

## APMA 1655 Chapter 10 Notes: Hypothesis Testing

### Basics of Hypothesis Testing

- Null hypothesis ( $H_0$ ): claim to be tested
- Alternative hypothesis ( $H_a$ ): the converse of the null hypothesis
- Test statistic: is a function of the sample measurements on which the statistical decision will be based, like an estimator
- Rejection region (RR): specifies the values of the test statistic for which the null hypothesis is to be rejected in favor of the alternative hypothesis (given a particular sample, if the computed value of the test statistic falls in RR, reject  $H_0$  and accept  $H_a$ , otherwise accept  $H_0$ )
- Approach One: report testing result via p-value
  - P-value: probability of observing test statistics that are as extreme or more extreme than the present empirical data assuming  $H_0$  is valid
- Approach Two: Report testing result via rejection/acceptance and significance level
  - Determine rejection region
  - Report Testing Result

### Errors and Power

- Type 1 Error: made if  $H_0$  is rejected when  $H_0$  is true; probability of type 1 error is  $\alpha$ , which is called the level of the test;  $P(RR|H_0) = \alpha$
- Type 2 Error: made if  $H_0$  is accepted when  $H_a$  is true; probability of type 2 error is denoted by  $\beta$ ;  $\beta = 1 - P(RR|H_a)$
- Power of test: defined to be  $1 - \beta = P(RR|H_a)$

### Testing Means of Normals

- Let  $\{X_1, \dots, X_n\}$  be iid samples from  $N(\mu, \sigma^2)$ , where  $\sigma^2$  is known but  $\mu$  is unknown; want to perform hypothesis testing on  $\mu$ ; there are three scenarios:
- One-sided test:  $H_0 : \mu = \mu_0, H_a : \mu > \mu_0$ 
  - Test statistic: sample mean  $\bar{X}$ , ( $\bar{x}$  is the observed sample mean)
  - P-Value Approach: Under  $H_0$ ,  $\bar{X}$  is  $N(\mu_0, \sigma^2/n)$ :

$$\text{P-value} = P(\bar{X} \geq \bar{x}) = \phi\left(-\frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}\right)$$

- Rejection Region:  $RR = \{\bar{X} > c\}$ ; want  $P(RR) = \alpha$  under  $H_0$ ; this leads to:

$$c = \mu_0 + z_\alpha \frac{\sigma}{\sqrt{n}}$$

- Conclusion:

- \* P-value  $\leq \alpha$  or  $\bar{x} \in RR$ : reject  $H_0$  and accept  $H_a$  because of statistically significant evidence in data
- \* P-value  $> \alpha$  or  $\bar{x} \notin RR$ : accept  $H_0$  and reject  $H_a$  because evidence in data not statistically significant
- Power: by definition is  $P(RR|H_a)$  but depends on value of  $\mu$ ; say  $\mu = \mu_a$ , then:

$$\text{Power} = P(\hat{X} > \mu_0 + z_\alpha \frac{\sigma}{\sqrt{n}} | \mu = \mu_a) = \phi\left(\frac{\mu_a - \mu_0}{\sigma/\sqrt{n}} - z_\alpha\right)$$

- One-Sided Test;  $H_0 : \mu = \mu_0, H_a : \mu < \mu_0$ 
  - Test statistic: sample mean  $\bar{X}$ , ( $\bar{x}$  is the observed sample mean)
  - P-Value Approach: Under  $H_0$ ,  $\bar{X}$  is  $N(\mu_0, \sigma^2/n)$ :

$$\text{P-value} = P(\bar{X} \leq \bar{x}) = \phi\left(\frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}\right)$$

- Rejection Region:  $RR = \{\bar{X} > c\}$ ; want  $P(RR) = \alpha$  under  $H_0$ ; this leads to:

$$c = \mu_0 - z_\alpha \frac{\sigma}{\sqrt{n}}$$

- Conclusion:
  - \* P-value  $\leq \alpha$  or  $\bar{x} \in RR$ : reject  $H_0$  and accept  $H_a$  because of statistically significant evidence in data
  - \* P-value  $> \alpha$  or  $\bar{x} \notin RR$ : accept  $H_0$  and reject  $H_a$  because evidence in data not statistically significant
- Power: by definition is  $P(RR|H_a)$  but depends on value of  $\mu$ ; say  $\mu = \mu_a$ , then:

$$\text{Power} = P(\hat{X} < \mu_0 - z_\alpha \frac{\sigma}{\sqrt{n}} | \mu = \mu_a) = \phi\left(\frac{\mu_0 - \mu_a}{\sigma/\sqrt{n}} - z_\alpha\right)$$

- Two-Sided Test;  $H_0 : \mu = \mu_0, H_a : \mu \neq \mu_0$ 
  - Test statistic: sample mean  $\bar{X}$ , ( $\bar{x}$  is the observed sample mean)
  - P-Value Approach: Under  $H_0$ ,  $\bar{X}$  is  $N(\mu_0, \sigma^2/n)$ :

$$\text{P-value} = P(|\bar{X} - \mu_0| \geq |\bar{x} - \mu_0|) = 2\phi\left(-\frac{|\bar{x} - \mu_0|}{\sigma/\sqrt{n}}\right)$$

- Rejection Region:  $RR = \{|\bar{X} - \mu_0| > c\}$ ; want  $P(RR) = \alpha$  under  $H_0$ ; this leads to:

$$c = z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

- Conclusion:
  - \* P-value  $\leq \alpha$  or  $\bar{x} \in RR$ : reject  $H_0$  and accept  $H_a$  because of statistically significant evidence in data
  - \* P-value  $> \alpha$  or  $\bar{x} \notin RR$ : accept  $H_0$  and reject  $H_a$  because evidence in data not statistically significant

- Power: by definition is  $P(RR|H_a)$  but depends on value of  $\mu$ ; say  $\mu = \mu_a$ , then:

$$\text{Power} = P(|\bar{X} - \mu_0| > z_{\alpha/2} \frac{\sigma}{\sqrt{n}} | \mu = \mu_a) =$$

$$\phi\left(\frac{\mu_a - \mu_0}{\sigma/\sqrt{n}} - z_{\alpha/2}\right) + \phi\left(\frac{\mu_0 - \mu_a}{\sigma/\sqrt{n}} - z_{\alpha/2}\right)$$

### Extensions to Large Samples: Testing Mean

- all discussions on testing means of normals can be extended to large sample test where the test statistic is approximately normally distributed
- For example: let  $\{X_1, \dots, X_n\}$  be iid samples from some population distribution with unknown mean  $\mu$ 
  - One sided test:  $H_0 : \mu = \mu_0, H_a : \mu > \mu_0$
  - One sided test:  $H_0 : \mu = \mu_0, H_a : \mu < \mu_0$
  - Two sided test:  $H_0 : \mu = \mu_0, H_a : \mu \neq \mu_0$
- Test statistic is sample mean  $\bar{X}$ , by central limit theorem,  $\bar{X}$  is approximately  $N(\mu, \sigma^2/n)$ ; all previous formulae are valid; when  $\sigma$  is unknown, can use sample standard deviation  $s$  in place of  $\sigma$

### Extensions to Large Samples: Testing Proportion

- let  $\{X_1, \dots, X_n\}$  be iid Bernoulli samples such that  $P(X_i = 1) = p, P(X_i = 0) = 1 - p$  where  $p$  is unknown
  - One sided test:  $H_0 : p = p_0, H_a : p > p_0$
  - One sided test:  $H_0 : p = p_0, H_a : p < p_0$
  - Two sided test:  $H_0 : p = p_0, H_a : p \neq p_0$
- Test statistic is sample mean  $\bar{X}$ , by central limit theorem,  $\bar{X}$  is approximately  $N(p, p(1-p)/n)$ ; all previous formulae are valid except for power, with  $p_0$  in place of  $\mu_0$  and  $\sqrt{p_0(1-p_0)}$  in place of  $\sigma$

### Neyman-Pearson Lemma

- Setup: Suppose  $\{X_1, \dots, X_n\}$  are iid samples with common density  $f(x)$ , consider the following hypotheses  $H_0 : f(x) = f_0(x), H_a : f(x) = f_a(x)$  where  $f_0$  and  $f_a$  are 2 given densities
- Question: fix an arbitrary significance level  $\alpha$ , among all possible rejection regions  $RR$  such that the type 1 error satisfies  $P(RR|H_0) \leq \alpha$ , which one gives the minimal type 2 error or maximal power?
- Theorem: Neyman-Pearson Lemma
  - Let  $\alpha \in (0, 1)$  be any given significance level
  - For each  $b > 0$  define a rejection region  $RR_b \subseteq \mathbb{R}^n$  of form

$$RR_b = (x_1, \dots, x_n) : \frac{f_a(x_1) \dots f_a(x_n)}{f_0(x_1) \dots f_0(x_n)} \geq b$$

- Let  $b^*$  be such that  $P(RR_{b^*}|H_0) = \alpha$ . Then  $RR_{b^*}$  yields the minimal type 2 error among all rejection regions whose type 1 error is bounded by  $\alpha$