Specification Error: Omitted and Extraneous Variables

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Omitted variable bias. Suppose that the "correct" model is

$$y = \alpha + \beta_1 X_1 + \beta_2 X_2 + \varepsilon$$

If we estimate

$$y = a + b_1 X_1 + b_2 X_2 + e$$

we know that $E(b_1) = \beta_1$ and $E(b_2) = \beta_2$ i.e. the regression coefficients are unbiased estimators of the population parameters.

Suppose, however, the researcher mistakenly believes

$$y = \alpha^* + \beta^*_1 X_1 + \varepsilon^*$$

and therefore estimates

$$y = a^* + b^*_1 X_1 + e^*$$

i.e. X2 is mistakenly omitted from the model. How does b_1 (the regression estimate from the correctly specified model) compare to b_1^* (the regression estimate from the mis-specified model)? What is $E(b_1^*)$? Is it a biased or unbiased estimator of β_1 ? If biased, how is it biased?

Note that b₁*

$$= \frac{\hat{C}ov(X_{1}, Y)}{\hat{V}(X_{1})}$$

$$= \frac{\hat{C}ov(X_{1}, a + b_{1}X_{1} + b_{2}X_{2} + e)}{\hat{V}(X_{1})}$$

$$= \frac{\hat{C}ov(X_{1}, a) + b_{1}\hat{C}ov(X_{1}, X_{1}) + b_{2}\hat{C}ov(X_{1}, X_{2}) + \hat{C}ov(X_{1}, e)}{\hat{V}(X_{1})}$$

$$= \frac{0 + b_{1}\hat{V}(X_{1}) + b_{2}\hat{C}ov(X_{1}, X_{2}) + 0}{\hat{V}(X_{1})}$$

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Single Proof of the content of the

Formula for bivariate regression coefficient

Substitute the formula for Y from the correctly specified model

Expectations rules: Cov(a+b,c+d) = Cov(a,c) + Cov(a,d) + Cov(b,c) + Cov(b,d)

Recall that Cov(variable, constant) = 0. Also, X's are uncorrelated with the residuals.

Simplify expression.

If your eyes glaze over when looking at equations, just make sure you get the conclusion. If X2 has mistakenly been omitted from the model, then, taking expectations, we get

$$E(b_1^*) = \beta_1 + \beta_2 \frac{\sigma_{12}}{\sigma_1^2}$$

Very Important: Hence, b_1^* is a biased estimator of β_1 . Further, this bias will not disappear as sample size gets larger, so the omission of a variable from a model also leads to an inconsistent estimator. In effect, x1 gets credit (or blame) for the effects of the variables that have been omitted from the model.

Note that there are two conditions under which b_1^* will not be biased:

- $\beta_2 = 0$. Of course, if $\beta_2 = 0$, this means that the model is not mis-specified, i.e. X2 does not belong in the model because it has no effect on Y.
- $\sigma_{12} = 0$. That is, if the 2 X's are uncorrelated, then omitting one does not result in biased estimates of the effect of the other.

Example 1. I will construct a data set where b1 = 3, b2 = 2, and x1 and x2 have a correlation of .5. The standard deviation of x1 is 4 and the standard deviation of x2 is 4. We will see what happens if x2 is omitted from the model.

```
. clear all
. matrix input corr = (1,.5,0 \setminus .5,1,0 \setminus 0,0,1)
. matrix input sds = (4\4\10)
. corr2data x1 x2 e, corr(corr) sd(sds) n(500)
. gen y = 3*x1 + 2*x2 + e
. corr y x1 x2
(obs=500)
          y x1 x2
        y | 1.0000
        x1 | 0.7960 1.0000
        x2 | 0.6965 0.5000 1.0000
. corr y x1 x2, cov
(obs=500)
               y x1 x2
        y | 404
x1 | 64 16
x2 | 56 8
        У
```

•	*	Correct	regres	sion

. reg y x1 x2

Source	SS	df	MS		Number of obs F(2, 497)	
Model Residual Total	151696 49899.9993 201595.999	497 100.	17.9998 402413 		Prob > F R-squared Adj R-squared Root MSE	= 0.0000 = 0.7525
У	Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
x1	3 2 -4.41e-09	.1294885 .1294885 .4481125	23.17 15.45 -0.00	0.000 0.000 1.000	2.745588 1.745588 8804284	3.254412 2.254412 .8804284

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. * Omitted variable bias

. reg y x1

Source	SS	df	MS		Number of obs F(1, 498)		500 861.41
Model Residual	127744 73851.9991	1 498	127744 148.297187		Prob > F R-squared	= =	0.0000 0.6337
Total	201595.999	499	403.999998		Adj R-squared Root MSE		0.6329
У	Coef.	Std. E	Err. t	P> t	[95% Conf.	Int	terval]
x1 _cons	4 7.29e-08	.13628		0.000	3.732231 -1.070006	_	.267769

We see that, when x2 is omitted from the model, the effect of x1 is over-estimated in this case. (In other situations it could be under-estimated). To confirm that Stata got it right,

$$b_1^* = b_1 + b_2 \frac{\hat{C}ov(X_1, X_2)}{\hat{V}(X_1)} = 3 + 2\frac{8}{16} = 4$$

Example 2. Here is an example of a special case where omitting a variable does NOT result in omitted variable bias. I construct a data set similar to what we had before, except x1 and x2 are uncorrelated.

- . clear all . matrix input corr = (1,0,0,0,1,0,0,0,1). matrix input sds = $(4\10)$. corr2data x1 x2 e, corr(corr) sd(sds) n(500) (obs 500) . gen y = 3*x1 + 2*x2 + e. corr y x1 x2 (obs=500) y | 1.0000 0.6838 1.0000 0.4558 **0.0000** 1.0000 x1 |
- . * Correct regression

x2

. reg y x1 x2

Source	SS	df	MS		Number of obs	=	500
	+				F(2, 497)	=	516.88
Model	103792	2	51896.0002		Prob > F	=	0.0000
Residual	49899.9994	497	100.402413		R-squared	=	0.6753
	+				Adj R-squared	=	0.6740
Total	153692	499	308		Root MSE	=	10.02
у	 Coef.	Std.	Err. t	P> t	[95% Conf.	Tni	tervall
x 1	3	.1121	403 26.7	5 0.000	2.779672	3	.220328
x2	2	.1121	403 17.8	3 0.000	1.779672	2	.220328
_cons	-4.71e-08	.4481	125 -0.0	0 1.000	8804285	. 8	3804284

. * X2 omitted but no bias in this case

. reg y x1

	Source	SS	df		MS		Number of obs F(1, 498)		500 437.27
_	Model Residual	71856.0006 81835.9992	1 498		56.0006 .329316		Prob > F R-squared	= =	0.0000 0.4675
_	Total	153692	499		308		Adj R-squared Root MSE		12.819
-	У	Coef.	Std.	Err.	t	P> t	[95% Conf.	In	terval]
	x1 _cons	3.71e-08	.1434		20.91	0.000	2.718128 -1.12636		.281872 1.12636

Inclusion of extraneous variables. Suppose that the "correct" model is

$$y = \alpha + \beta_1 X_1 + \varepsilon$$

If we estimate

$$y = \alpha + b_1 X_1 + e$$

we know that $E(b_1) = \beta_1$, i.e. the regression coefficients is an unbiased estimators of the population parameter.

Suppose, however, the researcher mistakenly believes

$$y = \alpha^* + \beta^*_{1}X_{1} + \beta^*_{2}X_{2} + \varepsilon^*$$

and therefore estimates

$$y = a^* + b_1^* X_1 + b_2^* X_2 + e^*$$

i.e. X2 is mistakenly added to the model. How does b_1 (the regression estimate from the correctly specified model) compare to b_1 * (the regression estimate from the mis-specified model)? What is $E(b_1^*)$? Is it a biased or unbiased estimator of β_1 ? If biased, how is it biased?

Here is an informal proof: We can think of the "correct" model as being a special case of the "incorrect" model, where $\beta_2 = 0$. It will therefore be the case that $E(b_1^*) = \beta_1$, and $E(b_2^*) = 0$. Hence, addition of extraneous variables does not lead to biased coefficients.

However, adding extraneous (or "junk") variables to the model will result in inflated standard errors and all the problems they create. Recall that, in the two IV case,

$$s_{b_k} = \sqrt{\frac{1 - R_{Y12}^2}{(1 - R_{12}^2) * (N - K - 1)}} * \frac{s_y}{s_{X_k}}$$

As the formula suggests, adding irrelevant variables will tend not to increase the numerator, because irrelevant variables will not substantially increase R^2 . However, irrelevant variables will

tend to increase the denominator. The tolerance will be smaller $(1 - R^2_{12})$ and N-K-1 will be smaller.

Example 3. This is similar to the first example, except that x2 has no effect on y.

- . * Correct regression
- . reg y x1

Source Model Residual Total	71856.0006 49899.9991	498 1	1856.0006 00.200801		Number of obs F(1, 498) Prob > F R-squared Adj R-squared Root MSE	= 717.12 = 0.0000 = 0.5902 = 0.5893
У	Coef.	Std. Er	r. t	P> t	[95% Conf.	Interval]
x1 _cons					2.779895 8795398	
	s variable add	 ed				
. reg y x1 x2 Source	ss	df	MS		Number of obs	= 500
Source Model Residual	71856.0006 49899.9991	2 3 497 1	5928.0003 00.402413		F(2, 497) Prob > F R-squared Adj R-squared	= 357.84 = 0.0000 = 0.5902 = 0.5885
Source Model	71856.0006 49899.9991 	2 3 497 1 499 2	5928.0003 00.402413 	 P> t	F(2, 497) Prob > F R-squared	= 357.84 = 0.0000 = 0.5902 = 0.5885 = 10.02

As you can see the coefficient for x1 did not change but the standard error increased and the t value went down.

Appendix: Another example of omitted variable bias

EXAMPLE: Consider our income/education/job experience example:

- . use http://www3.nd.edu/~rwilliam/statafiles/reg01.dta, clear
- . corr educ jobexp income, cov

(obs=20)

	educ	jobexp	income
jobexp	20.05 -2.61316 37.0676		95.8119

. reg income educ jobexp

Source	SS	df	MS		Number of obs	
Model Residual	1538.22521 282.200265		.112605 6000156		Prob > F R-squared Adj R-squared	= 0.0000 = 0.8450
Total	1820.42548	19 95.	8118671		Root MSE	= 4.0743
income	Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
educ jobexp _cons	1.933393 .6493654 -7.096855	.2099494 .1721589 3.626412	9.21 3.77 -1.96	0.000 0.002 0.067	1.490438 .2861417 -14.74792	2.376347 1.012589 .5542052

Note that, when both EDUC and JOBEXP are in the equation, $b_1 = 1.933393$, $b_2 = .649365$, Cov(Educ, Jobexp) = -.2613, V(Educ) = 20.05, V(Jobexp) = 29.818. Hence, if we omit Jobexp from the model, the new coefficient b_1^* is

$$b_1^* = b_1 + b_2 \frac{\hat{C}ov(X_1, X_2)}{\hat{V}(X_1)} = 1.933393 + .649365 \frac{-2.613}{20.050} = 1.848765$$

Stata confirms that this is correct:

. reg income educ

Source SS df MS Number of obs:	
	= 45.21
Model 1302 05369	
MOGCI 1302.03307 1 1302.03307 FIOD > 1	= 0.0000
Residual 518.371789	= 0.7152
Adj R-squared	= 0.6994
Total 1820.42548	= 5.3664
income Coef. Std. Err. t P> t [95% Conf.]	Interval]
educ 1.84876 .2749479 6.72 0.000 1.271116 cons 2.137446 3.523734 0.61 0.552 -5.265645	2.426404
	2.010007

Or, if we instead omit EDUC from the equation, for b₂* we get

$$b_1^* = b_2 + b_1 \frac{\hat{C}ov(X_1, X_2)}{\hat{V}(X_2)} = .649365 + .1.933393 \frac{-2.613}{29.818} = .479928616$$

Stata again confirms this:

. reg income jobexp

Source	SS	df		MS		Number of obs F(1, 18)		20 1.39
Model Residual Total	130.495675 1689.9298 1820.42548	1 18 19	93.8	.495675 8849889 		Prob > F R-squared Adj R-squared Root MSE	= =	0.2538 0.0717 0.0201
income	Coef.	Std.	Err.	t	P> t	[95% Conf.	In	terval]
jobexp _cons	.4799311 18.34387	.4070		1.18	0.254	3753106 6.606476		.335173

If we assume that the model with both EDUC and JOBEXP is correct, omitting one or the other results in the effects of the remaining variable being mis-estimated.

In more complicated models with omitted variables, it will continue to be the case that observed effects represent a confounding of the actual effect with other sources of association.