

Statistical Analysis

Point estimation

LU4 - “Step-by-Step”

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Topics

■ LU4 – Point estimation

- Notation and concepts
- Homework – Exercise 1
- Homework – Exercise 2
- Example 1
- Example 2
- Homework – Exercise 4

Notation and concepts

Populations can be characterized by **PARAMETERS**, which are fixed numbers.



μ	σ^2	p
$\mu_1 - \mu_2$	$\frac{\sigma_1^2}{\sigma_1^2}$	$P_1 - P_2$

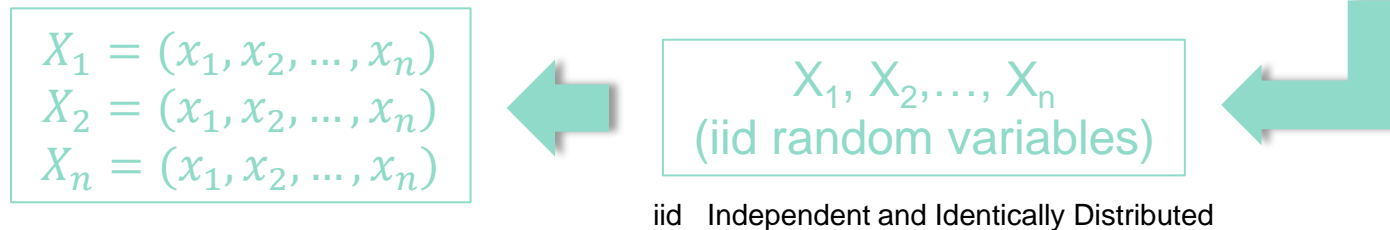
However, these parameters are generally non-observable, as they characterize large populations. That is why we need to **ESTIMATE** them.



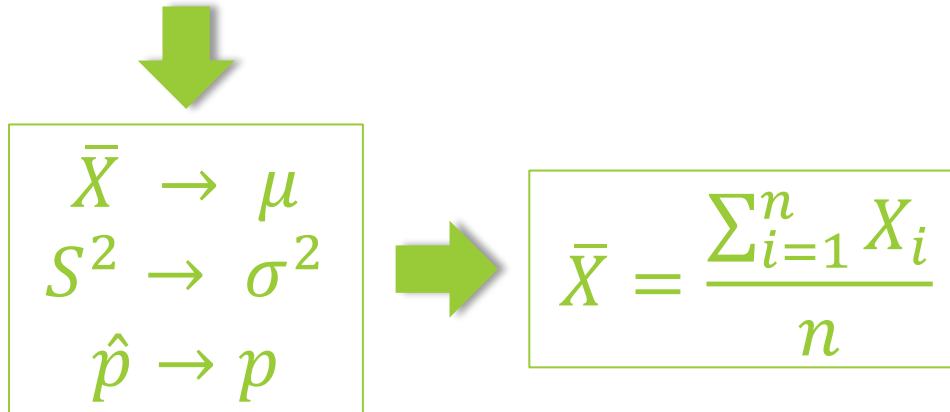
Find an approximate value that we can consider close enough to the unknown parameter.

Notation and concepts

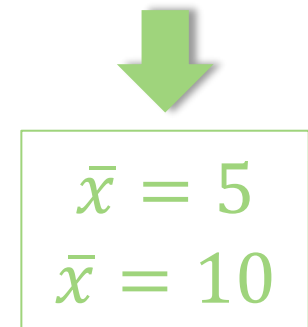
To estimate those parameters, we need to select **RANDOM SAMPLES**.



With the information collected from these Random Samples, we can use **ESTIMATORS** to estimate the parameters.



Estimators are **statistics** (functions of the r.v. X_i) that are used to produce **ESTIMATES**.



Notation and concepts

■ Issues addressed

- In case of more than one estimator of a parameter, how can we decide which one is better than another?
- What are the desirable properties of an estimator?

■ Desirable properties of estimators

- Sufficiency
- Unbiasedness
- Efficiency
- Consistency

Notation and concepts

■ Concepts

- **Sufficiency** → when the estimator takes all the relevant information about the population parameter from the sample.
- **Unbiasedness** → in medium terms, the estimator reaches the actual value of the parameter.
- **Efficiency** → the estimator is more efficient (i.e., the estimates are more *accurate*) the smaller the variability of its sampling distribution.
- **Consistency** → for *large* samples, the estimator should be *approximately* equal to the parameter.

Notation and concepts

■ Unbiasedness

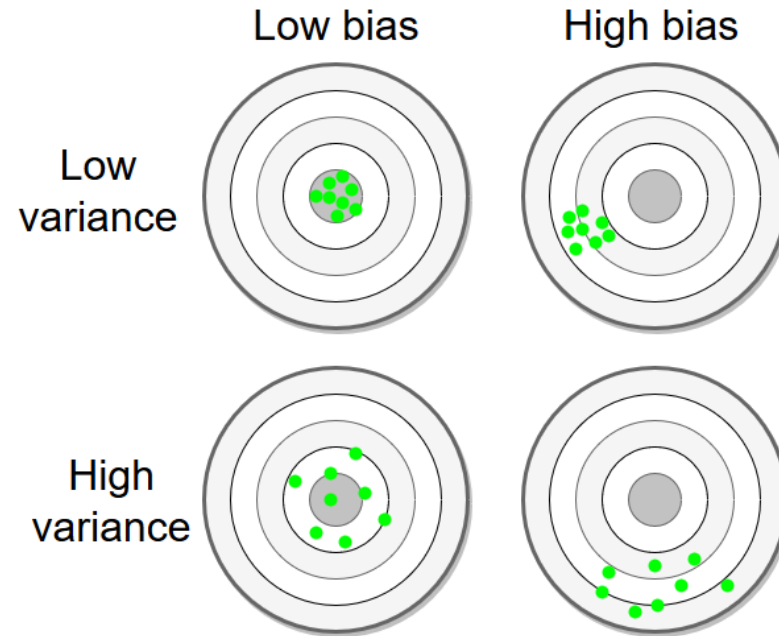


Is there a bias?

■ Efficiency



Is the variance low?



Source: <http://www.machinelearningtutorial.net/2017/01/26/the-bias-variance-tradeoff/>

Unbiasedness

- $\hat{\theta}$ is an **Unbiased** estimator of θ if

$$E(\hat{\theta}) = \theta$$

Otherwise, the estimator is said to be **biased**, and its **bias** is given by

$$bias(\hat{\theta}) = E(\hat{\theta}) - \theta$$

- $\hat{\theta}$ is an **Asymptotically Unbiased** estimator of θ if

$$\lim_{n \rightarrow +\infty} E(\hat{\theta}) = \theta$$

REMEMBER

$$\begin{aligned} E(a) &= a \\ E(aX) &= aE(X) \\ E(X + Y) &= E(X) + E(Y) \end{aligned}$$

Where:

- a is a constant
- X and Y are random variables

Efficiency

- The **Efficiency** of an estimator is its **Mean Squared Error (MSE)**

$$MSE(\hat{\theta}) = E[(\hat{\theta} - \theta)^2] = V(\hat{\theta}) + [bias(\hat{\theta})]^2$$

An estimator will be more efficient (more accurate)

- The smaller the bias.
- The smaller the variance.

REMEMBER

$$V(a) = 0$$

$$V(aX) = a^2 V(X)$$

$$V(X + Y) = V(X) + V(Y) \pm 2Cov(X, Y)$$

If X and Y are independent, then $Cov(X, Y) = 0$

Efficiency

- The **Relative Efficiency** of two estimators $\hat{\theta}_1$ and $\hat{\theta}_2$ of a parameter θ is given by

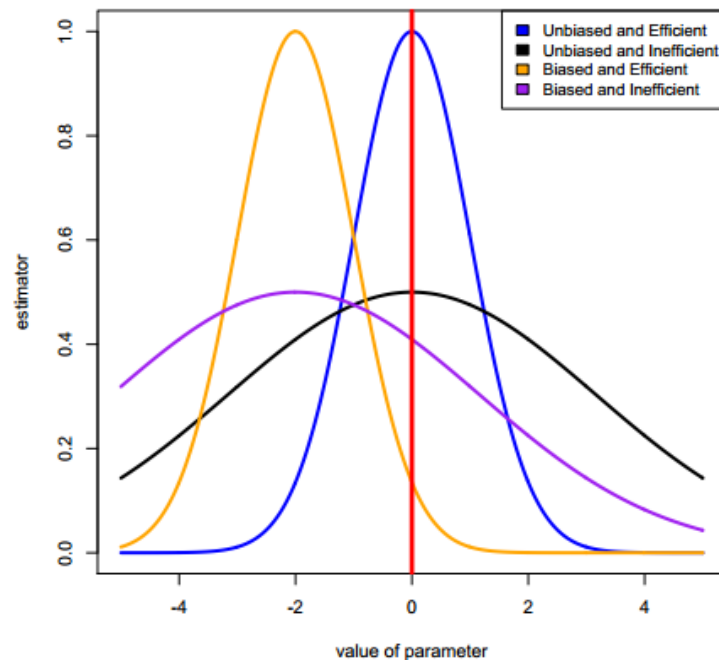
$$eff(\hat{\theta}_1, \hat{\theta}_2) = \frac{MSE(\hat{\theta}_1)}{MSE(\hat{\theta}_2)}$$

- $\hat{\theta}_1$ is more efficient than $\hat{\theta}_2$ when $eff(\hat{\theta}_1, \hat{\theta}_2) < 1$
- **Example**
 - If $eff(\hat{\theta}_1, \hat{\theta}_2) = 1.6$, then the variability associated to $\hat{\theta}_1$ is 1.6 higher than the variability associated to $\hat{\theta}_2$.

Efficiency

■ Efficiency and bias trade-off

- In practice, a biased estimator can be a better estimator than an unbiased one when its bias is small and its efficiency is higher



Source: <http://www.zorro-trader.com/manual/en/Lecture%205.htm>

Consistency

- $\hat{\theta}$ is a **Consistent** estimator of θ if and only if $\hat{\theta}$ converges in probability to θ :

$$\lim_{n \rightarrow +\infty} P(|\hat{\theta} - \theta| \leq \varepsilon) = 1 \quad \text{for all } \varepsilon > 0$$

- An estimator is consistent if increasing the sample size ($n \rightarrow +\infty$) implies an increase in the probability of the estimated value ($\hat{\theta}$) to be in a neighbourhood (ε) of the true value of the parameter (θ).
-
- An unbiased estimator $\hat{\theta}$ is **Consistent** if $\lim_{n \rightarrow +\infty} V(\hat{\theta}) = 0$
 - An estimator $\hat{\theta}$ is **Consistent in Mean Square Error** if $\lim_{n \rightarrow +\infty} MSE(\hat{\theta}) = 0$

Consistency

■ Properties

- If $\hat{\theta}$ and $\hat{\theta}'$ are consistent estimators of θ and θ' , respectively, then
 - $\hat{\theta} + \hat{\theta}'$ is a consistent estimator of $\theta + \theta'$
 - $\hat{\theta} \times \hat{\theta}'$ is a consistent estimator of $\theta \times \theta'$
 - $\hat{\theta}/\hat{\theta}'$ is a consistent estimator of θ/θ' , with $\theta' \neq 0$
 - If $g(\cdot)$ is a real continuous function in θ , then $g(\hat{\theta})$ is a consistent estimator of $g(\theta)$

Summary

■ If you are asked to

- Study the bias of an estimator or
- Evaluate if an estimator is biased/unbiased

1. Find the Expected Value of that Estimator: $E(\hat{\theta})$
2. Calculate its bias:
 $bias(\hat{\theta}) = E(\hat{\theta}) - \theta$

- Study the consistency of an estimator

1. Evaluate the estimator's bias.
2. If it is **unbiased**, then evaluate if $\lim_{n \rightarrow +\infty} V(\hat{\theta}) = 0$
3. If it is **biased**, then evaluate if $\lim_{n \rightarrow +\infty} MSE(\hat{\theta}) = 0$
4. An estimator is consistent if it is unbiased and $\lim_{n \rightarrow +\infty} V(\hat{\theta}) = 0$ or if $\lim_{n \rightarrow +\infty} MSE(\hat{\theta}) = 0$

- Evaluate the Relative Efficiency between two estimators or

- Compare the efficiency between two estimators

1. Calculate their bias.
2. Calculate their variance.
3. Calculate their MSE.
4. Calculate the Relative Efficiency:
$$eff(\hat{\theta}, \tilde{\theta}) = \frac{MSE(\hat{\theta})}{MSE(\tilde{\theta})}$$

Example 1

- a) Population with mean μ and variance σ^2 . Show that \bar{X} is an unbiased estimator of μ .

$$E(\bar{X}) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{1}{n} * n * \mu = \mu$$

$\therefore \bar{X}$ is an unbiased estimator of μ

Exercise 1

- **Population with mean μ and variance σ^2 . Show that \bar{X} is a consistent estimator of μ .**

- Step 1: Evaluate the estimator's bias.

- Step 1.1. Find the Expected Value of \bar{X} .

$$E(\bar{X}) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{1}{n} n\mu = \mu$$

- Step 1.2. Calculate the bias.

$$\text{bias}(\bar{X}) = E(\bar{X}) - \mu = \mu - \mu = 0 \Rightarrow \bar{X} \text{ is unbiased}$$

- Step 2: As \bar{X} is unbiased, then evaluate if $\lim_{n \rightarrow +\infty} V(\bar{X}) = 0$.

- Step 2.1. Find the Variance of \bar{X} .

$$V(\bar{X}) = V\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n^2} \sum_{i=1}^n V(X_i) = \frac{1}{n^2} n\sigma^2 = \frac{\sigma^2}{n}$$

- Step 2.2. Evaluate if $\lim_{n \rightarrow +\infty} V(\hat{\theta}) = 0$

When n increases and approaches $+\infty$, $\frac{\sigma^2}{n}$ gets close to 0.

$\therefore \bar{X}$ is a consistent estimator of μ as it is an unbiased estimator and its variance approaches zero as n increases.

Exercise 1

- Population with mean μ and variance σ^2 . Consider two alternative estimators of μ , both unbiased, and their variances:

$$\hat{\mu}_1 = \bar{X} = \frac{1}{n} \sum_{i=1}^n X_i \quad \hat{\mu}_2 = \frac{1}{n-1} \sum_{i=1}^{n-1} X_i \quad V(\hat{\mu}_1) = \frac{1}{n} \sigma^2 \quad V(\hat{\mu}_2) = \frac{1}{n-1} \sigma^2$$

Show that $\hat{\mu}_1$ is more efficient than $\hat{\mu}_2$.

- Step 1: Calculate their bias. In this case, both are unbiased:

$$\text{bias}(\hat{\mu}_1) = \text{bias}(\hat{\mu}_2) = 0$$

- Step 2: Calculate their variances.

$$V(\hat{\mu}_1) = \frac{1}{n} \sigma^2; \quad V(\hat{\mu}_2) = \frac{1}{n-1} \sigma^2$$

- Step 3: Calculate their MSE.

As they are unbiased, their MSE=Variance

- Step 4: Calculate the efficiency:
$$\text{eff}(\hat{\mu}_1, \hat{\mu}_2) = \frac{MSE(\hat{\mu}_1)}{MSE(\hat{\mu}_2)} = \frac{\frac{1}{n} \sigma^2}{\frac{1}{n-1} \sigma^2} = \frac{n-1}{n} < 1$$

$\therefore \hat{\mu}_1$ is more efficient than $\hat{\mu}_2$

Exercise 2

- Population with mean μ and variance σ^2

a) Find the value of a that makes $\hat{\mu}$ an unbiased estimator.

$$\hat{\mu} = 0,4X_1 + aX_3 + 0,3X_5 + 0,2X_n$$

- Step 1: Find the Expected Value of $\hat{\mu}$: $E(\hat{\mu})$

$$\begin{aligned} E(\hat{\mu}) &= E(0,4X_1 + aX_3 + 0,3X_5 + 0,2X_n) \\ &= 0,4E(X_1) + aE(X_3) + 0,3E(X_5) + 0,2E(X_n) \\ &= 0,4\mu + a\mu + 0,3\mu + 0,2\mu = (0,9 + a)\mu \end{aligned}$$

- Step 2: Find a so that $\hat{\mu}$ is unbiased: $bias(\hat{\mu}) = 0$.

$$\begin{aligned} bias(\hat{\mu}) &= 0 \\ \Leftrightarrow E(\hat{\mu}) - \mu &= 0 \\ \Leftrightarrow (0,9 + a)\mu - \mu &= 0 \\ \Leftrightarrow 0,9 + a &= 1 \\ \Leftrightarrow a &= 0,1 \end{aligned}$$

Exercise 2

- b) Study $\hat{\mu}$ as to its consistency in Mean Square Error (MSE).

$$\hat{\mu} = 0,4X_1 + 0,1X_3 + 0,3X_5 + 0,2X_n$$

- Step 1: Evaluate if $\hat{\mu}$ is biased/unbiased

We know $\hat{\mu}$ is unbiased from the previous exercise.

- Step 2: Find the Variance of $\hat{\mu}$ - $V(\hat{\mu})$

$$\begin{aligned} V(\hat{\mu}) &= V(0,4X_1 + 0,1X_3 + 0,3X_5 + 0,2X_n) \\ &= 0,4^2V(X_1) + 0,1^2V(X_3) + 0,3^2V(X_5) + 0,2^2V(X_n) \\ &= 0,16\sigma^2 + 0,01\sigma^2 + 0,09\sigma^2 + 0,04\sigma^2 = 0,30\sigma^2 \end{aligned}$$

- Step 3: Find the MSE of $\hat{\mu}$ - $MSE(\hat{\mu})$

$$MSE(\hat{\mu}) = V(\hat{\mu}) + bias(\hat{\mu}) = 0,30\sigma^2$$

$\therefore \hat{\mu}$ is not consistent in MSE because $MSE(\hat{\mu})$ does not depend on n and, therefore, $\lim_{n \rightarrow +\infty} MSE(\hat{\mu}) \neq 0$

Exercise 2

- c) Compare $\hat{\mu}$ with the following alternative estimator as to the Relative Efficiency.

$$\tilde{\mu} = \frac{1}{8}(2X_1 + 5X_2 - 3X_3 + 4X_n)$$

- Step 1: Evaluate if $\tilde{\mu}$ is biased/unbiased

$$\begin{aligned} E(\tilde{\mu}) &= E\left(\frac{1}{8}(2X_1 + 5X_2 - 3X_3 + 4X_n)\right) \\ &= \frac{1}{8}[2E(X_1) + 5E(X_2) - 3E(X_3) + 4E(X_n)] = \frac{1}{8}(2\mu + 5\mu - 3\mu + 4\mu) = \frac{1}{8}(8\mu) = \mu \end{aligned}$$

- Step 2: Find the Variance of $\tilde{\mu}$ - $V(\tilde{\mu})$

$$V(\tilde{\mu}) = V\left(\frac{1}{8}(2X_1 + 5X_2 - 3X_3 + 4X_n)\right) = \frac{1}{64}[4V(X_1) + 25V(X_2) + 9V(X_3) + 16V(X_n)] = \frac{1}{64}(54\sigma^2) = 0,84\sigma^2$$

- Step 3: Find the MSE of $\hat{\mu}$ - $MSE(\hat{\mu})$

$$MSE(\tilde{\mu}) = V(\tilde{\mu}) + bias(\tilde{\mu}) = 0,84\sigma^2$$

- Step 4: Calculate the Relative Efficiency

$$eff(\hat{\mu}, \tilde{\mu}) = \frac{MSE(\hat{\mu})}{MSE(\tilde{\mu})} = \frac{0,30\sigma^2}{0,84\sigma^2} = 0,36 < 1$$

$\therefore \hat{\mu}$ is more efficient than $\tilde{\mu}$

Example 1

b) Population $N(\mu, \sigma)$. Show that S^2 is an unbiased estimator of σ^2 .

From the form, is known that “If X_1, X_2, \dots, X_n are iid $N(\mu, \sigma)$ then $\frac{(n-1)S^2}{\sigma^2} \sim \chi^2_{(n-1)}$ ” and that the $E(\chi^2_n) = n$. Therefore:

$$E\left(\frac{(n-1)S^2}{\sigma^2}\right) = E(\chi^2_{(n-1)}) \Leftrightarrow \left(\frac{n-1}{\sigma^2}\right)E(S^2) = n-1 \Leftrightarrow E(S^2) = \frac{(n-1) * \sigma^2}{n-1} = \sigma^2$$

$\therefore S^2$ is an unbiased estimator of σ^2

Example 1

- c) Population $N(\mu, \sigma)$. Show that $M_2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2$ is an asymptotically unbiased estimator of σ^2 and derive its bias.

$$S^2 = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n - 1}$$

Therefore, $M_2 = \frac{n-1}{n} S^2$ and

$$E(M_2) = E\left(\frac{n-1}{n} S^2\right) = \left(\frac{n-1}{n}\right) E(S^2) = \frac{n-1}{n} \sigma^2 \xrightarrow{n \rightarrow +\infty} \sigma^2$$

M_2 is an asymptotically unbiased estimator of σ^2 . As n increases, the expected value of M_2 approaches the parameter's true value.

$$\text{bias}(M_2) = E(M_2) - \sigma^2 = \frac{n-1}{n} \sigma^2 - \sigma^2 = \left(\frac{n-1-n}{n}\right) \sigma^2 = \frac{-\sigma^2}{n}$$

$$\lim_{n \rightarrow +\infty} \text{bias}(M_2) = \lim_{n \rightarrow +\infty} \frac{-\sigma^2}{n} = 0$$

\therefore For large n , M_2 is approximately unbiased.

Example 2

Population $N(\mu, \sigma)$. Consider the two estimators of σ^2 :

$$S^2 \text{ and } M_2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2.$$

a) Show that $V(S^2) = \frac{2\sigma^4}{n-1}$ and $V(M_2) = \frac{n-1}{n^2} 2\sigma^4$

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi^2_{(n-1)} \quad V(\chi^2_{(n)}) = 2n$$

$$V\left(\frac{(n-1)S^2}{\sigma^2}\right) = V(\chi^2_{(n-1)}) \Leftrightarrow \left(\frac{n-1}{\sigma^2}\right)^2 V(S^2) = 2(n-1) \Leftrightarrow V(S^2) = \frac{2(n-1) * \sigma^4}{(n-1)^2} = \frac{2\sigma^4}{n-1}$$

$$M_2 = \frac{n-1}{n} S^2$$

$$V(M_2) = V\left(\frac{(n-1)}{n} S^2\right) = \frac{(n-1)^2}{n^2} V(S^2) = \frac{(n-1)^2}{n^2} * \frac{2\sigma^4}{n-1} = \frac{n-1}{n^2} 2\sigma^4$$

Example 2

b) Considering that n is large, determine the relative efficiency of S^2 and M_2 .

$$V(S^2) = \frac{2\sigma^4}{n-1} \quad V(M_2) = \frac{n-1}{n^2} 2\sigma^4$$

$$eff(M_2, S^2) = \frac{V(M_2)}{V(S^2)} = \frac{\frac{n-1}{n^2} 2\sigma^4}{\frac{2\sigma^4}{n-1}} = \frac{n-1}{n^2} 2\sigma^4 * \frac{n-1}{2\sigma^4} = \frac{(n-1)^2}{n^2} < 1$$

\therefore For large samples, M_2 is more efficient than S^2 .

Exercise 4

Population $N(\mu, \sigma)$.

a) Show that $E(\tilde{\sigma}^2) = \sigma^2$.

$$\tilde{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2$$

From the first page of the form - 1st Theorem: $X_i \sim N(\mu, \sigma)$ then $\sum_{i=1}^n \left(\frac{X_i - \mu}{\sigma}\right)^2 \sim \chi_n^2$

From the last page of the form – Mean and Variance of χ_n^2 : $E(\chi_n^2) = n$ $V(\chi_n^2) = 2n$

$$\begin{aligned} E \left[\sum_{i=1}^n \left(\frac{X_i - \mu}{\sigma} \right)^2 \right] &= E(\chi_n^2) = n \Leftrightarrow E \left[\frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu)^2 \right] = n \Leftrightarrow \frac{1}{\sigma^2} E \left[\sum_{i=1}^n (X_i - \mu)^2 \right] = n \\ \Leftrightarrow E \left[\sum_{i=1}^n (X_i - \mu)^2 \right] &= n\sigma^2 \Leftrightarrow \frac{1}{n} E \left[\sum_{i=1}^n (X_i - \mu)^2 \right] = \sigma^2 \Leftrightarrow E \left[\frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2 \right] = \sigma^2 \\ &\Leftrightarrow E(\tilde{\sigma}^2) = \sigma^2 \end{aligned}$$

$\therefore \tilde{\sigma}^2$ is an unbiased estimator of σ^2 .

Exercise 4

b) Compare the efficiency of $\tilde{\sigma}^2$ and $\hat{\sigma}^2 = S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$.

Since both $\tilde{\sigma}^2$ and S^2 are unbiased, we simply need to compare their variances to determine which one is more efficient. Determine $V(\tilde{\sigma}^2) = V\left[\frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2\right]$.

$$\begin{aligned} V\left[\sum_{i=1}^n \left(\frac{X_i - \mu}{\sigma}\right)^2\right] &= V(\chi_n^2) = 2n \Leftrightarrow V\left[\frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu)^2\right] = 2n \Leftrightarrow \frac{1}{\sigma^4} V\left[\sum_{i=1}^n (X_i - \mu)^2\right] = 2n \\ \Leftrightarrow V\left[\sum_{i=1}^n (X_i - \mu)^2\right] &= 2n\sigma^4 \Leftrightarrow \frac{1}{n^2} V\left[\sum_{i=1}^n (X_i - \mu)^2\right] = \frac{2n\sigma^4}{n^2} \Leftrightarrow V\left[\frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2\right] = \frac{2\sigma^4}{n} \\ &\Leftrightarrow V(\tilde{\sigma}^2) = \frac{2\sigma^4}{n} \end{aligned}$$

We have showned that $V(S^2) = \frac{2\sigma^4}{n-1}$ in example 2.

$$eff(\tilde{\sigma}^2, S^2) = \frac{V(\tilde{\sigma}^2)}{V(S^2)} = \frac{\frac{2\sigma^4}{n}}{\frac{2\sigma^4}{n-1}} = \frac{2\sigma^4}{n} * \frac{n-1}{2\sigma^4} = \frac{n-1}{n} < 1$$

$\therefore V(\tilde{\sigma}^2) < V(S^2)$,
therefore $\tilde{\sigma}^2$ is more
efficient than S^2 .