

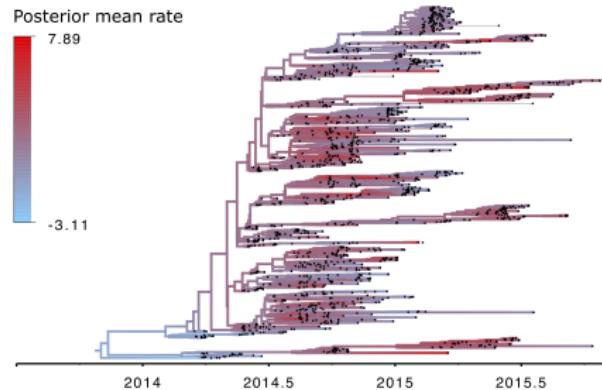
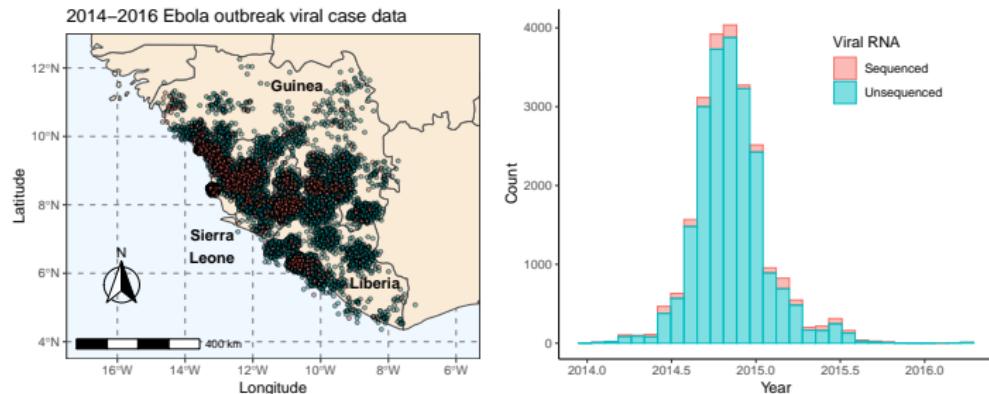
A quantum parallel Markov chain Monte Carlo

Andrew J. Holbrook

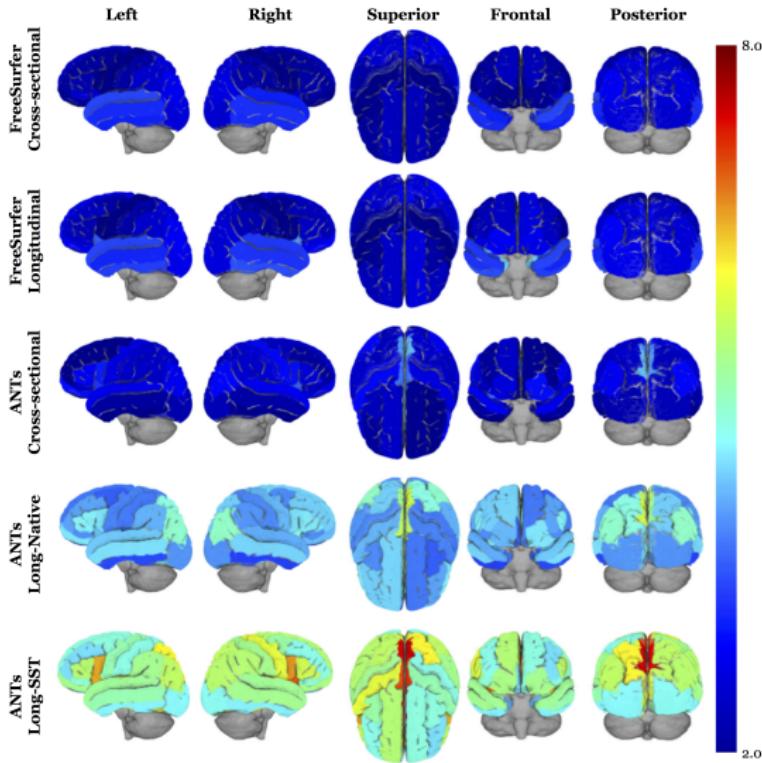
UCLA Biostatistics

March 11, 2022

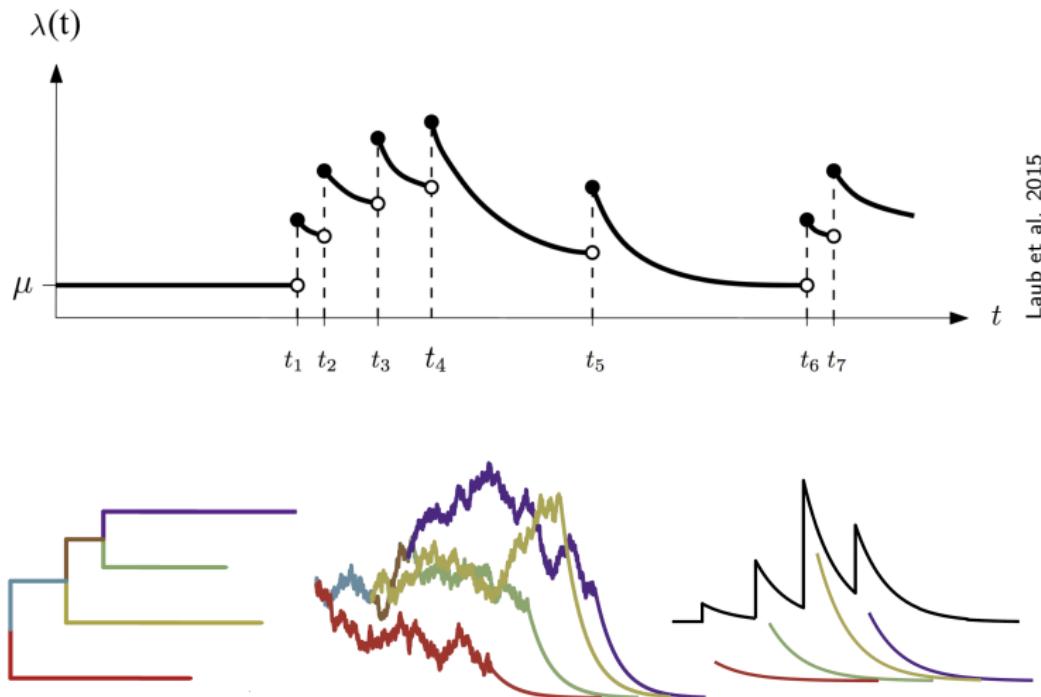
The data science perspective



The data science perspective

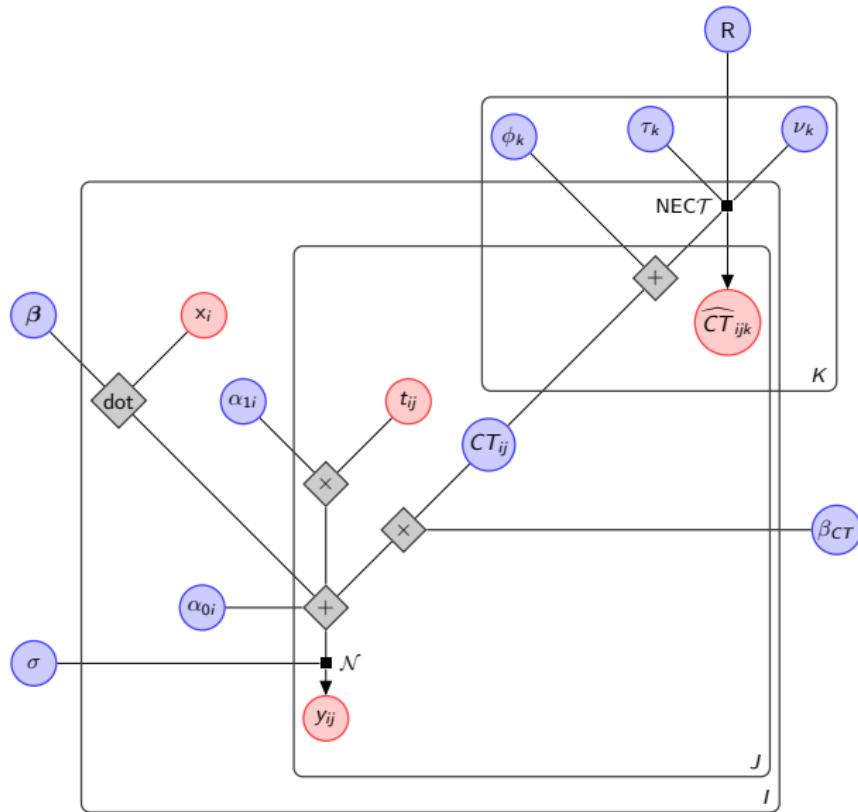


Hierarchical models for mixed-type data



Laub et al. 2015

Hierarchical models for mixed-type data



Bayesian data analysis

Assume data generated according to $y_n \stackrel{\perp}{\sim} f(y_n|\theta, z_n)$ with prior distributions $\theta \sim p_\theta(\theta)$ and $(z_1, \dots, z_N) = Z \sim p_z(Z)$.

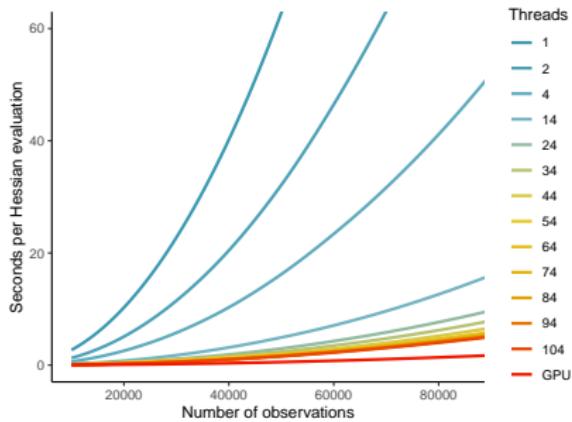
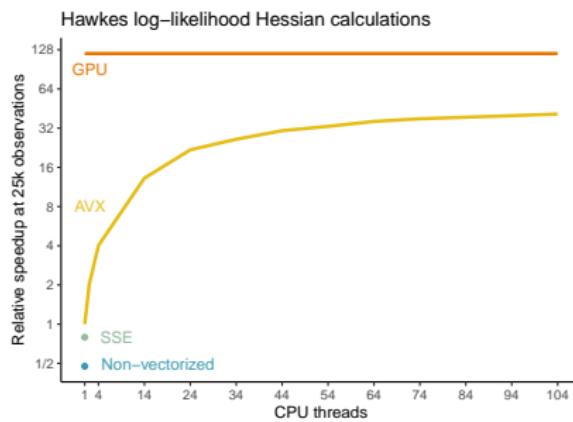
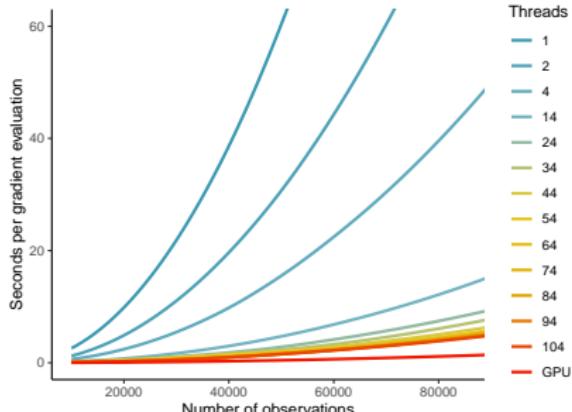
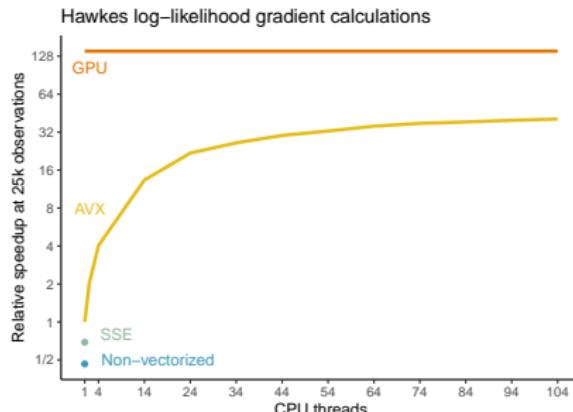
Bayes' theorem says:

$$p(\theta|Y) = \frac{f(Y|\theta) p_\theta(\theta)}{f(Y)} = \frac{\int_Z f(Y|Z, \theta) p_z(Z) dZ p_\theta(\theta)}{\int_{\Theta} \left(\int_Z f(Y|Z, \theta) p_z(Z) dZ \right) p_\theta(\theta) d\theta},$$

where $f(Y|\theta, Z) = \prod_n f(y_n|\theta, z_n)$ is the *likelihood* function and $f(Y|\theta)$ is the *marginal likelihood*.

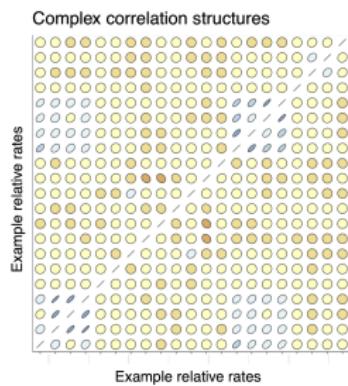
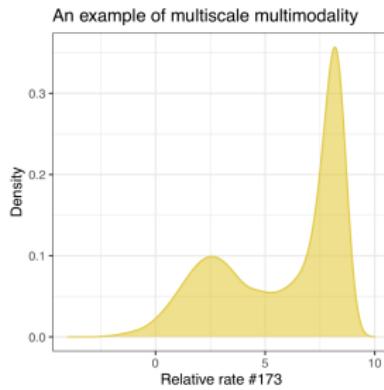
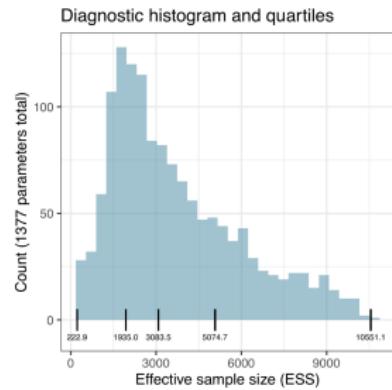
Instead of integrating, we use MCMC to generate samples from $p(Z, \theta|Y)$, but this is computationally demanding.

Parallelizing within-chain computations



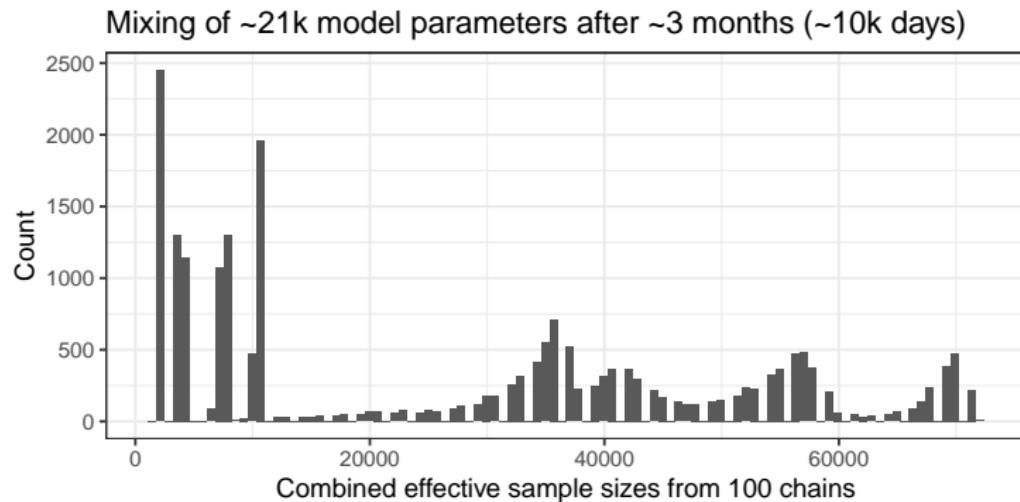
Parallelizing within-chain computations

I use Hamiltonian (hybrid) Monte Carlo with adaptive preconditioning to generate 100 million Markov chain states (~ 3.5 million samples/day on Nvidia GV100) in 1 month.



Parallelizing between-chain computations

Cortical thickness: 100 parallel chains using No-U-Turn with adaptive mass matrix.



Two questions

Neither within- nor between-chain parallelization is a “silver bullet.” Here, I am interested in two questions.

- ▶ Can we adjust the deeper *algorithmic* structure of MCMC to increase its parallelism?

- ▶ How can we leverage parallelism using emerging computational technologies?

Markov chain Monte Carlo

Consider a probability distribution $\pi(d\theta)$ defined on \mathbb{R}^D that admits a probability density $\pi(\theta)$ with respect to the Lebesgue measure, i.e., $\pi(d\theta) =: \pi(\theta)d\theta$.

To generate samples from the target distribution π , we craft a kernel $P(\theta_0, d\theta)$ that satisfies

$$\pi(A) = \int \pi(d\theta_0) P(\theta_0, A) \quad a.s.$$

for $A \subset \mathbb{R}^D$ for which $\pi(A) > 0$.

Markov chain Monte Carlo

The Metropolis-Hastings algorithm builds such a transition kernel $P(\theta_0, d\theta)$ by:

1. generating a proposal θ according to the distribution $Q(\theta_0, d\theta) =: q(\theta_0, \theta)d\theta$; and
2. accepting this proposal with probability

$$\pi = 1 \wedge \frac{\pi(\theta)q(\theta, \theta_0)}{\pi(\theta_0)q(\theta_0, \theta)}.$$

This kernel maintains *detailed balance* and therefore leaves $\pi(d\theta)$ invariant.

Parallel MCMC

The parallel MCMC algorithm ([Tjelmeland, 2004](#)) builds such a transition kernel $P(\theta_0, d\theta)$ by:

1. generating P proposals $\Theta_{-0} = (\theta_1, \dots, \theta_P)$ from a joint distribution $Q(\theta_0, d\Theta_{-0}) =: q(\theta_0, \Theta_{-0})d\Theta_{-0}$; and
2. selecting the next state with probabilities

$$\pi_p = \frac{\pi(\theta_p)q(\theta_p, \Theta_{-p})}{\sum_{p'=0}^P \pi(\theta_{p'})q(\theta_{p'}, \Theta_{-p'})}, \quad p = 0, \dots, P.$$

This kernel *also* maintains detailed balance and therefore leaves $\pi(d\theta)$ invariant.

Parallel MCMC

We can simplify these acceptance probabilities to remove a potentially $O(P^2)$ computational burden.

For example, the proposal mechanism

$$\boldsymbol{\theta}_1, \dots, \boldsymbol{\theta}_P \stackrel{\perp}{\sim} \text{Normal}_D(\bar{\boldsymbol{\theta}}, \Sigma), \quad \bar{\boldsymbol{\theta}} \sim \text{Normal}_D(\boldsymbol{\theta}_0, \Sigma)$$

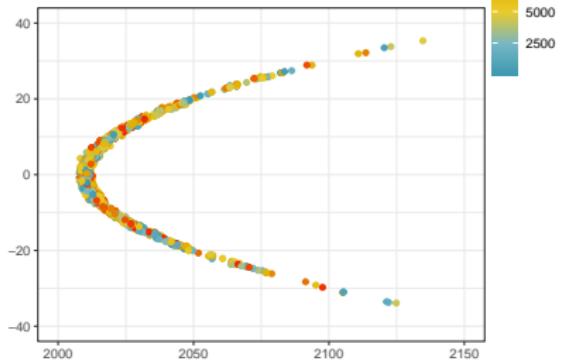
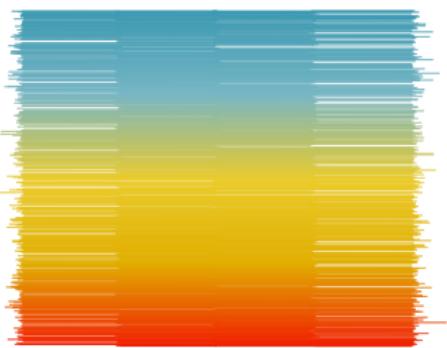
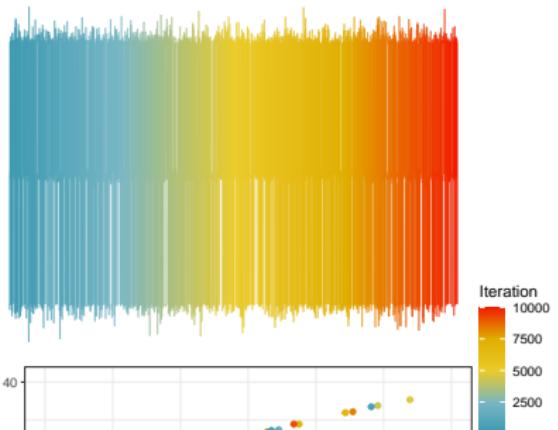
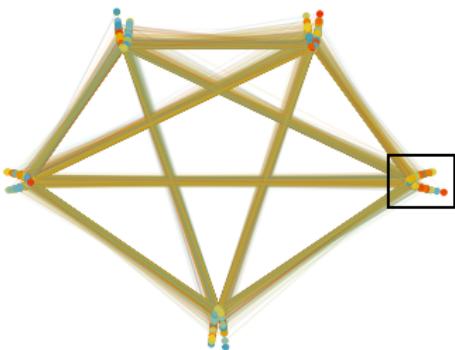
results in proposal densities that satisfy

$$q(\boldsymbol{\theta}_0, \Theta_{-0}) = q(\boldsymbol{\theta}_1, \Theta_{-1}) = \dots = q(\boldsymbol{\theta}_P, \Theta_{-P}).$$

Thus, these simplified probabilities preserve detailed balance:

$$\pi_p = \frac{\pi(\boldsymbol{\theta}_p)}{\sum_{p'=0}^P \pi(\boldsymbol{\theta}_{p'})}, \quad p = 0, \dots, P.$$

Parallel MCMC



Parallel MCMC

Pro: fast and flexible on an iteration-by-iteration basis.

Con: each iteration requires $O(P)$ evaluations of the target $\pi(\cdot)$ in

$$\pi_p = \frac{\pi(\theta_p)}{\sum_{p'=0}^P \pi(\theta_{p'})}, \quad p = 0, \dots, P.$$

But these evaluations do not depend on each other, so...
let's parallelize using a quantum computer!

$$|0\rangle^{\otimes n} |0\rangle \longrightarrow \left(\frac{1}{\sqrt{N}} \sum_{x=0}^{N-1} |x\rangle \right) |0\rangle \longrightarrow \frac{1}{\sqrt{N}} \sum_{x=0}^{N-1} |x\rangle |\pi(x)\rangle ???$$

Not so fast!

A quantum circuit could plausibly evaluate each $\pi(\theta_p)$, for $p = 0, 1, \dots, P$ at the same time using quantum parallelism.

But having done this, the best we could hope for is to obtain a single $\pi(\theta_p)$ upon measurement.

Hmmm, maybe we can also use the quantum circuit to sample

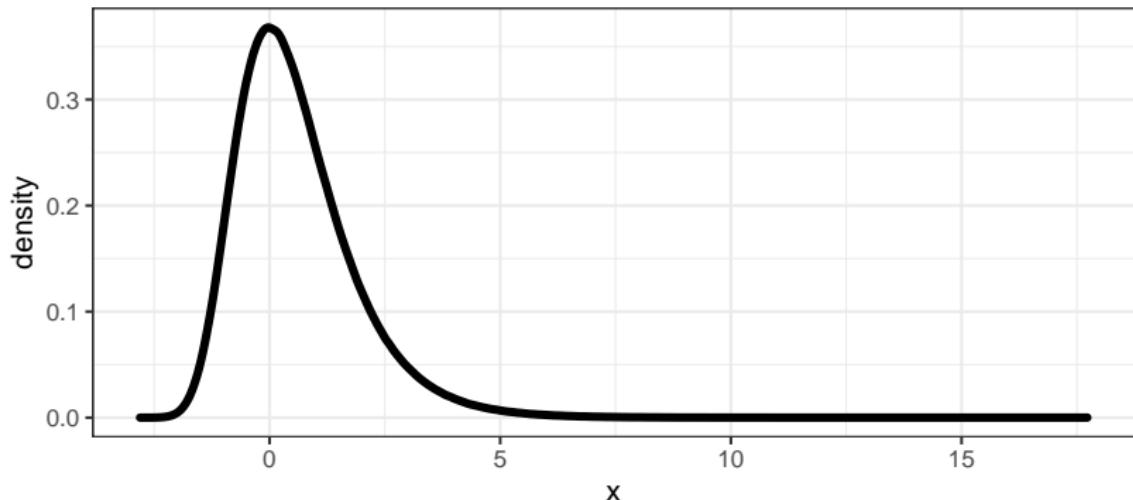
$$\hat{p} \sim \text{Discrete} \left(\frac{\pi(\theta_0)}{\sum_p \pi(\theta_p)}, \frac{\pi(\theta_1)}{\sum_p \pi(\theta_p)}, \dots, \frac{\pi(\theta_P)}{\sum_p \pi(\theta_p)} \right)$$

Maybe? But doesn't the normalization step violate quantum parallelism?

Detour: the Gumbel-max trick

The Gumbel distribution

Standard Gumbel distribution density



If $z \sim Gumbel(0, 1)$, then it has density and distribution functions

$$g(z) = \exp(-z - \exp(-z)) \quad \text{and} \quad G(z) = \exp(-\exp(-z)).$$

Gumbel-max trick

We wish to sample from the discrete distribution $\hat{p} \sim \text{Discrete}(\pi)$ for $\hat{p} \in \{0, 1, \dots, P\}$ and we only know $\pi^* = c\pi$ for some $c > 0$.

Define $\lambda^* = \log \pi^* = \log \pi + \log c$ and suppose
 $z_0, z_1, \dots, z_P \stackrel{\perp}{\sim} \text{Gumbel}(0, 1)$.

Finally, define $\alpha_p^* := \lambda_p^* + z_p$ and $\hat{p} = \arg \max_{p=0, \dots, P} \alpha_p^*$.

Then the following holds ([Papandreou and Yuille, 2011](#)):

$$\Pr(\hat{p} = p) = \pi_p, \quad p = 0, 1, \dots, P.$$

Data: A vector of unnormalized log-probabilities
 $\lambda^* = \log \pi + \log c$, for π a discrete probability vector with $P + 1$ elements.

Result: A single sample $\hat{p} \sim \text{Discrete}(\pi)$ satisfying
 $\hat{p} \in \{0, 1, \dots, P\}$.

for $p \in \{0, 1, \dots, P\}$ **do**

$z_p \leftarrow \text{Gumbel}(0, 1);$

$\alpha_p^* \leftarrow \lambda_p^* + z_p;$

end

$\hat{p} \leftarrow \arg \max_{p=0, \dots, P} \alpha_p^*$;

return \hat{p} .

Algorithm 1: The Gumbel-max trick

End of detour.

Data: Initial Markov chain state $\theta^{(0)}$; total length of Markov chain S ; total number of proposals per iteration P .

Result: A Markov chain $\theta^{(1)}, \dots, \theta^{(S)}$.

for $s \in \{1, \dots, S\}$ **do**

$\theta_0 \leftarrow \theta^{(s-1)}$;

$\bar{\theta} \leftarrow \text{Normal}_D(\theta_0, \Sigma)$;

$z_0 \leftarrow \text{Gumbel}(0, 1)$;

for $p \in \{1, \dots, P\}$ **do**

$\theta_p \leftarrow \text{Normal}_D(\bar{\theta}, \Sigma)$;

$z_p \leftarrow \text{Gumbel}(0, 1)$;

end

$\hat{p} \leftarrow \arg \min_{p=0, \dots, P} \left(f(p) := -(z_p + \log \pi(\theta_p)) \right)$;

$\theta^{(s)} \leftarrow \theta_{\hat{p}}$;

end

return $\theta^{(1)}, \dots, \theta^{(S)}$.

Algorithm 2: An equivalent parallel MCMC

The main idea of QPMCMC

Use a quantum circuit to obtain

$$\hat{p} = \arg \min_{p=0,\dots,P} \left(f(p) := -(z_p + \log \pi(\theta_p)) \right)$$

in time $O(\sqrt{P})$.

We use a Matryoshka doll of established quantum algorithms:

- ▶ Grover's search algorithm ([Grover, 1996](#)) embeds inside ...
- ▶ ... the exponential searching algorithm ([Boyer et al., 1998](#)) embeds inside ...
- ▶ the Quantum minimization algorithm ([Dürr and Høyer, 1996](#)).

Grover's search

Data: An oracle gate U_f taking $|x\rangle|y\rangle$ to $|x\rangle|y \oplus f(x)\rangle$ for a function $f(x) : \{0, \dots, N-1\} \rightarrow \{0, 1\}$ that satisfies $f(x_0) = 1$ for a single x_0 ; $n+1 = \log_2(N) + 1$ quantum states initialized to $|0\rangle^{\otimes n}|1\rangle$; an integer $R = \lceil \pi\sqrt{N}/4 \rceil$.

Result: An n -bit binary string x_0 satisfying $f(x_0) = 1$ with error less than $1/N$.

$$|0\rangle^{\otimes n}|1\rangle \longrightarrow H^{\otimes n+1}|0\rangle^{\otimes n}|1\rangle = \frac{1}{\sqrt{N}} \sum_{x=0}^{N-1} |x\rangle|- \rangle;$$

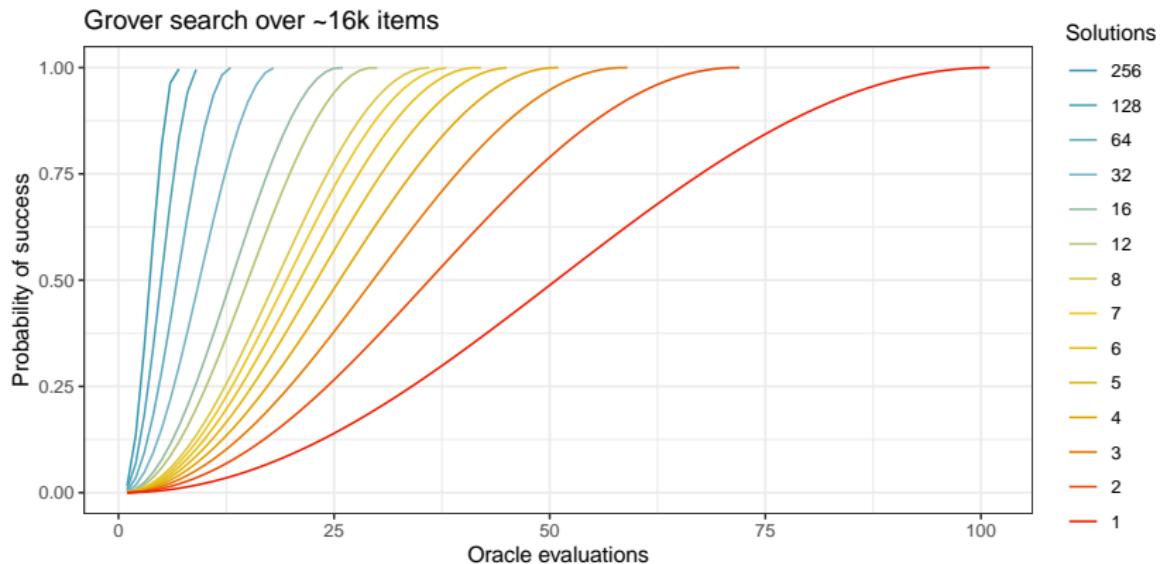
$$|h\rangle|- \rangle \longrightarrow ((2|h\rangle\langle h| - I)(I - 2|x_0\rangle\langle x_0|))^R |h\rangle|- \rangle \approx |x_0\rangle|- \rangle;$$

$$|x_0\rangle \longrightarrow x_0;$$

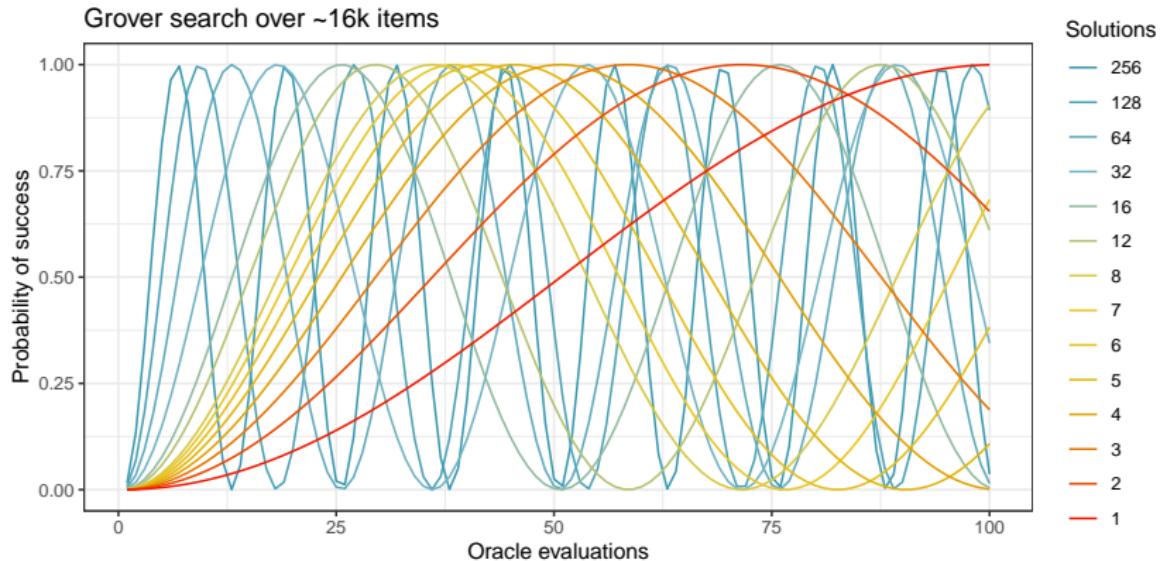
return x_0 .

Algorithm 3: Quantum search algorithm ([Grover, 1996](#))

Grover's search



Grover's search



What do we do when we don't know M ?

Data: An oracle gate U_f taking $|x\rangle|y\rangle$ to $|x\rangle|y \oplus f(x)\rangle$ for a function $f(x) : \{0, \dots, N-1\} \rightarrow \{0, 1\}$ with unknown number of solutions; $n = \log_2(N)$.

Result: If a solution exists, an n -bit binary string x_0 satisfying $f(x_0) = 1$; if no solution exists, the algorithm runs forever.

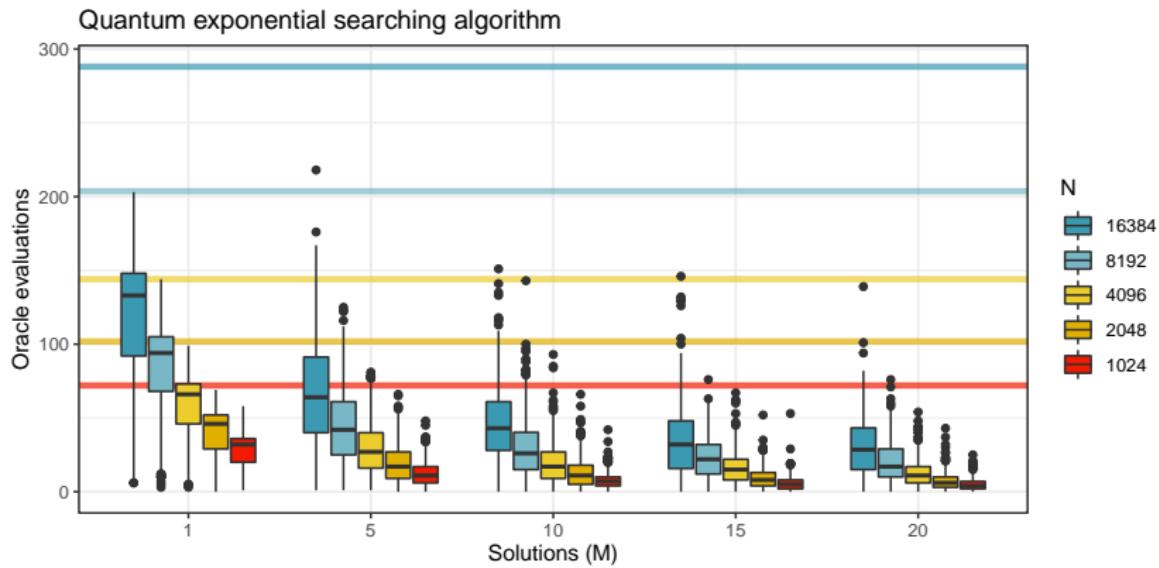
```

 $m \leftarrow 1;$ 
 $\gamma \leftarrow 6/5;$ 
success  $\leftarrow$  FALSE;
while success  $\neq$  TRUE do
     $j \leftarrow \text{Uniform}\{0, \dots, m-1\};$ 
     $|0\rangle^{\otimes n}|1\rangle \longrightarrow H^{\otimes n+1}|0\rangle^{\otimes n}|1\rangle = |h\rangle|-\rangle;$ 
     $|h\rangle|-\rangle \longrightarrow G^j|h\rangle|-\rangle = |x\rangle|-\rangle; \quad /* j Grover iterations. */$ 
     $|x\rangle \longrightarrow x; \quad /* \text{Measure and check. */}$ 
    if  $f(x) = 1$  then
         $x_0 \leftarrow x;$ 
        success  $\leftarrow$  TRUE;
    else
         $m \leftarrow \min(\gamma m, \sqrt{N}); \quad /* \text{If fail, increase } m. */$ 
    end
end
return  $x_0$ .

```

Algorithm 4: Exponential searching algorithm (Boyer et al., 1998)

Exponential searching algorithm



When $M \ll N$, the expected total number of Grover iterations is bounded above by $\frac{9}{4}\sqrt{\frac{N}{M}}$. Horizontal lines at $\frac{9}{4}\sqrt{N}$.

Data: A quantum sub-routine capable of evaluating a function $f(\cdot)$ over $\{0, \dots, N - 1\}$ with unique integer values; a maximum error tolerance $\epsilon \in (0, 1)$; expected total time to success $m_0 = \frac{45}{4}\sqrt{N} + \frac{7}{10}\log_2(N)$.

Result: A $\log_2(N)$ -bit binary string x_0 satisfying $f(x_0) = \min f$ with probability greater than $1 - \epsilon$.

```
s ← 0;  
x0 ← Uniform{0, ..., N − 1};  
while s < m0/ε do  
    Prepare initial state  $\frac{1}{\sqrt{N}} \sum_x |x\rangle |x_0\rangle$ ;  
    Mark all items x satisfying  $f(x) < f(x_0)$ ;  
    s ← s + log2(N);  
    Apply quantum exponential searching algorithm; /* / time steps */  
    s ← s + l;  
    Obtain x' by measuring first register;  
    if  $f(x') < f(x_0)$  then  
        | x0 ← x'  
    end  
end  
return x0 .
```

Algorithm 5: Quantum minimization algorithm ([Dürr and Høyer, 1996](#))

Quantum minimization and warm-starting

Proposition

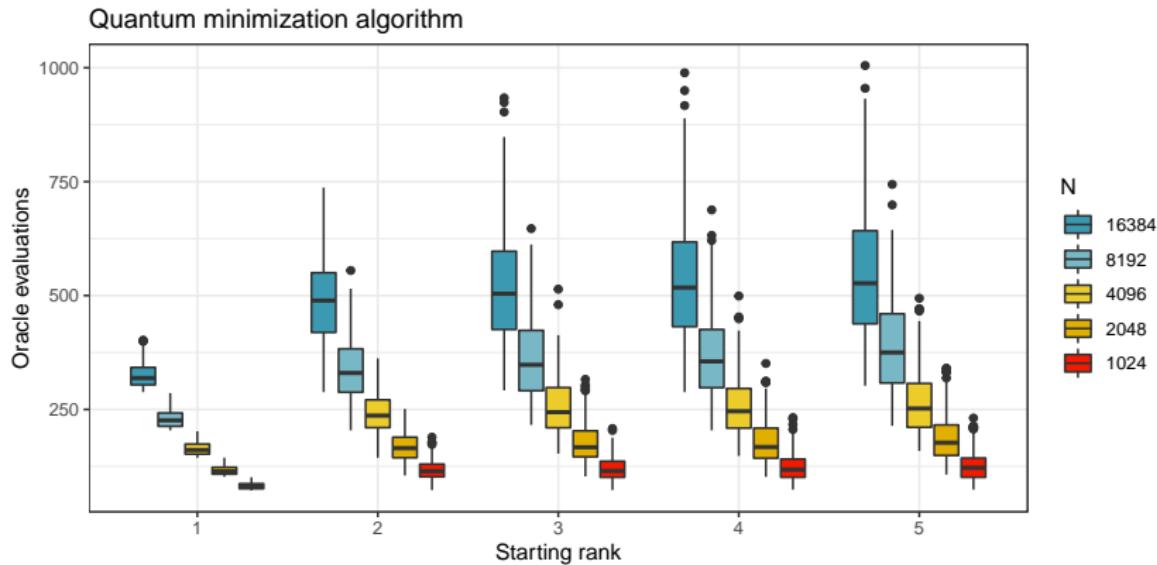
Suppose that the quantum minimization algorithm begins with a threshold F_0 such that $f(x) < F_0$ for only $K - 1$ items. Then the expected total number of time steps to find the minimizer is bounded above by

$$m_0^K = \left(\frac{5}{4} - \frac{1}{\sqrt{K-1}} \right) 9\sqrt{N} + \frac{7}{10} \log_2(K) \log_2(N),$$

and so the following rule relates the warm-started upper bound to the generic upper bound:

$$m_0^K = m_0 - 9\sqrt{\frac{N}{K-1}} + \frac{7}{10} \log_2 \left(\frac{K}{N} \right) \log_2(N).$$

Quantum minimization and warm-starting



Using the same early stopping threshold within the exponential search sub-routine, we observe a failure rate less than 1%.

Data: Initial Markov chain state $\theta^{(0)}$; total length of Markov chain S ; total number of proposals per iteration P .

Result: A Markov chain $\theta^{(1)}, \dots, \theta^{(S)}$.

for $s \in \{1, \dots, S\}$ **do**

$\theta_0 \leftarrow \theta^{(s-1)}$;

$\bar{\theta} \leftarrow \text{Normal}_D(\theta_0, \Sigma)$;

$z_0 \leftarrow \text{Gumbel}(0, 1)$;

for $p \in \{1, \dots, P\}$ **do**

$\theta_p \leftarrow \text{Normal}_D(\bar{\theta}, \Sigma)$;

$z_p \leftarrow \text{Gumbel}(0, 1)$;

end

$\hat{p} \leftarrow \arg \min_{p=0, \dots, P} \left(f(p) := -(z_p + \log \pi(\theta_p)) \right)$;

/* Quantum minimization starting at 0. */

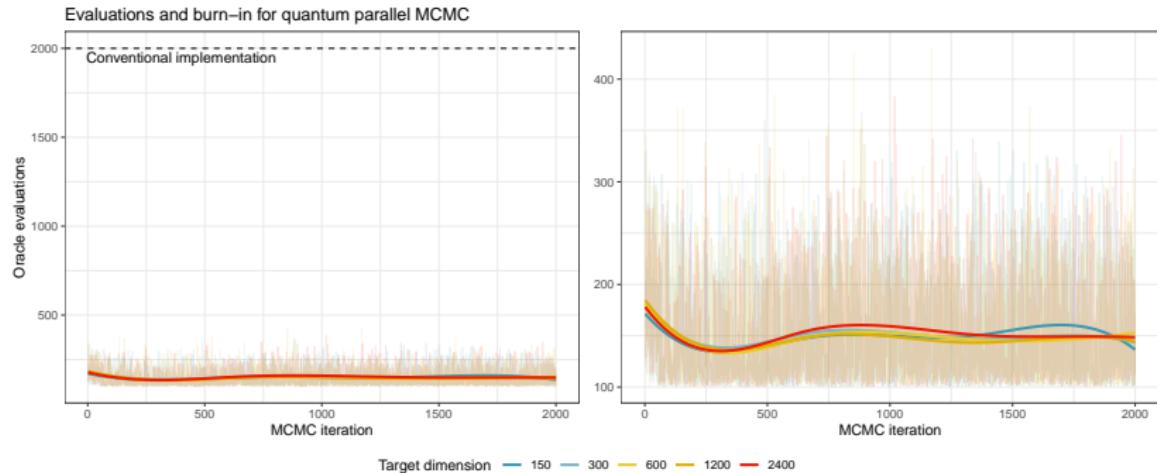
$\theta^{(s)} \leftarrow \theta_{\hat{p}}$;

end

return $\theta^{(1)}, \dots, \theta^{(S)}$.

Algorithm 6: Quantum parallel MCMC

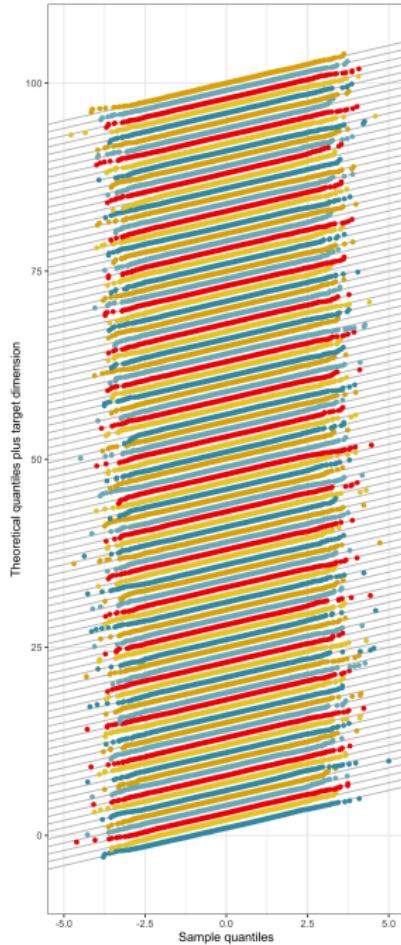
QPMCMC



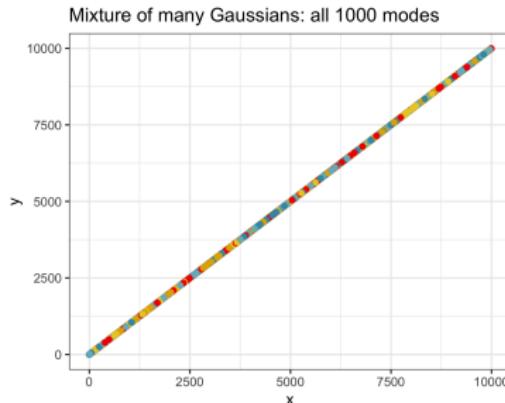
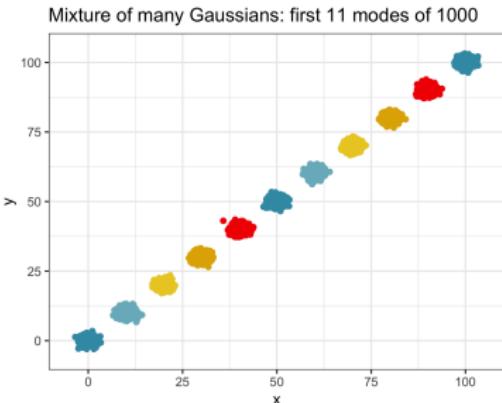
We use 7% of the conventional 2,000 target evaluations. Only 1 in 200 quantum minimizations fail.

QPMCMC

QQ plot for 100D Gaussian target



QPMCMC



Proposals	MCMC iterations	Target evaluations	Speedup	Efficiency gain
1,000	249,398 (200,998, 311,998)	24,988,963 (20,149,132, 31,265,011)	9.98 (9.98, 9.98)	1
5,000	14,398 (12,998, 16,998)	3,314,560 (2,989,418, 3,916,281)	21.72 (21.70, 21.74)	7.58 (6.25, 9.71)
10,000	5,998 (4,998, 6,998)	1,993,484 (1,662,592, 2,330,842)	30 (29.96, 30.26)	12.87 (8.64, 18.80)

Final thoughts

- ▶ Forget MCMC: just use quantum minimization and Gumbel-max to sample directly from the, e.g., Ising model.
- ▶ Don't forget MCMC: symmetric proposals can be made using discrete metrics for the, e.g., Ising model.
- ▶ Quantum minimization is “old” technology. Can we do better?

Thank you!