

## Verzani Problem Set

Next are considered the problems from Verzani's book on page 83.

### Problem 13.1

The cost of a home depends on the number of bedrooms in the house. Suppose the following data is recorded for homes in a given town

price (in thousands)	300	250	400	550	317	389	425	289	389	559
Number of bedrooms	3	3	4	5	4	3	6	3	4	5

Make a scatterplot, and fit the data with a regression line. On the same graph, test the hypothesis that an extra bedroom costs \$60,000 against the alternative that it costs more.

```
> price <- c(300000, 250000, 400000, 550000, 317000, 389000, 425000, 289000, 389000, 559000)
> bedrooms <- c(3, 3, 4, 5, 4, 3, 6, 3, 4, 5)
> lmResult <- lm(price ~ bedrooms)
> summary(lmResult)
```

Call:

lm(formula = price ~ bedrooms)

Residuals:

```
    Min     1Q  Median     3Q    Max
-108000 -53950 -5750  59775  99100
```

Coefficients:

```
            Estimate Std. Error t value Pr(>|t|)
(Intercept)  94400     97983   0.963  0.3635
bedrooms     73100     23764   3.076  0.0152 *
```

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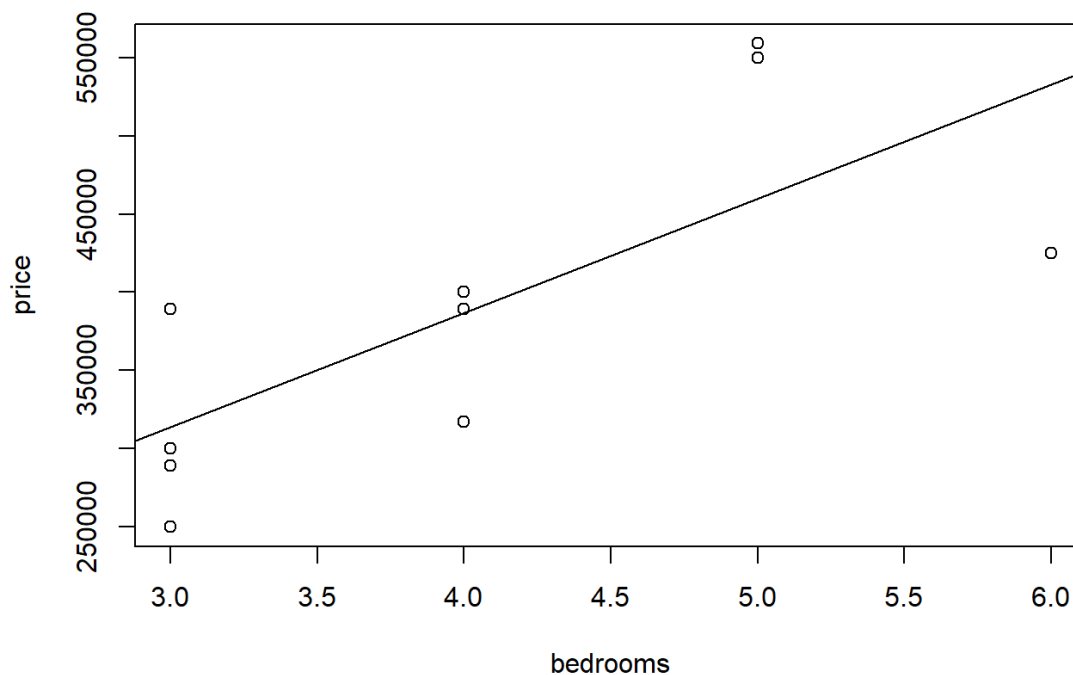
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 75150 on 8 degrees of freedom

Multiple R-squared: 0.5419, Adjusted R-squared: 0.4846

F-statistic: 9.462 on 1 and 8 DF, p-value: 0.01521

```
> plot.bedrooms, price)
> abline(lmResult)
```



$$H_0 : \beta_1 = 60000$$

$$H_1 : \beta_1 > 60000$$

$$S_\varepsilon^2 = \frac{\sum_{i=1}^n \varepsilon_i^2}{n-2}$$

$$SE(\beta_1) = \frac{S_\varepsilon}{\sqrt{\sum_{i=1}^n (X_i - \bar{X}_n)^2}}$$

$$\left( \frac{\hat{\beta}_1 - b_1}{SE(\beta_1)} \mid H_0 \right) \in t(n-2)$$

Given  $\alpha$  the critical area is

$$W_\alpha = \left\{ \frac{\hat{\beta}_1 - b_1}{SE(\beta_1)} \geq t_{1-\alpha; n-2} \right\}$$

```
> e <- resid(lmResult); e
```

```
 1    2    3    4    5    6    7    8    9   10
-13700 -63700 13200 90100 -69800 75300 -108000 -24700 2200 99100
```

```

> n <- length(e)
> Seps <- sqrt(sum(e^2) / (n - 2)); Seps
[1] 75149.43
> beta1hat <- coef(lmResult)[['bedrooms']]; beta1hat
[1] 73100
> SEbeta1 <- Seps / sqrt(sum((bedrooms - mean(bedrooms))^2)); SEbeta1
[1] 23764.34
> b1 <- 60000
> t <- (beta1hat - b1) / SEbeta1; t
[1] 0.5512462
> pvalue <- pt(t, n - 2, lower.tail = FALSE); pvalue
[1] 0.2982602

```

The  $p\text{-value} = 0.2982602 > 0.05 = \alpha$ , so we have no evidence to reject  $H_0$ . An extra bedroom costs \$60,000.

Some of the values, for example  $S_e$ ,  $\hat{\beta}_1$ ,  $SS(\beta_1)$  could be taken from the output of `summary(lmResult)`.

### Problem 13.2

It is well known that the more beer you drink, the more your blood alcohol level rises. Suppose we have the following data on student beer consumption

Student	1	2	3	4	5	6	7	8	9	10
Beers	5	2	9	8	3	7	3	5	3	5
BAL	0.10	0.03	0.19	0.12	0.04	0.095	0.07	0.06	0.02	0.05

Make a scatterplot and fit the data with a regression line. Test the hypothesis that another beer raises your BAL by 0.02 percent against the alternative that it is less.

```

> Beers <- c(5, 2, 9, 8, 3, 7, 3, 5, 3, 5)
> BAL <- c(0.10, 0.03, 0.19, 0.12, 0.04, 0.095, 0.07, 0.06, 0.02, 0.05)

```

$$H_0 : \beta_1 = 0.02$$

$$H_A : \beta_1 < 0.02$$

```

> lmResult <- lm(BAL ~ Beers)
> summary(lmResult)

```

Call:

`lm(formula = BAL ~ Beers)`

Residuals:

```

  Min      1Q  Median      3Q     Max
-0.0275 -0.0187 -0.0071  0.0194  0.0357

```

Coefficients:

```

      Estimate Std. Error t value Pr(>|t|)

```

```
(Intercept) -0.018500  0.019230 -0.962 0.364200
Beers       0.019200  0.003511  5.469 0.000595 ***
```

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

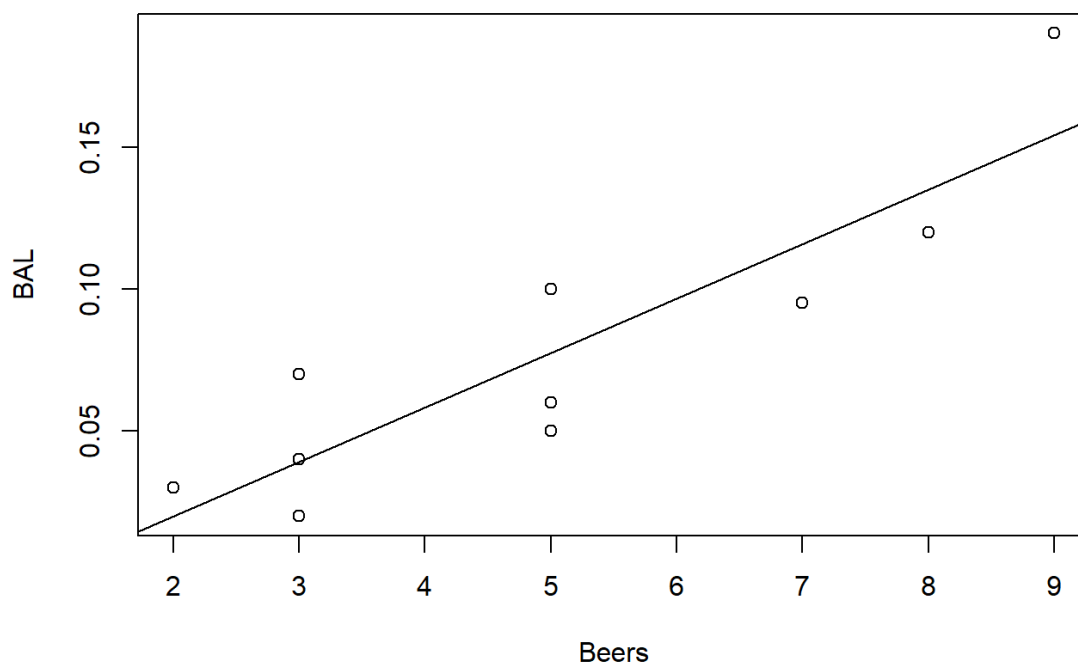
Residual standard error: 0.02483 on 8 degrees of freedom

Multiple R-squared: 0.789, Adjusted R-squared: 0.7626

F-statistic: 29.91 on 1 and 8 DF, p-value: 0.0005953

```
> plot(Beers, BAL)
```

```
> abline(lmResult)
```



```
> e <- resid(lmResult); e
```

```
1 2 3 4 5 6 7 8 9 10
0.0225 0.0101 0.0357 -0.0151 0.0009 -0.0209 0.0309 -0.0175 -0.0191 -0.0275
```

```
> n <- length(e)
```

```
> Seps <- sqrt(sum(e^2) / (n - 2)); Seps
```

```
[1] 0.02482564
```

```
> beta1hat <- coef(lmResult)[['Beers']]; beta1hat
```

```
[1] 0.0192
```

```
> SEbeta1 <- Seps / sqrt(sum((Beers - mean(Beers))^2)); SEbeta1
```

```
[1] 0.003510876
```

```
> b1 <- 0.02
```

```
> t <- (b1 - beta1hat) / SEbeta1; t
```

```
[1] 0.2278634
```

```
> pvalue <- pt(t, n - 2, lower.tail = TRUE); pvalue
```

```
[1] 0.5872658
```

The  $p\text{-value} = 0.5872658 > 0.05 = \alpha$ , so we have no evidence to reject  $H_0$ . Another beer raises your BAL by 0.02 percent or more.

### Problem 13.3

For the same Blood alcohol data, do a hypothesis test that the intercept is 0 with a two-sided alternative.

```
> Beers <- c(5, 2, 9, 8, 3, 7, 3, 5, 3, 5)
> BAL <- c(0.10, 0.03, 0.19, 0.12, 0.04, 0.095, 0.07, 0.06, 0.02, 0.05)
```

$$H_0 : \beta_0 = 0$$

$$H_1 : \beta_0 \neq 0$$

```
> lmResult <- lm(BAL ~ Beers)
> summary(lmResult)
```

Call:

```
lm(formula = BAL ~ Beers)
```

Residuals:

```
    Min     1Q  Median     3Q     Max
-0.0275 -0.0187 -0.0071  0.0194  0.0357
```

Coefficients:

```
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.018500   0.019230  -0.962 0.364200
Beers         0.019200   0.003511   5.469 0.000595 ***
```

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.02483 on 8 degrees of freedom

Multiple R-squared: 0.789, Adjusted R-squared: 0.7626

F-statistic: 29.91 on 1 and 8 DF, p-value: 0.0005953

The  $p\text{-value} = 0.3642 > 0.05 = \alpha$ , so we have no evidence to reject  $H_0$ .

### Problem 13.4

The lapse rate is the rate at which temperature drops as you increase elevation. Some hardy students were interested in checking empirically if the lapse rate of 9.8 degrees C/km was accurate for their hiking. To investigate, they grabbed their thermometers and their Suunto wrist altimeters and found the following data on their hike

elevation (ft)	600	1000	1250	1600	1800	2100	2500	2900
temperature (F)	56	54	56	50	47	49	47	45

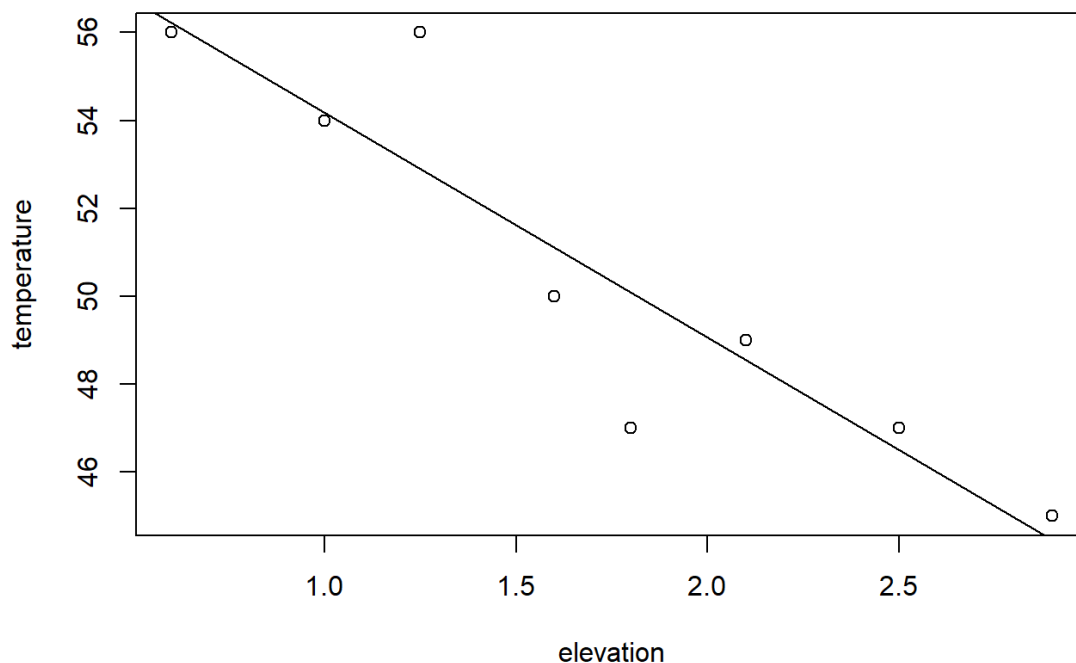
Draw a scatter plot with regression line, and investigate if the lapse rate is 9.8 C/km. (First, it helps to convert to the rate of change in Fahrenheit per feet with is 5.34 degrees per 1000 feet.) Test the hypothesis that the lapse rate is 5.34 degrees per 1000 feet against the alternative that it is less than this.

According to the conditions we have to check if the regression equation is

$$y = \beta_0 - 5.34x$$

Let us first build up our simple linear regression model and then to test the hypothesis if the slope is  $\beta_1 = -5.34$ .

```
> elevation <- c(600, 1000, 1250, 1600, 1800, 2100, 2500, 2900) / 1000
> temperature <- c(56, 54, 56, 50, 47, 49, 47, 45)
> lmResult <- lm(temperature ~ elevation)
> plot(elevation, temperature)
> abline(lmResult)
```



We test

$$H_0 : \beta_1 = -5.34$$

$$H_A : \beta_1 < -5.34$$

```
> summary(lmResult)
```

Call:

```
lm(formula = temperature ~ elevation)
```

Residuals:

Min	1Q	Median	3Q	Max
-3.0844	-0.4433	0.1369	0.5072	3.1025

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	59.2907	1.7173	34.525	3.93e-08 ***
elevation	-5.1146	0.9214	-5.551	0.00144 **

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.879 on 6 degrees of freedom  
Multiple R-squared: 0.837, Adjusted R-squared: 0.8099  
F-statistic: 30.81 on 1 and 6 DF, p-value: 0.001445

We can take  $\hat{\beta}_1$  and  $SE(\hat{\beta}_1)$  values from the output of the function summary however the rest of those in

$$W_\alpha = \left\{ \frac{\hat{\beta}_1 - b_1}{SE(\hat{\beta}_1)} < t_{\alpha; n-2} \right\}, SE(\hat{\beta}_1) := \frac{S_\varepsilon}{\sqrt{\sum_{i=1}^n (X_i - \bar{X}_n)^2}}$$

we have to compute them step by step

```
> e <- resid(lmResult)
> n <- length(e)
> beta1hat <- (coef(lmResult))['elevation']; beta1hat
[1] -5.114567
> Seps <- sqrt(sum(e^2) / (n-2))
> SEbeta1 <- Seps / sqrt(sum((elevation - mean(elevation))^2)); SEbeta1
[1] 0.9213715
> alpha <- 0.05
> tquantile <- qt(alpha, n - 2, lower.tail = TRUE); tquantile
[1] -1.94318
> const <- -5.34
> temp <- (const - beta1hat) / SEbeta1; temp
[1] -0.2446712
> pvalue <- pt(temp, n - 2, lower.tail = TRUE); pvalue
[1] 0.4074318
```

See also the outputs of `summary(lmResult)`.

The  $p\text{-value} = 0.4074318 > 0.05 = \alpha$ , so we have no evidence to reject  $H_0$  and according to the sample the difference between  $\hat{\beta}_1 = -5.114567$  and the tested value  $-5.34$  is not statistically significant.