

# Confidence Intervals for Binomial Proportions

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In []: import scipy.stats.distributions as dist
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Consider a discrete random variable  $X$  which indicates either “success” ( $X = 1$ ) or “failure” ( $X = 0$ ) as the outcome of a random experiment. Such an experiment is called a *Bernoulli trial*. The probability of success is  $p = P\{X = 1\}$ , and the probability of failure is  $P\{X = 0\} = 1 - p$ . Repeating a Bernoulli trial  $n$  times means drawing a sample of  $n$  independent and identically distributed random variables  $X_i$ . The probability mass function of observing  $k$  success is the binomial distribution

$$\binom{n}{k} p^k (1 - p)^{n-k}.$$

The expected number of successes is  $np$  with variance  $np(1 - p)$ . Hence, the success probability  $p$  is the expected *proportion* of successes  $\frac{k}{n}$ , with variance  $p(1 - p)/n$ .

Let  $\hat{p} := \frac{k}{n}$  denote the sample proportion which is the unbiased (and maximum likelihood) estimator for the success probability  $p$ , and let  $\hat{\sigma} := \hat{p}(1 - \hat{p})/n$  denote the sample variance. Then the normal  $1 - \alpha$  confidence interval for the binomial proportion  $\hat{p}$  is

$$\hat{p} \pm z_{\alpha/2} \hat{\sigma},$$

where  $z_{\alpha/2}$  is the  $1 - \frac{\alpha}{2}$  quantile of the standard normal distribution. This normal confidence interval is also called the *Wald confidence interval*.

As Cameron [1] puts it, this normal approximation “suffers a *systematic* decline in performance both for small  $n$  and towards extreme values of  $p$  near 0 and 1, generating binomial [confidence intervals] with effective coverage far below the desired level.” (see also [2] and [3])

A different approach to quantifying uncertainty is Bayesian inference. The normal (frequentist)  $1 - \alpha$  confidence interval derives from a procedure that produces  $1 - \alpha$  confidence intervals that contain the true parameter value  $100(1 - \alpha)\%$  of the times. The  $1 - \alpha$  credible interval of Bayesian inference is the interval in which the parameter lies with probability  $1 - \alpha$ . [4]

Specifically, Bayesian inference employs Bayes’ theorem

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}.$$

Associate  $A$  with the parameter and  $B$  with the outcome from an experiment (the data). Then  $P(A)$  is the *prior* probability of the parameter event  $A$ , with the *likelihood*  $P(B|A)$  of the outcome event  $B$  given the parameter event  $A$ . The *posterior*  $P(A|B)$  is the probability of the parameter event  $A$  given the outcome event  $B$ .

For probability density functions, this reads

$$f(\theta|x) = \frac{f(x|\theta)f(\theta)}{\int d\theta f(\theta|x)f(\theta)}$$

with parameter  $\theta$  and data  $x$ . A Bayesian interval estimate is the  $1 - \alpha$  *posterior interval* or *credible interval*  $(l, u)$  with  $\int_{-\infty}^l d\theta f(\theta|x) = \int_u^{+\infty} d\theta f(\theta|x) = \alpha/2$  such that  $P(\theta \in (l, u)|x) = 1 - \alpha$ .

For  $n$  independent Bernoulli trials with common success probability  $p$ , the *likelihood* to have  $k$  successes given  $p$  is the binomial distribution

$$P(k|p) = \binom{n}{k} p^k (1 - p)^{n-k} \equiv B(a, b),$$

where  $B(a, b)$  is the *Beta distribution* with parameters  $a = k + 1$  and  $b = n - k + 1$ . Assuming a uniform prior  $P(p) = 1$ , the *posterior* is [4]

$$P(p|k) = P(k|p) = B(a, b).$$

A point estimate is the posterior mean

$$\bar{p} = \frac{k + 1}{n + 2}$$

with  $1 - \alpha$  credible interval  $(p_l, p_u)$  given by

$$\int_0^{p_l} dp B(a, b) = \int_{p_u}^1 dp B(a, b) = \frac{\alpha}{2}.$$

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## References

- [1] E. Cameron, [Publications of the Astronomical Society of Australia](#) **28**, 128 (2011).
- [2] A. Agresti and B. A. Coull, [The American Statistician](#) **52**, 119 (1998).
- [3] A. DasGupta, T. T. Cai, and L. D. Brown, [Statistical Science](#) **16**, 101 (2001).
- [4] L. Wasserman, [All of Statistics](#) (Springer New York, 2004).