

MCMC:HW1

Qi Chen(qc586)

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Problem 1

Use rejection sampling to sample from the density function

$$f(x) \propto (-\log x)^2 x^3 (1-x)^2, \quad 0 < x < 1 \quad (1)$$

Carefully detail the method you use, and provide a figure of the histogram of the samples you obtained. What is the approximate, or actual if you can find it, probability of acceptance.

Solution: For rejection sampling, we can write

$$f(x) = l(x) h(x) \quad (2)$$

where

$$h(x) = \frac{\Gamma(3+3)}{\Gamma(3)\Gamma(3)} x^{3-1} (1-x)^{3-1} = 30 x^2 (1-x)^2 \quad (3)$$

is the Beta(3, 3) pdf. Then

$$l(x) = \frac{\mathcal{N}}{30} (-\log x)^2 x \quad (4)$$

where \mathcal{N} is the normalization constant. $l(x)$ is bounded by

$$M = l(e^{-2}) = \frac{4 \mathcal{N}}{30 e^2} = \frac{2 \mathcal{N}}{15 e^2} \quad (5)$$

As

$$\mathcal{N} = \frac{108000}{919} \quad (6)$$

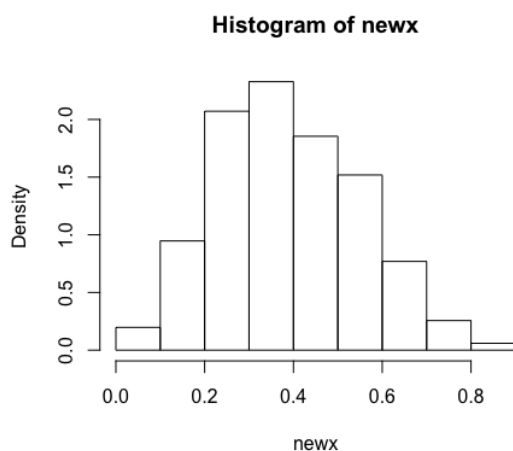
The rate of acceptance should be

$$P = 1/M \approx 0.471565 \quad (7)$$

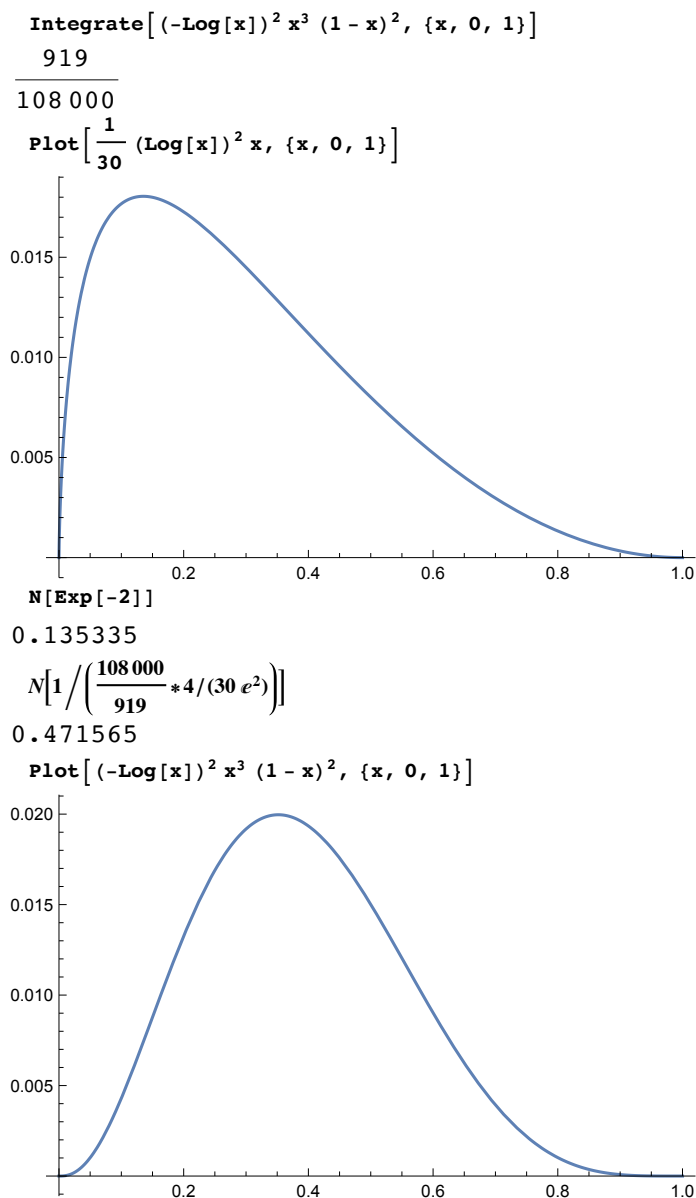
The simulation result ($N = 1000$) gives

$$P \approx 0.507 \quad (8)$$

The histogram is as follows:



Integrate[$30 x^2 (1-x)^2$, {x, 0, 1}]



Problem 2

Use Monte Carlo methods to evaluate the integral

$$I = \int_0^1 (-\log x)^2 x^3 (1-x)^{5/2} dx \quad (9)$$

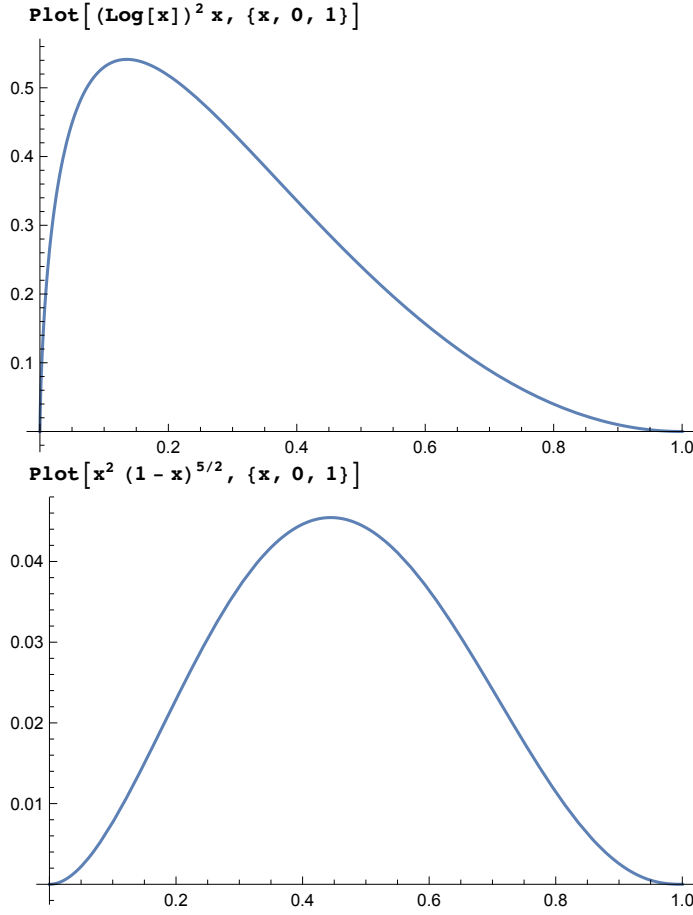
Describe in detail how you do this. Provide a graphical demonstration that your method has worked. If you fix the Monte Carlo sample size as $N = 1000$, is there a way to estimate the variance of your \hat{I}_N .

Solution:

```

Integrate[(-Log[x])^2 x^3 (1 - x)^{5/2}, {x, 0, 1}]
32 π^2 + 64 (7 857 230 824 + 45 045 Log[2] (-187 111 + 90 090 Log[2]))
-----
9009 6 093 243 231 075

```



The exact result is obtained in R as 0.006587444 with absolute error < 4.1e-07. First we consider direct evaluation:

$$I = \int_0^1 g(x) f(x) dx \quad (10)$$

where

$$f(x) = \frac{\Gamma(3 + 7/2)}{\Gamma(3) \Gamma(7/2)} x^{3-1} (1-x)^{7/2-1} \quad (11)$$

$$g(x) = \frac{\Gamma(3) \Gamma(7/2)}{\Gamma(3 + 7/2)} (-\log x)^2 x \quad (12)$$

The integral can be approximated by generating $X_1, X_2, \dots, X_N \stackrel{i.i.d}{\sim} \text{Beta}(3, 7/2)$ and calculate the average:

$$\hat{I}_N = \frac{1}{N} \sum_{i=1}^N g(X_i) \approx 0.006744701 \quad (13)$$

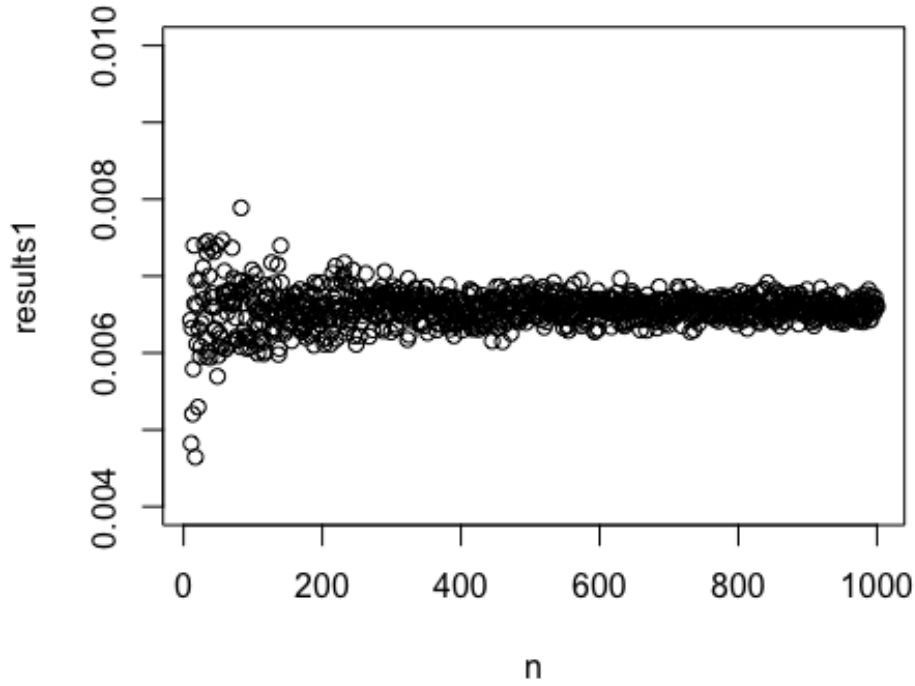
The variance is evaluated as

$$\text{Var}(\hat{I}_N) = \frac{\text{Var}(g(X_i))}{N} \quad (14)$$

which can be estimated by sample variance:

$$S(\hat{I}_N)^2 = \frac{S(g(X_i))^2}{N} = \frac{\sum_i (g(X_i) - \hat{I}_N)^2}{N(N-1)} \approx 1.198441 e - 08 \quad (15)$$

The demonstration of variance and convergence is shown in the following graph with the MC result as a function of sample size



Alternatively, we could use rejection sampling as

$$f(x) = (-\log x)^2 x^3 (1-x)^{5/2} = l(x) h(x) \quad (16)$$

$$h(x) = \frac{\Gamma(3 + 7/2)}{\Gamma(3) \Gamma(7/2)} x^{3-1} (1-x)^{7/2-1} \quad (17)$$

$$l(x) = \frac{\Gamma(3) \Gamma(7/2)}{\Gamma(3 + 7/2)} (-\log x)^2 x \quad (18)$$

$$I = \int_0^1 f(x) dx = \int_0^1 \int_0^1 M^{-1} \mathbf{1}(u < l(x)/M) h(x) du dx = \int_0^1 \int_0^1 f(x, u) du dx \quad (19)$$

where

$$M = \max \{l(x)\} = \frac{\Gamma(3) \Gamma(7/2)}{\Gamma(3 + 7/2)} 4 e^{-2} \quad (20)$$

The acceptance probability is given by

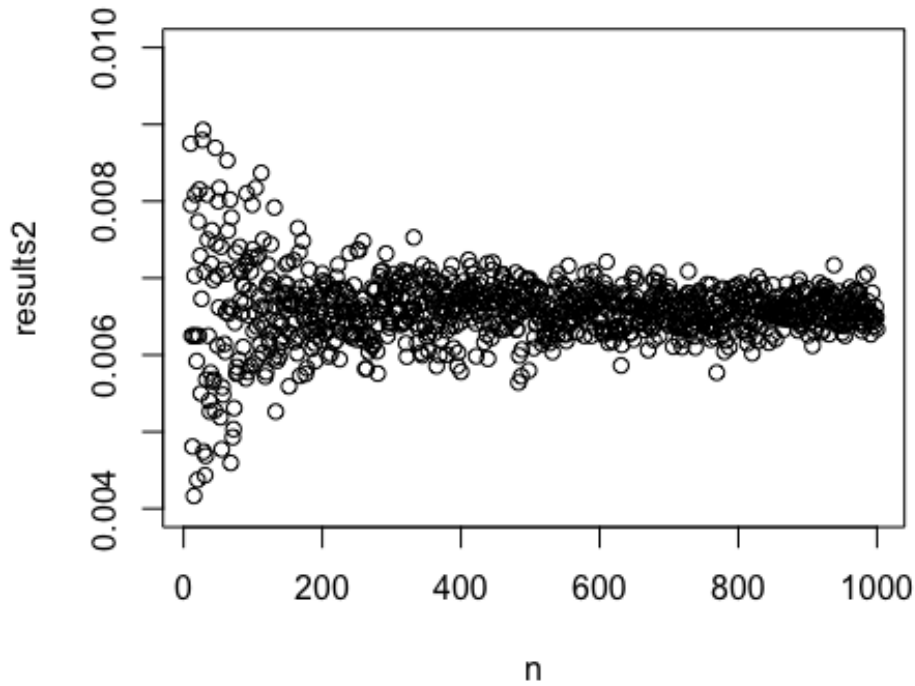
$$P = \frac{I}{M} \Rightarrow I = M \times P = \frac{\Gamma(3) \Gamma(7/2)}{\Gamma(3 + 7/2)} 4 e^{-2} \times P \approx 0.006424227 \quad (21)$$

where P is obtained by numerical simulation.

The variance is evaluated as

$$\begin{aligned} \text{Var}(\bar{I}_N) &= \frac{1}{N} \text{Var}(\mathbf{1}(u < l(x)/M)) = \frac{1}{N} [E(\mathbf{1}(u < l(x)/M)) - E(\mathbf{1}(u < l(x)/M))^2] \\ &= \frac{1}{N} \left[\frac{I}{M} - \left(\frac{I}{M} \right)^2 \right] = \frac{1}{N} (P - P^2) \approx 0.000249804 \end{aligned} \quad (22)$$

The demonstration of variance and convergence is shown in the following graph with the MC result as a function of sample size



As we can see, the variance is broader than the direct evaluation with beta distribution.

Problem 3

What is the acceptance probability when sampling a standard normal random variable with density

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \quad (23)$$

using a Cauchy density as proposal; i.e.

$$h(x) = \frac{1}{\pi(1+x^2)} \quad (24)$$

when using rejection sampling. Verify this using simulation and plot a histogram of 1000 accepted samples.

Solution: We can write $f(x)$ as

$$f(x) = h(x) l(x) \quad (25)$$

where

$$l(x) = \sqrt{\frac{\pi}{2}} e^{-x^2/2} (1+x^2) \quad (26)$$

which is bounded by

$$M = l(\pm 1) = \sqrt{2\pi} e^{-1/2} \quad (27)$$

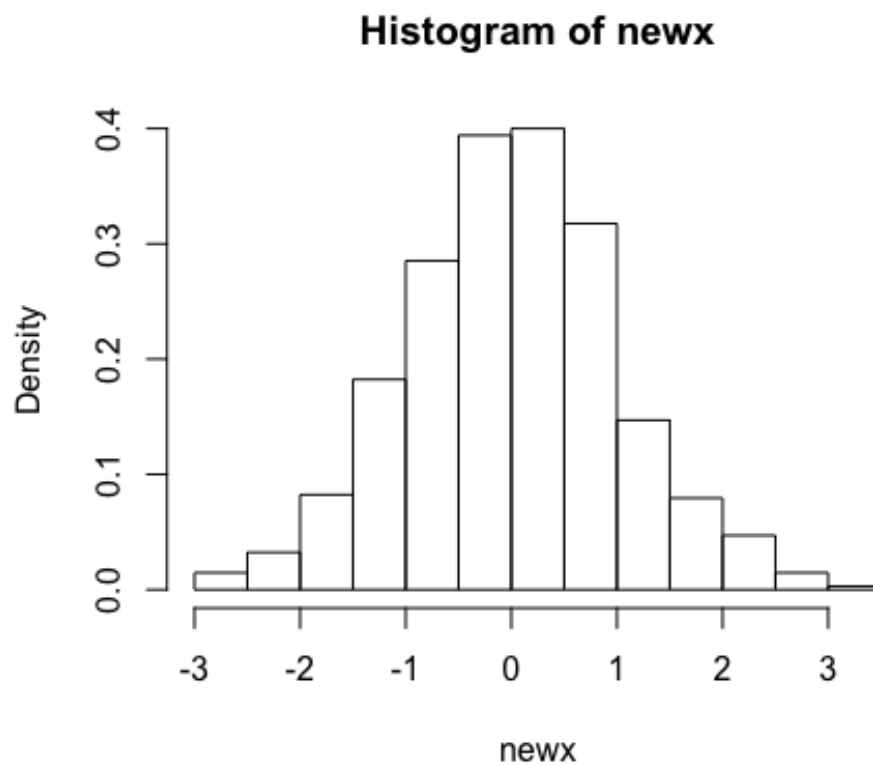
The acceptance probability is given by

$$P = 1/M = \frac{1}{\sqrt{2\pi} e^{-1/2}} \approx 0.657745 \quad (28)$$

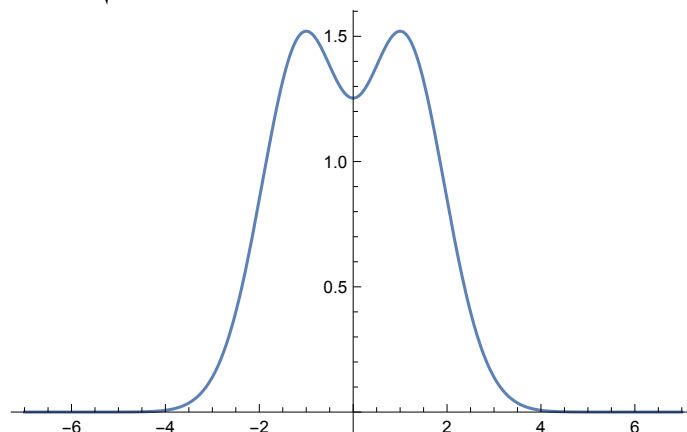
The simulation result gives

$$P \approx 0.68 \quad (29)$$

The histogram is as follows



```
Integrate[ $\frac{1}{\pi(1+x^2)}$ , {x, -∞, ∞}]
1
Plot[ $\sqrt{\frac{\pi}{2}} e^{-x^2/2} (1+x^2)$ , {x, -7, 7}, PlotRange → Full]
```



```
 $N\left[\frac{1}{\sqrt{2\pi}} e^{-1/2}\right]$ 
0.657745
```

Appendix: R-code

```
# Use rejection sampling to sample from the density function
#  $f(x) \propto (-\log x)^2 x^3 (1-x)^2, 0 < x < 1$ 
```

```

n <- 1000
x <- rbeta(n, shapel=3, shape2=3)
u <- runif(n, min=0, max=1)

sum(u<exp(2)*x*log(x)**2/4)/n

newx <- x[u<exp(2)*x*log(x)**2/4]
hist(newx,freq=FALSE)

# Use Monte Carlo methods to evaluate the integral
# I=\!\(
# \*SubsuperscriptBox[\(\int\), \{0\}, \{1\}]\(
# \*SuperscriptBox[\(\((-log\)\ x)\), \{2\}]
# \*SuperscriptBox[\(
# \*SuperscriptBox[\(x\), \{3\}](1 - x)\), \{5/2\}] dx\)\)

n<-1000

# analytical result
integrand <- function(x) {(log(x))**2*x**3*(1-x)**(5/2)}
integrate(integrand, lower = 0, upper = 1)

# Method I: sampling from beta distribution

x <- rbeta(n, shapel=3, shape2=7/2)
g <- gamma(3)*gamma(7/2)/gamma(3+7/2)*log(x)**2*x
I1 <- sum(g)/n
I1

# demonstration of convergence
I_N1 <- function(n){
  x <- rbeta(n, shapel=3, shape2=7/2)
  g <- gamma(3)*gamma(7/2)/gamma(3+7/2)*log(x)**2*x
  return(sum(g)/n)
}

n = 10:1000
results1 <- lapply(n,I_N1)
plot(n, results1, ylim=c(0.004, 0.01))

# variance estimation
varI1 = sum((g-mean(g))**2)/(n-1)/n
varI1

# Method II: rejection sampling

x <- rbeta(n, shapel=3, shape2=7/2)
u <- runif(n, min=0, max=1)

M=1/(sum(u<exp(2)*x*log(x)**2/4)/n)

I2 <- 4*gamma(3)*gamma(7/2)/(exp(2)*M*gamma(3+7/2))
I2

I_N2 <- function(n){
  x <- rbeta(n, shapel=3, shape2=7/2)
  u <- runif(n, min=0, max=1)
  M=1/(sum(u<exp(2)*x*log(x)**2/4)/n)
  return(4*gamma(3)*gamma(7/2)/(exp(2)*M*gamma(3+7/2)))
}

n = 10:1000
results2 <- lapply(n,I_N2)
plot(n, results2, ylim=c(0.004, 0.01))

# variance estimation

```

```
varI2 = 1/n*(1/M-(1/M)**2)
varI2

# What is the acceptance probability when sampling a standard normal random
variable with density
# f(x)=1/Sqrt[2π] E^(-x^2/2)
# using a Cauchy density as proposal; i.e.
# h(x)=1/π(1+x^2)
# when using rejection sampling. Verify this using simulation and plot a histogram
of 1000 accepted samples.

n <- 1000
x <- rcauchy(n, location=0, scale=1)
u <- runif(n, min=0, max=1)

sum(u < 1/2*exp(-(x**2-1)/2)*(1+x**2))/n

newx <- x[u<1/2*exp(-(x**2-1)/2)*(1+x**2)]
hist(newx,freq=FALSE)
```
