

1 (a)

$$\begin{aligned}
\frac{\partial E(w)}{\partial w_j} &= \sum_{i=1}^N (x_j)^{(i)} y^{(i)} (1 - \sigma(w^T x^{(i)})) - (x_j)^{(i)} \sigma(w^T x^{(i)}) + (x_j)^{(i)} y^{(i)} \sigma(w^T x^{(i)}) \\
&= \sum_{i=1}^N (x_j)^{(i)} y^{(i)} - (x_j)^{(i)} \sigma(w^T x^{(i)}) \\
\frac{\partial^2 E(w)}{\partial (w_j)^2} &= - \sum_{i=1}^N (x_j)^{(i)} \sigma(w^T x^{(i)}) (1 - \sigma(w^T x^{(i)})) (x_j)^{(i)} \\
\frac{\partial^2 E(w)}{\partial w_j \partial w_k} &= - \sum_{i=1}^N (x_j)^{(i)} \sigma(w^T x^{(i)}) (1 - \sigma(w^T x^{(i)})) (x_k)^{(i)}
\end{aligned}$$

Let  $X \in \mathbb{R}^{n \times m}$  be the design matrix and  $w \in \mathbb{R}^m$  be the weight vector, where  $n$  is the number of observations and  $m$  is the number of features. We can express the second-order partial derivatives in matrix form as

$$X^T \text{diag}[-\sigma(Xw)(1 - \sigma(Xw))]X$$

which gives us our Hessian matrix.

1 (b) Let  $x, z \in \mathbb{R}^m$ . Then

$$(x^T z)(x^T z) = \sum_{i=1}^N z_i x_i \sum_{j=1}^N x_j z_j = \sum_{i=1}^N \sum_{j=1}^N z_i x_i x_j z_j = (x^T z)^2$$

Consider

$$z^T X^T \text{diag}[\sigma(Xw)(1 - \sigma(Xw))]Xz$$

Let  $D$  represent  $X^T \text{diag}[\sigma(Xw)(1 - \sigma(Xw))]X$ .  $D \in \mathbb{R}^{m \times m}$ . Let  $D_i$  be a column of  $D$  and  $D^{(j)}$  be a row of  $D$ . we can now express the above as

$$\begin{aligned}
& - \sum_{i=1}^m \left[ \sum_{j=1}^m z_j (D_i)^{(j)} \right] z_i \\
&= - \sum_{i=1}^m \left[ \sum_{j=1}^m z_j ((D_i)^{(j)})^{\frac{1}{2}} ((D_i)^{(j)})^{\frac{1}{2}} \right] z_i = - \sum_{i=1}^m \sum_{j=1}^m z_j ((D_i)^{(j)})^{\frac{1}{2}} ((D_i)^{(j)})^{\frac{1}{2}} z_i
\end{aligned}$$

Since in general  $\sum_{i=1}^N \sum_{j=1}^N z_i x_i x_j z_j = (x^T z)^2 \geq 0$ ,

$$\sum_{i=1}^m \sum_{j=1}^m z_j ((D_i)^{(j)})^{\frac{1}{2}} ((D_i)^{(j)})^{\frac{1}{2}} z_i \geq 0$$

therefore

$$- \sum_{i=1}^m [\sum_{j=1}^m z_j (D_i)^{(j)}] z_i \leq 0 \quad \square$$

The Hessian matrix is negative semi-definite, and our original log-likelihood function is concave.

**1 (c)** We can invert the sign of the log-likelihood function to get a loss function. This gives us the Hessian matrix  $X^T \text{diag}[\sigma(Xw)(1 - \sigma(Xw))]X$ . For the gradient of the loss function with respect to  $w$ , we can take the partial derivative expression from **1 (a)**

$$\sum_{i=1}^N (x_j)^{(i)} y^{(i)} - (x_j)^{(i)} \sigma(w^T x^{(i)})$$

invert the sign, and express it in matrix form:

$$\nabla_w E = X^T \sigma(Xw) - X^T y$$

So we the update for Newton's method is

$$w_{new} = w_{old} - H^{-1} \nabla_w E$$

$$w_{new} = w_{old} - \{X^T \text{diag}[\sigma(Xw)(1 - \sigma(Xw))]X\}^{-1} [X^T \sigma(Xw) - X^T y]$$