An Approximate Fisher Scoring Algorithm for Finite Mixtures of Multinomials

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(`www.umbc.edu/hpcf`)

Background

- Morel and Neerchal (1991, 1993, 1998, 2005) proposed a multinomial model for overdispersion — Random Clumped Multinomial — and studied estimation under it.
- They obtained a large cluster approximation to the Fisher Information Matrix (FIM), and used it to formulate an Approximate Fisher Scoring Algorithm (AFSA).
- Liu (2005, PhD Thesis) extended the idea to general mixtures of multinomials, and found some interesting connections between AFSA and Expectation Maximization (EM).
- This work extends Liu (2005), further investigating the quality of the FIM approximation and the connection between AFSA and EM.

Mixture of Multinomials Example

Example: Housing satisfaction survey

Non-metropolitan area				Metropolitan area			
Neighborhood	US	S	VS	Neighborhood	US	S	VS
1	3	2	0	19	0	4	1
2	3	2	0	20	0	5	1
3	0	5	0	21	0	3	2
$\vdots$				$\vdots$			
17	4	1	0	35	4	1	0
18	5	0	0				

With labels, a reasonable likelihood is product of two multinomials

$$L(\boldsymbol{\theta}) = \left[\prod_{i=1}^{18} f(\mathbf{x}_i \mid \mathbf{p}_1, m) \right] \left[\prod_{i=19}^{35} f(\mathbf{x}_i \mid \mathbf{p}_2, m) \right], \qquad m = 5.$$

J. R. Wilson, Chi-Square Tests for Overdispersion with Multiparameter Estimates. Journal of the Royal Statistical Society (Series C), 38(3):441–453, 1989.

Mixture of Multinomials Example

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$\vdots$				$\vdots$			
17	4	1	0	35	4	1	0
18	5	0	0				

Without labels, a reasonable likelihood is mixture of two multinomials

$$L(\boldsymbol{\theta}) = \prod_{i=1}^{35} \left\{ \pi f(\mathbf{x}_i \mid \mathbf{p}_1, m) + (1-\pi) f(\mathbf{x}_i \mid \mathbf{p}_2, m) \right\}, \qquad m = 5.$$

J. R. Wilson, Chi-Square Tests for Overdispersion with Multiparameter Estimates. Journal of the Royal Statistical Society (Series C), 38(3):441–453, 1989.

Mixture of Multinomials

- Suppose we have s multinomial populations

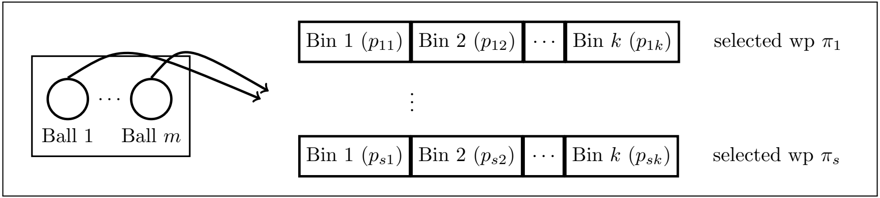
$$f(\mathbf{x} \mid \mathbf{p}_\ell, m) = \frac{m!}{x_1! \ldots x_k!} p_{\ell 1}^{x_1} \ldots p_{\ell k}^{x_k} \cdot I(\mathbf{x} \in \Omega), \qquad \ell = 1, \ldots, s$$

  which occur in the total population with probabilities $\pi_1, \ldots, \pi_s$.

- If we draw $\mathbf{T}$ from the mixed population,

$$\mathbf{T} \sim f(\mathbf{x} \mid \boldsymbol{\theta}) = \sum_{\ell=1}^{s} \pi_\ell f(\mathbf{x} \mid \mathbf{p}_\ell, m), \qquad \boldsymbol{\theta} = (\mathbf{p}_1, \ldots, \mathbf{p}_s, \boldsymbol{\pi})$$

  We'll write $\mathbf{T} \sim \text{MultMix}_k(\boldsymbol{\theta}, m)$.



Estimation Problem

- Suppose our sample is:
  $\mathbf{X}_i \overset{\text{ind}}{\sim} \text{MultMix}_k(\boldsymbol{\theta}, m_i), \qquad i = 1, \dots, n.$

  Then likelihood is

$$L(\boldsymbol{\theta}) = \prod_{i=1}^{n} \left\{ \sum_{\ell=1}^{s} \pi_\ell \left[\frac{m_i!}{x_{i1}! \dots x_{ik}!} p_{\ell 1}^{x_{i1}} \dots p_{\ell k}^{x_{ik}} \cdot I(\mathbf{x}_i \in \Omega) \right] \right\}.$$

- To find MLE $\hat{\boldsymbol{\theta}} = (\hat{\mathbf{p}}_1, \dots, \hat{\mathbf{p}}_s, \hat{\boldsymbol{\pi}})$, which maximizes the log-likelihood

- No nice closed form, so use iterative methods
  - Expectation Maximization (EM)
  - Newton-Raphson, **Fisher Scoring**, Quasi-Newton methods

$$\boldsymbol{\theta}^{(g+1)} = \boldsymbol{\theta}^{(g)} - \alpha \mathbf{H}^{-1} S(\boldsymbol{\theta}^{(g)}), \quad g = 1, 2, \dots$$

Fisher Scoring Algorithm

- The iterations become

$$\boldsymbol{\theta}^{(g+1)} = \boldsymbol{\theta}^{(g)} + \mathcal{I}^{-1}(\boldsymbol{\theta}^{(g)}) S(\boldsymbol{\theta}^{(g)}), \quad g = 1, 2, \ldots,$$

  but $\mathcal{I}(\boldsymbol{\theta})$ may not be easy to compute.

- Naive summation works when sample space Ω is small

$$\mathcal{I}(\boldsymbol{\theta}) := \sum_{\mathbf{x} \in \Omega} \left\{ -\frac{\partial^2}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}^T} \log f(\mathbf{x} \mid \boldsymbol{\theta}) \right\} f(\mathbf{x} \mid \boldsymbol{\theta}).$$

- Monte Carlo approximation

- For large m, an approximate FIM (shown for $\mathbf{X}_1 \sim \text{MultMix}_k(\boldsymbol{\theta}, m)$) is

$$\widetilde{\mathcal{I}}(\boldsymbol{\theta}) := \text{Blockdiag}\left(\pi_1 \mathbf{F}_1, \ldots, \pi_s \mathbf{F}_s, \mathbf{F}_\pi\right),$$
$$\mathbf{F}_\ell = m \left[\text{Diag}(p_{\ell 1}^{-1}, \ldots, p_{\ell, k-1}^{-1}) + p_{\ell k}^{-1} \mathbf{1}\mathbf{1}^T \right]$$
$$\mathbf{F}_\pi = \text{Diag}(\pi_\ell^{-1}, \ldots, \pi_{s-1}^{-1}) + \pi_s^{-1} \mathbf{1}\mathbf{1}^T$$

Approximate FIM Properties I

- **Result:** $\widetilde{\mathcal{I}}(\boldsymbol{\theta}) - \mathcal{I}(\boldsymbol{\theta}) \to \mathbf{0}$ as $m \to \infty$.

- **Result:** $\widetilde{\mathcal{I}}(\boldsymbol{\theta})$ is "complete data" FIM of $(\mathbf{X}, Z)$. Assume that

$$Z = \begin{cases} 1 & \text{wp } \pi_1 \\ & \vdots \\ s & \text{wp } \pi_s, \end{cases} \quad \text{and} \quad (\mathbf{X} \mid Z = \ell) \sim \text{Mult}_k(\mathbf{p}_\ell, m).$$

Then we have $\widetilde{\mathcal{I}}(\boldsymbol{\theta}) \equiv \mathrm{E}\left\{ -\frac{\partial^2}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}^T} \log f(\mathbf{x}, z \mid \boldsymbol{\theta}) \right\}$

- Note that EM is based on maximizing

$$Q(\boldsymbol{\theta}, \boldsymbol{\theta}') = \mathrm{E}_{\boldsymbol{\theta}'} \left[\log f(\mathbf{x}, z \mid \boldsymbol{\theta}) \mid \mathbf{x} \right].$$

Approximate FIM Properties II

- $\widetilde{\mathcal{I}}(\boldsymbol{\theta})$ is a block diagonal matrix of Multinomial FIMs.
  - Simple forms for inverse, trace, and determinant
- **Result:** $\widetilde{\mathcal{I}}(\boldsymbol{\theta}) - \mathcal{I}(\boldsymbol{\theta})$ is positive semi-definite.
  - Fairly general result for complete data FIMs, not just for multinomial mixtures
- **Result:** Standard errors derived from $\widetilde{\mathcal{I}}^{-1}(\boldsymbol{\theta})$ are systematically too optimistic (small)
- **Result:** The inverses also converge

$$\mathcal{I}^{-1}(\boldsymbol{\theta}) - \widetilde{\mathcal{I}}^{-1}(\boldsymbol{\theta}) \to \mathbf{0} \quad \text{as } m \to \infty.$$

Approximate FIM Properties III

Some recommendations for approximation FIM

- Large cluster size (m) needed for

$$\widetilde{\mathcal{I}}(\boldsymbol{\theta}) \approx \mathcal{I}(\boldsymbol{\theta}) \quad \text{and} \quad \widetilde{\mathcal{I}}^{-1}(\boldsymbol{\theta}) \approx \mathcal{I}^{-1}(\boldsymbol{\theta});$$

  inverses appear to converge faster.

- Approximate FIM and inverse are not recommended for general inference.

- But useful as a tool for estimation, as we will see.

Approximate Fisher Scoring Algorithm

- Using the approximate FIM in place of the true FIM gives AFSA

$$\boldsymbol{\theta}^{(g+1)} = \boldsymbol{\theta}^{(g)} + \widetilde{\mathcal{I}}^{-1}(\boldsymbol{\theta}^{(g)}) S(\boldsymbol{\theta}^{(g)}), \quad g = 1, 2, \ldots$$

until $\left| \log L(\boldsymbol{\theta}^{(g+1)}) - \log L(\boldsymbol{\theta}^{(g)}) \right| < \varepsilon$.

- **Result:** Under $\mathbf{X}_1, \ldots, \mathbf{X}_n \stackrel{\text{iid}}{\sim} \text{MultMix}_k(\boldsymbol{\theta}, m)$, EM and AFSA iterations are "equivalent", given the same starting place $\boldsymbol{\theta}^{(g)}$

$$\tilde{\pi}_\ell^{(g+1)} = \hat{\pi}_\ell^{(g+1)}, \quad \tilde{p}_{\ell j}^{(g+1)} = \left(\frac{\hat{\pi}_\ell^{(g+1)}}{\pi_\ell^{(g)}} \right) \hat{p}_{\ell j}^{(g+1)} + \left(1 - \frac{\hat{\pi}_\ell^{(g+1)}}{\pi_\ell^{(g)}} \right) p_{\ell j}^{(g)}.$$

- Don't get exact equality under the "independent but not iid" case.

Equivalence of AFSA and EM

- A more general connection is known between EM and iterations of the form

$$\boldsymbol{\theta}^{(g+1)} = \boldsymbol{\theta}^{(g)} + \mathcal{I}_c^{-1}(\boldsymbol{\theta}^{(g)})\, S(\boldsymbol{\theta}^{(g)}), \qquad g = 1, 2, \ldots.$$

- Titterington (1984) shows the two are approximately equivalent (under some regularity conditions)

- The equivalence is exact when the complete data likelihood is a regular exponential family

$$L(\boldsymbol{\mu}) = \exp\left\{ b(\mathbf{x}) + \boldsymbol{\eta}^T \mathbf{t} + a(\boldsymbol{\eta}) \right\},$$

$$\boldsymbol{\mu} = \mathrm{E}(\mathbf{t}(\mathbf{X})) : \text{the parameter of interest.}$$

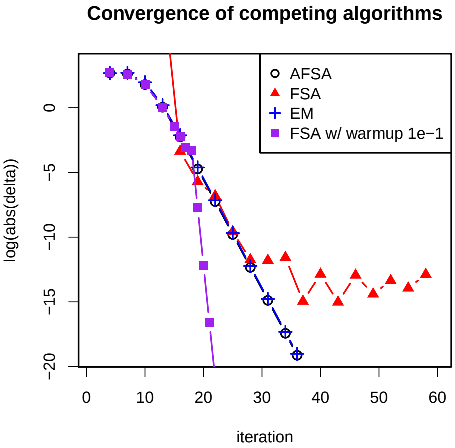
- For MultMix problem, equivalance is approximate, not exact.
  - AFSA's originally justified from $\widetilde{\mathcal{I}}(\boldsymbol{\theta})$ and Blischke (1964).
  - But this result justifies AFSA for finite mixtures other than multinomial.

Comparison between algorithms

Consider the mixture of two trinomials

$$\mathbf{X}_i \stackrel{\text{iid}}{\sim} \text{MultMix}_3(\boldsymbol{\theta}, m = 20), \qquad i = 1, \ldots, n = 500$$

$$\begin{pmatrix} \mathbf{p}_1^T \\ \mathbf{p}_2^T \end{pmatrix} = \begin{pmatrix} 1/3 & 1/3 & 1/3 \\ 0.1 & 0.3 & 0.6 \end{pmatrix}, \qquad \begin{pmatrix} \pi \\ 1-\pi \end{pmatrix} = \begin{pmatrix} 0.75 \\ 0.25 \end{pmatrix}.$$



method	tol	iter
AFSA	4.94×10^{-09}	36
EM	5.50×10^{-09}	36

Monte Carlo Comparison of EM and AFSA

Consider a scenario with varying cluster sizes

$$\mathbf{Y}_i \overset{\text{ind}}{\sim} \text{MultMix}_k(\boldsymbol{\theta}, m_i), \qquad i = 1, \ldots, n = 500, \qquad \boldsymbol{\pi} = (0.75, 0.25)$$
$$W_1, \ldots, W_n \overset{\text{iid}}{\sim} \text{Gamma}(\alpha, \beta), \qquad m_i = \lceil W_i \rceil.$$

Ran 1000 reps of nine scenarios and looked at the quantity

$$\overline{D} = \frac{1}{1000} \sum_{r=1}^{1000} D_r \quad \text{where} \quad D_r = \bigvee_{j=1}^{q} \left| \frac{\tilde{\theta}_j^{(r)} - \hat{\theta}_j^{(r)}}{\tilde{\theta}_j^{(r)}} \right|.$$

(kth probability not shown)		m_i equal	$\alpha = 100$	$\alpha = 25$
$\mathbf{p}_1$	$\mathbf{p}_2$	$m_i = 20$	$\text{Var}(m_i) \approx 4.083$	$\text{Var}(m_i) \approx 16.083$
0.1	0.5	2.178×10^{-6}	2.019×10^{-6}	2.080×10^{-6}
0.3	0.5	4.073×10^{-5}	3.501×10^{-5}	3.890×10^{-5}
0.35	0.5	8.683×10^{-4}	2.625×10^{-4}	2.738×10^{-4}
0.4	0.5	9.954×10^{-3}	6.206×10^{-2}	6.563×10^{-2}
$(0.1, 0.3)$	$(1/3, 1/3)$	1.342×10^{-3}	1.009×10^{-3}	1.878×10^{-3}
$(0.1, 0.5)$	$(1/3, 1/3)$	1.408×10^{-6}	1.338×10^{-6}	1.334×10^{-6}
$(0.3, 0.5)$	$(1/3, 1/3)$	3.884×10^{-6}	3.943×10^{-6}	3.885×10^{-6}
$(0.1, 0.1, 0.3)$	$(0.25, 0.25, 0.25)$	8.389×10^{-7}	8.251×10^{-7}	8.440×10^{-7}
$(0.1, 0.2, 0.3)$	$(0.25, 0.25, 0.25)$	1.523×10^{-6}	1.472×10^{-6}	1.408×10^{-6}

Conclusions

AFSA obtained as a Newton-type algorithm using approximate FIM.

- Nearly equivalent to EM iterations — similar solutions are obtained at similar rates of convergence.
- (EM advantange) M-step can be formulated so it won't wander outside parameter space.
- (AFSA advantange) May be easier to formulate when missing data structure is complicated.
  E.g. Random-Clumped Multinomial (Morel & Neerchal 1993).

Result of Titterington (1984) suggests AFSA approach is reasonable for finite mixtures in general.

- What about the FIM approximation?

Both EM and AFSA suffer from a slow convergence rate.

- Hybrid is recommended for fast convergence and robustness.
- ... if true FIM is feasible to compute.

Conclusions

For details and references, see our technical report:

`www.umbc.edu/hpcf` $\Rightarrow$ Research $\Rightarrow$ Publications

A. M. Raim, M. Liu, N. K. Neerchal & J. G. Morel. An Approximate Fisher Scoring Algorithm for Finite Mixtures of Multinomials. Technical Report HPCF–2012–14, UMBC High Performance Computing Facility, University of Maryland, Baltimore County, 2012.

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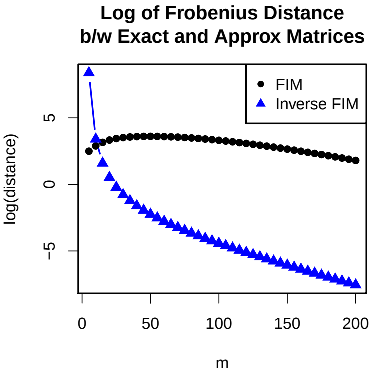
How good is the FIM approximation?

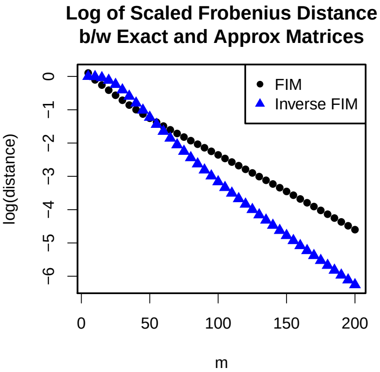
Consider a mixture $\text{MultMix}_2(\boldsymbol{\theta}, m)$ of three binomials, with parameters

$$\begin{pmatrix} p_1 & p_2 & p_3 \end{pmatrix} = \begin{pmatrix} 1/7 & 1/3 & 2/3 \end{pmatrix}, \qquad \boldsymbol{\pi} = \begin{pmatrix} 1/6 & 2/6 & 3/6 \end{pmatrix},$$

and two matrix distances

$$d(\mathbf{A}, \mathbf{B}) = \|\mathbf{A} - \mathbf{B}\|_F \qquad\qquad d(\mathbf{A}, \mathbf{B}) = \frac{\|\mathbf{A} - \mathbf{B}\|_F}{\|\mathbf{B}\|_F}$$





Large m is needed for a good approximation. Inverses are converging faster.