

Chapter 5 Some Discrete Probability Distributions

Binomial: Two possible outcomes from each trial.

Binomial Distribution: $b(x; n, p) = \binom{n}{x} p^x q^{n-x}$ x: num of successes. n: num of indep. trials. p: prob. of success. q: prob. of failure. $\mu = np$, $\sigma^2 = npq$

Multinomial: $f(x_1, x_2, \dots, x_k; p_1, p_2, \dots, p_k, n) = \binom{n}{x_1, x_2, \dots, x_k} p_1^{x_1} p_2^{x_2} \dots p_k^{x_k}$. $n = \sum_{i=1}^k x_i$, and $\sum_{i=1}^k p_i = 1$

Hypergeometric: Choosing successful items.

Hypergeometric Distribution: $h(x; N, n, k) = \frac{\binom{k}{x} \binom{N-k}{n-x}}{\binom{N}{n}}$. $\max(0, n - (N - k)) \leq x \leq \min(n, k)$. x: num of successes. N: num of items. n: num of selection. k: num of total successes. $\mu = \frac{nk}{N}$, $\sigma^2 = \frac{N-n}{N-1} n \frac{k}{N} (1 - \frac{k}{N})$

Estimating Hypergeometric using Binomial: If n is small compared to N: $(n/N) \leq 0.05$.

Chapter 6 Some Continuous Probability Distributions

Uniform Distribution: Equal Probability throughout interval. $f(x; A, B) = \frac{1}{B-A}$ if $A \leq x \leq B$, 0 otherwise. $\mu = \frac{A+B}{2}$, $\sigma^2 = \frac{(B-A)^2}{12}$.

Normal Distribution: Bell curve. $n(x; \mu, \sigma) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}$. x: select time. μ : mean. σ : standard deviation.

Standard Normal: A normal distribution where mean is 0 and variance is 1. $Z = \frac{X-\mu}{\sigma}$

Estimating Binomial with Normal: For large n. $P(X \leq x) \approx P(Z \leq \frac{x+0.5-np}{\sqrt{npq}})$

Gamma Function: $\Gamma(n) = (n-1)!$. $\Gamma(1) = 1$. $\Gamma(1/2) = \sqrt{\pi}$.

Gamma Distribution: Wait time, reliability. $f(x; \alpha, \beta) = \frac{1}{\beta^\alpha \Gamma(\alpha)} x^{\alpha-1} e^{-\frac{x}{\beta}}$, $x > 0$. or 0 otherwise. $\mu = \alpha\beta$, $\sigma^2 = \alpha\beta^2$.

Exponential Distribution: Special case of Gamma where $\alpha = 1$.

Chapter 8 Fundamental Sampling Distributions and Data Descriptions

Central Limit Theorem: $Z = \frac{\bar{X}-\mu}{\sigma/\sqrt{n}}$ as $n \rightarrow \infty$ is a standard normal distribution. \bar{X} : mean of random sample size. μ : mean of population. σ : standard deviation of population. n: sample size.

Difference of Means: Two populations, samples, means, and variances. $Z = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{\sqrt{(\sigma_1^2/n_1) + (\sigma_2^2/n_2)}}$. is approx. a standard normal variable.

$$\mu_{\bar{x}_1 - \bar{x}_2} = \mu_1 - \mu_2. \sigma_{\bar{x}_1 - \bar{x}_2}^2 = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}$$

Finding Chi-Squared from Variance: $\chi^2 = \frac{(n-1)S^2}{\sigma^2} = \sum_{i=1}^n \frac{(x_i - \bar{x})^2}{\sigma^2}$. Degrees of freedom is $v = n - 1$, n is sample size.

t-Distribution: $T = \frac{Z}{\sqrt{V/v}}$ or $T = \frac{\bar{x} - \mu}{S/\sqrt{n}}$. then $h(t) = \frac{\Gamma((v+1)/2)}{\Gamma(v/2)\sqrt{\pi v}} (1 + \frac{t^2}{v})$

Chapter 9 One- and Two-Sample Estimation Problems

CI on μ , σ^2 known: $\bar{x} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}} < \mu < \bar{x} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$. $100(1 - \alpha)\%$ confidence interval. $z_{\alpha/2}$ is the z-value leaving the area of $\alpha/2$ to the right. $100(1 - \alpha)\%$ confident that error will not exceed $z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$. Also confident error will not exceed size e as $n = (\frac{z_{\alpha/2}\sigma}{e})^2$. One sided bound? just take one side of the equation, + is upper.

CI on μ , σ^2 unknown: $\bar{x} - t_{\alpha/2} \frac{s}{\sqrt{n}} < \mu < \bar{x} + t_{\alpha/2} \frac{s}{\sqrt{n}}$ where $t_{\alpha/2}$ is the t-value with $v = n - 1$ degree of freedom leaving $\alpha/2$ area to the right.

Confidence limits on μ , σ^2 unknown: $\bar{x} \pm t_{\alpha/2} \frac{s}{\sqrt{n}}$. standard error is $\frac{s}{\sqrt{n}}$

CI for $\mu_1 - \mu_2$, σ_1^2 and σ_2^2 known: $(\bar{x}_1 - \bar{x}_2) - z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}} < \mu_1 - \mu_2 < (\bar{x}_1 - \bar{x}_2) + z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$

Examples

1. Suppose the yearly number of tornado occurrences can be modeled using a Poisson distribution and an average of 3.2 tornados were observed in a particular region yearly.

(a) probability of having exactly 4 tornados in that year?

$$P(X = 4) = \frac{e^{-3.2} 3.2^4}{4!} = 0.1781.$$

(b) at least 2 tornadoes?

$$P(X \geq 2) = 1 - P(X = 0) - P(X = 1) = 1 - \frac{e^{-3.2} 3.2^0}{0!} - \frac{e^{-3.2} 3.2^1}{1!} = 0.8288.$$

(c) 2 or more but at most 4 tornadoes?

$$P(2 \leq X \leq 4) = P(X = 2) + P(X = 3) + P(X = 4) = \frac{e^{-3.2} 3.2^2}{2!} + \frac{e^{-3.2} 3.2^3}{3!} + \frac{e^{-3.2} 3.2^4}{4!} = 0.6094.$$

(d) Assuming the number of tornado occurrences in that particular region are independent from year to year, what is the probability of having a total of 5 tornadoes in the next two years?

$$X_1 + X_2 = p(6.4), P(X_1 + X_2 = 5) = \frac{e^{-6.4} 6.4^5}{5!} = 0.1487.$$

(e) at least 2 tornadoes in exactly 3 of the next 5 years?

$$Y \sim b(5, 0.8288). P(Y = 3) = \binom{5}{3} (0.8288)^3 (0.1712)^2 = 0.1669.$$

2. The number of years a radio functions is exponentially distributed with parameter $\beta = 6$ from the first factory.

Multivariate: $f(x_1, x_2, \dots, x_k; a_1, a_2, \dots, a_k, N, n) = \frac{\binom{a_1}{x_1} \binom{a_2}{x_2} \dots \binom{a_k}{x_k}}{\binom{N}{n}}$. $n = \sum_{i=1}^k x_i$, $N = \sum_{i=1}^k a_i$.

Negative Binomial Distribution: Prob. the kth success will happen by the xth trial. $b^*(x; k, p) = \binom{x-1}{k-1} p^k q^{x-k}$. x: trial number. k: success number. p: prob. success. q: prob. failure.

Geometric Distribution: Prob. the xth trial is the first success. $g(x; p) = pq^{x-1}$. x: trial number. p: prob. success. q: prob. failure. $\mu = \frac{1}{p}$. $\sigma^2 = \frac{1-p}{p^2}$.

Poisson Distribution: Prob. something happens x times in t time. $p(x; \lambda t) = \frac{e^{-\lambda t} (\lambda t)^x}{x!}$. x: num of times. λ : average number of outcomes per time period. t: time interval.

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$f(x; \beta) = \frac{1}{\beta} e^{-\frac{x}{\beta}}$ where $x > 0$. 0 elsewhere. β : mean time between failures. α : number of events. $\mu = \beta$, $\sigma^2 = \beta^2$.

Chi-Squared Distribution: Special case of Gamma where $\alpha = v/2$ and $\beta = 2$. $f(x; v) = \frac{1}{2^{v/2} \Gamma(v/2)} x^{v/2-1} e^{-x/2}$, $x > 0$. 0 elsewhere. v: degrees of freedom. $\mu = v$, $\sigma^2 = 2v$.

Beta Function: $B(\alpha, \beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}$, $\alpha, \beta > 0$.

Beta Distribution: $f(x) = \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1}$, $0 < x < 1$. 0 elsewhere. $\mu = \frac{\alpha}{\alpha+\beta}$, $\sigma^2 = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$.

Lognormal Distribution: if $\ln(X)$ is a normal distribution. $f(x; \mu, \sigma) = \frac{1}{x\sigma\sqrt{2\pi}} e^{-1/2\sigma^2(\ln(x)-\mu)^2}$, $x \geq 0$. 0 if $x < 0$. mean = $e^{\mu+\sigma^2/2}$, variance = $e^{2\mu+2\sigma^2}(e^{\sigma^2} - 1)$.

from $-\infty < t < \infty$. Z: standard normal RV. V: chi2 RV. v: degrees of freedom.

F-Distribution: $h(f) = \frac{\Gamma((v_1+v_2)/2) \Gamma(v_1/2) v_1^{v_1/2}}{\Gamma(v_1/2) \Gamma(v_2/2)} \cdot \frac{f^{(v_1/2)-1}}{(1+v_1 f/v_2)^{(v_1+v_2)/2}}$. for $f > 0$, 0 if $f \leq 0$. $F = \frac{U/v_1}{V/v_2}$. V, U: indep. RV with chi2 distribution. v_1 ,

v_2 : degrees of freedom. $f_{1-\alpha}(v_1, v_2) = \frac{1}{f_{\alpha}(v_2, v_1)}$. $F = \frac{S_1^2/\sigma_1^2}{S_2^2/\sigma_2^2} = \frac{\sigma_2^2 S_1^2}{\sigma_1^2 S_2^2}$

Normal Q-Q Plot: Set of observations for normal distribution, will be straight if normal. 1) order data ascending. 2) split normal distribution to $n + 1$ parts. 3) match the data to the distribution x=data, y=normal. Match smallest with smallest.

$$(\bar{x}_1 - \bar{x}_2) + z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$$

Pooled Estimate of Variance: $S_p^2 = \frac{(n_1-1)S_1^2 + (n_2-1)S_2^2}{n_1 + n_2 - 2}$

CI for $\mu_1 - \mu_2$, $\sigma_1^2 = \sigma_2^2$ but both unknown:

$$(\bar{x}_1 - \bar{x}_2) - t_{\alpha/2} s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}} < \mu_1 - \mu_2 < (\bar{x}_1 - \bar{x}_2) + t_{\alpha/2} s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

CI for $\mu_1 - \mu_2$, $\sigma_1^2 \neq \sigma_2^2$ and both unknown:

$$(\bar{x}_1 - \bar{x}_2) - t_{\alpha/2} \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}} < \mu_1 - \mu_2 < (\bar{x}_1 - \bar{x}_2) + t_{\alpha/2} \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$$

CI for $\mu_D = \mu_1 - \mu_2$ for Paired Observations: For the table, find d_i which is the difference between items. $\bar{d} - t_{\alpha/2} \frac{s_d}{\sqrt{n}} < \mu_D < \bar{d} + t_{\alpha/2} \frac{s_d}{\sqrt{n}}$

(a) probability that the radio will stop working in less than 4 years?

$$P(X < 4) = 1 - e^{-1/6 \cdot 4} = 1 - e^{-2/3} = 0.4866.$$

(b) the radio has worked for 2 years, what is the probability that it will be working for an additional 6 years?

$$P(X > 6 + 2 | X > 2) = P(X > 6) \text{ memoryless property} = e^{-1/6 \cdot 6} = 0.3679.$$

(c) 40th percentile of the years a radio will be working?

$$0.4 = P(X < m_{0.4}) = 1 - e^{-1/6 \cdot m_{0.4}}, e^{-m_{0.4}/6} = 0.6, -\frac{m_{0.4}}{6} = \ln 0.6, m_{0.4} = 3.065$$

(d) 64 radios have been randomly selected from the first factory, what is the approximated probability that the average working time of these 64 radios will exceed 7 years?

$$X_1, \dots, X_{64} \sim \text{Exp}(6), \bar{X} \approx N(6, \sqrt{\frac{36}{64}}) \text{ by CLT. } P(\bar{X} > 7) \approx P(\frac{\bar{X}-6}{\sqrt{36/64}} > \frac{7-6}{\sqrt{36/64}}) = P(X > \frac{4}{3}) = 1 - \Phi(\frac{4}{3}) = 1 - 0.9082 = 0.0918.$$

(e) Suppose a second factory has started production, and 5 radios have been randomly selected, and their working time has been recorded as (in years): 4.5, 3.2, 7.8, 5.1, 4.4. Assuming that the number of years a radio functions is also exponentially distributed with an unknown parameter β . Give an estimate of β .

$$\text{Use } \bar{x} \text{ as an estimate for } \beta: \bar{x} = \frac{4.5+3.2+7.8+5.1+4.4}{5} = 5.$$

3. A communications channel in location I transmits the digits 0 and 1, where the digit transmitted is correctly received with probability 0.6. While another communications channel in location II also transmits the digits 0 and 1, and the digit transmitted is correctly received with probability 0.7.

(a) transmit a short message with 6 binary digits in location I, what is the probability that 4 digits will be received correctly?

$$X_1 = \text{number of correct digits received. } X_1 \sim b(6, 0.6). P(X_1 = 4) = \binom{6}{4} 0.6^4 0.4^2 = 0.311.$$

(b) calculate the mean and variance of the number of digits received correctly.

$$E(X_1) = 6 \cdot 0.6 = 3.6. \text{Var}(X_1) = 6 \cdot 0.6 \cdot 0.4 = 1.44.$$

(c) transmit an important message with 96 binary digits in location I, and denote the number of digits correctly received by X . Approximate the probability that at most 66 digits will be received correctly.

$$X \sim b(96, 0.6), X \approx N(57.6, \sqrt{23.04}) \text{ by the CLT. } P(X \leq 66) \approx P(\frac{X-57.6}{\sqrt{23.04}} \leq \frac{66.5-57.6}{\sqrt{23.04}}) = \Phi(1.85) = 0.9678.$$

(d) approximate the probability that X is more than 55.

$$P(X > 55) = P(\frac{X-57.6}{\sqrt{23.04}} > \frac{55.5-57.6}{\sqrt{23.04}}) = 1 - \Phi(-0.44) = \Phi(0.44) = 0.67.$$

(e) A similar message with 84 binary digits is transmitted in location II, and denote the number of digits correctly received by Y . Approximate the probability that $Y > X + 2$.

$$Y \approx N(84 \cdot 0.7, \sqrt{84 \cdot 0.7 \cdot 0.3}). P(Y > X + 2) = P(Y - X > 2). Y - X \approx N(58.8 - 57.6, \sqrt{17.64 + 23.04}). P(Y > X + 2) \approx P(\frac{Y-X-1.2}{\sqrt{40.68}} > \frac{2-1.2}{\sqrt{40.68}}) = P(X > \frac{0.8}{\sqrt{40.68}}) = 1 - \Phi(0.13) = 1 - 0.5517 = 0.4483.$$

4. For the following calculations, a random sample X_1, X_2, \dots, X_n are assumed to be independently and identically distributed with the population distribution.

(c) Assume the most recent Scholastic Aptitude Test (SAT) mathematics examination score is normally distributed with mean μ and variance σ^2 . Suppose that a random sample of 13 students whose most recent SAT examination in mathematics were recorded. The sample mean and sample standard deviation from these students were $\bar{x} = 520$ and $s = 125$. Based on the sample data, find a 95% (two-sided) confidence interval for the mean score of SAT mathematics examination μ .

$$95\% \text{ CI for } \mu: (\bar{x} - t_{0.025, 12} \frac{s}{\sqrt{13}}, \bar{x} + t_{0.025, 12} \frac{s}{\sqrt{13}}) = (520 - 2.179 \frac{125}{\sqrt{13}}, 520 + 2.179 \frac{125}{\sqrt{13}}) = (444.5, 595.5)$$

(d) another sample of 13 students was obtained, with sample mean and sample standard deviation of $\bar{x} = 516$ and $s = 130$. Find a 99% (two-sided) confidence interval for the mean score of SAT mathematics examination μ .

$$99\% \text{ CI for } \mu: (\bar{x} - t_{0.005, 12} \frac{s}{\sqrt{13}}, \bar{x} + t_{0.005, 12} \frac{s}{\sqrt{13}}) = (517 - 3.055 \frac{130}{\sqrt{13}}, 517 + 3.055 \frac{130}{\sqrt{13}}) = (406.9, 627.1).$$

(e) Compare the confidence intervals you obtained from (c) and (d). Which confidence interval has a higher probability of covering the true mean score μ of SAT mathematics examination?

A confidence interval either covers the true value or it does not, hence the probability of covering the true value is either 0 or 1. However, we do not know which is which. Therefore, we do not know.