Solutions to problem set 4, 3 March, 2024

Definition 1 A random variable has normal distribution with mean m and variance σ^2 if its density function is defined by

$$f(x) = \frac{1}{\sqrt{2\pi} \cdot \sigma} \cdot \exp\left(-\frac{(x-m)^2}{2\sigma^2}\right).$$

This is denoted by $X \sim N(m, \sigma^2)$

A random variable Z has standard normal distribution if it has normal distribution with mean m=0 and variance $\sigma=1$, that is, $Z \sim N(0,1)$. Its density function is given by

$$\varphi(x) = \frac{1}{\sqrt{2\pi}} \cdot \exp\left(-\frac{x^2}{2}\right).$$

That is, for every a < b we have

$$\mathbb{P}(a \le X \le b) = \int_a^b f(x) \, dx = \int_a^b \frac{1}{\sqrt{2\pi} \cdot \sigma} \cdot \exp\left(-\frac{(x-m)^2}{2\sigma^2}\right) dx.$$

The cumulative distribution function of the standard normal distribution is denoted by Φ :

$$\Phi(t) = \mathbb{P}(Z \le t) = \int_{-\infty}^{t} \frac{1}{\sqrt{2\pi}} \cdot \exp\left(-\frac{x^2}{2}\right) dx.$$

We can use that $\Phi(-x) = 1 - \Phi(x)$, and $\Phi(0.5) = 0.6915$, $\Phi(1) = 0.8413$, $\Phi(2) = 0.9772$, $\Phi(3) = 0.9987$.

Definition 2 Random variables X_1, \ldots, X_n are independent if

$$\mathbb{P}(X_1 \leq t_1, X_2 \leq t_2, \dots, X_n \leq t_n) = \mathbb{P}(X_1 \leq t_1) \cdot \mathbb{P}(X_2 \leq t_2) \cdot \dots \cdot \mathbb{P}(X_n \leq t_n)$$

holds for all real numbers t_1, t_2, \ldots, t_n .

Proposition 3 Let X and Y be independent random variables with normal distribution. Then

- a) for real numbers $c \neq 0$ and d, the random variable cX + d has normal distribution with mean $c \cdot \mathbb{E}(X) + d$ and variance $c^2 \text{Var}(X)$;
- b) for every real number t we have $(X-m)/\sigma$ has standard normal distribution, and

$$\mathbb{P}(X \le t) = \mathbb{P}\left(\frac{X - m}{\sigma} \le \frac{t - m}{\sigma}\right) = \Phi\left(\frac{t - m}{\sigma}\right).$$

- c) X + Y has normal distribution, and $\mathbb{E}(X + Y) = \mathbb{E}(X) + \mathbb{E}(Y)$, and Var(X + Y) = Var(X) + Var(Y).
- d) If $X_1, X_2, ..., X_n$ are independent normal variables with mean m and variance σ^2 , then their mean also has a normal distribution:

$$\overline{X} = \frac{X_1 + \ldots + X_n}{n} \sim N\left(m, \frac{\sigma^2}{n}\right).$$

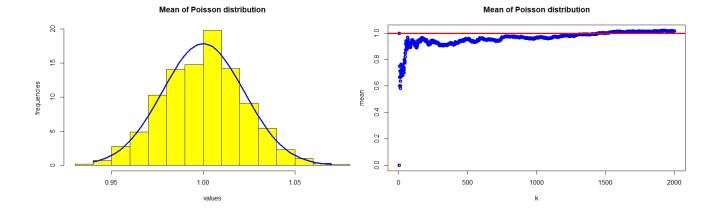
- 1. Generate m=1000 independent, different samples, each consisting of n=2000 independent random variables with
 - (a) Poisson distribution with parameter $\lambda = 1$;
 - (b) exponential distribution with parameter 1;
 - (c) uniform distribution on the interval [0, 2].
 - (d) uniform distribution on the interval [-3, 5]

Calculate the mean of each sample, and make a histogram (on four different figures) of these means.

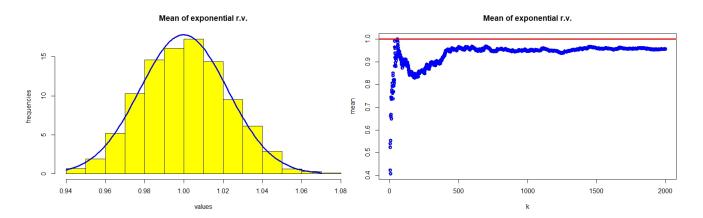
Then, let us take just one sample consisting of n=2000 elements (with the distributions given above), and plot the mean of the first k elements as a function of k (here $k=1,2,\ldots,2000$). Are the four figures similar? What are the differences?

Poisson distribution:

- > samplea=matrix(rpois(1000*2000, lambda=1), nrow=1000)
- > meana=rowMeans(samplea)



- > hist(meana, col="yellow", main="Mean of Poisson distribution",
 xlab="values", ylab="frequencies", freq=F)
- > curve(dnorm(x, mean=1, sd=1/sqrt(2000)), from=0.94, to=1.07, lwd="3", col="blue", add=T)
- > plot(cumsum(samplea[1,])/(1:2000), col="blue", lwd="3", xlab="k", ylab="mean",
 main="Mean of Poisson distribution")
- > lines(abline(a=1, b=0, col="red", lwd="3"))



Exponential distribution:

sampleb=matrix(rexp(1000*2000, rate=1), nrow=1000)

meana=rowMeans(sampleb)

hist(meana, col="yellow", main="Mean of exponential r.v.", xlab="values", ylab="frequencies", freq=F)

curve(dnorm(x, mean=1, sd=1/sqrt(2000)), from=0.94, to=1.07, lwd="3", col="blue", add=T)
plot(cumsum(sampleb[1,])/(1:2000), col="blue", lwd="3", xlab="k", ylab="mean",

lines(abline(a=1, b=0, col="red", lwd="3"))

main="Mean of exponential r.v.")

2. Suppose that the temperature in Budapest, on October 31 at midnight has normal distribution X with mean 1 and variance 4 (that is, $X \sim N(1,4)$), in Celsius degree).

a) What is the probability that the temperature at midnight on the given day is below 0° C?

$$\mathbb{P}(X < 0) = \mathbb{P}(\frac{X-1}{2} < -\frac{1}{2}) = \Phi(-0.5) = 1 - \Phi(0.5) = 1 - 0.6915 = 0.3085.$$

> pnorm(0, mean=1, sd=2)

[1] 0.3085375

b) What is the probability that the temperature is between $-1^{\circ}C$ and $3^{\circ}C$?

$$\mathbb{P}(Z > 1) = 1 - \Phi(1) \text{ if } Z \sim N(0, 1)$$

$$\mathbb{P}(-1 < X < 3) = \mathbb{P}\big(-1 < \frac{X-1}{2} < 1\big) = \mathbb{P}(X < 3) - \mathbb{P}(X < -1) = \Phi(1) - \Phi(-1) = \Phi(1) - (1 - \Phi(1)) = 2\Phi(1) - 1 = 2 \cdot 0.8413 - 1 = 0.6826, \text{ where we used that } \Phi(-x) = 1 - \Phi(x) \text{ holds for every } x.$$

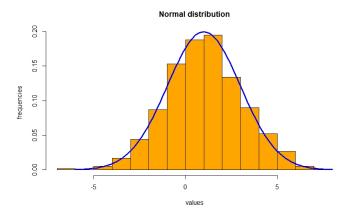
> pnorm(3, mean=1, sd=2)-pnorm(-1, mean=1, sd=2)

[1] 0.6826895

c) Generate a random independent sample of size 1000 from the normal distribution with mean 1 and variance 4. Make a histogram, calculate the mean and the standard deviation, and determine the proportion of the elements between -1 and 3.

- > temp=rnorm(1000, mean=1, sd=2)
- > hist(temp, col="orange", main="Normal distribution", xlab="values",
 ylab="frequencies", freq=F)
- > curve(dnorm(x, mean=1, sd=2), from=-6, to=8, lwd="3", col="blue", add=T)
- > length(temp[(-1 < temp)&(temp < 3)])/1000

[1] 0.704



3. Suppose that the volume of mineral water in a bottle has normal distribution, with standard deviation 0,01 (in litres). What is the mean of the distribution if the probability that the volume is less than 0,5 litres is 2%?

Let X be the volume of the water, its mean denoted by m (in litres). Then the condition says that

$$\mathbb{P}(X < 0.5) = 0.02.$$

Since X has normal distribution with mean m and standard deviation $\sigma = 0.01$, we obtain that

$$\mathbb{P}(X < 0.5) = \Phi\left(\frac{0.5 - m}{0.01}\right) = 0.02 \quad \Rightarrow \quad \frac{0.5 - m}{0.01} = \Phi^{-1}(0.02) = -\Phi^{-1}(0.98) = -2.05$$

$$0.5 - m = 0.01 \cdot -2.05 \quad \Rightarrow \quad m = 0.5 + 0.01 \cdot 2.05 = 0.52.$$

> qnorm(0.02)

[1] -2.053749

> qnorm(0.98)

[1] 2.053749

- 4. We install a programme which consists of n=68 components. The time of the installation of each file has mean m=10 and $\sigma=2$ (in seconds), and these are independent of each other.
 - a) Calculate the probability that the sum of the installation time is at most 12 minutes, supposing that the installation times have normal distribution.
 - b) Approximate the same probability without assuming normal distribution.
 - c) The next version consists of k files, with the same rules as above. We also know that the probability that the total time is less than 10 minutes is 95% percent. Determine the value of k.

Let us denote by S_n the sum of the installation times of n components (n = 68), in seconds. Since the sum of independent normal distribution has normal distribution, S_n has normal distribution, with mean $n\mu$ and standard deviation $\sqrt{n}\sigma$. Hence

$$\mathbb{P}(S_{68} \le 720) = \mathbb{P}(S_n < 720) = \mathbb{P}\left(\frac{S_n - n\mu}{\sqrt{n}\sigma} < \frac{720 - 680}{2\sqrt{68}}\right) = \Phi(2.42) = 99.2\%$$

b) Postponed till next week.

$$0.95 = \mathbb{P}(S_n < 600) = \mathbb{P}\left(\frac{S_n - n\mu}{\sqrt{n}\sigma} < \frac{600 - 10n}{2\sqrt{n}}\right) = \Phi\left(\frac{600 - 10n}{2\sqrt{n}}\right)$$

Since $\Phi(1.645) = 0.95$, we have $1.645 = \frac{600 - 10n}{2\sqrt{n}}$. By solving this, we have that $n \le 57.5$, hence the number of components has to be at most 58.

- 5. Suppose that X_1, X_2, \ldots , are independent random variables with Poisson distribution, and the variance of each X_j is equal to 3.
 - a) Find the almost sure limit of $\frac{X_1+X_2+...+X_n}{n}$.
 - b) Find the limit of $\mathbb{P}\left(\frac{X_1+\ldots+X_n-3\cdot n}{\sqrt{n}}\leq 1\right)$ as n tends to infinity.

$$\mathbb{P}\left(\lim_{n\to\infty}\frac{X_1+X_2+\ldots+X_n}{n}=3\right)=1.$$

- 6. Suppose that X_1, X_2, \ldots , are independent random variables with exponential distribution, and we also know that $\mathbb{P}(X_j \geq 1) = e^{-3}$.
 - a) Find the almost sure limit of $\frac{X_1+X_2+...+X_n}{n}$.
 - b) Find the limit of $\mathbb{P}\left(\frac{X_1+\ldots+X_n-\frac{1}{3}\cdot n}{\sqrt{n}}\leq 1\right)$ as n tends to infinity.

$$f(x) = \begin{cases} \lambda e^{-\lambda x}; & x > 0; \\ 0; & x < 0. \end{cases}$$

Random variables X, Y are identically distributed if their cumulative distribution function is the same: $\mathbb{P}(X \leq t) = \mathbb{P}(Y \leq t)$ holds for every t.

Theorem 4 (Law of large numbers) Let $X_1, X_2, ...$ be independent identically distributed random variables. Suppose furthermore that $m = \mathbb{E}(X_1) < \infty$. Then

$$\overline{X}_n = \frac{X_1 + X_2 + \ldots + X_n}{n} \to \mathbb{E}(X_1) = m$$

holds with probability 1 as $n \to \infty$.

Theorem 5 (Central limit theorem) Let X_1, X_2, \ldots be independent and identically distributed random variables, for which $\mathbb{E}(X_1) = m$ and $s.d.(X_1) = \sigma < \infty$, that is, their variance is finite. Then for every real number t we have

$$\mathbb{P}\left(\frac{X_1 + X_2 + \ldots + X_n - n \cdot m}{\sigma \sqrt{n}} \le t\right) \to \mathbb{P}(Z \le t) \qquad (n \to \infty),$$

where Z has standard normal distribution, that is.

$$\mathbb{P}(Z \le t) = \Phi(t) = \int_{-\infty}^{t} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx.$$