

1 (a) Prove that $w_0 = \bar{Y} - w_1\bar{X}$:

Let

$$f'(g(w_0)) = \frac{\partial L}{\partial \sum_{i=1}^N [y^{(i)} - (w_0 + w_1 x^{(i)})]}$$

$$g'(w_0) = \frac{\partial \sum_{i=1}^N [y^{(i)} - (w_0 + w_1 x^{(i)})]}{\partial w_0}$$

then

$$\frac{\partial L}{\partial w_0} = f'(g(w_0)) \cdot g'(w_0)$$

$$f'(g(w_0)) = \sum_{i=1}^N [y^{(i)} - w_0 - w_1 x^{(i)}] = N\bar{Y} - Nw_0 - Nw_1\bar{X}$$

$$g'(w_0) = \sum_{i=1}^N -1 = -N$$

$$\frac{\partial L}{\partial w_0} = -N(N\bar{Y} - Nw_0 - Nw_1\bar{X}) = -N^2\bar{Y} + N^2w_0 + N^2w_1\bar{X}$$

$$0 = -N^2\bar{Y} + N^2w_0 + N^2w_1\bar{X}$$

$$w_0 = \bar{Y} - w_1\bar{X} \quad \square$$

Prove that

$$w_1 = \frac{\frac{1}{N} \sum_{i=1}^N x^{(i)} y^{(i)} - \bar{Y} \bar{X}}{\frac{1}{N} \sum_{i=1}^N (x^{(i)})^2 - \bar{X}^2}$$

By the chain rule,

$$\frac{\partial L}{\partial w_1} = \sum_{i=1}^N [-x^{(i)}(y^{(i)} - (w_0 + w_1 x^{(i)}))]$$

Setting the partial derivative to 0, we get

$$0 = - \sum_{i=1}^N x^{(i)} y^{(i)} + \sum_{i=1}^N w_0 x^{(i)} + \sum_{i=1}^N [w_1 (x^{(i)})^2]$$

Since $w_0 = \bar{Y} - w_1\bar{X}$

$$0 = - \sum_{i=1}^N x^{(i)} y^{(i)} + \sum_{i=1}^N [(\bar{Y} - w_1\bar{X}) x^{(i)}] + \sum_{i=1}^N [w_1 (x^{(i)})^2]$$

$$\begin{aligned}
&= - \sum_{i=1}^N x^{(i)} y^{(i)} + \bar{Y} \sum_{i=1}^N x^{(i)} - \bar{X} w_1 \sum_{i=1}^N x^{(i)} + \sum_{i=1}^N [w_1 (x^{(i)})^2] \\
&w_1 [\bar{X} \sum_{i=1}^N x^{(i)} - \sum_{i=1}^N (x^{(i)})^2] = - \sum_{i=1}^N x^{(i)} y^{(i)} + \bar{Y} \sum_{i=1}^N x^{(i)} \\
&w_1 = \frac{- \sum_{i=1}^N x^{(i)} y^{(i)} + \bar{Y} \sum_{i=1}^N x^{(i)}}{\bar{X} \sum_{i=1}^N x^{(i)} - \sum_{i=1}^N (x^{(i)})^2} \\
&w_1 = \frac{\sum_{i=1}^N x^{(i)} y^{(i)} - N \bar{Y} \bar{X}}{\sum_{i=1}^N (x^{(i)})^2 - N \bar{X}^2} \\
&w_1 = \frac{\frac{1}{N} \sum_{i=1}^N x^{(i)} y^{(i)} - \bar{Y} \bar{X}}{\frac{1}{N} \sum_{i=1}^N (x^{(i)})^2 - \bar{X}^2} \quad \square
\end{aligned}$$

1 (b) i. Let us first show that if $\lambda_i > 0$ for all i , then A must be PD.

For any $z \neq 0 \in \mathbb{R}^d$, $z^T A z = z^T (U \Lambda U^T) z$. Let $y = U^T z$. Then

$$z^T A z = y^T \Lambda y = y_1^2 \lambda_1 + y_2^2 \lambda_2 + \dots + y_d^2 \lambda_d$$

Since U is an orthogonal matrix, no row or column of U can consist entirely of zeros, since each row and column must have a norm of 1. The entries of y can be written as

$$y_1 = u_1^T z, y_2 = u_2^T z, \dots, y_d = u_d^T z$$

Since $z \neq 0$, then for at least one $i = \{1, 2, \dots, d\}$, $y_i \neq 0$. We have assumed that for all i $\lambda_i > 0$. In the expression

$$y_1^2 \lambda_1 + y_2^2 \lambda_2 + \dots + y_d^2 \lambda_d$$

each term will be 0 if $y_i = 0$ and greater than 0 if $y_i \neq 0$. So

$$z^T A z > 0 \quad \square$$

Now let us show that if A is PD then for all i $\lambda_i > 0$.

We have it that, for all values of i , $A u_i = \lambda_i u_i$, where u_i is a column of U . By multiplying both sides of the equation by u_i^T on the left, we get

$$u_i^T A u_i = u_i^T \lambda_i u_i$$

Because A is PD and $u_i \neq 0 \in \mathbb{R}^d$, $u_i^T A u_i > 0$. We can write $u_i^T \lambda_i u_i$ as

$$\lambda_i \sum_{j=1}^d u_{ji}^2$$

where u_{ji} is entry of U at the j th row and i th column. If, for any value of i , $\lambda_i = 0$, then

$$\lambda_i \sum_{j=1}^d u_{ji}^2 = 0 = u_i^T A u_i$$

which would contradict A being PD. Suppose instead that $\lambda_i < 0$. For all values of i and j , if $u_{ji} = 0$ then $u_{ji}^2 = 0$, and if $u_{ji} \neq 0$ then $u_{ji}^2 > 0$. Since the column vector u_i is orthonormal, for some value of j , $u_{ji}^2 > 0$. In this case,

$$\lambda_i \sum_{j=1}^d u_{ji}^2 = u_i^T A u_i < 0$$

which would also contradict A being PD. So it must be the case that if A is PD then, for all values of i ,

$$\lambda_i > 0 \quad \square$$

1 (b) ii. Let us start with the eigenvalues of $\Phi^T \Phi + \beta I$.

In effect, $\Phi^T \Phi + \beta I$ differs from $\Phi^T \Phi$ by having diagonal values shifted by β . So if, for $i = \{1, 2, \dots, d\}$, the eigenvalues of $\Phi^T \Phi$ are λ_i , then the eigenvalues of $\Phi^T \Phi + \beta I$ are $\lambda_i + \beta$. We can see this in the following way. Let $A = \Phi^T \Phi$ and $B = \Phi^T \Phi + \beta I$, and μ_i expresses the eigenvalues of B . Let M be the diagonal matrix $\text{diag}(\mu_i)$. Analogously to $AU = U\Lambda$,

$$BU = UM$$

Since $B = A + \beta I$,

$$(A + \beta I)U = UM$$

$$AU + \beta U = UM$$

Since $AU = U\Lambda$,

$$U\Lambda + \beta U = UM$$

$$U\beta = UM - U\Lambda = U(M - \Lambda)$$

$$\beta = M - \Lambda \quad \square$$

Therefore, the difference between the diagonal values of M and Λ is given by β , and the eigenvalues of B (that is, $\Phi^T \Phi + \beta I$) are given by $\lambda_i + \beta$.

Now let us show that A and B have the same eigenvectors. Let z be an eigenvector of B . Since the eigenvalues of B are given by $\lambda_i + \beta$, we can write

$$Bz = (\lambda_i + \beta)z = \lambda_i z + \beta z$$

$$Az + \beta Iz - \beta Iz = \lambda_i z$$

$$Az = \lambda_i z \quad \square$$

By the definition of eigenvectors and eigenvalues, this means that z is also an eigenvector of A . If u_i is an eigenvector of $\Phi^T \Phi$, it is also an eigenvector of $\Phi^T \Phi + \beta I$.

To see that $\Phi^T \Phi + \beta I$ is PD if $\beta > 0$, we can first show that $\Phi^T \Phi$ is PSD. In general, for any matrix X , $X^T X$ is PSD. For any vector $z \neq 0 \in \mathbb{R}^d$,

$$z^T (X^T X) z = (Xz)^T Xz = \|Xz\|_2^2 \geq 0$$

From the proof in **1 (b) i.** we can see that for a PSD matrix, for all values of i , $\lambda_i \geq 0$. In this case, if $\beta > 0$,

$$\lambda_i + \beta > 0 \quad \square$$

As we have seen, if all the eigenvalues of $\Phi^T \Phi + \beta I$ are positive, then $\Phi^T \Phi + \beta I$ is PD.

1 (c) We can write $\sum_{n=1}^N \log P(y^{(n)} | x^{(n)})$ as

$$\sum_{n=1}^N \{ \mathbb{I}(y^{(n)} = 1) \log P(y^n = 1 | x^{(n)}) + \mathbb{I}(y^{(n)} = -1) \log P(y^{(n)} = -1 | x^{(n)}) \}$$

For the probabilities of the class labels $\{-1, 1\}$, we can use the standard sigmoid function for logistic regression $\sigma(w^T x)$ and treat y as $y \in \{0, 1\}$:

$$\sum_{n=1}^N \{ y^{(n)} \log[\sigma(w^T \phi(x^{(n)}))] + (1 - y^{(n)}) \log[1 - \sigma(w^T \phi(x^{(n)}))] \}$$

and then set the partial derivative $f'(w)$ to 0, switching the expression's sign to make it a convex optimization problem:

$$0 = - \sum_{n=1}^N \{ y^{(n)} \frac{1}{\sigma[w^T \phi(x^{(n)})]} \sigma[w^T \phi(x^{(n)})] [1 - \sigma(w^T \phi(x^{(n)}))] \phi(x^{(n)}) \}$$

$$\begin{aligned}
& + (1 - y^{(n)}) \frac{1}{1 - \sigma[w^T \phi(x^{(n)})]} [-\sigma(w^T \phi(x^{(n)}))][1 - \sigma(w^T \phi(x^{(n)}))]\phi(x^{(n)})\} \\
0 & = - \sum_{n=1}^N \{y^{(n)}[1 - \sigma(w^T \phi(x^{(n)}))]\phi(x^{(n)}) \\
& \quad - (1 - y^{(n)})[\sigma(w^T \phi(x^{(n)}))]\phi(x^{(n)})\} \\
0 & = - \sum_{n=1}^N \{[y^{(n)} - \sigma(w^T \phi(x^{(n)}))]\phi(x^{(n)})\}
\end{aligned}$$

By setting the partial derivative of expression (4) with respect to w to 0 we get

$$\begin{aligned}
0 & = - \sum_{n=1}^N \left\{ \frac{\exp[-yw^T \phi(x^{(n)})]}{1 + \exp[-yw^T \phi(x^{(n)})]} y \phi(x^{(n)}) \right\} \\
0 & = - \sum_{n=1}^N \left\{ \frac{1}{\exp[yw^T \phi(x^{(n)})] + 1} y \phi(x^{(n)}) \right\}
\end{aligned}$$

If the derivative equation for the negative log likelihood above has $y = 0$, then the equation becomes

$$0 = - \sum_{n=1}^N \{[-\sigma(w^T \phi(x^{(n)}))]\phi(x^{(n)})\}$$

which is also what we get from the partial derivative equation for the loss function we derived from (4) if $y = -1$. So maximizing the partial derivative of the log-likelihood function with respect to w is equivalent to minimizing the partial derivative of loss function (4) with respect to w when $y = -1$. If $y = 1$, then the derivative equation for the negative log likelihood becomes

$$0 = - \sum_{n=1}^N \{[1 - \sigma(w^T \phi(x^{(n)}))]\phi(x^{(n)})\}$$

Since $1 - \sigma(w^T \phi(x^{(n)})) = \frac{\exp[-w^T \phi(x^{(n)})]}{1 + \exp[-w^T \phi(x^{(n)})]}$

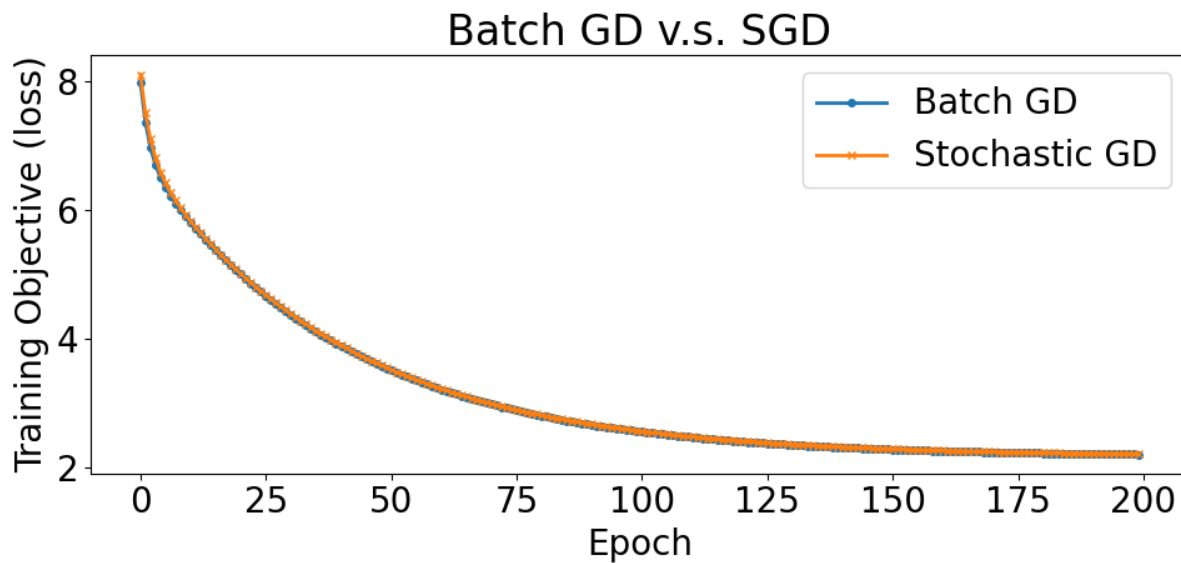
$$0 = - \sum_{n=1}^N \left\{ \frac{\exp[-w^T \phi(x^{(n)})]}{1 + \exp[-w^T \phi(x^{(n)})]} \phi(x^{(n)}) \right\}$$

We can use the same expression the partial derivative equation for the loss function in the same way if $y = 1$:

$$0 = - \sum_{n=1}^N \left\{ \frac{1}{\exp[w^T \phi(x^{(n)})] + 1} \phi(x^{(n)}) \right\} = - \sum_{n=1}^N \left\{ \frac{\exp[-w^T \phi(x^{(n)})]}{1 + \exp[-w^T \phi(x^{(n)})]} \phi(x^{(n)}) \right\}$$

Therefore, since w has the same value at the maximum of the log-likelihood function and the minimum of loss function (4) for all values of y , maximizing the log-likelihood is equivalent to minimizing this loss function.

2.1 (b)



GD took less time (0.00 seconds as opposed to 0.03 seconds for SGD, as measured by the time module) but SGD showed a lower test objective (2.6796 for SGD vs. 2.7017 for GD). SGD updates weights more frequently than GD. Hence it is not surprising that more frequent computation made SGD take longer. In this case, it also seems that the more frequent updates to the weights with SGD allowed SGD to minimize the loss on the training set more effectively over a set number of epochs.