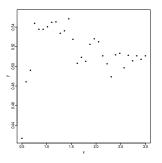
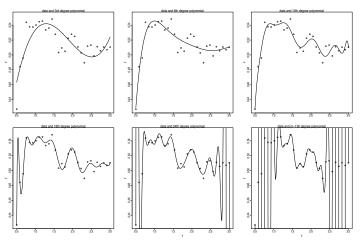
Bias/variance trade-off



- Yesterday we observed n couples (x_i, y_i) , for i = 1, ..., n, of data (n = 30).
- ▶ These data are artificially generated by the law y = f(x) + error where f(x) is a unspecified smooth and regular function.
- ▶ We wish to obtain a rule (model), like $\hat{y} = \hat{f}(x)$, that enables us to predict y once we know x; a rule that allows us to predict y as new observations of x become available, say tomorrow.

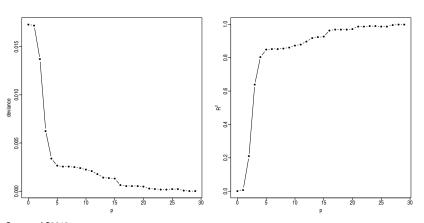
- a simple possibility is to interpolate data with a polynomial
- \blacktriangleright but, of which degree? $0, 1, 2, \ldots, 29$?
- Let's try to use polynomials of degree p (with $p=0,1,\ldots,n-1=29$). We need to estimate p parameters $(+\sigma^2)$.



Source: AS2012

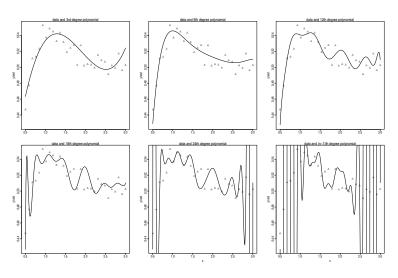
By growing of \boldsymbol{p} the fitting of the polynomials is getting better.

We measure the goodness of fit by obtaining, for each p the residual deviance and the coefficient of determination \mathbb{R}^2 .



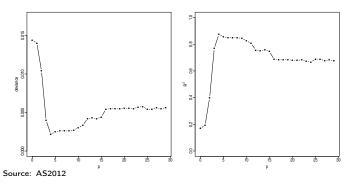
Source: AS2012

- ▶ Tomorrow we will receive a new set of n data $\{y_i, i=1,\ldots,n\}$, generated by the same phenomenon of the yesterday data, that is, the same function f(x)
- We want to predict these new observations, by assuming (for simplicity) that the new y_i are associated to the same x_i of the yesterday data.
- We compare our predictions (one for each polynomial) with the new data observed tomorrow.



Source: AS2012

▶ Goodness of fit for each p: residual deviance and coefficient of determination R^2 on the new (tomorrow) data.



Residual deviance first decreases, then increases, while R^2 reaches a maximum value and then decreases.

If we knew f(x)...

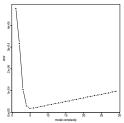
- ▶ We want to estimate f(x) using a generic estimator $\hat{y} = \hat{f}(x)$ (in our example, can be one of the 30 fitted polynomials)
- We start by considering a specific value x' of x, among the n observed.
- If we knew the mechanism used to generate the data precisely, we knew also f(x'), and we could calculate some quantities of interest to evaluate the estimator \hat{y} .
- For example, an important goodness-of-fit indicator is the mean squared error (with respect to the random variable y)

$$\mathbb{E}_y\big\{[\hat{y} - f(x')]^2\big\}$$

Since we are not interested only on the single point x', we consider the sum of the mean squared errors for all the n values of x,

$$\sum_{i=1}^{n} \mathbb{E}_{y} \left\{ \left[\hat{y} - f(x_i) \right]^2 \right\}$$

▶ If we do it for all the possible choices of *p*, which is an indicator of the model complexity, we may obtain the plot



Even if the true f(x) is not a polynomial, there exists a degree p which is better than the others

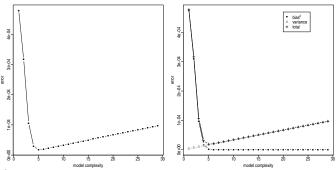
A trade-off

The mean squared error may be divided in two components

$$\mathbb{E}\left\{ \left[\hat{y} - f(x') \right]^2 \right\} = \mathbb{E}\left\{ \left[\hat{y} \pm \mathbb{E}\left\{ \hat{y} \right\} - f(x') \right]^2 \right\}$$
$$= \left[\mathbb{E}\left\{ \hat{y} \right\} - f(x') \right]^2 + \operatorname{var}\left\{ \hat{y} \right\}$$
$$= \operatorname{bias}^2 + \operatorname{variance}$$

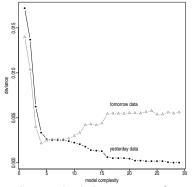
A trade-off

If we knew f(x), we could plot separately bias and variance



Source: AS2012

▶ But as we do not know f(x), we only may compute the residual variance for the new (tomorrow) data:



This plot gives the residual deviance as function of the degree p, by using the model obtained with the *yesterday* data to predict the *tomorrow* data

- When p (the model complexity indicator) increases, the fit improves on the yesterday data, but this is not true for the tomorrow data.
- goodness-of-fit measure is not a good indicator of the quality of the model
- ▶ When p increases too much, we 'overfit' the data and this indicates an excess of optimism!
- ▶ This happens because the model (the polynomial in the example) follows random fluctuations in yesterday's data not observed in the new sample (and not characteristic of the studied phenomenon), and it mistakes local (random) regularity with a systematic pattern.

- So that...do not evaluate a model by using the same data used to fit it (the *yesterday* ones).
- If we want a more reliable evaluation, we need to use other data (the tomorrow ones)
- ► How?!

- ▶ We need tools in order to select models:
 - 1. training set and test (evaluation) set
 - 2. cross-validation
 - 3. information criteria

Training set, test set

- If we have n data, and n is large, we can divide it in two groups randomly chosen:
 - a training set used to fit the various candidate models and a test set (sometime called *evaluation set*) used to evaluate the performance of the available models and to choose the most accurate one.
- We compare results obtained with different models on the test set.
- ▶ This scheme reduces the sample size used for fitting the model, but this is not a problem when n is huge.
- Because the same test set can be used to evaluate many different models, there is a risk that the final assessment is still somewhat biased and too optimistic. Sometimes a third set of data, called validation set, is often created and used for final evaluation of the prediction error
- Examples of proportions for the sizes of the sets are:

training set	test set	validation se	t
50%	25%	25%	
75%	25%	_	

training and test sets are somehow similar to what was done with yesterday and tomorrow data.

Information criteria

- The residual variance (or the deviance) is an unreliable indicator of the quality of the model, because it is too optimistic in evaluating the prediction error.
- We can penalize the deviance $D = \sum_{i=1}^{n} (y_i \hat{y}_i)^2$
- ... or a monotonic transform:
 - $-2\log L = n\log(D/n) + (costant)$
- with a suitable quantity quantifying the model complexity
- ▶ The $\log L$, for the gaussian model, has an interpretation as log-likelihood.
- Criteria that follow this logic can be traced back to objective functions such as

$$IC(p) = -2\log L + \mathsf{penalty}(p)$$

► The choice of the specific penalty function identifies a particular criterion.

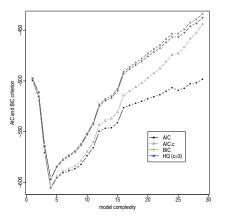
Information criteria

► Some possibile penalty are in the following table

criterion	author	penalty(p)
AIC	Akaike	2p
AIC_c	Sugiura, Hurvich-Tsay	$2p + \frac{2p(p+1)}{n - (p+1)}$
BIC/SIC	Akaike, Schwarz	$p \log n$
HQ	Hannan-Quinn	$c p \log \log n, (c > 2)$

► These criteria are applied also to *not nested* and *not gaussian* models.

Information criteria – example



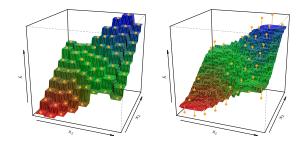
We choose p minimising IC(p) using some criteria in the previous table;

in our example all choices for penalty suggest p=4.

Non parametric regression

Given a value k and a prediction point x_0 , the KNN regression identifies in the training set the k nearest observations, N_0

$$\hat{f}(x_0) = \frac{1}{k} \sum_{x_i \in N_0} y_i$$

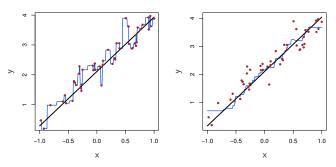


KNN with p=2, k=1 (left) and k=9 (right). With small k high variance and low bias, since prediction is performed on a single observation.

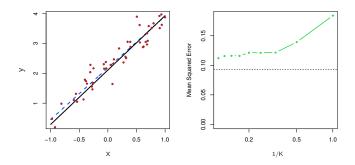
The optimal value of k is related to the trade-off viarance-bias.

- ightharpoonup small k o high variance and low bias
- big $k \to \text{low variance (smoother prediction)}$ and high bias local structure of f(X) may not be captured-

Parametric approach may be preferred to the non parametric if the parametric form is close to the 'real' f.

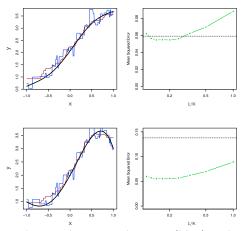


Comparison between KNN with k=1 (left) e k=9 (right). Since the true relationship is linear the non parametric approach will have a worse performance.

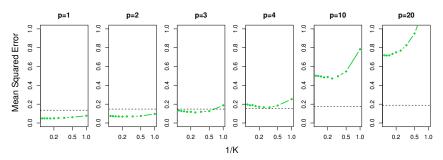


Regression line (dashed line) Test MSE for regression line (dashed) and KNN (green) as function of 1/k.

Best results for KNN are with high value of k.



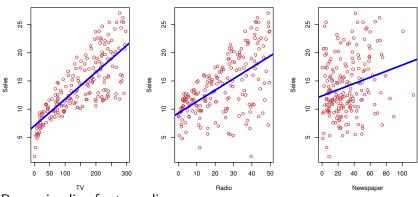
Nonlinear relationships and KNN with k=1 (blue) and k=9 (red). Conditional to nonlinearity of f the KNN performance changes with respect to LM. As the nonlinearity becomes more evident, the performance of KNN with high k will increase.



By increasing the number of variables p, the KNN performance will rapidly decrease in terms of MSE test.

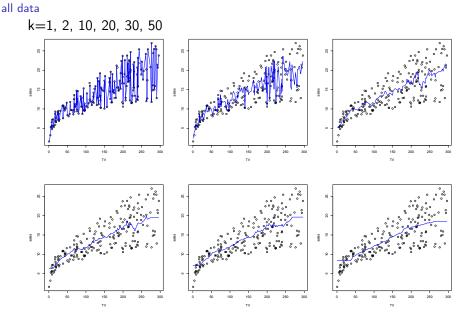
It is more difficult to find the 'nearest neighbours' . . . curse of dimensionality

Sales of a product in thousands of units as function of budget in tv, radio, newspapers for 200 different markets.

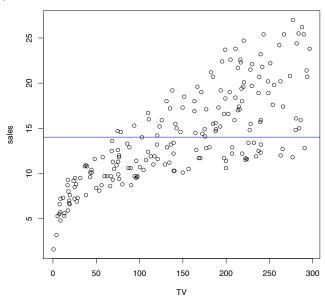


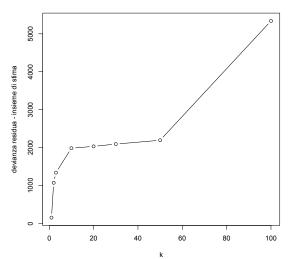
Regression line for tv, radio, newspapers.

We wish to study the performance of KNN for some values of \boldsymbol{k} with the only variable $\ensuremath{\text{tv}}$



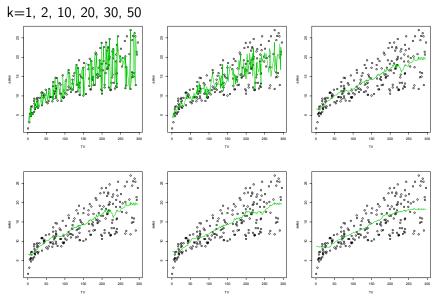
k = 200

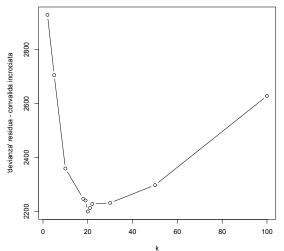




KNN performance decreases as \boldsymbol{k} increases.

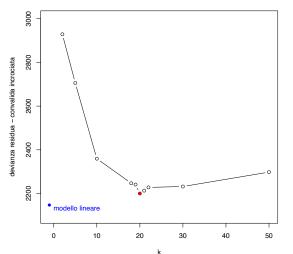
Cross validation leave-one-out





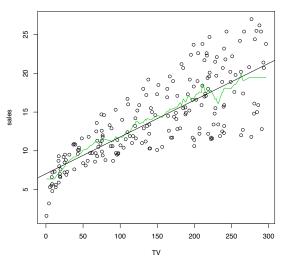
A minimum has reached ... trade-off between variance and bias

Performance of linear model and KNN with variable tv

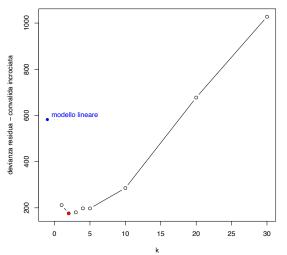


In the case of tv the linear model performs better than the KNN for each value of k.

Linear model and KNN-20 with variable tv

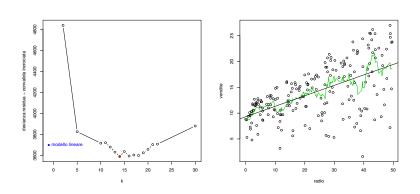


Performance of linear model and KNN with variable tv and radio

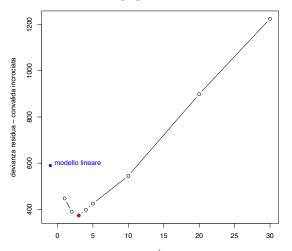


Adding the variable radio highly increases the performance of KNN. The minimum is reached for $k=2\,$

Performance of linear model and KNN with variable radio

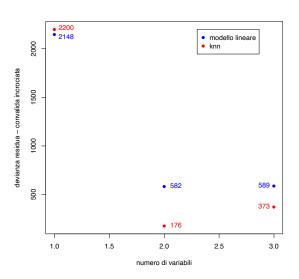


Let us add the variable newspapers



Adding the variable newspapers does not increase the performance of the model.

The KNN is better than the linear model in any case.



Local regression and loess

▶ If f(x) is a derivable function in x_0 then, the Taylor's approximation says that it is locally approximated by a line passing through $(x_0, f(x_0))$, i.e.,

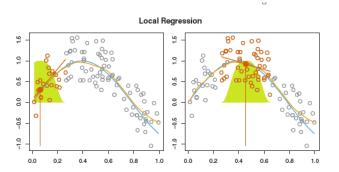
$$f(x) = \underbrace{f(x_0)}_{\alpha} + \underbrace{f'(x_0)}_{\beta} (x - x_0) + \text{error}$$

We introduce the weighted least squares by weighting observations x_i with their distance from x_0 :

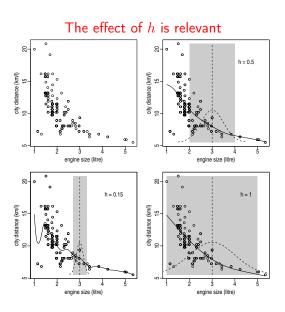
$$\min_{\alpha,\beta} \sum_{i=1}^{n} \left\{ y_i - \alpha - \beta(x_i - x_0) \right\}^2 w_h(x_i - x_0)$$

- ▶ h (h > 0) is a scale factor, called bandwidth or smoothing parameter, and
- $ightharpoonup w_h(\cdot)$ is a symmetric density function around 0, said kernel.

- ▶ By varying x_0 , we obtain a whole estimated curve $\hat{f}(x)$.
- ▶ The most important component is *h*, which regulates the smoothness of the curve, while the choice of *w* is less relevant.
- We could think to w as the density of the normal distribution N(0,1)



Local regression: blue curve represents the real f(x), light orange curve corresponds to the local regression estimate $\hat{f}(x)$. The orange colored points are local to the target point x_0 , represented by the orange vertical line. The yellow bell-shape superimposed on the plot indicates weights assigned to each point, decreasing to zero with distance from the target point. The fit $\hat{f}(x)$ at x_0 is obtained by fitting a weighted linear regression (orange line segment), and using the fitted value at x_0 (orange solid dot) as the estimate $\hat{f}(x_0)$



Variable bandwidths and loess

- in many cases, there is an advantage in using a non constant bandwidth along the x-axis, according it to the level of sparseness of observed points
- ightharpoonup variable bandwidth: it is reasonable to use larger values of h when x_i are more scattered
- ► Good idea! ... but how do we modify *h*?
- loess: express the smoothing parameter defining the fraction of effective observations for estimating f(x) at a certain point x_0 on the x-axis;
- this fraction is kept constant
- this implies automatically a setting of the bandwidth related to the sparsity of data