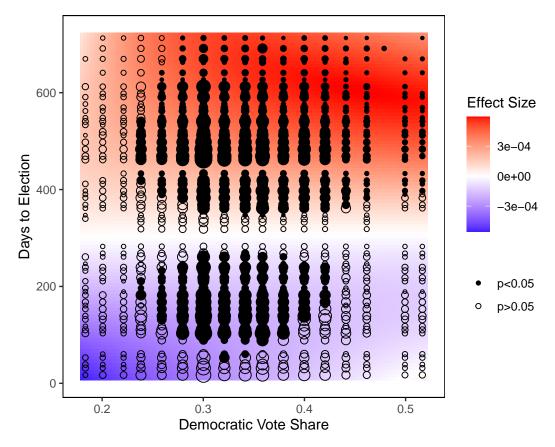


Two modifications are needed here. First, since about half of these effects are positive, we need to use a divergent palette so that the intensity would represent the magnitude of the effect, and the hue would

represent the direction. We can use the scale_fill_gradient2() function to specify the endpoints and midpoint of the color scale.

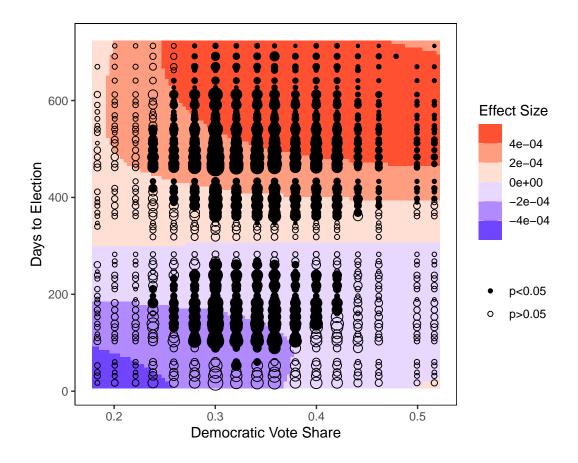
Second, to reduce the number of markers and subsequent overplotting, we can round the values shown on the axes:

```
newdt <- within(dt, {
   daystoelection <- round(daystoelection, -1) ## round to the nearest multiple of 10
   dvprop <- round(dvprop/0.02)*0.02 ## round to the nearest multiple of 0.02
})
plot_me(model = m, data=newdt, x = "daystoelection", over = "dvprop", vcov=vc) +
   scale_fill_gradient2(low="blue", mid="white", high="red") +
   labs(y="Days to Election", x="Democratic Vote Share")</pre>
```



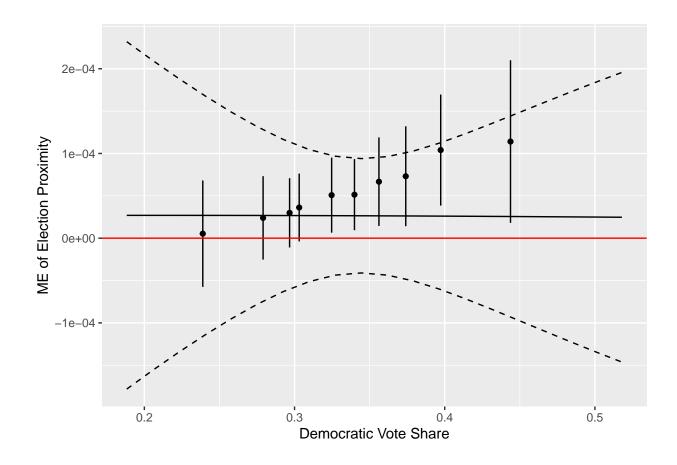
To obtain a contour-plot with a divergent scale, replace scale_fill_gradient2() with scale_fill_steps2(): plot_me(model = m, data=newdt, x = "daystoelection", over = "dvprop", vcov=vc) +

```
plot_me(model = m, data=newdt, x = "daystoelection", over = "dvprop", vcov=vc) +
    scale_fill_steps2(low="blue", mid="white", high="red", n.breaks=7) +
    labs(y="Days to Election", x="Democratic Vote Share")
```



Computing and plotting DAME of election proximity

Do not forget to specify the variance-covariance matrix when computing DAME and MEM:



Foreign direct investment and labor protest (RT)

Robertson and Teitelbaum (2011) study the response of the local labor to foreign direct investment. The article argues that FDI flows lead to more labor protests, and this effect is more substantial when there are fewer democratic means for resolving such conflicts.

Load the data and estimate the model

Since the dependent variable is a count of protests, we use a negative binomial regression. The right-hand side of the model equation includes an interaction of political regime (Polity 2 score) and the natural log of FDI flow: we expect the latter's effect to be conditional on the values of the former.

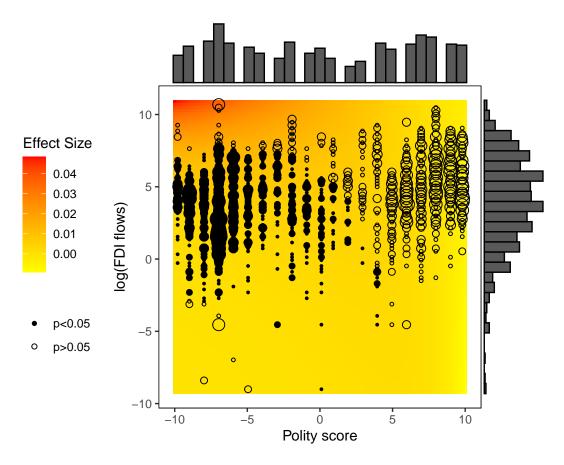
```
gdp_grth + inflation_1 + urban + xratchg + l_pop + time,
         data=dt )
summary(m)
##
## Call:
  glm.nb(formula = dispute ~ l_l_flows * l_polity2 + l_dispute +
##
       open_penn + l_gdp_pc_penn + gdp_grth + inflation_1 + urban +
##
       xratchg + l_pop + time, data = dt, init.theta = 0.3997142679,
##
       link = log)
##
## Deviance Residuals:
      Min
                 1Q
                      Median
                                   3Q
                                           Max
## -1.8660 -0.4005 -0.2286 -0.1056
                                        6.3269
##
## Coefficients:
##
                         Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                       -1.724e+01 1.950e+00 -8.839 < 2e-16 ***
## l_l_flows
                        1.508e-01 6.132e-02
                                              2.460 0.013896 *
## l_polity2
                        1.872e-01 4.030e-02
                                              4.645 3.40e-06 ***
                        3.391e-01 3.205e-02 10.580 < 2e-16 ***
## 1 dispute
## open penn
                       -5.637e-04 1.880e-03 -0.300 0.764346
## l_gdp_pc_penn
                        2.981e-01 1.829e-01
                                              1.630 0.103171
## gdp grth
                       -4.039e-04 1.755e-02 -0.023 0.981640
## inflation_1
                        9.812e-05 7.645e-05
                                              1.284 0.199305
## urban
                        2.267e-02 6.359e-03
                                               3.565 0.000364 ***
## xratchg
                       -7.718e-04 4.401e-03 -0.175 0.860797
## 1_pop
                        6.341e-01 7.802e-02
                                              8.127 4.38e-16 ***
                       -1.171e-02 1.408e-02 -0.832 0.405371
## time
## l_l_flows:l_polity2 -1.948e-02 6.104e-03 -3.192 0.001415 **
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for Negative Binomial(0.3997) family taken to be 1)
##
       Null deviance: 1501.51 on 2347 degrees of freedom
## Residual deviance: 634.28 on 2335 degrees of freedom
## AIC: 1607.2
##
## Number of Fisher Scoring iterations: 1
##
##
##
                 Theta: 0.3997
##
             Std. Err.: 0.0546
##
   2 x log-likelihood: -1579.1580
The article employs a more complicated version of this model with country-specific random effects. As earlier,
we can estimate such a model with the glmmTMB() function from the glmmTMB package.
library(glmmTMB)
m2 <- glmmTMB(dispute ~ l_l_flows*l_polity2 + l_dispute + open_penn + l_gdp_pc_penn +
                gdp_grth + inflation_1 + urban + xratchg + l_pop + time + (1|isocode),
              family=nbinom1, data=dt)
summary(m2)
```

```
## Family: nbinom1 (log)
## Formula:
## dispute ~ 1_1_flows * 1_polity2 + 1_dispute + open_penn + 1_gdp_pc_penn +
##
       gdp_grth + inflation_1 + urban + xratchg + l_pop + time +
##
       (1 | isocode)
## Data: dt
##
##
        AIC
                BIC
                       logLik deviance df.resid
##
     1484.7
              1571.1
                       -727.3
                               1454.7
                                           2333
##
## Random effects:
##
## Conditional model:
   Groups Name
                        Variance Std.Dev.
##
   isocode (Intercept) 1.576
                                 1.255
## Number of obs: 2348, groups:
                                isocode, 131
## Dispersion parameter for nbinom1 family (): 1.19
##
## Conditional model:
##
                         Estimate Std. Error z value Pr(>|z|)
## (Intercept)
                       -1.744e+01 3.236e+00 -5.390 7.07e-08 ***
## 1_1_flows
                        3.382e-01 8.686e-02
                                              3.894 9.87e-05 ***
## 1 polity2
                        2.551e-01 6.007e-02
                                              4.247 2.17e-05 ***
                       5.996e-02 1.648e-02
                                              3.637 0.000275 ***
## l_dispute
## open_penn
                       -1.445e-03 3.568e-03 -0.405 0.685382
## l_gdp_pc_penn
                       9.334e-02 2.741e-01
                                              0.340 0.733487
                       -2.570e-02 1.626e-02
                                            -1.580 0.114100
## gdp_grth
## inflation_1
                        1.147e-04 6.852e-05
                                              1.674 0.094076
## urban
                        1.692e-02 1.094e-02
                                              1.546 0.122083
## xratchg
                        1.383e-03 4.104e-03
                                              0.337 0.736148
## 1_pop
                        6.491e-01 1.468e-01
                                              4.422 9.80e-06 ***
                        2.143e-02 1.703e-02
                                               1.258 0.208211
## time
## 1_1_flows:1_polity2 -2.973e-02 8.618e-03 -3.450 0.000560 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

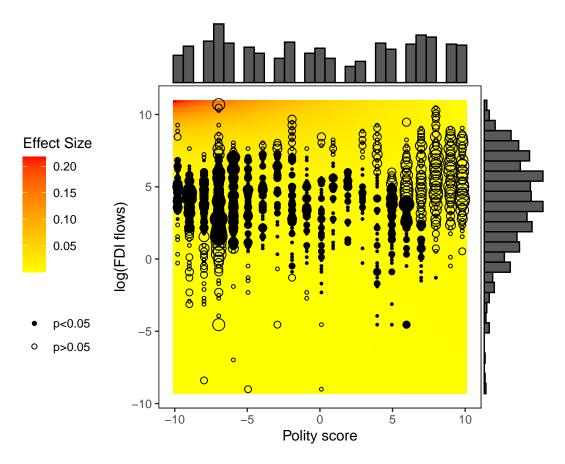
Plotting marginal effects of logged FDI flows

To plot the marginal effects from the negative binomial regression estimated using glm.nb(), we can use the standard syntax of plot_me(), with the exception of the dataset, which needs to be specified directly.

```
g <- plot_me(model = m, data=dt, x = "l_l_flows", over = "l_polity2")
gt <- g +
    scale_fill_gradient(low="yellow", high="red") +
    labs(x="Polity score", y="log(FDI flows)") +
    theme(legend.position="left")
ggExtra::ggMarginal(gt, type="histogram", data=dt, x=l_polity2, y=l_l_flows)</pre>
```

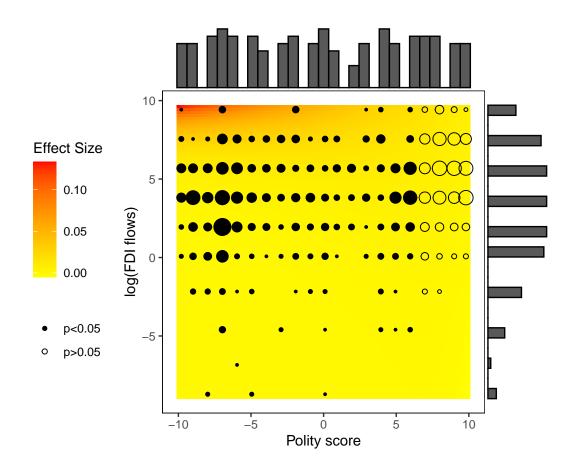


To plot a heatmap for the random effects model, we need to extract the coefficients and the variance-covariance matrix.



The figure in Zhirnov, Moral, and Sedashov (2022) uses binned values of logged FDI flows to reduce overplotting and applies simulations instead of the delta method for computing the standard errors. We can replicate these calculations in R by binning the values of 1_1_flows and setting mc=TRUE in the plot_me() function (please note that the parameter estimates produced with glmmTMB and lme4 are slightly different from Stata's estimates shown in the paper, which also leads to varied marginal effect estimates):

```
mcut <- cut(dt$1 1 flows, breaks=seq(from=-11.22, to=11.22, by=2))</pre>
binned <- tapply(dt$1_1_flows, INDEX=mcut, FUN=mean, na.rm=TRUE)</pre>
dt2 <- within(dt, {
  1_1_flows <- binned[mcut]</pre>
})
g <- plot_me(formula = dispute ~ l_l_flows*l_polity2 + l_dispute + open_penn +
               l_gdp_pc_penn + gdp_grth + inflation_1 + urban + xratchg + l_pop + time,
        data=dt2,
        coefficients=coef,
        vcov=vc,
        link="log",
        x = "l_l_flows", over = "l_polity2", mc=TRUE)
gt <- g +
  scale_fill_gradient(low="yellow", high="red") +
  labs(x="Polity score", y="log(FDI flows)") +
  theme(legend.position="left")
ggExtra::ggMarginal(gt, type="histogram", data=dt, x=l_polity2, y=l_l_flows)
```

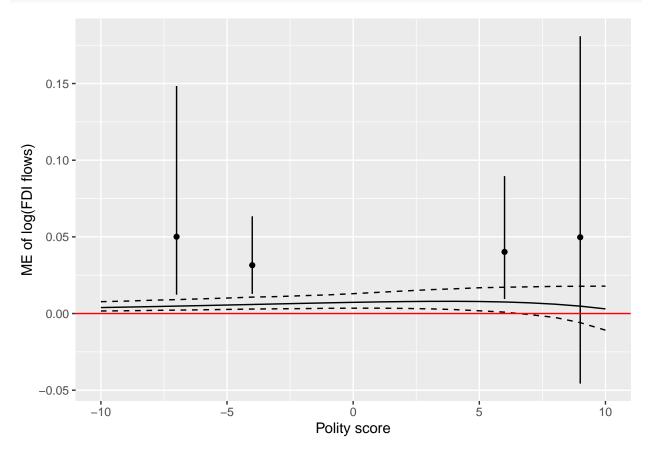


Computing and plotting DAME of logged FDI flows

As earlier, we will include the group-specific components of intercepts as the offset.

```
re <- ranef(m2)$cond$isocode
dt$reff <- re[dt$isocode,"(Intercept)"]</pre>
d <- dame(formula = dispute ~ 1_1_flows*1_polity2 + 1_dispute + open_penn +</pre>
            l_gdp_pc_penn + gdp_grth + inflation_1 + urban + xratchg + l_pop +
            time + offset(reff),
        data=dt,
        coefficients=coef,
        vcov=vc,
        link="log",
        x="l_l_flows", over="l_polity2", nbins=4, mc=TRUE)
mem <- mem(formula = dispute ~ 1_1_flows*1_polity2 + 1_dispute + open_penn +
             l_gdp_pc_penn + gdp_grth + inflation_1 + urban + xratchg + l_pop +
             time + offset(reff),
        data=dt,
        coefficients=coef,
        vcov=vc,
        link="log", x="l_1flows", mc=TRUE, at=list("l_polity2"=seq(from=-10,to=10, by=1)))
ggplot(data=d, aes(x=bin_id, y=est, ymin=lb, ymax=ub)) +
  geom_point() +
  geom_linerange() +
```

```
geom_line(aes(x=1_polity2, y=est), data=mem) +
geom_line(aes(x=1_polity2, y=lb), data=mem, linetype="dashed") +
geom_line(aes(x=1_polity2, y=ub), data=mem, linetype="dashed") +
geom_hline(yintercept=0, color="red") +
labs(x="Polity score", y="ME of log(FDI flows)")
```



This document

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