# The statistical model behind simmr (and SIAR)

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#### Learning outcomes:

- Understand the statistical model behind simmr/SIAR
- ► Know how to run a model in simmr/SIAR and check that it works
- ▶ Be able to follow the technical details of the 2010 SIAR Plos ONE paper

# Our simple SIMM

In the last class we had a simple SIMM defined via:

$$y_i \sim N\left(\sum_{k=1}^2 p_k s_k, \sigma^2\right)$$

with  $s_k \sim N(\mu_{s_k}, \sigma_{s_k}^2)$ ,  $p_1 \sim U(0, 1)$  and  $\sigma \sim U(0, 100)$ 

- ▶ Here  $y_i$  is the isotope value, s are the source values, p are the dietary proportions, and  $\sigma$  is the residual standard deviation
- ▶ The goal is to estimate the *p* and its uncertainty. The other parameters can be considered nuisance parameters

## Expanding the simple SIMM

- This SIMM is currently too simplistic. We need to expand it by:
  - increasing the number of food sources
  - including trophic enrichment factors (TEFs)
  - including concentration dependence
  - allowing for multiple isotopes
  - allowing for richer source sampling by consumers
- ▶ If we include all of these factors we end up with the simmr/SIAR model
- ▶ We will take them in turn and add them into our JAGS code

## Reminder: the SIAR geese data data(geese data day1) str(geese\_data\_day1)

```
## List of 8
  $ mixtures : num [1:9, 1:2] -11.4 -11.9 -10.6 -11.2 -11.7 ...
##
  ..- attr(*, "dimnames")=List of 2
##
## ...$ : NULL
## ....$ : chr [1:2] "d13C Pl" "d15N Pl"
##
   $ tracer names : chr [1:2, 1] "d13C Pl" "d15N Pl"
```

\$ source means : num [1:4, 1:2] -11.17 -30.88 -11.17 -14.06 6.49 ## ..- attr(\*, "dimnames")=List of 2 ## ## ...\$ : NULL

\$ source names : chr [1:4] "Zostera" "Grass" "U.lactuca" "Entero

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....\$ : chr [1:2] "meand13CP1" "meand15NP1" ## \$ source sds : num [1:4, 1:2] 1.215 0.641 1.959 1.172 1.459 .. ##

.. ..\$ : NULL

..- attr(\*, "dimnames")=List of 2

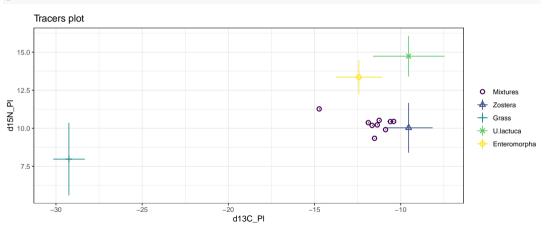
##

## ##

# Plotting the data

A plot in isotope space:

plot(simmr\_in)



### Including multiple sources

▶ Adding in multiple sources to the likelihood means having more terms in the sum:

$$y_i \sim N\left(\sum_{k=1}^K p_k s_k, \sigma^2\right)$$

- ▶ In the above we have K sources and hence K dietary proportions
- ▶ We also now need K source prior distributions
- ▶ The tricky part about adding in multiple proportions is the prior distribution

## Priors for constrained dietary proportions

- We must have  $\sum_{k=1}^{K} p_k = 1$  so any prior distribution we place on the ps must satisfy this restriction
- ► (You will often hear values restricted in sum referred to as a *simplex*)
- ► Luckily there is a distribution known as the *Dirichlet* which is suitable for restricted sum parameters
- The Dirichlet has one parameter for each proportion  $\alpha_1, \ldots, \alpha_K$ . The larger the  $\alpha$  value the larger prior weight that dietary proportion will be given
- lacktriangleright Setting all the lpha values to 1 is equivalent to the simplex uniform distribution, i.e. a prior assumption that all sources are consumed equally

#### JAGS SIMM with a Dirichlet prior

```
model code ='
model {
  for(i in 1:N) { y[i] ~ dnorm(inprod(p,s),sigma^-2) }
  p ~ ddirch(alpha)
  for (k \text{ in } 1:K) \{ s[k] \sim dnorm(s mean[k], s sd[k]^{-2}) \}
  sigma \sim dunif(0,100)
data=with(geese_data_day1,
          list(v=mixtures[,1],
                s mean=source means[,1],
                s sd=source sds[.1].
          N = nrow(mixtures), K=nrow(source_means),
          alpha=rep(1,nrow(source_means))))
set.seed(123)
model_run = jags(data = data,
                  parameters.to.save = c("p", "sigma"),
                  model.file = textConnection(model code))
```

#### Results

- ► We can explore/plot results with summary(output), plot(output), and also run multiple chains, form predictive distributions, check convergence, etc
- ▶ One important thing to note is that the fitting method (MCMC) produces a joint posterior distribution of the dietary proportions. This means that each set of samples will sum to 1:

#### head(model\_run\$BUGSoutput\$sims.matrix,4)

```
## deviance p[1] p[2] p[3] p[4] sigma
## [1,] 38.24013 0.2931735 0.037865208 0.6128807 0.05608053 2.666854
## [2,] 29.94426 0.2925945 0.049061414 0.4754383 0.18290585 1.237191
## [3,] 29.19475 0.3280195 0.051608762 0.4583445 0.16202724 1.385694
## [4,] 29.35735 0.5883123 0.004139584 0.1634236 0.24412454 1.037509
```

► The key implication of this is that, aside from exploring the *marginal* posterior distributions (with means, sds, etc) we can explore the *joint* uncertainty of the dietary proportions

## A joint plot of the posterior dietary proportions

```
out 2 = model run$BUGSoutput$sims.list$p
colnames(out_2) = geese_data_day1$source_names
pairs(out_2, lower.panel = panel.smooth,
      upper.panel = panel.cor)
                    0.00
                         0.05
                              0.10
                                   0.15
                                           -0.72
                                                               -0.46
       Zostera
                           Grass
                                               0 14
                                                                 -0.21
                                            Ulactuca
                                                                 -0.28
                                                             Enteromorpha
```

### Trophic enrichment factors and concentration dependence

- Trophic enrichment factors (c) and concentration dependence (q) represent adjustments to the source values to account for various measurement effects
- ▶ We can include them by expanding the likelihood:

$$y_i \sim N\left(\frac{\sum_{k=1}^K p_k q_k(s_k + c_k)}{\sum_{k=1}^K p_k q_k}, \sigma^2\right)$$

- ► The extra part on the denominator is needed so that the dietary proportions still sum to 1
- ightharpoonup The prior for  $c_k$  comes from external data and are given normal distributions like the source values
- ▶ In SIAR the concentration dependencies must be less than 1 (given as proportions) and are treated as fixed. You could use a strong Dirichlet prior on these instead

# Including TEFs and CD - JAGS model

```
model code = '
model {
 for(i in 1:N) {
   y[i] ~ dnorm(inprod(p*q,s+c)/inprod(p,q),sigma^-2)
  p ~ ddirch(alpha)
 for(k in 1:K) {
    s[k] ~ dnorm(s mean[k],s sd[k]^-2)
   c[k] ~ dnorm(c mean[k],c sd[k]^-2)
  sigma ~ dunif(0,100)
data=with(geese_data_day1,
         list(y=mixtures[,1], s_mean=source_means[,1],
              s_sd=source_sds[,1], c_mean = correction_means[,1],
              c sd = correction sds[,1], q = concentration means[,1],
         N = nrow(mixtures), K=nrow(source means),
         alpha=rep(1,length(source names))))
```

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#### Notes on the TEF and CD model

- ► If you run this, you'll find that convergence isn't quite as neat and it starts to get a bit slower
- ► Although it's a nuisance parameter, saving sigma is often a good idea because a large value indicates a poorly fitting model (usually also seen in the iso-space plot)
- ▶ The model will also create posterior distributions for s and c, though these are usually pretty similar to the prior, as there isn't much information about their values in the data

### Adding extra isotopes

- ▶ If we have extra isotopes we can just list the likelihood twice, once for each value of the isotope. Only the dietary proportions are 'shared' between the isotopes
- Now write  $y_{ij}$  as the consumer values for observation i on isotope j, where  $j=1,\ldots,J$
- We now have source values  $s_{jk}$ , TEF values  $c_{jk}$ , concentration dependencies  $q_{jk}$ , and each isotope has its own residual standard deviation  $\sigma_i$
- ► The likelihood is now:

$$y_{ij} \sim N\left(\frac{\sum_{k=1}^{K} p_k q_{jk}(s_{jk} + c_{jk})}{\sum_{k=1}^{K} p_k q_{jk}}, \sigma_j^2\right)$$

### Richer source sampling

- The model we've been fitting up to now assumes that all individuals sample the same source value  $s_k$  for each source and isotope. This is unrealistic
- A better model has each individual sampling a different source value from the source prior distribution, i.e. we now have  $s_{ik}$  (or  $s_{iki}$  with multiple isotopes)
- ▶ The JAGS code becomes:

```
for(k in 1:K) {
  for(i in 1:N) {
    s[i,k] ~ dnorm(s_mean[k],s_sd[k]^-2)
  }
}
```

- ▶ We can do the same with the concentration dependence values
- ▶ In fact with a bit of clever maths we can remove (marginalise over) the  $s_{ik}$  values to get a simpler model with fewer parameters.

# The full simmr/SIAR model

Using the trick mentioned on the last slide, we end up with a full model which looks like this:

$$y_{ij} \sim N\left(\frac{\sum_{k=1}^{K} p_k q_{jk} (\mu_{s,jk} + \mu_{c,jk})}{\sum_{k=1}^{K} p_k q_{jk}}, \frac{\sum_{k=1}^{K} p_k^2 q_{jk}^2 (\sigma_{s,jk}^2 + \sigma_{c,jk}^2)}{(\sum_{k=1}^{K} p_k q_{jk})^2} + \sigma_j^2\right)$$

► This model has a more complicated likelihood, but removes the extra s and c parameters

# Full SIAR model: JAGS code

```
model code ='
model {
              for (i in 1:N) {
                              for (j in 1:J) {
                                             y[i,j] ~ dnorm(inprod(p*q[,j], s_mean[,j]+c_mean[,j]) /
                                                             inprod(p,q[,j]), var_y[j]^-1)
               p ~ ddirch(alpha)
               for(j in 1:J) {
                             var y[j] < -inprod(pow(p*q[,j],2),s sd[,j]^2+c sd[,j]^2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j],2)/pow(inprod(p,-sd[,j
                                            + pow(sigma[i].2)
              for(j in 1:J) { sigma[j] ~ dunif(0,100) }
```

# Full simmr/SIAR model: R code

### Summary of posterior dietary proportions

0.05 0.86 2.45

0.02 0.36 1.40

## SD1

## SD2

Some of these proportions are quite imprecise: perhaps see better with matrix plot?

# Running SIAR/simmr

- ▶ The SIAR/simmr R packages run exactly this model with a few extra tweaks
- ▶ It contains a slightly optimised algorithm as JAGS sometimes gets a bit stuck on harder data sets. It's also much faster than JAGS for complicated problems
- ▶ It allows for direct plotting of the data in isotope space and *p*-space (i.e. dietary proportion space pairs plots)
- lacktriangle It allows for changing the lpha values to put in proper prior information
- ▶ It includes convergence checking
- Most of this covered in the practical this afternoon

#### simmr version

- ▶ simmr is a much more elegantly written version of SIAR with neater plots and many more features
- Four steps to run a simmr model
  - 1. Call simmr\_load to load in the data
  - 2. Call plot to see the iso-space plot
  - 3. Call simmr\_mcmc to run the model
  - 4. Call plot or summary to access the output
- simmr has further features to combine sources and to compare dietary proportions

#### simmr code

```
# I.oa.d.
data("geese_data_day1")
simmr_in = with(geese_data_day1,
                simmr load(mixtures = mixtures,
                      source names = source names,
                      source means = source means,
                      source sds = source sds,
                      correction means = correction means,
                      correction sds = correction sds,
                       concentration means = concentration means))
# Iso-space plot
plot(simmr_in)
# MCMC run
simmr out = simmr mcmc(simmr in)
# Box-plots
plot(simmr out, type = 'boxplot')
```

## Summary

- ▶ The simmr and SIAR models are just complicated versions of linear regression
- ► The response is multivariate and the prior distributions on some of the parameters have to be constrained to sum to 1
- ▶ It used to be the case that JAGS was slow and couldn't run SIMM-type models. This is no longer true. You can fit much richer models in JAGS (and now MixSIAR) than with SIAR/simmr
- ▶ More details on running simmr in the practical next