Comments on Blair and Sambanis, 2020, "Forecasting Civil Wars: Theory and Structure in an Age of 'Big Data' and Machine Learning", JCR

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We highlight some issues with the analysis in the recent Blair and Sambanis (2020) paper that undermine their main empirical finding – that a theoretically-informed escalation model can forecast civil war onsets better than more inductive alternatives. Their paper has two key results:

- 1. A theoretically-informed model outperforms alternative specifications in predicting civil war onset.
- 2. Structural variables do not add, and sometimes detract, from the forecast accuracy of the escalation model they advance.

These results are highlighted in Table 1 and Table 2 (Blair and Sambanis 2020, pg. 13 and 19 respectively). They also evaluate forecasts for the first half of 2016 and conclude good accuracy in forecasting civil war (Table 4 on pg.23).

We find three main issues in their empirical analyses:

- Their primary finding that the escalation model they advance consistently out-performs the alternation models depends entirely on the unorthodox use of smoothed ROC curves when calculating the AUC-ROC values in Table 1 and Table 2. When using the standard (not smoothed) ROC curves, this finding is reversed. That is, the escalation performs no better and often worse than the other models.
- Regarding the forecast assessment in Table 4, the forecasts for 2016-H1 are erroneously assessed using civil war *incidence*, not *onset*, even though the models from which they are generated are clearly fitted to civil conflict onset. This vastly overstates the precision of the forecasts.
- The way the random forest models are fitted is very unorthodox.¹ The number of tree the model grow, the number of nodes in these trees, and the sample size used for each tree deviates drastically for traditional approaches (e.g., they set the number of trees to 100,000 where the default is 500).

We demonstrate the first point through model replication: we use the replication files that Blair and Sambanis provide to calculate the AUC-ROC scores using smoothed ROC curves (the published results) and then calculate the AUC-ROC scores using the standard approach – without the unorthodox smoother.² We demonstrate the second point using the replicated results from the original model Blair and Sambanis use to create Table 4 and then compare these results to the observed outcomes in the 2016 data.

For the final point, we evaluate two questions: (1) even if we allow for the use of smoothed ROC curves, does the key model's superior test forecast performance generalize or is it a result of inadvertent fitting of the model hyperparameters to test forecast performance, and (2) does the inferior performance of the alternative models, and especially the CAMEO model with 1,159 covariates, reflect a lack of proper hyperparameter tuning for these models? (work still in progress).

¹This is in fact what first led us to examine the analysis in more detail.

²In this draft, we focus our replication efforts on what Blair and Sambanis call their "Base specification". This set of models are what Blair and Sambanis use when discussing their main findings and the implications of their results.

Review of the Blair and Sambanis paper

Blair and Sambanis (2020) argue that theory contributes to forecasting accuracy even when using non-parametric machine learning models (in this case, the random forest model) and alternative specifications that are not informed by theory. They arrive at this conclusion by examining the problem of predicting civil war onset, and find that a parsimonious model using a small number of covariates derived from escalation theories of conflict can forecast civil war onset better than alternative specifications based on generic covariates not specifically informed by theory and a kitchen sink model with more than 1,000 covariates.

Their data consist of country-month observations covering 2001 to 2015. The main outcome is civil war onset, measured using Sambanis' civil war dataset. The independent variables for the first set of analysis reported in Table 1 in the paper are all derived from the ICEWS event data, using domestic events between actors within a country. The key results are based on a comparison of the forecast accuracy of a random forest (RF) model with 4 covariate specifications as well as an ensemble model. These models are:

- Escalation: a set of 10 theoretically-informed indicators.
- Quad: ICEWS quad counts, i.e. material conflict, material cooperation, verbal conflict, verbal cooperation.
- Goldstein: -10 (conflictual) to 10 (cooperative) scores derived from the ICEWS data for interactions between the government one one side and opposition or rebel actors on the other. These are directed, thus making for 4 total covariates.
- CAMEO: counts for all CAMEO event codes, thus a total of 1,159 covariates.
- Average: unweighted average of the predictions from the 4 models above.

To assess forecast accuracy, Blair and Sambanis conduct the following procedures:

- 1. Split the training data into training (2001 2007) and test (2008 2015) sets.
- 2. Estimate 4 random forest models with the specifications from above as well as a unweighted average of the models
- 3. Create out-of-sample (OOS) predictions from each model using the test set.
- 4. Calculate AUC-ROC measures for each set of OOS predictions.

This is done for both 1-month ahead and 6-month ahead forecasts. Blair and Sambanis also conduct a series of robustness tests that (1) vary hyperparameters of the RF model, (2) change the split year for the training/test data split, and (3) alternative codings of the dependent variable.

The results for this analysis are summarized in Table 1 in Blair and Sambanis (2020), which we copied in here:

Table 1. Out-of-sample Area under the Receiver Operating Characteristic Curves for Onemonth and Six-month Forecasts.

Model	Escalation	Quad	Goldstein	CAMEO	Avg.
One-month forecasts					
Base specification	.85	.80	.79	.82	.82
Terminal nodes	.85	.80	.78	.83	.82
Sample size	.85	.81	.71	.86	.84
Trees per forest	.85	.80	.78	.83	.82
Training/test sets I	.86	.78	.76	.81	.80
Training/test sets 2	.81	.79	.73	.77	.78
Training/test sets 3	.79	.81	.69	.75	.76
Coding of DV I	.86	.81	.79	.84	.83
Coding of DV 2	.92	.80	.81	.81	١8.
Six-month forecasts					
Base specification	.82	.78	.82	.76	.79
Terminal nodes	.80	.76	.81	.76	.78
Sample size	.83	.78	.78	.79	.79
Trees per forest	.82	.78	.82	.77	.79
Training/test sets I	.79	.78	.81	.76	.78
Training/test sets 2	.73	.73	.76	.73	.75
Training/test sets 3	.88	.71	.81	.68	.79
Coding of dependent variable I	.83	.78	.82	.78	.80
Coding of dependent variable 2	.83	.77	.83	.78	.79

Note: AUCs for our five random forests models. The top row in each panel reports AUCs for the base specification. We also report results with 10 rather than five terminal nodes (second row); 500 rather than 100 observations per tree (third row); 1,000,000 rather than 100,000 trees per forest (fourth row); a test set that begins January 1, 2009, January 1, 2010, or January 1, 2011 (fifth, sixth, and seventh rows, respectively); and alternate codings of the dependent variable (eighth and ninth rows) as described in the Online Appendix.

The escalation model generally out-performs the alternative models/specifications, and on that basis Blair and Sambanis conclude that theory does indeed provide better forecast accuracy when compared to models not informed by theory.

After establishing that the escalation model they advance out-performs all other covariate specifications, Blair and Sambanis move to show how the inclusion of structural risk variables – the set of structural variables identified by the Political Instability Task Force (PITF) that can contribute to an increased risk of civil conflict. The figure below is Table 2 from Blair and Sambanis (2020), which presents these findings.

Table 2. Out-of-sample Area under the Receiver Operating Characteristic Curves for Models Including PITF.

	Escalation	With PITF	Weighted	PITF Split	PITF
	Only	Predictors	by PITF	Population	Only
One-month forecasts	.85	.78	.53	.84	.76
Six-month forecasts	.82	.86	.52	.83	.74

Note: AUCs for the escalation model incorporating Political Instability Task Force (PITF) predictors. The top and bottom panels report results from one-month and six-month forecasts, respectively.

Issues

Smoothed ROC curves

One issue that stood out as we read Blair and Sambanis (2020) was how smooth their Receiver Operating Characteristic curves where in Figure 1 on page 14. A Receiver Operating Characteristic curve displays the balance between the true positive rate and the false positive rate, across a range of discrimination thresholds – the value in which a zero to one probability is assumed to be an "onset". Given that the DV is civil conflict onset (1/0) and that this is a rare event, these curves should resemble a step-like function. Looking through the Blair/Sambanis replication code, we noticed that they use a smoothing function when producing their Receiver Operating Characteristic curves.

This is important because Blair and Sambanis use a common performance metric – Area Under the Receiver Operating Characteristic Curves (AUC-ROC) – as evidence that their escalation model out preforms all others. However, by smoothing the Receiver Operating Characteristic Curves, they are misrepresenting the true AUC-ROC scores of their various models. To help make this point clear, we replicate Figure 1 below. However, rather than a two panel figure (one ROC plot their one-month-ahead forecasts and one for their six-months-ahead forecasts, as is presented in the article), we now include four panels: one panel for the two forecast windows with the smoothed curve, as they produce, and one panel for the two forecast windows that are not smoothed.

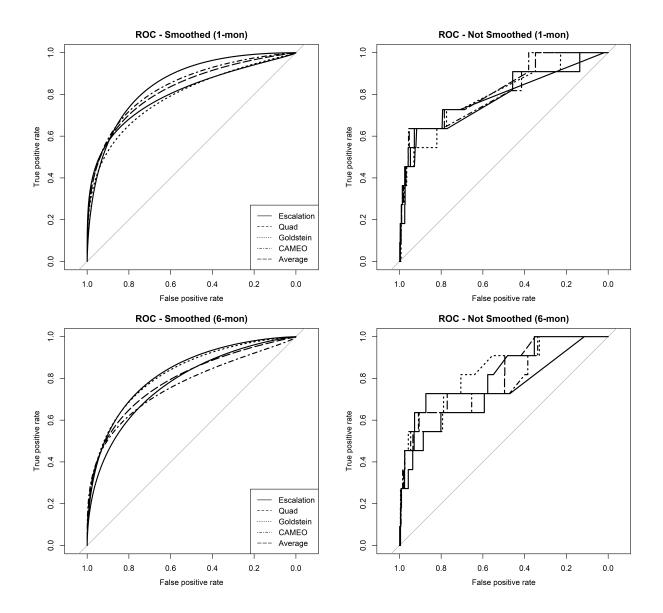


Table 1: Comparison of smoothed and not smoothed AUC-ROC values

	horizon	ROC Smoothed	Escalation	Quad	Goldstein	CAMEO	Avg
Bas	Base Specification						
	1 month	Yes	0.85	0.80	0.79	0.84	0.82
		No	0.79	0.79	0.79	0.81	0.82
	6 month	Yes	0.82	0.77	0.82	0.77	0.79
		No	0.77	0.79	0.83	0.78	0.81

Using these smoothed and non-smoothed curves, we then replicate Table 1 (page 13), but now the table reports the AUC-ROC scores for each model specification with and without the use of a smoother. If the non-smoothed AUC-ROC scores are drastically different from those of the smoothed AUC-ROC score (a difference of ≥ 0.05) or if the non-smoothed AUC-ROC scores are < 0.8, would suggest the findings presented in Blair and Sambanis (2020) are fragile to the use of a non-standard procedure.

As the table above shows, the results presented in Blair and Sambanis (2020) are fragile to the use of a smoothing option when calculating the AUC-ROC performance metric. Further, even when we rely solely on the replication files provided by Blair and Sambanis, we cannot replicate their findings exactly. For example, in our replication, the AUC-ROC score for the 1-month CAMEO model with the use of the smoother is 0.84; however, in the paper, the reported score is 0.82. This is most likely a function of the version of R Blair and Sambanis used to conduct their analysis. The current replication project was done using R version 4.0.0, which updated their random number seeding procedure, so small differences should be expected.

Assessment of the 2016 forecasts

The predictions for the first half of 2016 are assessed in Table 4:

Table 4. 2016 Confusion Matrices for Six-month Escalation Model.

Assuming Persistence					
			Predic		
			0	ı	
	Observed	0	132	17	
	Observed	I	2	13	
Assuming Change					
			Predi	cted	
			0	I	
	Observed	0	132	16	
	Observed	1	2	14	

Note: Confusion matrices based on predicted probabilities from the six-month escalation model. We code as I the top 30 countries at highest risk of civil war in the first half of 2016. The top panel assumes that all ongoing civil wars as of December 31, 2015, continued in the first half of 2016 and that no new civil wars began. The bottom panel codes the ongoing civil war in Colombia ending in the first half of 2016 and codes new civil wars beginning in Turkey and Burundi.

The table presents 15 or 16 positive cases for 2016-H1, depending on a minor coding decision. These numbers correspond to a positive rate of around 9.5% for the first half of 2016.

By comparison, the 6-month version of the full data from 2001 to 2015 contains only 20 positive cases, with a positive rate of around 0.5%.

The relevant variables in the data are "incidence_civil_ns" and "incidence_civil_ns_plus1", which appears to be a 1-period lead version of the DV that is used in the actual prediction models. The incidence DV contains both 0/1 and missing values. By examining the pattern of missing values, it seems clear that this was originally an incidence variable indicating whether a country was at civil war in a given year or not, and which was converted to an onset version so that onsets retain the value of 1 but continuing civil war years are coded as missing.

The main data set covering 2001 - 2015 contains 20 civil war onsets. But there are 15 "Persistence" cases (top of table 4) and 16 "Change" cases (bottom of table 4)

By examining the code used to generated Table 4, we were able to confirm that the onset forecasts are assessed using incidence, not onset. In the file <code>6mo_make_confusion_matrix.do</code> on line 52, missing values in "incidence_civil_ns" are recoded to 1, thus reverting the onset coding of this variable back to incidence.

The actual data contains 0 new civil war onsets in 2016-H1. Using those data, the 2016-H1 forecasts would have retained a recall value of 1, but precision would have decreased from 13/30 = 0.76 to 1/30 = 0.03.

The second part of the Table 4 script manually codes onset in two countries, Turkey and Burundi. With this coding, the correct recall (bottom of Table 4) is 1 (not 14/16 = 0.88) and the precision changes from 14/30 = 0.47 to 2/30 = 0.07.

Another, minor issue or rather coding error, is related to using a lead version of the DV. With the lead version of the DV, "incidence_civil_ns_plus1", which is what the models are predicting, the predicted value

for 2016-H1 actually indicates the risk of civil war onset in 2016-H2. In the Table 4 script referenced above, the 2016-H1 predictions (for 2016-H2) are assessed using the raw DV, "incidence_civil_ns", not the lead version. Essentially, the forecasts for 2016-H2 are assessed using observed outcomes for 2016-H1. In this case it probably doesn't make a difference since both the raw DV and lead version for 2016-H1 do not have any positive values.

Random forest model tuning

3. Unorthodox RF hyperparameter choices that may be driving the OOS test prediction results. There are two issues here. One is that the RF models are so unorthodox. I'm not sure that this in itself is a problem. But the second potential problem is that the hyperparameter values may not have been tuned in a way that works the best for the non-escalation model, and especially the CAMEO model with 1,100+ features.

The table below shows the randomForest package default hyperparameter values for a binary classification problem like the one at hand and compares them to the Blair and Sambanis base settings. Some hyperparameters have heuristics that determine the default value based on characteristics of the input data; we note these in the second column and what the realized default setting would be for the basic Escalation model (first row in Table 1 in the paper) in the 3rd column. The last column has the settings Blair and Sambanis use.

Hyperparameter	Default heuristic	Default values (Escalation)	B&S value
type ntree mtry replace sampsize nodesize	<pre>floor(sqrt(ncol(x))) nrow(x) if replace, else ceiling(.632*nrow(x)) 1 for classification</pre>	classification 500 3 true 11,869	regression 100,000 3 false 100 1
maxnodes		null	5

It is worth noting that commonly the ntree, mtry, and nodesize parameters are the main parameters tuned; of these three in the Blair and Sambanis specification only 1, the ntree parameter, deviates from the default settings.

In any case, there is a stark contrast in the default RF settings and the way Blair and Sambanis use the RF models. The default approach is to train a relatively small number (ntree; 500) of classification trees, but where each tree is fairly big in that it is trained on data that has the same number of rows as the training data, albeit sampled with replacement (replace is true; sampsize is 11,869), and allowed to grow fairly deep (this is governed by nodesize (1), which is the minimum size a terminal node must have). In contrast, Blair and Sambanis grow very extensive forests with a large number of trees (100,000 compared to 500), but each tree is very small and shallow; only a 100 rows are sampled from the training data for each tree, and the trees are constrained to at most 5 terminal nodes (maxnodes).

This approach only works due to the other unorthodox choice, which is to use regression, not classification, trees. Trying to use classification trees with the other parameter settings in fact does not work at all because it is almost guaranteed that a sample of 100 from the 11,869 training data rows with 9 positive cases will only include 0 (negative) outcomes in the sample. As it is, using regression with a 0 or 1 outcome produces warnings when estimating the models:

Warning message:

```
In randomForest.default(y = as.integer(train_df$incidence_civil_ns_plus1 == :
    The response has five or fewer unique values. Are you sure you want to do regression?
```

While unorthodox, the approach does work. It produces predictions that are within the 0 to 1 interval. The concern is that the settings work well only for the escalation model, and only for the particular test set at

hand. Specifically, we wonder (1) if the hyperparameter settings only work well for the particular test set chosen, and (2) if the CAMEO model with 1,100+ features, compared to 10 or less for the other specifications, would perform better with a more explicit tuning procedure.

Are the RF hyperparameters overfit to the test set?

Does the base specification's good performance in the test set generalize?

One initial piece of evidence is already available from Table 1. In addition to the test set starting 2008, Blair and Sambanis also evaluate test sets starting in 2009, 2010, and 2011 (sets 1, 2, and 3 in Table 1). For the 1 month models, when we compare the base specification performance of each model to the performance in the alternate test sets, the AUC-ROC increases in 2 (of 15) cases, is the same in 0 cases, and decreases in 13 cases. For the 6 month models, it increases in 1 case, is the same in 3 cases, and decreases in 11 cases. Altering the test set thus generally shows reduced performance. The table below shows cross-validated training and OOS test AUC-ROC for the original 1 month escalation model, a RF model with default hyperparameter settings, and a tuned RF model with 10,000 trees and otherwise the default settings. The first set of results are out of sample results from repeated cross-validations performed on the training data. We show the average AUC-ROC, it's lower and upper 95% CI, obtained via bootstrapping, and the standard deviation of the distribution of resampled AUC-ROC values.

Because there are only 9 positive cases in the whole training data, and thus to ensure that any given split will include at least 1 positive case in each data partition, we use 2-fold CV, i.e. splitting the original training data into equally-sized new training and validation sets. This is repeated 21 times for a total of 42 OOS performance samples for each model. The last column shows the OOS performance on the original test set. The first value in this column corresponds to the base escalation model results reported Table 1 in the Blair and Sambanis paper when *not* using smoothed ROC curves.

Model	Avg_CV_ROC_AUC	ci_lower	ci_upper	SD_CV_ROC_AUC	Test_ROC_AUC
Escalation, 1mo	0.68	0.63	0.72	0.15	0.78
Modified Escalation, 1mo	0.62	0.58	0.66	0.12	0.59
Tuned Escalation, 1mo	0.67	0.63	0.71	0.12	0.58

The results show that:

- Both of the alternative RF models are able to achieve roughly similar OOS performance in the training data split. T-tests comparing model 1 to the other models fail to reject the null hypothesis at a 95% confidence level. (A t-test comparing model 1 and 2 average AUC-ROC just barely fails to reject the null hypothesis (p slightly above 0.05).)
- Only the base escalation model is able to achieve good test performance; the other two models, despite achieving similar training data performance, have significantly lower test performance.

This suggests that the base RF specifications are (over-)fit to the test data.

References

Blair, Robert A., and Nicholas Sambanis. 2020. "Forecasting Civil Wars: Theory and Structure in an Age of 'Big Data' and Machine Learning." *Journal of Conflict Resolution*.