

# Bivariate conditional spatial models: Case study in Section 5

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## Setting up

As part of this vignette we will be needing part of the **INLA** package and installation instructions for this can be found on the [R-INLA homepage](#). Once this is installed we can load the required packages. As with the other vignette (simulation example in Section 3.2), we will also need **dplyr**, **tidyr**, and **Matrix** for core operations and **ggplot2**, **gridExtra**, **grid**, **extrafont** for plotting purposes. In addition, in this vignette we will need **maptools**, **mapproj**, and **RandomFields** that provides the data used in this problem. and the **verification** package that contains a handy routing for computing CRPSs. For parallel operations we will be requiring **foreach** and **doParallel** and possibly **doMPI** if an MPI backend is available.

```
library(INLA)
library(dplyr)
library(tidyr)
library(Matrix)
library(ggplot2)
library(gridExtra)
library(grid)
library(extrafont)
library(maptools)
library(mapproj)
library(RandomFields)
library(verification)
library(foreach)
library(doParallel)
```

Finally, we will also need the package **bicon** to facilitate some of the operations used here.

```
library(bicon)
```

```
## Warning: replacing previous import by 'scales::alpha' when loading 'bicon'
```

As detailed in Cressie & Zammit-Mangion (2015), we consider four models that vary only through the interaction function  $b_o(h)$ . Each model is given by

$$\begin{array}{ll} \text{Model 1 (independent Matérns):} & b_o(h) \equiv 0, \\ \text{Model 2 (pointwise dependence):} & b_o(h) \equiv A\delta(h), \\ \text{Model 3 (diffused dependence):} & \text{Model 4 with } \Delta = 0 \\ \text{Model 4 (asymmetric dependence):} & b_o(h) \equiv \begin{cases} A\{1 - (\|h - \Delta\|/r)^2\}^2, & \|h - \Delta\| \leq r \\ 0, & \text{otherwise,} \end{cases} \end{array}$$

where  $\Delta = (\Delta_1, \Delta_2)^T$  is a shift parameter vector that captures asymmetry,  $r$  is the aperture parameter, and  $A$  is the amplitude. In Models 3 and 4,  $b_o(h)$  is a shifted bisquare function in  $\mathbb{R}^2$ . The covariance functions  $C_{11}(\cdot)$  and  $C_{2|1}(\cdot)$  are Matérn covariance functions. For each model we also consider a *reversed* dependence, where we switch  $Y_2$  and  $Y_1$ . This gives us a total of eight models to fit and compare.

The first thing we do is set some options. `is` to indicate whether we want to run the full program. We use a flag `L00_analysis` to indicate whether we want to run the leave-one-out cross-validation study (without re-fitting) while the flag `LK_analysis` indicates whether we want to carry out a standard likelihood fit of all 8 models using the entire dataset (requires about 30 minutes of computing time). In this vignette we set both flags to zero; the results are instead loaded from cache.

```
### Model choice
model_names <- c("independent","pointwise","moving_average_delta0","moving_average")
img_path <- "../paper/art"                ## Where to save the figures
show_figs <- 1                            ## Show the figures in document
print_figs <- 0                           ## Print figures to file (leave =0)
LK_analysis <- 0                          ## Carry out likelihood analysis
L00_analysis <- 0                          ## Carry out L00 analysis
Shifted_Pars_estimation <- 0               ## Fit shifted parimonious Matern
RF_estimation <- 0                        ## Carry out L00 with RFields
useMPI <- 0                               ## MPI backend available?
```

## The data

The data were made available through the package `RandomFields` and was studied at length in the paper of T. Gneiting, Kleiber, & Schlather (2010). We load the data and pre-process it as follows:

```
data(weather,package = "RandomFields")
weather <- weather %>% data.frame()
print(head(weather))
```

```
##   pressure temperature      lon      lat
## 1 200.4844  0.60537720 -131.000  46.000
## 2 384.8516 -0.02233887 -124.400  41.900
## 3 156.8984 -0.26644897 -124.500  46.100
## 4 248.4297 -1.30670166 -124.700  47.300
## 5 253.2266  0.14398193 -124.500  44.600
## 6 159.2031 -0.27954102 -124.985  49.907
```

The `weather` table contains four fields, with latitude, longitude, pressure forecasting errors and temperature forecasting errors for December 13, 2003 at 4 p.m. in the North American Pacific Northwest. Since pressure and temperature have different units, we find a scaling factor by taking the ratio of the variances of the two variates, and computing its square root:

```
P_scale <- (var(weather$pressure)/var(weather$temperature)) %>%
  sqrt() %>%
  as.numeric()
```

From this data frame we extract  $Z_1$  and  $Z_2$  and concatenate them into one long vector  $Z$  through a function `form_Z`. The vectors  $Z_1$  and  $Z_2$  are inverted if the model being analysed is greater than 4 (reversed model). We also define `m1` as the number of observations of  $Y_1$ , `m2` as the number of observations of  $Y_2$  and `m` as the total number of observations.

```
form_Z <- function(model_num,scale=T) {
  Z1 <- matrix(weather$temperature)
  Z2 <- matrix(weather$pressure)
  if(scale) Z2 <- Z2 / P_scale # Change pressure to have similar scale
```

```

if(model_num > 4) {
  temp <- Z1
  Z1 <- Z2
  Z2 <- temp
}
Z <- rbind(Z1,Z2)
}

m1 <- m2 <- nrow(weather)
m <- m1 + m2
I_m1 <- Diagonal(m1) # identity matrix of size m1 x m1

```

## Process discretisation

We approximate the processes as a sum of elemental basis functions (tent functions) constructed on a triangulation. The triangulation is formed using the mesher in the INLA package, while we provide a tailored function `initFEbasis` which takes information from the INLA mesher and casts it into a `Mesh` object. We provide several methods associated with the `Mesh` class which will be useful for plotting later on. Importantly, the `Mesh` object also contains information on the areas of the elements in the Voronoi tessellation, which will be used to approximate the integrations.

```

### Construct mesh
###-----
mesh <- inla.mesh.2d(loc= weather[c("lon","lat")],
                    cutoff=0,
                    max.edge=0.75,offset = 4) ## Fine mesh
mesh_locs <- mesh$loc[,1:2]

## Compute distances as in Gneiting (2010)
D <- as.matrix(RFearth2dist(as.matrix(mesh_locs)))
Dvec <- as.double(c(D))
Dobs <- as.matrix(RFearth2dist(as.matrix(weather[c("lon","lat")]))))
Dobsvec <- c(Dobs)

Mesh <- initFEbasis(p = mesh_locs,
                  t = mesh$graph$tv,
                  K = mesh$graph$vv)

```

```

##
##      PLEASE NOTE: The components "delsgs" and "summary" of the
## object returned by deldir() are now DATA FRAMES rather than
## matrices (as they were prior to release 0.0-18).
## See help("deldir").
##
##      PLEASE NOTE: The process that deldir() uses for determining
## duplicated points has changed from that used in version
## 0.0-9 of this package (and previously). See help("deldir").

```

We next establish the dimension of our grids. Since we will be evaluating  $Y_1$  and  $Y_2$  on the same grid,  $n1 = n2$ .

```
### Process models
###-----
n1 <- nrow(mesh_locs)
n2 <- nrow(mesh_locs)
n <- n1 + n2
```

As in the first vignette (simulation example in Section 3.2), we will approximate the integration using the rectangular rule. When using finite elements, this reduces to using the area of the Voronoi tessellation as the weight for the function values.

We first compute the vector of displacements  $h$  which will be of length  $(n2 \times n1)$  and with each element associate an integration weight equal to the area of the Voronoi tessellation associated with the element:

```
### Mesh integration points
###-----
h <- matrix(0,n1*n2,2)
areas <- rep(0,n1*n2)
for(i in 1:n2) {
  h[((i-1)*n1+1):(i*n1),] <- t(t(mesh_locs) - mesh_locs[i,])
  areas[((i-1)*n1+1):(i*n1)] <- Mesh["area_tess"]
}
h1_double <- as.double(h[,1])
h2_double <- as.double(h[,2])
```

The displacements  $(h1,h2)$  and the areas `areas` can then be used to construct the matrix  $B$  using the function `bisquare_B`.

## Organising the observations

In order to map the process to the observations we construct an incidence matrix, which contains a 1 wherever the observation coincides with a vertex on the triangulation and a 0 otherwise. The dimension of this incidence matrix is  $(m1 + m2) \times (n1 + n2)$ , where  $m1, m2$ , are the number of observations in  $Z_1, Z_2$ , respectively. Since in this problem we have collocated observations, we find the incidence matrix for one of the observations,  $Z_1$ , and then form the whole incidence matrix by simply carrying out `bdiag` (block diagonal) of the first matrix with itself. We find the points with which the observation locations coincide by using the function `left_join`, which returns an NA if no observation coincides with the vertex.

```
mesh_locs <- data.frame(lon=mesh_locs[,1],lat=mesh_locs[,2])      ## mesh locations
idx <- which(!(is.na(left_join(mesh_locs,weather)$temperature))) ## index of coincidence
C1 <- sparseMatrix(i=1:m1,j=idx,x=1,dims=c(m1,n1))               ## incidence matrix of Z1
C <- bdiag(C1,C1)         ## incidence matrix
```

## Maximum likelihood estimation

Since the optimisation algorithm requires a parameter vector of the same length (irrespective of the model number) we first define a function `append_theta` that takes the parameter vector associated with the model in question and appends it so it is of the required size (in this case of length 11)

```
append_theta <- function(theta,model_num) {
  if(model_num %in% c(1,5)) {
    theta <- c(theta,rep(0,4))
  }
```

```

    theta[10] <- 0.001
  } else if(model_num %in% c(2,6)) {
    theta <- c(theta,rep(0,3))
    theta[10] <- 0.001
  } else if(model_num %in% c(3,7)) {
    theta <- c(theta,rep(0,2))
  }
  theta
}

```

Next, we require a function that, given the parameter vector `theta` and the model number `model_num`, returns the required matrices and vectors used in fitting. These are the process covariance matrix

$$\text{cov}((Y_1^T, Y_2^T)^T) = \begin{bmatrix} \Sigma_{11} & \Sigma_{11}B^T \\ B\Sigma_{11} & \Sigma_{2|1} + B\Sigma_{11}B^T \end{bmatrix}, \quad (1)$$

the observation error covariance matrix and the vector  $Z$  that, recall, is equal to  $[Z_1^T, Z_2^T]^T$  or  $[Z_2^T, Z_1^T]^T$ , depending on the model number. If `whole_mesh` is `TRUE`, then the process covariance matrix is evaluated over the entire mesh (used for co-kriging at unobserved locations).

```

construct_mats <- function(theta,model_num,whole_mesh=F) {

  nu1 <- theta[7]
  nu2 <- theta[8]

  B <- theta[9]*Diagonal(n1) # Automatically zero if Model 1

  if(model_num %in% c(3,4,7,8)) {
    B <- theta[9]*bisquare_B(h1_double,h2_double,
                           delta=theta[11:12], # Automatically zero for Model with no shift
                           r=theta[10],
                           n1 = n1,
                           n2 = n2,
                           areas = areas)
  }
  C1B <- C1 %*% B

  ## Form matrices (scaled pressure)
  S11 <- makeS(r = Dobsvec,var = theta[3],
              kappa = theta[5],nu = nu1)
  S2_1 <- makeS(r = Dobsvec,var = theta[4],
               kappa = theta[6],nu = nu2)
  if(model_num %in% c(3,4,7,8) | whole_mesh==TRUE) {
    S11_big <- makeS(r = Dvec,var = theta[3],
                    kappa = theta[5],nu = nu1)
    S21 <- C1B %*% (S11_big %*% t(C1))
    S12 <- t(S21)
    S22 <- S2_1 + forceSymmetric(C1B %*% forceSymmetric(S11_big) %*% t(C1B))
  } else {
    S21 <- S12 <- theta[9]*S11
    S22 <- S2_1 + theta[9]^2 * S11
  }
}

```

```

if(whole_mesh) {
  S11 <- S11_big
  S2_1 <- makeS(r = Dvec,var = theta[4],
               kappa = theta[6],nu =nu2)
  S21 <- B %%% S11_big
  S12 <- t(S21)
  S22 <- S2_1 + Matrix::crossprod(chol(S11_big) %%% t(B))
}

## Form matrices (Unscaled pressure)
S11_true <- ifelse(model_num > 4,P_scale^2,1) * S11
S12_true <- P_scale * S12
S21_true <- P_scale * S21
S22_true <- ifelse(model_num < 5, P_scale^2,1) *S22
SY_true <- rBind(cBind(S11_true,S12_true),
                cBind(S21_true,S22_true)) %>% as("dgeMatrix")
So_true <- bdiag(ifelse(model_num < 5,1,P_scale^2) * theta[1]*I_m1,
                ifelse(model_num < 5,P_scale^2,1) * theta[2]*I_m1)
if(whole_mesh) So_true <- t(C) %%% So_true %%% C

list(SY = SY_true, So = So_true, Z = form_Z(model_num,scale=F))
}

```

Now we're in place to define the log-likelihood function. This is the usual Gaussian log-likelihood function. In the function we allow the dropping of certain observations for cross-validation purposes. The indices of the observations we wish to drop are stored in the parameter `i`.

```

loglik_Model <- function(theta,model_num,i=NULL) {
  # theta1: sigma2e1
  # theta2: sigma2e2
  # theta3: sigma211
  # theta4: sigma22_1
  # theta5: kappa11
  # theta6: kappa2_1
  # theta7: nu11
  # theta8: nu2_1
  # theta9: A
  # theta10: r
  # theta11: d1
  # theta12: d2

  theta <- append_theta(theta,model_num)

  if(theta[1] <= 0 | theta[2] <= 0 | theta[3] <= 0 |
     theta[4] <= 0 | theta[5] <= 0.001 | theta[6] <= 0.001 |
     theta[7] <= 0.05 | theta[8] <= 0.05 | theta[10] < 0.0005) {
    return(Inf)
  } else {

    ## Construct matrices
    X <- construct_mats(theta,model_num)

    ## Drop observations if required for CV

```

```

    if(is.null(i)) {
      SY <- X$SY
      So <- X$So
      Z <- X$Z
    } else {
      SY <- X$SY[-i,-i]
      So <- X$So[-i,-i]
      Z <- X$Z[-i,,drop=F]
    }

    ## Evaluate log-likelihood function
    cholYo <- chol(SY + So)
    loglik <-
      -(-0.5 * logdet(cholYo) -
        0.5 * t(Z) %*% chol2inv(cholYo) %*% Z -
        0.5 * nrow(Z)*log(2*pi)) %>% as.numeric()

    return(loglik)
  }
}

```

For optimising we will use the R function `optim` (BFGS). We allow for 3000 maximum iterations and set `trace=6` for detailed output. We choose to not compute the Hessian since this is not required in our analysis. Recall that the parameter `i` here contains the indices of the observations we do not wish to include in the fit. If `i = NULL` then all observations are included. The function `optim_loglik` is called for each model in the program later on.

```

optim_loglik <- function(par,model_num, i = NULL){
  optim(par=par,
    fn = loglik_Model,
    model_num=model_num,
    i=i,
    hessian=FALSE,
    control=list(trace=6,
      pgtol=0,
      parscale=rep(0.1,length(par)),
      maxit=3000))
}

```

The last function we need to define is one that fits all the models, possibly with a set of observations in `i` removed. Note that in the function we first fit Model 1 using realistic starting values and store the results in `fit.Model1`, and then fit the reversed version (with pressure as  $Y_1$ ) and store that in `fit.Model1_rev`. Model 2 is then fit using the results of Model 1 as starting values. Model 3 uses the maximum likelihood estimates of Model 2 and so on. The reversed version of Model 2 uses the results of the reversed version of Model 1 as starting parameters and so on.

```

fit_all_models <- function(i) {
  fit.Model1 <- optim_loglik(par=c(0.01,1,5,15,0.01,0.01,0.6,1.5), model_num=1, i = i)
  fit.Model1_rev <- optim_loglik(par=c(1,0.01,15,5,0.01,0.01,1.5,0.6), model_num=5, i = i)
  fit.Model2 <- optim_loglik(par=c(fit.Model1$par,-0.2), model_num=2, i = i)
  fit.Model2_rev <- optim_loglik(par=c(fit.Model1_rev$par,-0.2),model_num=6, i = i)
  fit.Model3 <- optim_loglik(par=c(fit.Model2$par,0.1), model_num=3, i = i)
}

```

```

fit.Model3_rev <- optim_loglik(par=c(fit.Model2_rev$par,0.1), model_num=7, i = i)
fit.Model4      <- optim_loglik(par=c(fit.Model3$par,0,0),      model_num=4, i = i)
fit.Model4_rev  <- optim_loglik(par=c(fit.Model3_rev$par,0,0), model_num=8, i = i)

list(Model1 = fit.Model1,
      Model2 = fit.Model2,
      Model3 = fit.Model3,
      Model4 = fit.Model4,
      Model5 = fit.Model1_rev,
      Model6 = fit.Model2_rev,
      Model7 = fit.Model3_rev,
      Model8 = fit.Model4_rev)
}

```

With all functions in place we now call `fit_all_data <- fit_all_models(NULL)`. All this does is fit all the models using all the observations (since `i = NULL`). If `LK_analysis = 1` then this is done from scratch (takes about 30 minutes), otherwise the data is loaded from cache.

```

## First we carry out the analysis with all data in
if(LK_analysis) {
  fit_all_data <- fit_all_models(NULL)
  save(fit_all_data, file=paste0("../inst/extdata/temp_pressure/LK_fits.rda"))
} else {
  #load("../inst/extdata/temp_pressure/LK_fits.rda")
  load(system.file("extdata/temp_pressure", "LK_fits.rda", package = "bicon"))
}

```

The log-likelihoods and AICs given by our fit are given in the table below. Note that Model 5 is Model 1 reversed (i.e., with pressure as  $Y_1$ ), Model 6 is Model 2 reversed and so on.

```
print("Log-likelihood for all models trained with complete dataset")
```

```
## [1] "Log-likelihood for all models trained with complete dataset"
```

```
sapply(fit_all_data,function(x) x$value) ## Negative LL
```

```
##      Model1      Model2      Model3      Model4      Model5      Model6      Model7      Model8
## 1276.770 1269.922 1264.901 1258.212 1276.770 1266.826 1268.983 1268.486
```

```
sapply(fit_all_data,function(x) x$value)*2 + 2*c(8,9,10,12) ## AIC
```

```
##      Model1      Model2      Model3      Model4      Model5      Model6      Model7      Model8
## 2569.541 2557.844 2549.803 2540.425 2569.541 2551.651 2557.967 2560.972
```

The parameters are listed below (the output is in LaTeX for direct use in paper). Note that since `P_scale` was used to put pressure on the same scale as temperature, we scale the fitted marginal standard deviation of the pressure fields so that they are on the original scale.

```
print("Estimated parameters for all models")
```

```
## [1] "Estimated parameters for all models"
```



```

par_est <- plyr::rbind.fill(sapply(fit_all_data,function(x) data.frame(t(x$par))))
par_est1 <- par_est[1:4,]
par_est1[,c(1,3)] <- sqrt(par_est1[,c(1,3)])
par_est1[,c(2,4)] <- sqrt(par_est1[,c(2,4)]) * P_scale
par_est1[,9] <- par_est1[,9] * P_scale

colnames(par_est1) <- c("\sigma_1$", "\sigma_2$", "\sigma_{11}$",
                        "\sigma_{2|1}$", "\kappa_{11}$", "\kappa_{2|1}$",
                        "\nu_{11}$", "\nu_{2|1}$",
                        "$A$", "$r$", "\Delta_1$", "\Delta_2$")
rownames(par_est1) <- c("Model 1", "Model 2", "Model 3", "Model 4")
print(xtable::xtable(par_est1,digits=c(rep(2,5),3,3,rep(2,6))),
      sanitize.text.function=function(x){x},
      hline.after=NULL)

```

```

## % latex table generated in R 3.2.0 by xtable 1.7-4 package
## % Wed Jun 22 15:37:28 2016
## \begin{table}[ht]
## \centering
## \begin{tabular}{rrrrrrrrrrrrrrrr}
##   & $\sigma_1$ & $\sigma_2$ & $\sigma_{11}$ & $\sigma_{2|1}$ & $\kappa_{11}$ & $\kappa_{2|1}$ & $\nu_{11}$ & $\nu_{2|1}$ & $A$ & $r$ & $\Delta_1$ & $\Delta_2$ \\
## Model 1 & 0.00 & 68.47 & 2.60 & 275.34 & 0.011 & 0.010 & 0.60 & 1.56 & & & & \\
## Model 2 & 0.00 & 67.78 & 2.60 & 242.04 & 0.011 & 0.011 & 0.60 & 1.58 & -14.30 & & & \\
## Model 3 & 0.00 & 70.16 & 2.68 & 243.77 & 0.011 & 0.010 & 0.61 & 1.84 & -40.83 & 1.46 & & \\
## Model 4 & 0.01 & 69.79 & 3.02 & 199.86 & 0.007 & 0.004 & 0.56 & 1.24 & -65.58 & 1.18 & 0.76 & -1.42 \\
## \end{tabular}
## \end{table}

```

## Prediction

We predict the temperature and pressure fields at the unobserved locations using cokriging. Since we assume zero mean, this is simple cokriging; the predictive mean and variance can thus be obtained by simple conditioning with a joint multivariate Gaussian distribution. If  $i = \text{NULL}$  then the data is used to predict at all (observed and unobserved) locations. Otherwise prediction is only carried out at the locations in  $i$  with the observations in  $i$  removed. Note that when  $i$  is specified it is assumed that only the covariance matrices associated with the observation locations are supplied. This enables us to use the same function for cross-validation (see below).

```

cokrige <- function(X,i=NULL) {
  SS <- X$SY + X$So

  if(is.null(i)) {
    Z <- X$Z
    Q <- chol2inv(chol(C%*% SS %*% t(C))) %>% as("dgeMatrix")
    mu_pred <- SS%*% t(C) %*% Q %*% Z %>% as.numeric()
    var_pred <- diag(SS - SS %*% t(C) %*% Q %*% C %*% SS) %>% as.numeric()
    data.frame(mu_pred = mu_pred,
               var_pred = var_pred)
  } else {
    cholSS <- chol(SS[-i,-i]) ## this was SS[-i,-i]
    SSinv <- chol2inv(cholSS) %>% as("dgeMatrix")
  }
}

```

```

mu_pred <- SS[i,-i] %*% SSinv %*% X$Z[-i,,drop=FALSE] %>% as.numeric()
var_pred <- diag(SS[i,i] - SS[i,-i] %*% SSinv %*% SS[-i,i]) %>% as.numeric()
data.frame(mu_pred = mu_pred,
           var_pred = var_pred,
           Z = X$Z[i,],
           i=i)
}
}

```

Below we predict at all the mesh locations using Model 1 and Model 4. First we construct the required matrices and store them in X1 and X4. Then we carry out cokriging and add the mean predictions to the mesh.

```

X1 <- construct_mats(theta = append_theta(fit_all_data[[1]]$par,model_num = 1),
                    model_num = 1,whole_mesh = T)

```

```

## Note: method with signature 'matrix#Matrix' chosen for function '+',
## target signature 'matrix#ddiMatrix'.
## "ANY#ddiMatrix" would also be valid

```

```

X4 <- construct_mats(theta = append_theta(fit_all_data[[4]]$par,model_num = 4),
                    model_num = 4,whole_mesh = T)
ALL1 <- cokrige(X=X1,i=NULL)
ALL4 <- cokrige(X=X4,i=NULL)

Mesh["y1_Model1"] <- ALL1$mu_pred[(1:n1)]
Mesh["y1_Model4"] <- ALL4$mu_pred[(1:n1)]
Mesh["y2_Model1"] <- ALL1$mu_pred[-(1:n1)]
Mesh["y2_Model4"] <- ALL4$mu_pred[-(1:n1)]

```

## LOO Analysis

Unlike Tilmann Gneiting, Raftery, Westveld III, & Goldman (2005), here we carry out leave-one-out cross validation (LOOCV) without re-fitting the model each time. If we have an MPI cluster available we carry out the LOOCV over MPI, otherwise we parallelise using the machine's multiple cores. The two loops below iterate over the observations and models.

```

## Now we do a LOO analysis
if(LOO_analysis) {

  if(useMPI) {
    library(doMPI)
    cl <- startMPIcluster(count=80)
    registerDoMPI(cl)
  } else {
    library(doParallel)
    cl <- makePSOCKcluster(4,outfile="cores_output.txt")
    registerDoParallel(cl)
  }
}

```

```

## Loop over each observation location
pred <- foreach(i = 1:m1,.combine = "rbind",
               .packages = c("Matrix","bicon","dplyr","foreach")) %dopar% {

  fit.Model <- fit_all_data

  ## Loop over each model (not parallelised)
  pred <- foreach(j = seq_along(fit.Model),.combine = "rbind") %do% {

    ## Construct matrices
    X <- construct_mats(theta = append_theta(fit.Model[[j]]$par,model_num = j),
                      model_num = j)

    ## Cokrige, leaving out the ith observation (for both temperature and pressure)
    cbind(cokrige(X=X,i=c(i,i+m1)),
          model_num = j)
  }
  pred
}

if(useMPI) {
  closeCluster(cl)
} else {
  stopCluster(cl)
}

### Should only be run with path set as vignette source directory
save(pred, file=paste0("../inst/extdata/temp_pressure/all_predictions.rda"))
}

```

If we decided not to run the LOOCV, we load the results from cache.

```

if(!LOO_analysis) {
  load(system.file("extdata/temp_pressure/all_predictions.rda", package = "bicon"))
  #load("../inst/extdata/temp_pressure/all_predictions.rda")
}

```

## Displaying and plotting results

The following code is only documented in-line since it just involves data manipulation for obtaining the results shown in the paper. For verification we find the mean absolute error (MAE), the root mean-squared prediction error (RMSPE) and the continuous probability rank score (CRPS) as described by Tilman Gneiting et al. (2005).

```

### Analyse results

## put data set into long format
weather_long <- mutate(weather,loc_num = 1:nrow(weather)) %>%
  mutate(sum_D = apply(Dobs,1,function(x) sum(sort(x)[1:2]))) %>%
  gather(process,z,temperature,pressure,convert = TRUE)

## Utility wrapper around the crps function

```

```

crps_wrapper <- function(Z,mu,sd) {
  crps(Z,cbind(mu,sd))$crps
}

sanitise_results <- function(pred) {
  ## put data set into long format
  pred2 <- mutate(pred,
    process = ifelse((model_num < 5 & i <= m1) |      # Take our LOOCV results
                     (model_num >=5 & i > m1) ,      # Assign process name to row
                     "temperature",
                     "pressure"),
    loc_num = ((i-1) %% m1)+1) %>%                  # Assign loc ID
  left_join(weather_long) %>%                        # Join up with data
  dplyr::select(-i,-z)                             # Remove unwanted columns

  results <- pred2 %>%                               # Take the predictions
  group_by(process,model_num) %>%                   # Group by process and model
  summarise(MAE = mean(abs(mu_pred - Z)),           # Find MAE
    MAE_se = sd(mu_pred - Z)/sqrt((m1-1)),         # Find MAE standard error
    Bias = mean(mu_pred - Z),                      # Find mean bias
    Bias_se = sd(mu_pred - Z)/sqrt((m1-1)),        # Find bias standard error
    Bias_norm = mean((mu_pred - Z)/sqrt(var_pred)), # Find mean normalised bias
    RMSPE = sqrt(mean((mu_pred - Z)^2)),           # Find RMSPE
    CRPS = mean(crps_wrapper(Z,mu_pred,sqrt(var_pred))), # Find CRPS
    CRPS_se = sd(crps_wrapper(Z,mu_pred,sqrt(var_pred)))/sqrt(m1-1)) # Find CRPS se

  results
}
results <- sanitise_results(pred)
print(results)

```

```

## Source: local data frame [16 x 10]
## Groups: process [?]
##
##   process model_num    MAE    MAE_se    Bias    Bias_se
##   (chr)      (int)    (dbl)    (dbl)    (dbl)    (dbl)
## 1 pressure      1 69.557498 9.8831976 -8.723436484 9.8831976
## 2 pressure      2 70.190275 9.9657301 -9.135887569 9.9657301
## 3 pressure      3 70.317490 9.8578626 -8.036973575 9.8578626
## 4 pressure      4 66.068990 9.2041738 -4.212511072 9.2041738
## 5 pressure      5 69.557498 9.8831976 -8.723436484 9.8831976
## 6 pressure      6 67.020147 9.8313416 -4.821590904 9.8313416
## 7 pressure      7 66.808978 9.7410402 -5.364313591 9.7410402
## 8 pressure      8 66.577363 9.6881394 -5.095728792 9.6881394
## 9 temperature  1  1.143941 0.1305158  0.040281435 0.1305158
##10 temperature  2  1.143934 0.1305225  0.040190622 0.1305225
##11 temperature  3  1.095204 0.1228428  0.037770689 0.1228428
##12 temperature  4  1.080266 0.1176788  0.017517495 0.1176788
##13 temperature  5  1.143941 0.1305158  0.040281435 0.1305158
##14 temperature  6  1.118945 0.1272823 -0.003432577 0.1272823
##15 temperature  7  1.104627 0.1255866 -0.024560013 0.1255866
##16 temperature  8  1.101850 0.1255400 -0.023512694 0.1255400
## Variables not shown: Bias_norm (dbl), RMSPE (dbl), CRPS (dbl), CRPS_se

```

```
## (dbl)
```

```
## Deprecated:
pred2 <- mutate(pred,
  process = ifelse((model_num < 5 & i <= m1) |
    (model_num >= 5 & i > m1),
    "temperature",
    "pressure"),
  loc_num = ((i-1) %% m1)+1) %>%
left_join(weather_long) %>%
dplyr::select(-i,-z)
results2 <- pred2 %>%
  mutate(MAE = abs(mu_pred - Z),
    Bias = mu_pred - Z,
    SR = (mu_pred - Z)/sqrt(var_pred),
    CRPS = crps_wrapper(Z,mu_pred,sqrt(var_pred)))
```

*# Take our LOOCV results*  
*# Assign process name to row*  
*# Assign loc ID*  
*# Join up with data*  
*# Take original (ungrouped) predictions*  
*# Find MAE of each obs*  
*# Find bias of each obs*  
*# Find normalised residual*  
*# Find CRPS*

Next, we compare our results to the *shifted* parsimonious Mat{e}rn model, obtained by applying the method of Li & Zhang (2011) to the standard parsimonious model. First we define a function that constructs the matrices based on the usual parameters:

```
sh_pars_mats <- function(theta) {
  ## Now create the shifted locations for the cross-covariances
  new_locs <- weather[,3:4] + matrix(theta[1:2],ncol=2,nrow=nrow(weather),byrow=TRUE)

  X <- rbind(new_locs,weather[,3:4])
  tot_D <- as.matrix(RFearth2dist(as.matrix(X)))
  D12 <- tot_D[-(1:m1),1:m1]
  D21 <- t(D12)

  # Dobs <- as.matrix(RFearth2dist(as.matrix(weather[,3:4])))
  # Dobsvec <- c(Dobs)

  sigma2_21 <- sqrt(theta[9] * theta[10])*theta[6]
  SY11 <- makeS(r=Dobsvec,var= theta[9],kappa = theta[3],nu=theta[4])
  SY22 <- makeS(r=Dobsvec,var= theta[10],kappa = theta[3],nu=theta[5])
  SY12 <- makeS(r=c(D21),var=sigma2_21, kappa = theta[3], nu = (theta[4] + theta[5])/2)
  SY21 <- t(SY12)
  So11 <- theta[7]^2 * diag(m1)
  So22 <- theta[8]^2 * diag(m1)
  SY <- cbind(rbind(SY11,SY12),rbind(SY21,SY22))
  So <- bdiag(So11,So22)
  list(SY = SY, So = So,Z = form_Z(1L,scale=F))
}
```

Next, we define the likelihood function; note that the parameter definitions for the parsimonious Mat{e}rn are different than for the conditional approach:

```
loglik_sh_pars_model <- function(theta,shift=FALSE) {
  # theta1: delta1
  # theta2: delta2
  # theta3: kappa
  # theta4: nu1
```

```

# theta5: nu2
# theta6: rho
# theta7: tau1
# theta8: tau2/100
# theta9: sigma2_11
# theta10: sigma2_22/10000
if(theta[3] < 0.00001 | theta[4] < 0.1 | theta[5] < 0.1 |
  abs(theta[6]) >= 1 | theta[7] <= 0 | theta[8] <= 0 |
  theta[9] <= 0 | theta[10] <= 0) {
  return(Inf)
} else if (abs(theta[6]) > sqrt(theta[4] * theta[5]) / 0.5*(theta[4] + theta[5]))
{
  return(Inf)
} else {

  if(!shift) {
    theta[1] <- theta[2] <- 0
  }

  theta[8] <- theta[8] * 100    ## These were divided by 100 and 1000
  theta[10] <- theta[10] * 10000 ## in the initial call, respectively

  ## Get the data
  Z = form_Z(1L,scale=F)

  ## Get the matrices
  Matrices <- sh_pars_mats(theta)
  S <- Matrices$SY + Matrices$So
  cholS <- chol(S)

  loglik <-
    -as.numeric(-0.5 * determinant(S)$modulus -
      0.5 * t(Z) %*% chol2inv(cholS) %*% Z -
      0.5 * nrow(Z)*log(2*pi))

  return(loglik)
}
}

```

Finally, we estimate the parameters:

```

## Set shift=TRUE below to estimate the shift parameters; otherwise they are fixed to zero
## and we get the same likelihood and parameter estimates of Gneiting (2010)
if(Shifted_Pars_estimation) {
  optim_est_sh_pars <- optim(c(0.5,-1,1/95.88,
    0.6,1.6,-0.5,
    0.019,69.66/100,
    6.95,
    67191/10000),
    loglik_sh_pars_model,
    control=list(trace=6,
      maxit=3000),
    shift=TRUE)
  save(optim_est_sh_pars,file = "../inst/extdata/temp_pressure/Shifted_Pars_est_results.rda")
}

```

```

} else {
  load(system.file("extdata/temp_pressure", "Shifted_Pars_est_results.rda", package = "bicon"))
}

print(paste0("Log-lik. and AIC of shifted parsimonious Matern are -",
  optim_est_sh_pars$value,
  " and ", optim_est_sh_pars$value*2 + 20,
  ", respectively.)) # 10 parameters

```

```
## [1] "Log-lik. and AIC of shifted parsimonious Matern are -1260.87294870203 and 2541.74589740405, resp"
```

For co-kriging (for cross-validation) we follow the same approach as earlier:

```

theta_lk <- optim_est_sh_pars$par
theta_lk[8] <- theta_lk[8] * 100    ## These were divided by 100 and 1000
theta_lk[10] <- theta_lk[10] * 10000 ## in the initial call, respectively
X <- sh_pars_mats(theta = theta_lk)

## Loop over each observation location
print("Running CV for shifted parsimonious Matern...")

```

```
## [1] "Running CV for shifted parsimonious Matern..."
```

```

pred_sh_pars <- foreach(i = 1:m1, .combine = "rbind") %do% {
  X <- sh_pars_mats(theta = theta_lk)
  cbind(cokrige(X=X, i=c(i, i+m1)),
        model_num = 1)}

## Extract our results
sh_pars_results <- sanitise_results(pred_sh_pars) %>%           # Only consider models with Y1 temp.
  dplyr::select(model_num, MAE, RMSPE, CRPS)

```

```
## Joining by: c("process", "loc_num")
```

```
sh_pars_results$model_num <- "Shifted_Pars"
```

For completeness we also carry out LOOCV on the parsimonious and full Matérn models using the RandomFields package. This closely follows the approach illustrated in M. Schlather, Malinowski, Menck, Oesting, & Strokorb (2015).

```

Dist.mat <- as.vector(RFearth2dist(as.matrix(weather[, 3:4]))) # Compute distances
PT <- as.matrix(weather[, 1:2])                               # Change data into matrix
## Note the below code only works with RandomFields v3.0.62
if(RF_estimation) {

  ## Parsimonious Matern model
  nug <- RMmatrix(M = matrix(nc = 2, c(NA, 0, 0, NA)), RMnugget()) # nugget model
  pars.model <- nug + RMbiwm(nudiag = c(NA, NA), scale = NA,         # parsimonious model
                            cdiag = c(NA, NA), rhored = NA)
  RFpars <- RFfit(pars.model, distances = Dist.mat, dim = 3, data = PT) # fit model
  print(RFpars)
}

```

```

CVresults.pars <- RFcrossvalidate(RFpars,                # carry out CV
                                x = as.matrix(weather[, 3:4]),
                                data = PT,
                                full = TRUE)           # on all the data (no re-fitting)

## Full Matern model
whole.model <- nug + RMbiwm(nudiag = c(NA, NA), nured = NA,
                           s = rep(NA, 3), cdiag = c(NA, NA), rhored = NA)
RFwhole <- RFfit(whole.model, distances = Dist.mat, dim = 3, data = PT)
CVresults.whole <- RFcrossvalidate(RFwhole,
                                   x = as.matrix(weather[, 3:4]),
                                   data = PT,
                                   full = TRUE)

## cache results
save(CVresults.pars, RFpars, CVresults.whole, RFwhole, file = "../inst/extdata/temp_pressure/RF_CV_results.rda")

} else {
  ## otherwise load
  #load("../inst/extdata/temp_pressure/RF_CV_results.rda")
  load(system.file("extdata/temp_pressure", "RF_CV_results.rda", package = "bicon"))
}

```

```

## Combine results into one long data frame
RFPred2 <- rbind(data.frame(Z = weather$pressure,
                            mu_pred = CVresults.pars$user's model$predicted[,1],
                            var_pred = CVresults.pars$user's model$krige.var[,1],
                            loc_num = 1:m1,
                            model_name = "Pars",
                            process = "pressure"),
  data.frame(Z = weather$temperature,
            mu_pred = CVresults.pars$user's model$predicted[,2],
            var_pred = CVresults.pars$user's model$krige.var[,2],
            loc_num = 1:m1,
            model_name = "Pars",
            process = "temperature"),
  data.frame(Z = weather$pressure,
            mu_pred = CVresults.whole$user's model$predicted[,1],
            var_pred = CVresults.whole$user's model$krige.var[,1],
            loc_num = 1:m1,
            model_name = "Whole",
            process = "pressure"),
  data.frame(Z = weather$temperature,
            mu_pred = CVresults.whole$user's model$predicted[,2],
            var_pred = CVresults.whole$user's model$krige.var[,2],
            loc_num = 1:m1,
            model_name = "Whole",
            process = "temperature")) %>%
left_join(weather_long) %>%
dplyr::select(-z)

```

```
## Joining by: c("loc_num", "process")
```

```
## Warning in left_join_impl(x, y, by$x, by$y): joining character vector and
```



```

## factor, coercing into character vector

## Get out the diagnostics
RFresults <- RFpred2 %>%
  group_by(process,model_name) %>% # Group by process and model name
  summarise(MAE = mean(abs(mu_pred - Z)),
            MAE_se = sd(mu_pred - Z)/sqrt((m1-1)),
            Bias = mean(mu_pred - Z),
            Bias_se = sd(mu_pred - Z)/sqrt((m1-1)),
            Bias_norm = mean((mu_pred - Z)/sqrt(var_pred)),
            RMSPE = sqrt(mean((mu_pred - Z)^2)),
            CRPS = mean(crps_wrapper(Z,mu_pred,sqrt(var_pred))),
            CRPS_se = sd(crps_wrapper(Z,mu_pred,sqrt(var_pred)))/sqrt(m1-1))

## Extract our results
select_results <- filter(results,model_num < 5) %>% # Only consider models with Y1 temp.
  dplyr::select(model_num,MAE,RMSPE,CRPS)

## Extract RandomFields results
select_resultsRF <- dplyr::select(RFresults,model_name,MAE,RMSPE,CRPS)

## Relabel columns
colnames(select_results) <- colnames(select_resultsRF) <- colnames(sh_pars_results) <-
  c("Process","Model","MAE","RMSPE","CRPS")

## Join our results with those from RandomFields and the shifted parsimonious Matern
all_results <- rbind(select_results,select_resultsRF,sh_pars_results) %>%
  as.data.frame() %>% arrange(Process)

## Print the LaTeX table
print(xtable::xtable(all_results,digits=3),
      each = "column", max = c(F,NA,NA,T,NA),
      sanitize.text.function=function(x){x},
      hline.after=NULL,include.rownames=FALSE)

## % latex table generated in R 3.2.0 by xtable 1.7-4 package
## % Wed Jun 22 15:38:19 2016
## \begin{table}[ht]
## \centering
## \begin{tabular}{llrrr}
## Process & Model & MAE & RMSPE & CRPS \\
## pressure & 1 & 69.557 & 123.356 & 55.327 \\
## pressure & 2 & 70.190 & 124.411 & 55.640 \\
## pressure & 3 & 70.317 & 122.995 & 55.187 \\
## pressure & 4 & 66.069 & 114.671 & 51.725 \\
## pressure & Pars & 70.150 & 122.970 & 55.349 \\
## pressure & Whole & 66.189 & 122.758 & 55.225 \\
## pressure & Shifted_Pars & 67.009 & 114.978 & 52.479 \\
## temperature & 1 & 1.144 & 1.625 & 0.813 \\
## temperature & 2 & 1.144 & 1.626 & 0.814 \\
## temperature & 3 & 1.095 & 1.530 & 0.780 \\
## temperature & 4 & 1.080 & 1.465 & 0.767 \\
## temperature & Pars & 1.110 & 1.562 & 0.790

```

```
## temperature & Whole & 1.109 & 1.576 & 0.792 \\
## temperature & Shifted_Pars & 1.091 & 1.478 & 0.771 \\
## \end{tabular}
## \end{table}
```

```
best_performing <- pred2 %>%
  filter(model_num %in% 1:4) %>%
  group_by(lon,lat,process) %>%
  summarise(best_model_MAE = model_num[which.min(abs(mu_pred - Z))],
            best_model_CRPS = model_num[which.min(crps_wrapper(Z,mu_pred,sqrt(var_pred)))] %>%
  mutate(reversed = best_model_MAE > 4)

best_summary <- group_by(best_performing,process,best_model_MAE) %>%
  summarise(tot_count = length(best_model_MAE),
            reverse = best_model_MAE[1] > 4)
```

```
if(show_figs) {

  g1 <- LinePlotTheme() +
    stat_qq(data=filter(pred2,model_num %in% c(1,4)),
            aes(sample=(mu_pred - Z)/sqrt(var_pred),
                shape=process,
                colour=as.factor(model_num)),
            size=4) +
    geom_abline()

  g2 <- LinePlotTheme() + geom_bar(data=best_performing,
                                   aes((best_model_MAE - 1)%%4 + 1,fill=c(process)),
                                   position="dodge",
                                   binwidth=0.5) +
    scale_fill_grey(guide_legend(title="process")) +
    xlab("best model for MAE")

  g3 <- LinePlotTheme() + geom_bar(data=best_performing,
                                   aes((best_model_CRPS - 1)%%4 + 1,fill=c(process)),
                                   position="dodge",
                                   binwidth=0.5) +
    scale_fill_grey(guide_legend(title="process")) +
    xlab("best model for CRPS")

  plot(arrangeGrob(g1,g2,g3,nrow=1))
}
```

```
## Warning: `geom_bar()` no longer has a `binwidth` parameter. Please use
## `geom_histogram()` instead.
```

```
## Warning: `geom_bar()` no longer has a `binwidth` parameter. Please use
## `geom_histogram()` instead.
```

```
table(best_performing$best_model_MAE)
```

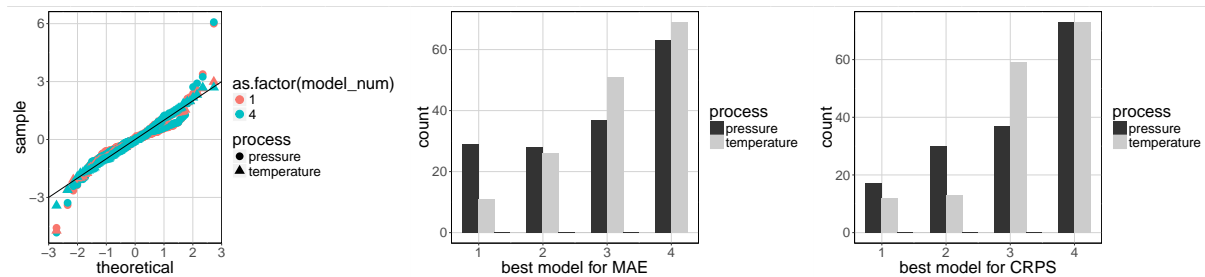


Figure 1: QQ plots and bar charts illustrating the number of times each model does best at different spatial locations

```
##
##    1    2    3    4
##  40   54   88  132
```

```
table(best_performing$best_model_CRPS)
```

```
##
##    1    2    3    4
##   29   43   96  146
```

```
# Load shapefiles
shape1 <- system.file("extdata", "cb_2013_us_state_5m.shp", package = "bicon")
shape2 <- system.file("extdata", "Canada_provinces.SHP", package = "bicon")
#shape1 <- "../inst/extdata/cb_2013_us_state_5m.shp"
#shape2 <- "../inst/extdata/Canada_provinces.SHP"
US_States <- maptools::readShapeSpatial(shape1)
Ca_States <- maptools::readShapeSpatial(shape2)

US_States_fort <- fortify(US_States) %>% mutate( id= as.numeric(id))
```

```
## Regions defined for each Polygons
```

```
Ca_States_fort <- fortify(Ca_States) %>% mutate( id= as.numeric(id)+100)
```

```
## Regions defined for each Polygons
```

```
US_State_names <- data.frame(Name = US_States$NAME, id = 0:(length(US_States$NAME)-1))
Ca_State_names <- data.frame(Name = Ca_States$NAME, id = 0:(length(Ca_States$NAME)-1) + 100)

US_States_fort <- left_join(US_States_fort,US_State_names)
```

```
## Joining by: "id"
```

```
Ca_States_fort <- left_join(Ca_States_fort,Ca_State_names)
```

```
## Joining by: "id"
```

```

All_States <- rbind(US_States_fort,Ca_States_fort) %>%
  filter(Name %in% c("British Columbia","Alberta",
                    "Washington","Oregon",
                    "Idaho","California",
                    "Nevada","Montana"))

conv_hull <- Mesh[c("x","y")][chull(Mesh[c("x","y")]),]
conv_hull <- rbind(conv_hull,conv_hull[1,])

Stateplot <- LinePlotTheme() +
  geom_path(data=All_States,aes(long,lat,group=group,label=Name),linetype="solid") +
  coord_fixed(xlim=c(-137,-110),ylim=c(35,58)) +
  geom_path(data=conv_hull,aes(x,y),size=1,colour="black",linetype="dashed") +
  theme(plot.margin = grid::unit(c(2, 2, 2, 2),units="mm")) + xlab("lon")

meshplot <- function(g,include_obs=1L) {
  p <- plot(Mesh,g=g,plot_dots=F)
  p <- p +
    xlab('lon') + ylab('lat') +
    coord_fixed(xlim=c(-137,-110),ylim=c(35,58)) +
    geom_path(data=All_States,aes(long,lat,group=group,label=Name),linetype="solid") +
    geom_path(data=conv_hull,aes(x,y),size=1,colour="black",linetype="dashed")
  if(include_obs) p <- p + geom_point(data = weather,aes(lon,lat),size=3,col="red")
  p
}

if(show_figs) {
  States_mesh <- meshplot(LinePlotTheme())
  print(States_mesh)
}

```

```

g1M1 <- plot_interp(Mesh,"y1_Model1",150,max=5,min=-5,leg_title="degC") %>%
  meshplot(include_obs = 0L)
g1M4 <- plot_interp(Mesh,"y1_Model4",150,max=5,min=-5,leg_title="degC") %>%
  meshplot(include_obs = 0L)

g2M1 <- plot_interp(Mesh,"y2_Model1",150,max=500,min=-500,leg_title="Pa") %>%
  meshplot(include_obs = 0L)
g2M4 <- plot_interp(Mesh,"y2_Model4",150,max=500,min=-500,leg_title="Pa") %>%
  meshplot(include_obs = 0L)
if(show_figs) {
  plot(arrangeGrob(g1M1,g1M4,g2M1,g2M4,ncol=2))
}

```

```

par1 <- fit_all_data[[4]]$par
par2 <- fit_all_data[[8]]$par
x <- seq(-3,3,length=100)
XY <- expand.grid(h1=x,h2=x)
XY$b1 <- bisquare_2d(h1=XY[,1],h2=XY[,2],delta=par1[10:11],r=par1[9],A=par1[8])
XY$b2 <- bisquare_2d(h1=XY[,1],h2=XY[,2],delta=par2[10:11],r=par2[9],A=par2[8])

axes <- geom_line(data=data.frame(x=c(0,0,-5,5),y=c(-5,5,0,0),grp=c(1,1,2,2)),

```

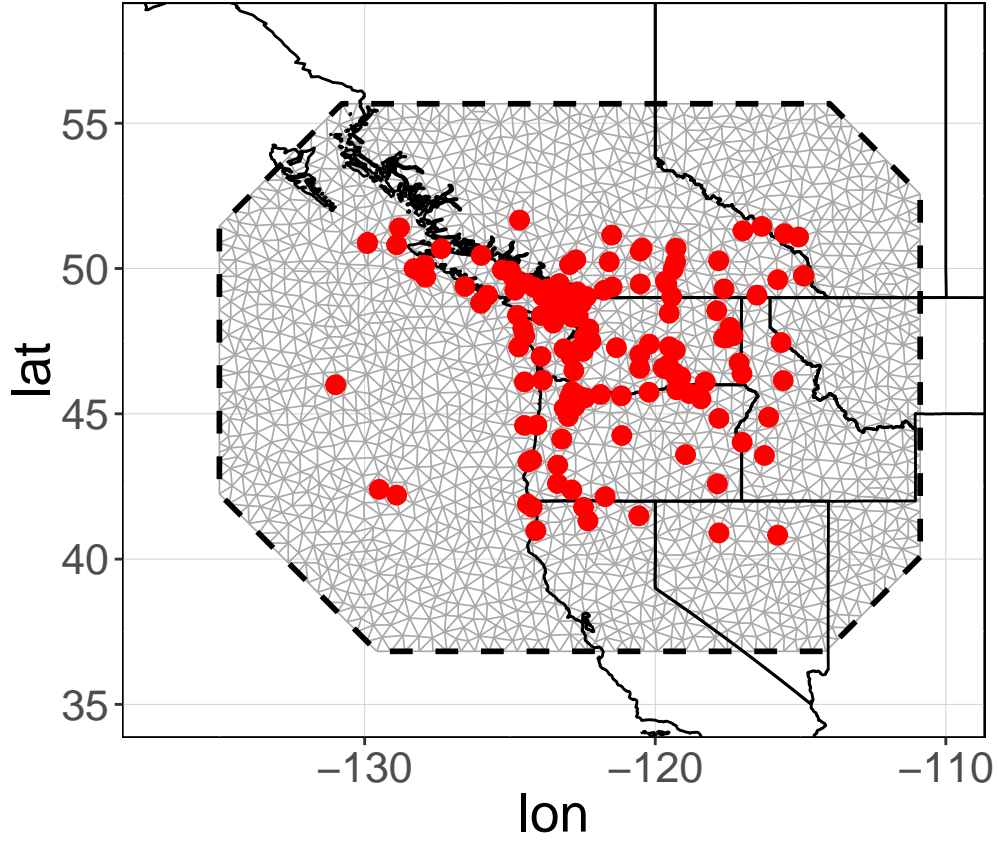


Figure 2: State boundaries and province boundaries of a region of the U.S.A. and Canada (dark solid lines), with the domain of interest enclosed by a bounding polygon (dashed line). The irregular triangular grid used for discretizing  $D$  (light solid lines) and the observation locations given by  $D^O$  (dots) are also shown. The discretized spatial domain  $D^L$  consists of the vertices of the triangular grid.

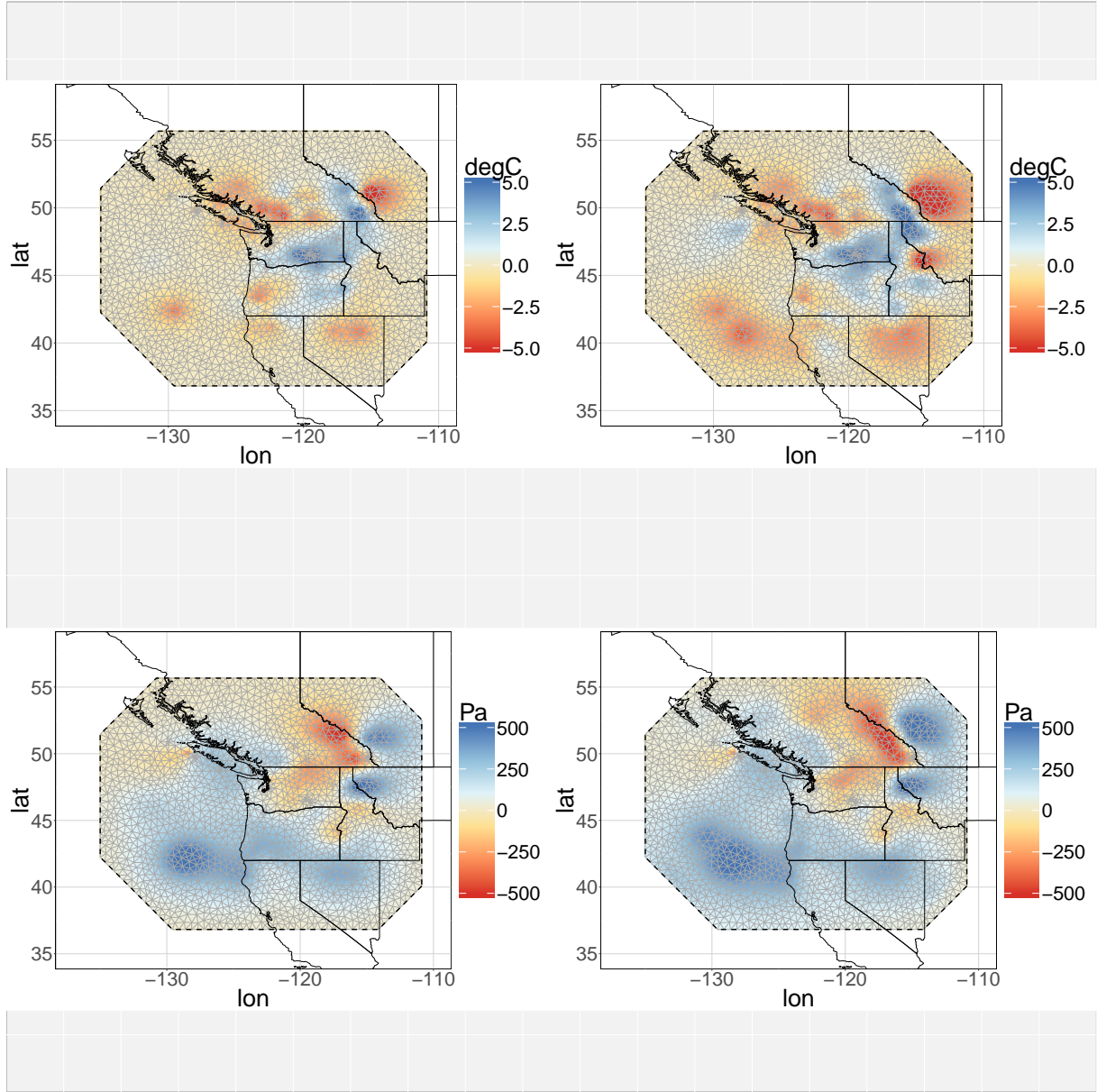


Figure 3: The cokriged surface using maximum likelihood estimates for the parameters with Model 1 (left panels) and Model 4 (right panels) for the temperature (top panels) and pressure (bottom panels) error fields.

```

aes(x,y,group=grp),colour="black")
library(scales)
greys <- scale_fill_gradient(low="black",high="white")
bluered <- scale_fill_gradient2(low=muted("blue"),high=muted("red"))
g1 <- LinePlotTheme() + geom_tile(data=XY,aes(h1,h2,fill=b1)) +
  greys + axes + coord_fixed() + ggtitle("Model 4")
g2 <- LinePlotTheme() + geom_tile(data=XY,aes(h1,h2,fill=b2)) +
  axes + greys + coord_fixed() + ggtitle("Model 8")
if(show_figs) plot(arrangeGrob(g1,g2,nrow=1))

```

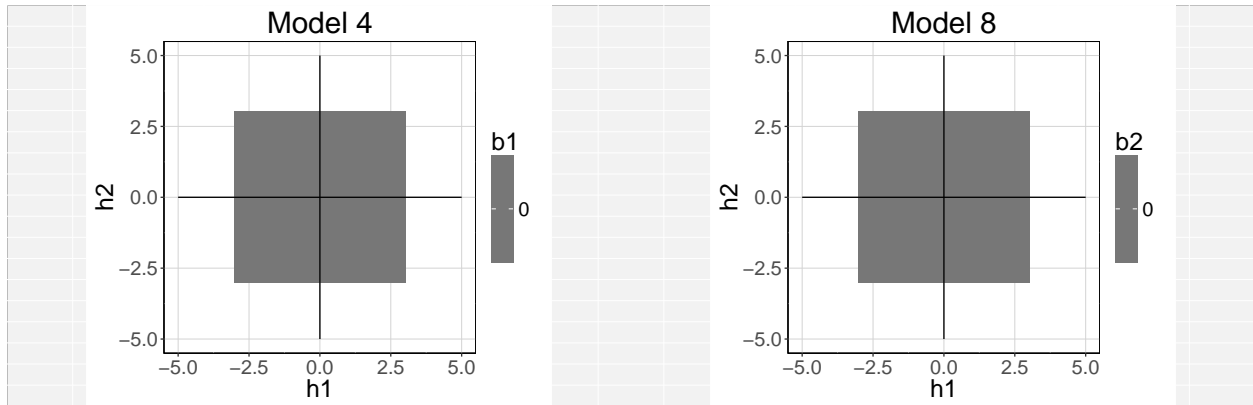


Figure 4: Cross-covariance functions for Model 4 and Model 8 (that is, Model 4 with temperature and pressure reversed)

```

h1_grid<- seq(-4,4,by=0.2)
h2_grid <- seq(-4,4,by=0.2)
Disp <- expand.grid(h1 = h1_grid,h2 = h2_grid)
xo <- seq(-127,-113,by=0.2)
yo <- seq(40,55,by=0.2)

doMC::registerDoMC(6)
Disp$corr <- foreach(i = 1:nrow(Disp),.combine="c") %dopar% {
  Temp <- akima::interp(mesh_locs[,1]-Disp$h1[i],mesh_locs[,2]-Disp$h2[i],
    Mesh["y1_Model11"],xo,yo)
  Pres <- akima::interp(mesh_locs[,1],mesh_locs[,2],
    Mesh["y2_Model11"],xo,yo)
  cor(c(Pres$z),c(Temp$z),"na.or.complete")
}
axes <- geom_line(data=data.frame(x=c(0,0,-4,4),y=c(-4,4,0,0),grp=c(1,1,2,2)),
  aes(x,y,group=grp),colour="black")
corr_plot <- LinePlotTheme() + geom_tile(data=Disp,aes(h1,h2,fill=corr)) + axes + bluered +
  geom_point(data=data.frame(d1 = par1[10],d2 = par1[11]),
    aes(d1,d2),pch=9,size=5,colour="yellow") +
  coord_fixed(xlim=c(-2.5,2.5),ylim=c(-2.5,2.5))

if(show_figs) print(corr_plot)

```

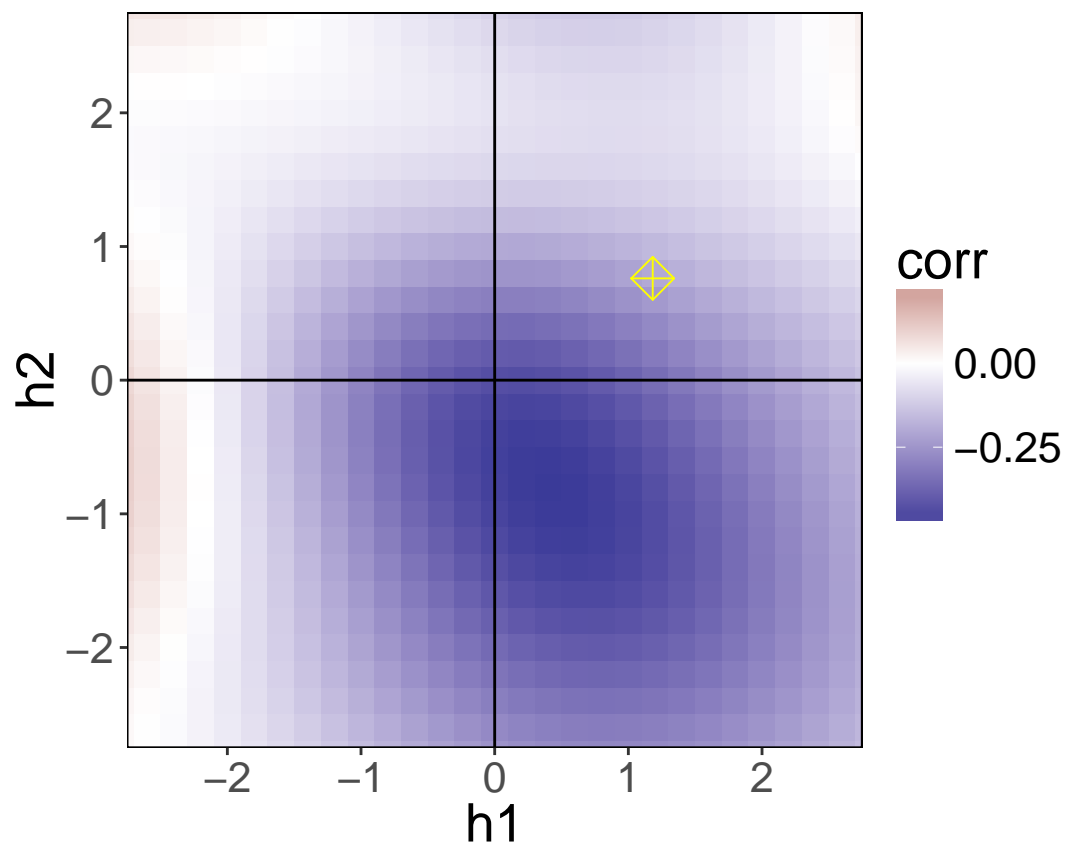


Figure 5: Correlation between interpolated gridded maps of temperature and pressure as a function of displacement of the temperature field in latitude/longitude degrees,  $h$ . The yellow symbol indicates the maximum-likelihood estimate of the shift parameter with Model 4.



```

if(print_figs) ggsave(corr_plot,
                      filename = file.path(img_path,"T-P-corr.png"),
                      width=8,height=7,family="Arial")

#delta <- 0.2
#x_grid<- seq(-5,5,by=delta)
#y_grid <- seq(-5,5,by=delta)
delta <- 0.25
x_grid<- seq(-131,-115,by=delta)
y_grid <- seq(40,50,by=delta)
grid_locs <- expand.grid(x = x_grid, y = y_grid) %>% as.matrix()

# D_grid <- as.matrix(dist(grid_locs))
D_grid <- as.matrix(RFearth2dist(grid_locs))
Dvec_grid <- as.double(c(D_grid))

# Dvec_grid <- fields::rdist(grid_locs,matrix(c(0,0),1,2))
# h_grid <- t(t(grid_locs) -grid_locs[i,])

n2_grid <- n1_grid <- nrow(grid_locs)
h_grid <- matrix(0,n1_grid^2,2)
areas_grid <- rep(delta^2,n1_grid^2)
for(i in 1:n2_grid) {
  h_grid[((i-1)*n1_grid+1):(i*n1_grid),] <- t(t(grid_locs) -grid_locs[i,])
}

theta = fit_all_data$Model4$par
B <- theta[9]*bisquare_B(h_grid[,1],h_grid[,2],
                        delta=theta[11:12], # Automatically zero for Model with no shift
                        r=theta[10],
                        n1 = n1_grid,
                        n2 = n2_grid,
                        areas = areas_grid)

S11 <- makeS(r = Dvec_grid,var = theta[3],
            kappa = theta[5],nu = theta[7])
S11 <- S11 + theta[1]^2*diag(nrow(S11))
S2_1 <- makeS(r = Dvec_grid, var = theta[4],
            kappa = theta[6],nu = theta[8])
S21 <- B %*% S11
S12 <- t(S21)
S22 <- S2_1 + Matrix::crossprod(chol(S11) %*% t(B))
S22 <- S22 + theta[2]^2*diag(nrow(S22))

#centre_node <- which(grid_locs[,1] == 0 & grid_locs[,2] == 0)
centre_node <- which(grid_locs[,1] == -123 & grid_locs[,2] == 45)
# h_centre <- t(t(grid_locs) -c(0,0))
h_centre <- t(t(grid_locs) -c(-123,45))
H <- data.frame(h1 = h_centre[,1],
                h2 = h_centre[,2],
                S11 = (S11)[centre_node,],
                S12 = (S12)[centre_node,],
                S21 = (S21)[centre_node,],

```

```

      S22 = (S22)[centre_node,]) %>%
gather(cov_mat,C,-h1,-h2) %>%
mutate(Cgrp1 = ifelse(cov_mat %in% c("S11","S12"),"Y1","Y2"),
      Cgrp2 = ifelse(cov_mat %in% c("S11","S21"),"Y1","Y2")) %>%
group_by(cov_mat) %>%
mutate(corr = C / max(abs(C)))

axes <- geom_line(data=data.frame(x=c(0,0,-5,5),y=c(-5,5,0,0),grp=c(1,1,2,2)),
      aes(x,y,group=grp),colour="black")
corr_fn_plot <- LinePlotTheme() + geom_tile(data=H,aes(h1,h2,fill=corr)) +
      geom_contour(data=H,aes(h1,h2,z=corr),binwidth=0.2,colour="black",lty="dashed") +
      facet_grid(Cgrp1~Cgrp2) + axes +
      scale_fill_gradient2(low="blue",high="red") +
      coord_fixed(xlim=c(-4,4),ylim=c(-4.2,4.2)) +
      theme(panel.margin = grid::unit(3, "lines"))
if(print_figs) ggsave(corr_fn_plot,
      filename = file.path(img_path,"T-P-cov.png"),
      width=8,height=7,family="Arial")

if(print_figs) {
  g <- arrangeGrob(States_mesh,corr_fn_plot,ncol=2)
  ggsave(g,
      filename = file.path(img_path,"Fig2.eps"),
      width=14,height=6,family="Arial")
}

```

```

Mesh["M1T_errors"] <- sqrt(pmax(ALL1$var_pred,0))[1:n1]
Mesh["M4T_errors"] <- sqrt(pmax(ALL4$var_pred,0))[1:n1]
Mesh["diffT_errors"] <- Mesh["M1T_errors"] - Mesh["M4T_errors"]
Mesh["ratioT_errors"] <- Mesh["M1T_errors"]/pmax(Mesh["M4T_errors"],1e-9)
std_diff <- (plot_interp(Mesh,"diffT_errors",150) +
      scale_fill_gradient2(low=muted("yellow"),
      mid="white",
      high=muted("magenta"),
      guide = guide_legend(title="diff (deg. C)")) %>%
      meshplot(include_obs = 1L)

```

## Scale for 'fill' is already present. Adding another scale for 'fill',  
## which will replace the existing scale.

```

# std_diff <- (plot_interp(Mesh,"ratioT_errors",150) +
#       scale_fill_gradient(low="green", high="red",limits=c(0.77,1.5),
#       guide =guide_legend(title="diff (deg. C)")) %>%
#       meshplot(include_obs = 1L); std_diff

X <- data.frame(M1_errors = Mesh["M1T_errors"],
      M4_errors = Mesh["M4T_errors"])
std_diff_scatter <- LinePlotTheme() +
      geom_point(data=X,aes(M1_errors,M4_errors,fill = M1_errors - M4_errors),
      colour="black",size=4,shape=21) +
      scale_fill_gradient2(low=muted("green"),mid="white", high=muted("magenta"),
      guide=guide_legend(title="diff (deg. C)")) +

```

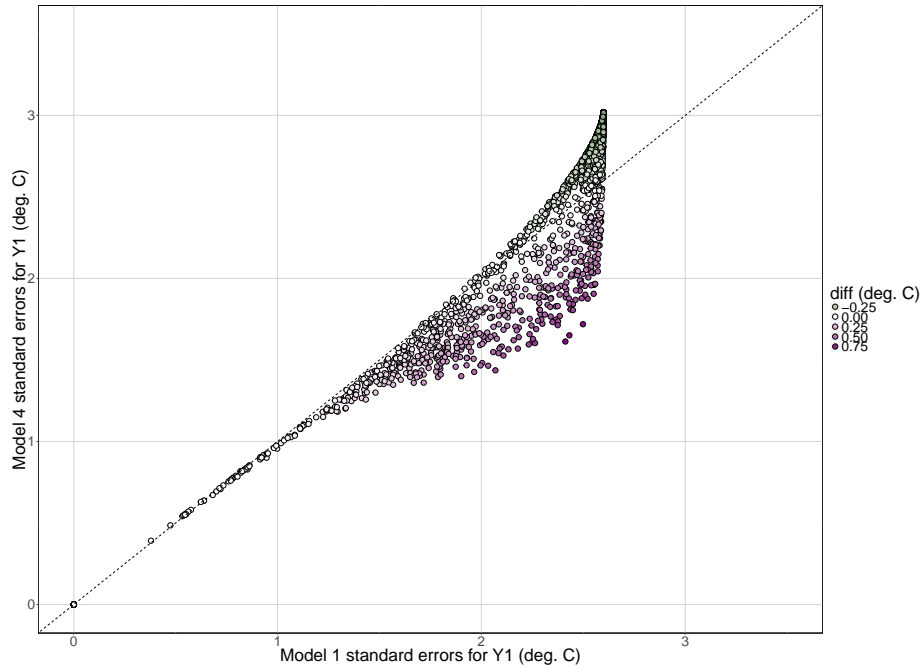


Figure 6: A scatter plot of the kriging standard errors of  $Y_1$  obtained with Model 4 against those obtained with Model 1 at each of the mesh vertices. The colour illustrates the difference between the two, with green denoting the higher standard error of Model 4 and purple the higher standard error of Model 1.

```
geom_line(data=data.frame(x=c(-5,5),y=c(-5,5)),aes(x,y),linetype="dashed") +
xlab("Model 1 standard errors for Y1 (deg. C)") +
ylab("Model 4 standard errors for Y1 (deg. C)") +
  coord_fixed(xlim=c(0,3.5),ylim=c(0,3.5),ratio=0.8) +
  theme(text = element_text(size = 30), axis.title.y = element_text(vjust=2));
if (show_figs) print(std_diff_scatter)
```

```
if(print_figs) {
  g <- arrangeGrob(g1M4,(std_diff_scatter +
                        theme(plot.margin = grid::unit(c(10, 10, 10, 10),
                                                          units="mm"))),
                  g2M4,std_diff,ncol=2)

  ggsave(g,
    filename = file.path(img_path,"Fig3.eps"),
    width=28,height=16,family="Arial")
}
```

## References

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- Gneiting, T., Kleiber, W., & Schlather, M. (2010). Matérn cross-covariance functions for multivariate random fields. *Journal of the American Statistical Association*, 105, 1167–1177.
- Gneiting, T., Raftery, A. E., Westveld III, A. H., & Goldman, T. (2005). Calibrated probabilistic forecasting

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Li, B., & Zhang, H. (2011). An approach to modeling asymmetric multivariate spatial covariance structures. *Journal of Multivariate Analysis*, 102, 1445–1453.

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