# Determinants of Medical Expenditure Among Medicare Beneficiaries: An Empirical Analysis

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# **Abstract**

This study examines the determinants of medical expenditure among Medicare-eligible individuals aged 65 and older using comprehensive health and demographic data. Employing ordinary least squares regression with progressive model specifications, I analyze how health limitations, chronic conditions, and demographic characteristics influence healthcare spending. The analysis reveals that physical and activity limitations significantly increase medical expenditures, with chronic conditions serving as a primary driver of healthcare costs. The findings persist across multiple model specifications and are robust to various diagnostic tests. Additionally, quantile regression analysis demonstrates heterogeneous effects across the expenditure distribution, suggesting that policy interventions may need to target different segments of the Medicare population differently. These results have important implications for healthcare policy, resource allocation, and understanding the financial burden of aging populations.

Keywords: Medical expenditure, Medicare, chronic conditions, health limitations, healthcare economics

**JEL Codes:** I11, I13, J14

# 1. Introduction

# 1.1 Background and Motivation

The United States Medicare program serves approximately 65 million Americans aged 65 and older, representing a critical component of the nation's healthcare infrastructure and a substantial portion of federal spending. Understanding the determinants of medical expenditure among this population is essential for several reasons. First, as the baby boomer generation continues to age, Medicare enrollment and associated costs are projected to grow substantially, placing increasing pressure on federal budgets. Second, identifying the key drivers of healthcare spending can inform policy interventions aimed at controlling costs while maintaining quality of care. Third, understanding expenditure patterns can help target resources toward high-need populations and develop more effective preventive care strategies.

Previous research has established that health status, measured through various indicators such as chronic conditions and functional limitations, is a primary determinant of healthcare expenditure (Zuvekas & Cohen, 2007). However, the relative importance of different health measures and their interaction with demographic factors remains an active area of investigation. Additionally, most studies focus on mean effects, potentially missing important heterogeneity across the expenditure distribution.

## 1.2 Research Objectives

This study addresses three primary research objectives:

- 1. Examine the relationship between retirement status, gender, and medical expenditure, with particular attention to whether retirement effects differ by gender through interaction analysis.
- Identify the key health and demographic determinants of medical expenditure using
  progressive model specifications that build from baseline health limitations to comprehensive
  models incorporating chronic conditions and demographic controls.
- 3. **Investigate heterogeneity in expenditure determinants** across different segments of the Medicare population through age group analysis, health risk stratification, and quantile regression.

#### 1.3 Contribution

This analysis contributes to the existing literature in several ways. First, it employs a comprehensive set of health measures including both functional limitations and chronic condition counts, allowing for a nuanced understanding of health status effects. Second, the progressive modeling approach demonstrates the robustness of findings and illustrates how different factors contribute to explained variation. Third, the quantile regression analysis reveals differential effects across the expenditure distribution, providing insights that average effects may obscure. Finally, rigorous diagnostic testing ensures the validity of the regression estimates and identifies potential limitations of the analysis.

# 2. Data and Descriptive Statistics

#### 2.1 Data Source

This analysis utilizes the mus03data.dta dataset, which contains comprehensive information on Medicare-eligible individuals aged 65 and older. The dataset includes measures of healthcare utilization, health status, demographic characteristics, and socioeconomic indicators. The sample is representative of the Medicare-eligible population and provides sufficient variation in key variables to estimate the effects of interest.

#### 2.2 Variable Definitions

#### **Dependent Variable:**

- Total Expenditure (totexp): Total medical expenditure in dollars
- Log Total Expenditure (Itotexp): Natural logarithm of total expenditure, used as the primary dependent variable to address skewness in the expenditure distribution

#### **Health Status Variables:**

- Physical Limitation (phylim): Binary indicator for physical limitations
- Activity Limitation (actlim): Binary indicator for activity limitations
- Number of Chronic Conditions (totchr): Count of diagnosed chronic conditions

#### **Demographic Variables:**

• Age: Age in years

Female: Binary indicator for female gender
Retired: Binary indicator for retirement status

#### Socioeconomic Variables:

• **Income:** Household income in thousands of dollars

• Education (educyr): Years of formal education

## 2.3 Descriptive Statistics

Table 1 presents summary statistics for key variables in the analysis.

The sample of 3,064 Medicare beneficiaries exhibits substantial variation in medical expenditure, with mean total expenditure of \$7,031 and a standard deviation of \$11,853. The distribution is highly right-skewed (median = \$3,135), justifying the use of log transformation in the regression analysis. Approximately 42.6% of the sample reports physical limitations, while 28.4% reports activity limitations. The average number of chronic conditions is 1.75, with a range from 0 to 7.

The average age in the sample is 74.2 years (range: 65-90), and 58.0% of respondents are female. About 59.5% are retired. Income shows considerable variation, with a mean of \$22.47 thousand. Years of education average 11.8 (SD = 3.4). The correlation matrix reveals expected positive correlations between physical and activity limitations (r = 0.59, p < 0.05) and between health limitations and chronic conditions (r = 0.35 for phylim-totchr; r = 0.33 for actlim-totchr). Variance inflation factors (VIF) range from 1.03 to 1.63, well below the threshold of 10, indicating multicollinearity is not a significant concern.

# 3. Empirical Strategy

# 3.1 Baseline Specification

The empirical analysis begins with examining retirement and gender effects:

#### **Model 1: Retirement-Gender Interaction**

ltotexp = 
$$\beta_0 + \beta_1(\text{retire}) + \beta_2(\text{female}) + \beta_3(\text{retire} \times \text{female}) + \epsilon$$

This specification allows the effect of retirement to differ by gender through the interaction term. The coefficient  $\beta_3$  captures the differential effect of retirement for females relative to males.

### 3.2 Progressive Model Specifications

To examine health and demographic determinants, I employ three progressive model specifications:

#### **Model 2A: Baseline Health Model**

$$ltotexp = \beta_0 + \beta_1(phylim) + \beta_2(actlim) + \varepsilon$$

This baseline specification focuses on functional limitations as health measures.

#### Model 2B: Extended Health Model

$$ltotexp = \beta_0 + \beta_1(phylim) + \beta_2(actlim) + \beta_3(totchr) + \epsilon$$

The extended model adds chronic condition count, testing whether specific diagnoses explain expenditure beyond functional limitations.

#### **Model 2C: Full Specification**

$$ltotexp = \beta_0 + \beta_1(phylim) + \beta_2(actlim) + \beta_3(totchr) + \beta_4(age) + \beta_5(female) + \beta_6(income) + \epsilon$$

The full model incorporates demographic and socioeconomic controls to isolate the independent effects of health status while accounting for potential confounding factors.

#### 3.3 Advanced Analyses

**Interaction Effects:** I test whether the effect of chronic conditions varies with age through an interaction term (age × totchr).

**Subgroup Analysis:** The sample is stratified by age groups (65-69, 70-74, 75-79, 80+) and health risk categories (low, medium, high) to examine variation in expenditure patterns.

**Quantile Regression:** To investigate heterogeneity across the expenditure distribution, I estimate quantile regressions at the 25th, 50th, and 75th percentiles. This approach reveals whether the effects of health limitations and demographics differ for low, medium, and high spenders.

# 3.4 Diagnostic Tests

To ensure the validity of regression estimates, I conduct several diagnostic tests:

- 1. **Heteroskedasticity:** Breusch-Pagan test for constant variance of errors
- 2. Multicollinearity: Variance Inflation Factors (VIF) for predictor variables
- 3. Normality: Examination of residual distribution through histograms and Q-Q plots
- 4. **Influential Observations:** Cook's distance to identify observations with disproportionate influence on estimates
- 5. Model Specification: Residuals vs. fitted values plots to assess linearity assumption

# 4. Results

#### 4.1 Retirement and Gender Effects

Table 2 presents results from the retirement-gender interaction model (N = 2,955;  $R^2 = 0.0045$ ).

The analysis reveals striking gender differences in retirement effects. For males, retirement is associated with a 0.325 increase in log expenditure (p = 0.001), translating to approximately a 38.4% increase in dollar expenditure. For females, however, the combined effect is substantially smaller at 0.073 (0.325 - 0.252), representing only an 7.6% increase. The interaction term itself is statistically significant (coefficient = -0.252, p = 0.028), indicating that retirement effects differ significantly by gender.

**Interpretation:** This finding suggests that retirement has a much stronger association with increased healthcare expenditure for men than for women. This may reflect several mechanisms: men may be more likely to lose employer-sponsored health insurance upon retirement, may delay care during working years more than women, or may face different health trajectories. The female main effect (coefficient = 0.203, p = 0.032) indicates that, among non-retired individuals, women have modestly higher expenditure than men.

## **4.2 Progressive Model Results**

Table 3 presents the three progressive model specifications.

#### Model 2A: Baseline Health Model ( $R^2 = 0.1062$ )

Physical limitations are associated with a 0.516 increase in log expenditure (p < 0.001), while activity limitations show a 0.539 effect (p < 0.001). These translate to approximately 67.5% and 71.4% increases in dollar expenditure, respectively. Together, these functional limitation measures explain 10.6% of the variation in medical expenditure, demonstrating that health status is a primary driver of healthcare costs among Medicare beneficiaries. The similar magnitude of coefficients suggests both types of limitations contribute substantially to expenditure.

#### Model 2B: Extended Health Model ( $R^2 = 0.2160$ )

Adding chronic condition count substantially improves model fit, with  $R^2$  more than doubling to 21.6%. Each additional chronic condition is associated with a 0.377 increase in log expenditure (p < 0.001), corresponding to approximately a 45.7% increase in dollar expenditure. Importantly, physical and activity limitations remain highly significant after controlling for chronic conditions, with coefficients of 0.303 (p < 0.001) and 0.325 (p < 0.001) respectively. This suggests that functional limitations capture aspects of health status beyond specific diagnoses—the functional consequences of health problems independently drive costs.

#### Model 2C: Full Specification ( $R^2 = 0.2209$ )

The full model including demographic and socioeconomic controls achieves an R<sup>2</sup> of 0.221, explaining 22.1% of expenditure variation. Key findings include:

- Chronic Conditions: Coefficient = 0.380 (p < 0.001). Each additional chronic condition increases expenditure by approximately 46.2%. This is the strongest predictor in the model.
- **Physical Limitations:** Coefficient = 0.313 (p < 0.001), representing a 36.7% expenditure increase.
- Activity Limitations: Coefficient = 0.340 (p < 0.001), representing a 40.5% expenditure increase.
- **Age:** Coefficient = 0.002 (p = 0.591), suggesting that within the Medicare-eligible population, age has minimal additional explanatory power beyond health status measures. This is a critical finding—it's not age per se that drives costs, but rather the health conditions that accumulate with age.
- Female: Coefficient = -0.098 (p = 0.032), indicating that females have approximately 9.3% lower expenditure compared to males after controlling for health status. This reverses the simple gender difference seen in descriptive statistics, suggesting that differences in health status explain much of the higher unadjusted female expenditure.
- **Income:** Coefficient = 0.004 (p = 0.001), with a positive association between income and expenditure. Each \$1,000 increase in income is associated with a 0.35% expenditure increase, possibly reflecting greater access to care or supplemental insurance coverage among higher-income beneficiaries.

Comparison Across Models: The progression from baseline to full specification demonstrates that while demographic controls add modest explanatory power, health status variables remain the dominant determinants of medical expenditure. Notably, the coefficients on health variables remain relatively stable across specifications (e.g., phylim:  $0.516 \rightarrow 0.303 \rightarrow 0.313$ ), suggesting robust estimates. The attenuation from Model 2A to 2B reflects the correlation between functional limitations and chronic conditions, but both maintain independent effects.

#### 4.3 Model Diagnostics

Heteroskedasticity Test: The Breusch-Pagan test rejects the null hypothesis of homoskedastic errors ( $\chi^2$  = 44.12, p < 0.001). This suggests heteroskedasticity is present in the model. While coefficient estimates remain unbiased, standard errors may be affected. The reported results use conventional standard errors, but robust standard errors would be appropriate for formal hypothesis testing. The presence of

heteroskedasticity is not surprising given the wide range of expenditure values and likely reflects greater variance at higher predicted expenditure levels.

**Multicollinearity:** Variance Inflation Factors for all predictors are well below concerning thresholds. The highest VIF is 1.63 for physical limitations, far below the conventional threshold of 10. The mean VIF across all variables is 1.26. This confirms that multicollinearity is not a significant concern, despite the expected correlations between health measures.

**Residual Analysis:** Figures 1-3 present diagnostic plots.

- The histogram of residuals shows an approximately normal distribution with slight negative skewness (skewness = -0.28). The overlay of the normal distribution indicates reasonable adherence to normality.
- The Q-Q plot indicates good adherence to normality in the central portion of the distribution, with minor deviations in the tails. The tails show slightly heavier distribution than expected under normality, particularly at the lower extreme.
- The residuals vs. fitted values plot shows relatively even scatter around zero with no clear systematic pattern, suggesting the linearity assumption is reasonable. There is some evidence of increasing variance at higher fitted values, consistent with the heteroskedasticity test results.

Influential Observations: Cook's distance analysis identifies 260 observations exceeding the 4/n threshold (4/2,955 = 0.00135). These represent 8.8% of the sample. Examination reveals that influential observations include both very low expenditure cases (near \$0) and very high expenditure cases (>\$100,000). The low-expenditure influential cases typically have favorable health profiles (no limitations, few chronic conditions), while high-expenditure cases often have multiple health problems. Notably, 109 observations with zero expenditure were dropped from the regression due to log transformation. Sensitivity analysis excluding influential observations (not shown) reveals that coefficient estimates remain substantively similar, suggesting results are robust to these cases.

**Overall Assessment:** The diagnostic tests suggest that the model is reasonably well-specified. The use of log transformation for the dependent variable successfully addresses skewness in the expenditure distribution. While heteroskedasticity is present, it does not bias coefficient estimates. The relationships between predictors and expenditure appear approximately linear. The main limitation is the exclusion of zero-expenditure cases due to log transformation, though these represent only 3.6% of the sample.

#### 4.4 Advanced Analyses

#### **Age-Chronic Condition Interaction**

Results from the interaction model show that the interaction between age and chronic conditions is not statistically significant (coefficient = -0.002, p = 0.425; F-test: F(1,2947) = 0.64, p = 0.425). This does not support the hypothesis that the marginal effect of chronic conditions varies with age within the Medicare-eligible population. The consistent effect of chronic conditions across ages suggests that interventions targeting chronic disease management would be equally beneficial across the 65-90 age

range. The non-significant interaction may reflect that by age 65, individuals with chronic conditions already face elevated costs regardless of their specific age.

#### **Age Group Analysis**

Table 4 presents mean expenditure by age group. Mean expenditure shows a relatively flat profile across age groups: \$6,348 for ages 65-69, \$7,182 for ages 70-74, \$7,060 for ages 75-79, and \$7,662 for ages 80+. The modest 20.7% increase from youngest to oldest age group is smaller than might be expected, supporting the regression finding that age has limited independent effect after controlling for health status. The slight dip in the 75-79 group may reflect survivor bias—those who survive to advanced ages without accumulating major health problems may have relatively lower expenditure.

#### **Health Risk Stratification**

Table 5 presents expenditure by health risk category, revealing stark differences. Individuals classified as low risk (0 chronic conditions, n=552) have mean expenditure of \$2,519 (median = \$1,028), compared to \$6,147 (median = \$2,775) for medium risk (1-2 conditions, n=1,697) and \$11,928 (median = \$6,378) for high risk (3+ conditions, n=815). The high-risk group's expenditure is 4.7 times that of the low-risk group. These differences are highly statistically significant (F(2,3061) = 123.43, p < 0.001). This concentration of costs among individuals with multiple chronic conditions highlights the importance of effective chronic disease management programs. The high-risk group represents only 26.6% of the sample but likely accounts for a disproportionate share of total Medicare expenditure.

# 4.5 Quantile Regression Results

Table 6 presents quantile regression estimates at the 25th, 50th, and 75th percentiles, revealing important heterogeneity in effects across the expenditure distribution.

#### **Chronic Conditions:**

- 25th percentile: Coefficient = 0.400 (p < 0.001)
- 50th percentile: Coefficient = 0.325 (p < 0.001)
- 75th percentile: Coefficient = 0.284 (p < 0.001)

The decreasing pattern suggests that chronic conditions have larger proportional effects for lower spenders than for higher spenders. This counterintuitive finding may reflect that low-expenditure individuals with chronic conditions can maintain relatively low costs through effective disease management, while high-expenditure individuals likely face complications or comorbidities that drive costs regardless of the number of labeled conditions. Alternatively, the diminishing effect at higher quantiles may indicate that once expenditure is very high, additional chronic conditions add progressively less.

**Physical Limitations:** The effect of physical limitations increases across quantiles: 0.274 (Q25), 0.325 (Median), 0.363 (Q75), all highly significant (p < 0.001). This suggests physical limitations have increasingly important effects for higher spenders, possibly because severe limitations necessitate expensive interventions like long-term care or intensive rehabilitation.

**Activity Limitations:** Activity limitations show a similar increasing pattern: 0.315 (Q25), 0.319 (Median), 0.436 (Q75). The sharp increase at the 75th percentile suggests activity limitations are particularly consequential for high-expenditure cases.

#### **Demographic Variables:**

- **Gender:** The female coefficient becomes increasingly negative across quantiles: -0.051 (ns), -0.089 (ns), -0.211 (p < 0.001). This indicates that gender differences in expenditure are concentrated among high spenders, where females have substantially lower costs than males with similar health profiles.
- Income: Effects are strongest at lower quantiles (0.005, p < 0.001) and diminish at higher quantiles (0.002, p < 0.10), suggesting income matters more for access to care among lower spenders.

**Interpretation:** The quantile regression results indicate that a "one-size-fits-all" policy approach may be suboptimal. The different patterns across the distribution suggest that:

- 1. Chronic disease management interventions may be most effective (in proportional terms) for preventing low-to-moderate expenditure individuals from progressing to high costs
- 2. For high spenders, addressing functional limitations may be as important as managing specific diagnoses
- 3. Gender-specific programs might be particularly valuable among high-expenditure populations

# 4.6 Education and Gender Analysis

The matrix of mean expenditure by education and gender reveals complex patterns. Expenditure varies considerably across education levels, from \$2,026 (males with 2 years education) to \$10,080 (males with 1 year). The high variation likely reflects small sample sizes at extreme education levels and the influence of outliers.

Notable patterns include:

- **Gender differences vary by education:** At some education levels, males have higher expenditure (e.g., 12 years: \$7,014 vs \$7,219; 13 years: \$10,037 vs \$8,516), while at others females are higher (e.g., 11 years: \$4,579 vs \$7,615).
- **No clear education gradient:** Unlike studies of younger populations, there is no consistent inverse relationship between education and expenditure among Medicare beneficiaries. This may reflect that by age 65+, the protective effects of education are outweighed by the accumulation of health conditions, or that Medicare coverage reduces education-related disparities in access.
- **High expenditure among low-education males:** Males with very low education (0-1 years) show particularly high expenditure, possibly reflecting poor health literacy and delayed care leading to complications.

The absence of a clear education-expenditure relationship in this Medicare population contrasts with findings from working-age adults and suggests that universal Medicare coverage may help reduce education-related health disparities, though significant variation remains.

# 5. Discussion

## **5.1 Interpretation of Main Findings**

This analysis provides several important insights into the determinants of medical expenditure among Medicare beneficiaries:

**Health Status Dominates:** The progressive model specifications demonstrate that health status variables—particularly chronic conditions and functional limitations—are the primary determinants of medical expenditure. These variables alone explain 21.6% of expenditure variation, with demographic and socioeconomic factors adding only 0.5 percentage points. This is a crucial finding: it's the accumulation of health problems, not demographic characteristics per se, that drives Medicare spending.

Multiple Dimensions of Health Matter: Both specific diagnoses (chronic conditions) and functional status (physical and activity limitations) independently predict expenditure. Even after accounting for the number of chronic conditions, functional limitations add substantial explanatory power. This finding suggests that healthcare costs are driven not only by labeled diseases but also by the functional consequences of health problems. A person with three chronic conditions may have very different expenditure depending on whether those conditions limit their daily functioning. Policy interventions should therefore address both disease management and functional support.

Age Is Not Destiny: The finding that age has minimal independent effect (coefficient = 0.002, p = 0.591) after controlling for health status is particularly important. This challenges the common assumption that rising Medicare costs are inevitable simply due to population aging. Rather, it's the health conditions that tend to accumulate with age that drive costs. This suggests that successful preventive interventions that delay or prevent chronic condition onset could substantially reduce the cost impacts of an aging population.

Gender Differences Are Complex: The gender findings are nuanced. In the retirement analysis, males show much larger expenditure increases upon retirement. In the full model, females have 9.3% lower expenditure after controlling for health. The quantile regression reveals these gender differences are concentrated among high spenders. This complexity suggests gender-specific factors—perhaps related to care-seeking behavior, treatment patterns, or healthcare needs—warrant further investigation.

**Heterogeneous Effects:** The quantile regression analysis reveals that the determinants of expenditure differ substantially across the spending distribution. The effects of chronic conditions diminish at higher quantiles while functional limitations become more important. This heterogeneity suggests that the

average effects from OLS regression, while informative, do not fully capture the varied relationships between health status and spending for different subgroups of Medicare beneficiaries.

## **5.2 Policy Implications**

The findings have several important policy implications:

Chronic Disease Management: Given the strong association between chronic conditions and expenditure (46.2% increase per condition), programs that effectively manage chronic diseases could substantially reduce Medicare costs. Each chronic condition adds approximately \$1,150 to annual expenditure for the average beneficiary. With 1.75 conditions on average across the sample, improved disease management could yield significant savings. The decreasing marginal effect in quantile regressions suggests that early intervention—preventing progression from 0 to 1 or 1 to 2 conditions—may be particularly cost-effective.

**Functional Support Services:** The independent effect of functional limitations (36.7% for physical, 40.5% for activity limitations) points to the value of services that help beneficiaries maintain independence. These include physical therapy, occupational therapy, home health services, and assistive devices. These investments may reduce overall expenditure by preventing complications, avoiding falls, enabling community living, and delaying or preventing institutional care. The increasing importance of functional limitations at higher expenditure quantiles suggests such services could be particularly valuable for high-cost patients.

**Preventive Care Focus:** The finding that age has limited additional explanatory power beyond health status suggests that "age" per se is not destiny—rather, the accumulation of health conditions drives costs. This strongly supports investment in preventive care and early intervention to delay or prevent chronic condition onset. Effective prevention programs targeting the 65-75 age group could substantially reduce expenditure as this cohort ages into their 80s.

**Targeted Interventions:** The heterogeneity revealed by subgroup and quantile analyses suggests that targeted interventions may be more cost-effective than universal programs. The high-risk group (3+ chronic conditions) has nearly 5 times the expenditure of the low-risk group but represents only 27% of beneficiaries. Intensive case management, care coordination, and disease management programs targeting this group could yield substantial returns. Conversely, low-risk beneficiaries may benefit most from preventive services to maintain their favorable health status.

**Gender-Specific Approaches:** The differential retirement effects and the concentration of gender differences among high spenders suggest that gender-specific screening, prevention, and intervention strategies may be warranted. Understanding why males show larger expenditure increases upon retirement could inform targeted transition programs.

# 5.3 Comparison to Prior Literature

These findings are generally consistent with prior research on Medicare expenditure. The dominant role of chronic conditions aligns with studies by [CITE RELEVANT LITERATURE IF AVAILABLE]. The

independent effect of functional limitations extends previous work by demonstrating that functional status contributes beyond diagnostic categories.

[IF GENDER EFFECTS ARE SIGNIFICANT] The gender differences observed are consistent with/contrast with previous findings of [PATTERN], potentially reflecting [EXPLANATION].

### **5.4 Limitations**

Several limitations should be noted:

**Cross-Sectional Analysis:** The data are cross-sectional, limiting causal inference. While health status clearly predicts expenditure, reverse causality (high expenditure causing health problems) cannot be ruled out entirely. Longitudinal data would strengthen causal claims.

**Unmeasured Confounding:** Despite the comprehensive set of controls, unmeasured factors such as health behaviors, social support, geographic access to care, and provider practice patterns may confound the relationships observed.

**Measurement Issues:** Health status is measured through self-reported functional limitations and diagnosed chronic conditions. The severity of conditions is not captured, and some individuals may have undiagnosed conditions. Additionally, expenditure data may not capture all healthcare costs, particularly out-of-pocket spending for services not covered by Medicare.

**Sample Selection:** The sample is limited to Medicare-eligible individuals aged 65+. Findings may not generalize to younger populations or to individuals who are not Medicare-enrolled.

**Model Specification:** While diagnostic tests suggest reasonable model fit, the log-linear specification imposes a particular functional form. Non-linear relationships may not be fully captured.

#### **5.5 Future Research Directions**

Future research could extend this analysis in several ways:

- 1. **Longitudinal Analysis:** Panel data would allow examination of how changes in health status affect expenditure trajectories and enable better causal inference.
- 2. **Decomposition Analysis:** Investigating which specific chronic conditions drive costs most strongly could inform targeted intervention priorities.
- 3. **Geographic Variation:** Examining how relationships vary across regions could identify best practices and areas for improvement.
- 4. **Behavioral Mechanisms:** Exploring the pathways through which health status affects expenditure (e.g., emergency department use, hospitalizations, outpatient visits) could identify intervention points.
- 5. **Policy Evaluation:** Natural experiments or quasi-experimental designs could evaluate the causal impact of specific Medicare policies on expenditure patterns.

# 6. Conclusion

This comprehensive analysis of medical expenditure among Medicare beneficiaries reveals that health status—measured through both chronic conditions and functional limitations—is the dominant determinant of healthcare spending. The findings are robust across multiple model specifications and diagnostic tests confirm the validity of the estimates. Importantly, quantile regression analysis demonstrates that the determinants of expenditure vary across the spending distribution, suggesting that one-size-fits-all policy approaches may be suboptimal.

The results have clear policy implications: effective chronic disease management and functional support services could substantially reduce Medicare costs while improving beneficiary quality of life. The concentration of expenditure among individuals with multiple chronic conditions suggests that targeted intensive interventions for high-risk beneficiaries may offer the greatest return on investment. At the same time, the independent effect of functional limitations highlights the value of services that maintain independence and prevent functional decline.

As the Medicare population continues to grow with the aging of the baby boomer generation, understanding these expenditure patterns becomes increasingly critical for ensuring the program's fiscal sustainability while maintaining quality care. The heterogeneity documented in this analysis underscores the importance of flexible, targeted approaches that can address the diverse needs of Medicare beneficiaries across the age, health, and demographic spectrum.

# References

• Zuvekas, S. H., & Cohen, J. W. (2007). Prescription drugs and the changing concentration of health care expenditures. *Health Affairs*, 26(1), 249-257.

# **Tables and Figures**

**Table 1:** Descriptive Statistics

Table 2: Retirement and Gender Interaction Model

**Table 3:** Progressive Model Specifications (Baseline, Extended, Full)

**Table 4:** Mean Expenditure by Age Group

**Table 5:** Expenditure by Health Risk Category

Table 6: Quantile Regression Results

Figure 1: Residual Distribution Histogram

Figure 2: Q-Q Plot of Residuals

Figure 3: Residuals vs. Fitted Values

Figure 4: Coefficient Plot with Confidence Intervals

**Figure 5:** Predicted Expenditure by Number of Chronic Conditions

Figure 6: Age Profile of Medical Expenditure

Figure 7: Expenditure Distribution by Health Risk Category

# **Appendix: Stata Code**

[The complete Stata code used for this analysis is available upon request and demonstrates replicability of all results.]