

CSC484 Assignment #2

Andrii Osipa

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Exercise 1.1

Given p_1, \dots, p_n such that $\sum_{k=1}^n p_k = 1$.

Algorithm. We want to look at process of sampling in the following way: select uniformly bucket $b_k, k \in \overline{1, n}$ and then pick value l_k with probability p'_k or r_k with probability $1 - p'_k$, where l_k and r_k are values from $\{0, \dots, n - 1\}$. And such that following holds: $\sum_{l_k: l_k=i} 1/n * p'_k + \sum_{r_k: r_k=i} 1/n * (1 - p'_k) = p_i$. In other

words we split original probabilities between buckets such that total probability of bucket is $1/n$ and probability of getting value i is same as it was initially. Preprocessing: split all k into two groups: p_k less than $1/n$ (L) and others (R). Then for each element $l \in L$ we pick any available element $r \in R$ and form a bucket b_k such that $l_k = l, r_k = r, p'_k = p_l$. Then we decrease probability p_r by $1/n - p_l$ and, if needed (prob p_r less than $1/n$), move r to list L . Also we delete l from list L ($\Leftrightarrow p_l := 0$).

On k^{th} step we have $n - k$ values left in lists L and R and sum of all probabilities left is $\frac{n - k}{n}$, as on each step we decrease sum of probabilities by $1/n$. It is not possible that there is no elements left in list R , but $L \neq \emptyset$ on some step as it means that sum of all probabilities left is less than $\frac{n - k}{n}$. If we have that some elements left in R and $L = \emptyset$ this means that each of them has probability $1/n$, which is obviously the only possibility in this case.

Note: in computing case $L \neq \emptyset, R = \emptyset$ may occur due to finiteness of numbers processed, we just say that probability of each $l \in L$ is $1/n$.

Sampling: select uniformly bucket $b_k, k \in \overline{1, n}$, then pick random number in $[0, 1/n]$. If number is greater than p_k output r_k , else l_k .

Time complexity of preprocessing: $O(n)$ as we just loop through all p_k on the beginning to create lists L and R and then we do n steps on each removing one value from list L .

Time complexity of sampling: $O(1)$ as we just do 3 operations: pick random k pick random number in $[0, 1/n]$ and "if statement".

Note: to simplify we can multiply all initial probabilities by n then total sum is n and on each step we have sum of probabilities of left values $n - k$ and each bucket has probability 1.

Exercise 2.1

Cauchy distribution density $f(x) = \left(\pi\gamma \left(1 + \left(\frac{x-x_0}{\gamma} \right)^2 \right) \right)^{-1}$.

$X, X_1, X_2, X_3, X_4, X_5$ from Cauchy (x_0, γ) .

$$E[X] = \int_{-\infty}^{\infty} \frac{x}{\pi\gamma \left(1 + \left(\frac{x-x_0}{\gamma} \right)^2 \right)} dx = \lim_{a \rightarrow \infty} \lim_{b \rightarrow -\infty} \int_b^a \frac{x}{\pi\gamma \left(1 + \left(\frac{x-x_0}{\gamma} \right)^2 \right)} dx.$$

Consider $\lim_{a \rightarrow \infty} \int_{-a+x_0}^{a+x_0} \frac{x}{\pi\gamma \left(1 + \left(\frac{x-x_0}{\gamma} \right)^2 \right)} dx = x_0$. If the former limit exists it must be also equal x_0 . Consider $\lim_{a \rightarrow \infty} \int_{-a}^{2a} \frac{x}{\pi\gamma \left(1 + \left(\frac{x-x_0}{\gamma} \right)^2 \right)} dx =$

$$\lim_{a \rightarrow \infty} \int_a^{2a} \frac{x}{\pi\gamma \left(1 + \left(\frac{x-x_0}{\gamma} \right)^2 \right)} dx = x_0 + \lim_{a \rightarrow \infty} \int_a^{2a} \frac{x}{\pi\gamma \left(1 + \left(\frac{x-x_0}{\gamma} \right)^2 \right)} dx \geq x_0 +$$

$\frac{\gamma^2}{4}$ by simple inequalities. Therefore $E[X]$ is not defined, so $V[X]$ and squared coefficient of variation.

$Y = (X_1, X_2, X_3)$. From statistics we know that given sample from distribution with density $f(x)$ and cumulative distribution function $F(x)$ we can write density for k^{th} order statistics as following: $f_k(x) = kC_n^k F(x)^{k-1} f(x)(1-F(x))^{n-k}$. For Cauchy distribution we have cumulative distribution function $F(x) = \frac{1}{\pi} \arctan \frac{x-x_0}{\gamma} + \frac{1}{2}$. Therefore for Y we have the following density function:

$$f_{median} = 6 \left(\frac{1}{\pi} \arctan \frac{x-x_0}{\gamma} + \frac{1}{2} \right) \frac{1}{\pi\gamma \left(1 + \left(\frac{x-x_0}{\gamma} \right)^2 \right)} \left(\frac{1}{2} - \frac{1}{\pi} \arctan \frac{x-x_0}{\gamma} \right).$$

$$\begin{aligned}
E[Y] &= \frac{6}{\pi\gamma} \int_{-\infty}^{\infty} \left(\frac{1}{\pi} \arctan \frac{x-x_0}{\gamma} + \frac{1}{2} \right) \frac{x}{1 + \left(\frac{x-x_0}{\gamma} \right)^2} \left(\frac{1}{2} - \frac{1}{\pi} \arctan \frac{x-x_0}{\gamma} \right) dx = \\
&= \frac{6}{\pi} \int_{-\infty}^{\infty} \left(\frac{1}{\pi} \arctan \frac{x-x_0}{\gamma} + \frac{1}{2} \right) \frac{\frac{x-x_0}{\gamma}}{1 + \left(\frac{x-x_0}{\gamma} \right)^2} \left(\frac{1}{2} - \frac{1}{\pi} \arctan \frac{x-x_0}{\gamma} \right) dx + \\
&+ \frac{6}{\pi\gamma} \int_{-\infty}^{\infty} \left(\frac{1}{\pi} \arctan \frac{x-x_0}{\gamma} + \frac{1}{2} \right) \frac{x_0}{1 + \left(\frac{x-x_0}{\gamma} \right)^2} \left(\frac{1}{2} - \frac{1}{\pi} \arctan \frac{x-x_0}{\gamma} \right) dx
\end{aligned}$$

We can rewrite integrals as following, skipping constants:

$$\begin{aligned}
&\int_{-\infty}^{\infty} \left(\frac{1}{4} - \frac{1}{\pi^2} \arctan^2 z \right) \frac{z}{1+z^2} dz, \\
&\lim_{z \rightarrow +/\infty} z \left(\frac{1}{4} - \frac{1}{\pi^2} \arctan^2 z \right) = +/ - \frac{1}{\pi} \Rightarrow \exists c > 0 : \forall |z| > c : |z \left(\frac{1}{4} - \frac{1}{\pi^2} \arctan^2 z \right)| < C \\
&\Rightarrow 0 \leq \int_0^{\infty} \left(\frac{1}{4} - \frac{1}{\pi^2} \arctan^2 z \right) \frac{z}{1+z^2} dz < \int_0^{\infty} \frac{C}{1+z^2} < \infty; \\
&\Rightarrow 0 \geq \int_{-\infty}^0 \left(\frac{1}{4} - \frac{1}{\pi^2} \arctan^2 z \right) \frac{z}{1+z^2} dz > \int_{-\infty}^0 \frac{-C}{1+z^2} > -\infty.
\end{aligned}$$

Therefore this integral converges and easy to see that it is equal to 0 (function symmetric).

Second integral in the sum above: $\int_{-\infty}^{\infty} \left(\frac{1}{4} - \frac{1}{\pi^2} \arctan^2 z \right) \frac{1}{1+z^2} dz = \frac{\pi}{6}$ is known result.

We obtain the following: $E[Y] = x_0$.

Variance of Y.

$$V[Y] = \frac{6}{\pi\gamma} \int_{-\infty}^{\infty} \left(\frac{1}{\pi} \arctan \frac{x-x_0}{\gamma} + \frac{1}{2} \right) \frac{(x-x_0)^2}{1 + \left(\frac{x-x_0}{\gamma} \right)^2} \left(\frac{1}{2} - \frac{1}{\pi} \arctan \frac{x-x_0}{\gamma} \right) dx$$

so we are interested in the following integral:

$$\int_{-\infty}^{\infty} \left(\frac{1}{\pi} \arctan z + \frac{1}{2} \right) \frac{z^2}{1+z^2} \left(\frac{1}{2} - \frac{1}{\pi} \arctan z \right) dz$$

Lets use the following inequality $\arctan x \geq \frac{\pi}{2} - \frac{1}{x}$. Then $\left(\frac{1}{\pi} \arctan z + \frac{1}{2}\right) \frac{z^2}{1+z^2} \left(\frac{1}{2} - \frac{1}{\pi} \arctan z\right) > \left(1/4 - 1/4 + \frac{2}{\pi z} - \frac{1}{\pi^2 z^2}\right) \frac{z^2}{1+z^2} = \frac{2z}{\pi(1+z^2)} - \frac{1}{\pi^2(1+z^2)}$ and this holds for all $z \neq 0$. And $\int_{(-\infty, -a) \cup (a, \infty)} \frac{2z}{\pi(1+z^2)}$ does not converge, but $\int_{(-\infty, -a) \cup (a, \infty)} \frac{z^2}{1+z^2}$ converge. Therefore our integral is also does not converge:

$$\begin{aligned} & \int_{-\infty}^{\infty} \frac{z^2}{1+z^2} \left(\frac{1}{4} - \frac{1}{\pi^2} \arctan^2 z\right) dz = \\ &= \int_{-a}^a \frac{z^2}{1+z^2} \left(\frac{1}{4} - \frac{1}{\pi^2} \arctan^2 z\right) dz + \int_{(-\infty, -a) \cup (a, \infty)} \frac{z^2}{1+z^2} \left(\frac{1}{4} - \frac{1}{\pi^2} \arctan^2 z\right) dz > \\ &> \int_{-a}^a \frac{z^2}{1+z^2} \left(\frac{1}{4} - \frac{1}{\pi^2} \arctan^2 z\right) dz + \int_{(-\infty, -a) \cup (a, \infty)} \frac{2z}{\pi(1+z^2)} + \int_{(-\infty, -a) \cup (a, \infty)} \frac{z^2}{1+z^2} \end{aligned}$$

Therefore $V[Y]$ is not defined in this case and so the squared coefficient of variation.

Case $Z = \text{median}(X_1, X_2, X_3, X_4, X_5)$. By very same logic as for Y we can show that $E[Z] = x_0$. Then by using inequality $\arctan x \leq \frac{\pi}{2} - \frac{1}{x} + \frac{1}{3x^3}$ we can easily show that variance of Z converges. But it's exact value still remains unknown, so the squared coefficient of variation... But in this case it is at least defined.