

International Investing

Diversification and Beyond

Soohun Kim
Robert Korajczyk
Andreas Neuhierl

KAIST
Northwestern University
Purdue University

Purdue Brownbag

Motivation

- Is international investing useful? In what sense?
- Conventional wisdom suggests “good for diversification”
- Recent studies question this (Bae, Elkamhi, and Simutin (2019)), due to **increased integration**
- Often studies (e.g. Griffin (2002); Fama and French (2017)) *assume* a version of Fama and French (1993) for each country
- Aim for a more agnostic approach to factor structure
- **Develop latent factor models in a cross-country setting**

Literature

- International Diversification
 - Grubel (1968); Solnik (1974); Quinn and Voth (2008); Bae et al. (2019) and many more
- Integration (broadly defined)
 - Cho et al. (1986); Harvey (1991); De Santis and Gerard (1997); Bekaert and Harvey (1995); Karolyi and Stulz (2003); Carrieri et al. (2007); Pukthuanthong and Roll (2009); Bekaert et al. (2011); Chaieb et al. (2021) and many more
- Factor models
 - Chamberlain and Rothschild (1983); Connor and Korajczyk (1986, 1988); Connor et al. (2012); Fan et al. (2016); Kelly et al. (2019); Andreou et al. (2019); Kim et al. (2021), Zaffaroni (2025), Ferson et al. (2025) and many more

Simple Example I

- Excess returns follow a factor model
- Investor is allocated to country 1 – “home country” (USA)

$$R_{1it} = \underbrace{\beta_{1i}(\lambda_1^c + f_t)}_{\text{common to the pair}} + \underbrace{\delta_{1i}(\lambda_1^s + g_{1t})}_{\text{country specific}} + e_{1it}$$

- Similarly for country 2

$$R_{2it} = \underbrace{\beta_{2i}(\lambda_2^c + f_t)}_{\text{common to the pair}} + \underbrace{\delta_{2i}(\lambda_2^s + g_{2t})}_{\text{country specific}} + e_{2it}$$

Simple Example II

- How should an investor allocate toward country 2 (if possible)?
- Exposure to foreign country risk g_{2t} ?
 - Expand investment opportunity set
- Possibly exploit different prices of risk for $f_t - \lambda_1^c$ v.s. λ_2^c ?
 - Segmentation premia
- Throughout assume large economies, i.e. $N_1, N_2 \rightarrow \infty$
- How can we achieve this?
 - Single common factor, single country specific factor
 - Factors loadings are known
 - Factor loadings are cross-sectionally orthogonal ($\beta' \delta = 0$)

Simple Example III

- Consider $w_i^s = \frac{\delta_{2i}}{N}$, then (up to normalization)

$$\sum_{i=1}^N w_i^s R_{2it} = \sum_{i=1}^N \frac{\delta_{2i}\beta_{2i}}{N} (\lambda_2^c + f_t) + \sum_{i=1}^N \frac{\delta_{2i}^2}{N} (\lambda_2^s + g_{2t}) + \sum_{i=1}^N \frac{\delta_{2i}e_{2it}}{N} \xrightarrow{P} (\lambda_2^s + g_{2t}).$$

- Interesting if $\lambda_2^s \neq 0$ and g_t not spanned by f_t
- Consider $w_i^c = \frac{\beta_{2i}}{N}$, then (up to normalization)

$$\sum_{i=1}^N w_i^c R_{2it} = \sum_{i=1}^N \frac{\beta_{2i}^2}{N} (\lambda_2^c + f_t) + \sum_{i=1}^N \frac{\beta_{2i}\delta_{2i}}{N} (\lambda_2^s + g_{2t}) + \sum_{i=1}^N \frac{\beta_{2i}e_{2it}}{N} \xrightarrow{P} (\lambda_2^c + f_t)$$

- Segmentation premia if $\lambda_1 \neq \lambda_2$ (segmentation portfolio)

Challenges

- Factors not observed – f_t and g_t not given
- Factor loadings not observed – β and δ not given
- Generally, $\beta'\delta \neq 0$ even if we assume f and g known
 - Local Fama and French (1993) factors and their global counterparts strongly correlated
- **Can we still implement these portfolios?**
- **What are their empirical properties?**

This Paper

- Cross-country *latent factor model*
- Develop feasible estimation for
 - Common factors f_t
 - Country specific factors g_t
 - Common factor loadings β_1 and β_2
 - Country specific loadings δ_2
- Large N , but fixed T
- Implement portfolios empirically

Preview of Results

- **Country-specific portfolios** (exposure to foreign-specific risk)
 - Sharpe ratios from 0.49 (Japan) to 1.16 (Germany) among G7
 - Nearly **quadruple** the US market Sharpe ratio when combined
 - Virtually uncorrelated with US market
- **Segmentation portfolios** (exploiting risk premia differences)
 - Sharpe ratios 0.67–1.32 even among G7 countries
 - Challenges view of integrated developed markets
- **Naive diversification** (country indices) yields SR of only 0.46–0.54
- **Combined approach:** SR up to 2.23

Setup: Returns within a Country I

- Consider a pair of two countries, indexed by $g = 1, 2$
- Returns follow a K_g ($= K^c + K_g^s$)-factor model

$$\mathbf{R}_g = \underbrace{\mathbf{B}_g (\boldsymbol{\lambda}_g^c \mathbf{1}'_T + \mathbf{F}')}_{\text{Common factors}} + \underbrace{\mathbf{D}_g (\boldsymbol{\lambda}_g^s \mathbf{1}'_T + \mathbf{G}'_g)}_{\text{Country-Specific factors}} + \mathbf{E}_g$$

- Can we separately identify common and country specific factors?
- Allow for segmentation and country specific risks
- Show how to isolate
 - Possible differences in $\boldsymbol{\lambda}_g^c$
 - Distinct premium for each country $\boldsymbol{\lambda}_g^s$

Setup: Returns within a Country II

- Consider a pair of two countries, indexed by $g = 1, 2$
- Returns follow a $K_g (= K^c + K_g^s)$ -factor model

$$\mathbf{R}_g = \underbrace{\mathbf{B}_g (\boldsymbol{\lambda}_g^c \mathbf{1}'_T + \mathbf{F}')}_{\text{Common factors}} + \underbrace{\mathbf{D}_g (\boldsymbol{\lambda}_g^s \mathbf{1}'_T + \mathbf{G}'_g)}_{\text{Country-Specific factors}} + \mathbf{E}_g$$

- Premia, factors, and factor loadings **not observed**
- If we knew loadings or could estimate them all problems solved
- Intuition as in simple example (with more complicated notation)

Characteristic Based Factor Model

Characteristics map to factor loadings $[B_g \ D_g]$:

$$\underbrace{B_g = X_g \Theta_g^c + \Gamma_g^c}_{\text{common factor loadings}} \quad \underbrace{D_g = X_g \Theta_g^s + \Gamma_g^s}_{\text{country specific loadings}}$$

- Use $X_g \Theta_g^{c,s}$ instead of B_g and D_g
- Learn loadings from cross-section rather than time series
- Γ_g^c and Γ_g^s sources of beta not attributable to characteristics
- Mapping can be different across countries and over time
- Common factors are orthogonal to country-specific factors (identification restriction)

Estimation Roadmap

Recall Objectives

- **Isolate Country Specific Risk**
 - Learn Θ_g^s for each country to obtain D_g^s
- **Segmentation Portfolio**
 - Learn Θ_g^c for each country to obtain B_g
 - Rather than pursue separate estimation, joint estimation easier

Estimation Steps

- ① Estimate common and country specific factors (**F** and **G**)
- ② Estimate common and country specific factor loadings (Θ_g^c and Θ_g^s)
- ③ Form portfolios
 - Segmentation portfolio ($r^{\Delta\lambda} = \mathbf{w}'_1 \mathbf{R}_1 + \mathbf{w}'_2 \mathbf{R}_2$)
 - Country specific portfolio ($r^{cs} = \mathbf{w}'_2 \mathbf{R}_2$)

Step 1 – Estimate All Factors

- De-mean returns ($\mathbf{J}_T = \mathbf{I}_T - \frac{1}{T} \mathbf{1}_T \mathbf{1}'_T$)
- Information in characteristics (Fan et al. (2016); Kim et al. (2021))
 $(\mathbf{P}_g = \mathbf{X}_g (\mathbf{X}'_g \mathbf{X}_g)^{-1} \mathbf{X}'_g)$

$$\begin{aligned}\widehat{\mathbf{R}}_g &\equiv \mathbf{P}_g \mathbf{R}_g \mathbf{J}_T \\ &= (\mathbf{P}_g \mathbf{X}_g \Theta_g^c + \mathbf{P}_g \Gamma_g^c) \mathbf{F}' \mathbf{J}_T + (\mathbf{P}_g \mathbf{X}_g \Theta_g^s + \mathbf{P}_g \Gamma_g^s) \mathbf{G}'_g \mathbf{J}_T + \mathbf{P}_g \mathbf{E}_g \mathbf{J}_T \\ &\approx \mathbf{X}_g \Theta_g^c \mathbf{F}' \mathbf{J}_T + \mathbf{X}_g \Theta_g^s \mathbf{G}'_g \mathbf{J}_T\end{aligned}$$

Consistency for $[\mathbf{F}' \mathbf{J}_T \mathbf{G}'_g \mathbf{J}_T]$

- Apply PCA to $\frac{\widehat{\mathbf{R}}'_g \widehat{\mathbf{R}}_g}{N_g}$
- eigenvectors $\rightarrow^P \mathcal{O}_g [\mathbf{F}' \mathbf{J}_T \mathbf{G}'_g \mathbf{J}_T]$ for a $(K_g \times K_g)$ matrix \mathcal{O}_g
- F and G are not separately identified here

Step 2: Separate $\mathbf{F}'\mathbf{J}_T$ and $\mathbf{G}'_g\mathbf{J}_T \quad \mathbf{I}$

- Consistent estimators *for each country*

$$\text{eigenvectors of } \frac{\widehat{\mathbf{R}}'_1 \widehat{\mathbf{R}}_1}{N_1} = [\widehat{\mathbf{FG}}]_1' \xrightarrow{P} \mathcal{O}_1 [\mathbf{F}'\mathbf{J}_T \mathbf{G}'_1\mathbf{J}_T]$$

and

$$\text{eigenvectors of } \frac{\widehat{\mathbf{R}}'_2 \widehat{\mathbf{R}}_2}{N_2} = [\widehat{\mathbf{FG}}]_2' \xrightarrow{P} \mathcal{O}_2 [\mathbf{F}'\mathbf{J}_T \mathbf{G}'_2\mathbf{J}_T]$$

- Note that $\mathbf{F}'\mathbf{J}_T$ are common!
 - Exploit commonality
- ⇒ **Canonical correlation analysis**

Step 2: Separate $\mathbf{F}'\mathbf{J}_T$ and $\mathbf{G}'_g\mathbf{J}_T$ II

- Under commonality there exists \mathbf{w}_1 and \mathbf{w}_2 such that

$$\rho \left(\widehat{[\mathbf{FG}]}_1 \mathbf{w}_1, \widehat{[\mathbf{FG}]}_2 \mathbf{w}_2 \right) \xrightarrow{P} 1$$

- This correlation is known as canonical correlation
- Find $\mathbf{W}_1 (K_1 \times K_1)$ and $\mathbf{W}_2 (K_2 \times K_2)$ such that

$$\max_{\mathbf{W}_1, \mathbf{W}_2} \text{tr} \left(\text{diag} \left(\mathbf{W}'_1 \left(\frac{\widehat{[\mathbf{FG}]}'_1 \widehat{[\mathbf{FG}]}_2}{T} \right) \mathbf{W}_2 \right) \right) \quad \text{subject to}$$

$$\mathbf{W}'_1 \left(\frac{\widehat{[\mathbf{FG}]}'_1 \widehat{[\mathbf{FG}]}_1}{T} \right) \mathbf{W}_1 = \mathbf{I} \quad \text{and} \quad \mathbf{W}'_2 \left(\frac{\widehat{[\mathbf{FG}]}'_2 \widehat{[\mathbf{FG}]}_2}{T} \right) \mathbf{W}_2 = \mathbf{I}$$

Step 2: Separate $\mathbf{F}'\mathbf{J}_T$ and $\mathbf{G}'_g\mathbf{J}_T$ III

- From canonical correlations, first K_c elements for \mathbf{F} rest for \mathbf{G}

$$\mathbf{W}'_g [\widehat{\mathbf{FG}}]'_g = [\mathbf{W}^c_g \mathbf{W}^s_g]' [\widehat{\mathbf{FG}}]'_g = [\widehat{\mathbf{F}}'_g \widehat{\mathbf{G}}'_g]$$

- Separately identified \mathbf{F} and \mathbf{G}
- **Consistency for $\mathbf{F}'\mathbf{J}_T$ and $\mathbf{G}'_g\mathbf{J}_T$**

$$[\widehat{\mathbf{F}}'_g \widehat{\mathbf{G}}'_g] \xrightarrow{P} [\mathcal{S}^{c'} \mathbf{F}' \mathbf{J}_T \mathcal{S}_g^{s'} \mathbf{G}'_g \mathbf{J}_T]$$

where

$$\mathcal{S}^c \mathcal{S}^{c'} = \Sigma_F^{-1} \quad \text{and} \quad \mathcal{S}_g^s \mathcal{S}_g^{s'} = \Sigma_{G_g}^{-1}$$

- Factors are normalized to have unit variance

Step 3: Estimate Loadings through Θ_g^c and Θ_g^s

De-meanned & projected returns

$$\begin{aligned}\widehat{\mathbf{R}}_g &\equiv \mathbf{P}_g \mathbf{R}_g \mathbf{J}_T \\ &\approx \mathbf{X}_g \Theta_g^c \mathbf{F}' \mathbf{J}_T + \mathbf{X}_g \Theta_g^s \mathbf{G}' \mathbf{J}_T\end{aligned}$$

- Use consistent factor estimates from steps 1 & 2
 - $\widehat{\mathbf{F}}'_g \xrightarrow{P} \mathcal{S}^{c'} \mathbf{F}' \mathbf{J}_T$
 - $\widehat{\mathbf{G}}'_g \xrightarrow{P} \mathcal{S}^{s'} \mathbf{G}' \mathbf{J}_T$
- Now can estimate *factor loading functions*
 - Θ_g^c for common factor loadings
 - Θ_g^s for country specific factor loadings

Step 3: Estimate Loadings through Θ_g^c and Θ_g^s II

Solve quadratic program

$$(\widehat{\Theta}_g^c, \widehat{\Theta}_g^s) = \arg \min_{(\Theta_g^c, \Theta_g^s)} \|\widehat{\mathbf{R}}_g - (\mathbf{X}_g \Theta_g^c) \widehat{\mathbf{F}}'_g - (\mathbf{X}_g \Theta_g^s) \widehat{\mathbf{G}}'_g\|$$

- $\widehat{\Theta}_g^c \rightarrow^P \Theta_g^c (\mathcal{S}^{c'})^{-1}$ where $\mathcal{S}^c \mathcal{S}^{c'} = \Sigma_F^{-1}$
- $\widehat{\Theta}_g^s \rightarrow^P \Theta_g^s (\mathcal{S}_g^{s'})^{-1}$ where $\mathcal{S}_g^s \mathcal{S}_g^{s'} = \Sigma_{G_g}^{-1}$
- Solution: ▶ Closed Form
- Consistent estimators for
 - Factors: \mathbf{F}, \mathbf{G}
 - Factor loadings: Θ_g^c and Θ_g^s

⇒ Build portfolios

Step 4a: Country-Specific Portfolio I

- Original motivation of international investing
- Estimated loadings

$$\widehat{\mathbf{B}}_g := \mathbf{X}_g \widehat{\boldsymbol{\Theta}}_g^c \quad \text{and} \quad \widehat{\mathbf{D}}_g := \mathbf{X}_g \widehat{\boldsymbol{\Theta}}_g^s$$

- Consider the portfolio weight

$$\widehat{\mathbf{w}}_g^b = \underbrace{[\widehat{\mathbf{B}} \ \widehat{\mathbf{D}}] \left([\widehat{\mathbf{B}} \ \widehat{\mathbf{D}}]' [\widehat{\mathbf{B}} \ \widehat{\mathbf{D}}] \right)^{-1}}_{\text{somewhat like } (\widehat{\beta}' \widehat{\beta})^{-1} \widehat{\beta}' \text{ in FMB}}$$

- Portfolio weights capture
 - Country specific exposures only
 - Differences in risk premia across countries

Step 4a: Country-Specific Portfolio II

- Partition the weights into *common* and *country specific*

$$\widehat{\mathbf{w}}_g^{b'} \mathbf{R}_g = \begin{bmatrix} \widehat{\mathbf{w}}_g^{c'} \mathbf{R}_g \\ \widehat{\mathbf{w}}_g^{s'} \mathbf{R}_g \end{bmatrix} \xrightarrow{P} \begin{bmatrix} \mathcal{S}^{c'} (\lambda_g^c \mathbf{1}'_T + \mathbf{F}') \\ \mathcal{S}_g^{s'} (\lambda_g^s \mathbf{1}'_T + \mathbf{G}'_g) \end{bmatrix}$$

- Exposure to **country 2 risk only**

$$\widehat{\mathbf{w}}_2^{s'} \mathbf{R}_2 \xrightarrow{P} \mathcal{S}_2^{s'} (\lambda_2^s \mathbf{1}'_T + \mathbf{G}'_2)$$

- Can be extended to multiple sources of country 2 risk
- This portfolio captures the core of international diversification
- Out-of-sample implementation

$$r_{T+1}^{cs} = \widehat{\mathbf{w}}'_{2,T} \mathbf{R}_{2,T+1}$$

- Multivariate Construction

Step 4b: Segmentation Portfolio I

- Explore $\lambda_1^c \neq \lambda_2^c$
- Could run Fama and MacBeth (1973) for each g

$$\bar{\mathbf{R}} = \widehat{\mathbf{B}}\boldsymbol{\lambda}_g^c + \widehat{\mathbf{D}}\boldsymbol{\lambda}_g^s + \mathbf{U}$$

- Estimate prices of risk $\widehat{\boldsymbol{\lambda}}_g^c$ (and $\widehat{\boldsymbol{\lambda}}_g^s$)
- Consistency only for $T \rightarrow \infty$
- T potentially small
- Project returns onto $\mathbf{X}\widehat{\boldsymbol{\Theta}}_g^{c,s}$ each period
- Recall

$$\widehat{\mathbf{w}}_g^{b'} \mathbf{R}_g = \begin{bmatrix} \widehat{\mathbf{w}}_g^{c'} \mathbf{R}_g \\ \widehat{\mathbf{w}}_g^{s'} \mathbf{R}_g \end{bmatrix} \xrightarrow{P} \begin{bmatrix} \mathcal{S}^{c'} (\lambda_g^c \mathbf{1}'_T + \mathbf{F}') \\ \mathcal{S}_g^{s'} (\lambda_g^s \mathbf{1}'_T + \mathbf{G}'_g) \end{bmatrix}$$

- Study $\widehat{\mathbf{w}}_1^c$ and $\widehat{\mathbf{w}}_2^c$ (common risks in *each country*)
- Problem **pollution by factor realization**

Step 4b: Segmentation Portfolio II

- Assume for simplicity that $K_c = 1$ (one common factor)

$$\widehat{\mathbf{w}}_2' \mathbf{r}_{2,t} - \widehat{\mathbf{w}}_1' \mathbf{r}_{1,t} \xrightarrow{P} (\lambda_2 + f_t) - (\lambda_1 + f_t) = \lambda_2 - \lambda_1 + (\mathbf{f}_t - f_t)$$

- Run regression

$$\underbrace{\widehat{\mathbf{w}}_2^{c'} \mathbf{R}_2}_{1 \times T} = \beta_0 + \beta_1 \left(\underbrace{\widehat{\mathbf{w}}_1^{c'} \mathbf{R}_1}_{1 \times T} \right) + \mathbf{e}$$

- Intercept captures difference in risk premia, i.e.

$$\widehat{\beta}_0 \xrightarrow{P} (\lambda_2^c - \lambda_1^c) \quad \text{and} \quad \widehat{\beta}_1 \xrightarrow{P} 1$$

- $\widehat{\beta}_0 > 0$ implement (out-of-sample)

$$\mathbf{r}_{T+1}^{\Delta\lambda} = \underbrace{\widehat{\mathbf{w}}_{2,T}^{c'} \mathbf{r}_{2,T+1}}_{\text{common factor in country 2}} - \underbrace{\widehat{\beta}_1 \widehat{\mathbf{w}}_{1,T}^{c'} \mathbf{r}_{1,T+1}}_{\text{common factor in country 1}}$$

- Multivariate Construction

Simulation

- We validate our method, enabling investors in Country 1 to:
 - Diversify into Country 2-specific factors
 - Capitalize on differences in risk premia
- Simulation investigates: Impact of
 - Total number of estimated factors
 - Choice between common and country-specific factors

Calibration

- Return-generating process calibrated using:
 - USA (Country 1) and Canada (Country 2) returns
- Baseline specification:
 - Six systematic factors per country
 - Three common factors across both countries
 - Three country-specific factors
- Using the most recent 60-month window, we estimate:
 - Betas
 - Residual variance
 - Common and country-specific factors

Canonical Correlations for Different Numbers of Estimated Factors

| Canonical Correlations | 1 | 2 | 3 | 4 | 5 | 6 | 7 |
|---------------------------|------|------|------|------|------|------|------|
| $K = 5$, Underestimation | 0.99 | 0.94 | 0.77 | 0.23 | 0.07 | - | - |
| $K = 6$, Correct | 0.99 | 0.96 | 0.94 | 0.33 | 0.18 | 0.06 | - |
| $K = 7$, Overestimation | 0.99 | 0.97 | 0.94 | 0.41 | 0.27 | 0.15 | 0.05 |

- Average canonical correlations over 10,000 repetitions.
- Overestimation is safer than underestimation.

Sharpe Ratios of Segmentation and Country-Specific Portfolios

| | 1 | 2 | 3 | 4 | 5 |
|---|-------|-------|-------|-------|-------|
| K_c (Common Factors) | 1 | 2 | 3 | 4 | 5 |
| K_s (Country-Specific Factors) | 5 | 4 | 3 | 2 | 1 |
| Panel A: Neither | | | | | |
| Segmentation Portfolio SR | -0.01 | -0.02 | -0.04 | -0.05 | -0.05 |
| Country-Specific Portfolio SR | 0.01 | 0.01 | 0.01 | 0.00 | -0.01 |
| Panel B: Country-Specific Premia Only | | | | | |
| Segmentation Portfolio SR | 0.01 | 0.05 | 0.04 | 0.29 | 0.46 |
| Country-Specific Portfolio SR | 0.87 | 0.87 | 0.87 | 0.68 | 0.44 |
| Panel C: Segmentation in Common Factor Premia Only | | | | | |
| Segmentation Portfolio SR | 0.55 | 0.69 | 0.79 | 0.78 | 0.78 |
| Country-Specific Portfolio SR | 0.23 | 0.14 | 0.01 | 0.01 | -0.02 |
| Panel D: Both | | | | | |
| Segmentation Portfolio SR | 0.71 | 0.86 | 0.96 | 1.06 | 1.13 |
| Country-Specific Portfolio SR | 0.92 | 0.90 | 0.88 | 0.68 | 0.44 |

- True K_c and K_s : reveals the state of the international market

Sharpe Ratios of Segmentation and Country-Specific Portfolios

| | 1 | 2 | 3 | 4 | 5 |
|---|------|------|------|------|-------|
| K_c (Common Factors) | 1 | 2 | 3 | 4 | 5 |
| K_s (Country-Specific Factors) | 5 | 4 | 3 | 2 | 1 |
| Panel B: Country-Specific Premia Only | | | | | |
| Segmentation Portfolio SR | 0.01 | 0.05 | 0.04 | 0.29 | 0.46 |
| Country-Specific Portfolio SR | 0.87 | 0.87 | 0.87 | 0.68 | 0.44 |
| Panel C: Segmentation in Common Factor Premia Only | | | | | |
| Segmentation Portfolio SR | 0.55 | 0.69 | 0.79 | 0.78 | 0.78 |
| Country-Specific Portfolio SR | 0.23 | 0.14 | 0.01 | 0.01 | -0.02 |

- Overestimating K_c :
Country-specific premia → Segmentation premia (Panel B)
- Overestimating K_s :
Segmentation premia → Country-specific premia (Panel C)

Data

- Build on Jensen et al. (2023)
- Consider 27 international markets (usa, twn, tur, tha, swe, sgp, pol, nor, mex, kor, jpn, ita, ind, hkg, gbr, fra, esp, dnk, deu, can, bra, aut, chn, chl, bel, aus, arg)
- Drop smallest 20% of the firms in local currency for each country
- Impute missing characteristics Freyberger, Höppner, Neuhierl, and Weber (2024)
- Contrary to many other studies **we do not winsorize returns**
- Excess returns from US investor perspective, i.e. returns in USD

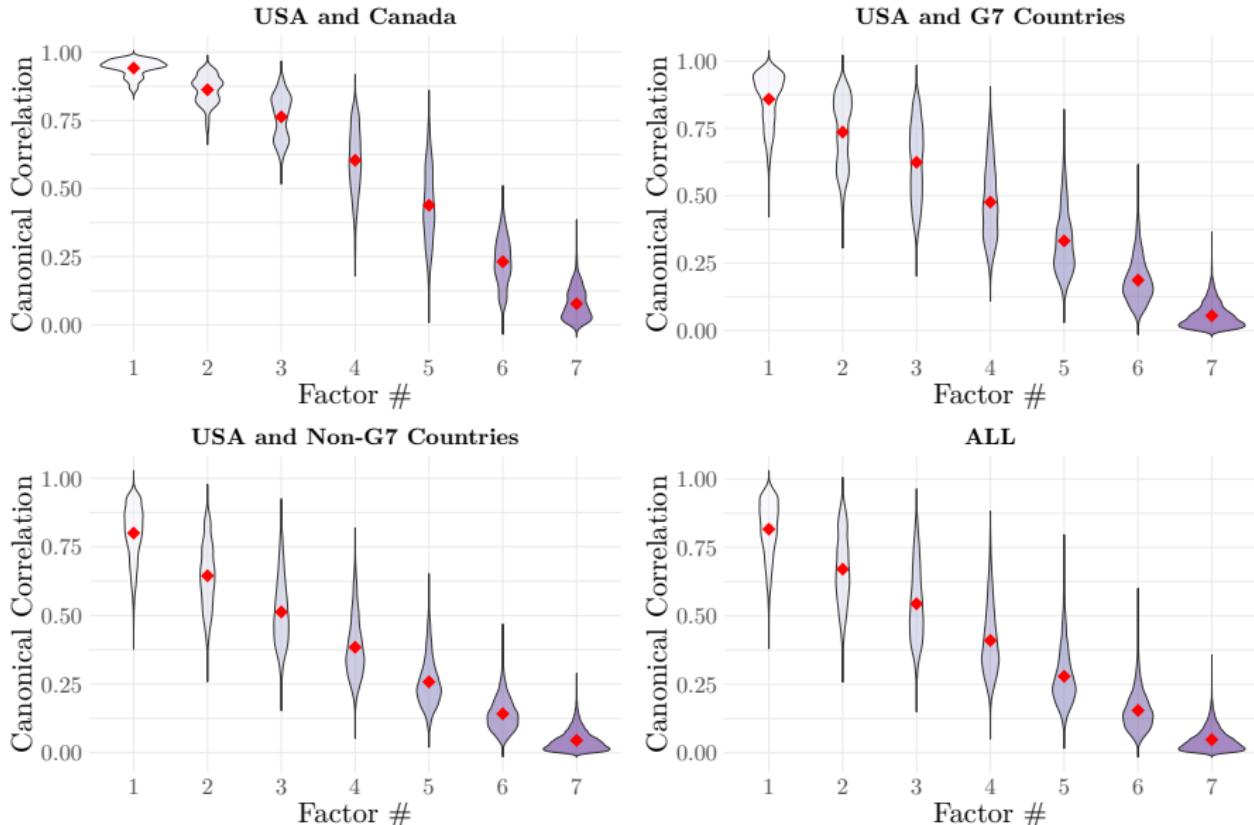
Country Overview

| Country | | | # Firms | Market Cap. | | | Cross Section | | | Excess Return | |
|---------|----------|----------|---------|-------------|---------|---------|------------------|---------|------------------|---------------|-------|
| | | | | q25 | Avg. | q75 | N _{min} | ̄N | N _{max} | ̄R | σ(̄R) |
| arg | Jul-2005 | Dec-2022 | 95 | 21.55 | 907.08 | 479.11 | 46 | 54.88 | 69 | 0.18 | 0.38 |
| aus | Feb-1986 | Dec-2022 | 3721 | 8.93 | 650.58 | 125.35 | 117 | 842.87 | 1583 | 0.08 | 0.27 |
| aut | Nov-1999 | Dec-2022 | 153 | 44.49 | 1226.49 | 1226.17 | 44 | 61.02 | 79 | 0.08 | 0.20 |
| bel | Jan-1991 | Dec-2022 | 286 | 35.52 | 1590.81 | 585.27 | 44 | 103.57 | 133 | 0.07 | 0.16 |
| bra | Dec-2006 | Dec-2022 | 336 | 323.76 | 2419.44 | 2223.03 | 61 | 143.95 | 216 | 0.09 | 0.35 |
| can | Jan-1983 | Dec-2022 | 3191 | 4.02 | 505.00 | 81.45 | 179 | 751.27 | 1392 | 0.11 | 0.23 |
| chl | Feb-1997 | Dec-2022 | 229 | 153.78 | 1623.10 | 1414.77 | 52 | 111.80 | 144 | 0.31 | 1.08 |
| chn | Jan-1995 | Dec-2022 | 4883 | 335.38 | 1255.28 | 1074.47 | 40 | 1430.51 | 3793 | 0.12 | 0.32 |
| deu | Jan-1987 | Dec-2022 | 1774 | 6.96 | 1220.62 | 195.40 | 75 | 502.61 | 830 | 0.03 | 0.18 |
| dnk | Jan-1994 | Dec-2022 | 379 | 19.62 | 813.37 | 221.08 | 65 | 126.32 | 171 | 0.08 | 0.18 |
| esp | May-1988 | Dec-2022 | 417 | 175.73 | 5091.85 | 2978.30 | 36 | 115.99 | 180 | 0.05 | 0.21 |
| fra | Jan-1987 | Dec-2022 | 1814 | 19.91 | 2386.31 | 453.87 | 97 | 480.59 | 687 | 0.07 | 0.19 |
| gbr | Feb-1986 | Dec-2022 | 5548 | 12.77 | 1382.21 | 254.38 | 258 | 1225.09 | 1819 | 0.05 | 0.20 |
| hkg | Apr-1987 | Dec-2022 | 2872 | 54.34 | 1504.17 | 608.74 | 38 | 814.20 | 1963 | 0.10 | 0.29 |
| ind | May-1989 | Dec-2022 | 5096 | 11.94 | 658.76 | 194.24 | 41 | 1298.94 | 3299 | 0.15 | 0.35 |
| ita | Jan-1987 | Dec-2022 | 781 | 72.85 | 2296.99 | 1092.67 | 43 | 192.64 | 310 | 0.02 | 0.23 |
| jpn | Jan-1987 | Dec-2022 | 5648 | 30.43 | 945.28 | 369.38 | 1107 | 2537.70 | 3242 | 0.05 | 0.22 |
| kor | Mar-1989 | Dec-2022 | 3309 | 22.52 | 442.11 | 121.12 | 91 | 1012.31 | 2027 | 0.05 | 0.34 |
| mex | Mar-2001 | Dec-2022 | 195 | 207.32 | 3095.61 | 2506.78 | 56 | 82.48 | 105 | 0.12 | 0.21 |
| nor | Jun-1993 | Dec-2022 | 664 | 39.91 | 1100.65 | 563.90 | 54 | 151.54 | 269 | 0.10 | 0.26 |
| pol | Mar-1996 | Dec-2022 | 1176 | 11.26 | 349.57 | 126.11 | 51 | 345.52 | 684 | 0.08 | 0.29 |
| sgp | Feb-1989 | Dec-2022 | 1143 | 21.55 | 620.93 | 187.04 | 42 | 370.89 | 660 | 0.07 | 0.28 |
| swe | Jan-1992 | Dec-2022 | 1336 | 9.41 | 956.02 | 265.10 | 52 | 300.46 | 653 | 0.10 | 0.24 |
| tha | Jan-1989 | Dec-2022 | 1121 | 14.93 | 327.42 | 139.11 | 88 | 376.24 | 713 | 0.12 | 0.31 |
| tur | Dec-1993 | Dec-2022 | 606 | 30.43 | 834.56 | 436.42 | 50 | 232.79 | 419 | 0.23 | 0.47 |
| twn | Jan-1989 | Dec-2022 | 2549 | 35.50 | 461.45 | 248.47 | 46 | 846.08 | 1691 | 0.08 | 0.32 |
| usa | Jan-1966 | Dec-2022 | 25603 | 62.35 | 3010.13 | 1269.86 | 1579 | 4008.30 | 6448 | 0.07 | 0.20 |

Baseline Implementation

- All results are out-of-sample
- Rolling estimation window of 60 months
- $K = 7$, i.e. seven factors in total
- Use the test by Gonçalves et al. (2025) to test for the number of common factors
- Normalize all zero-investment portfolios' in-sample standard deviation to 20% annualized
- Rank normalize all characteristics to $[0, 1]$ each period

Canonical Correlations



Country-Specific Portfolio: G7

| Country | Mean (%) | Standard Deviation (%) | t | Sharpe Ratio | Maximum Drawdown | Worst Month | Best Month | ρ_{US} |
|---------|----------|------------------------|------|--------------|------------------|-------------|------------|-------------|
| can | 14.53 | 20.48 | 3.99 | 0.71 | 50.59 | -20.06 | 25.41 | 0.06 |
| deu | 20.53 | 17.76 | 6.08 | 1.16 | 44.21 | -16.69 | 17.58 | -0.02 |
| fra | 12.85 | 18.77 | 3.60 | 0.68 | 50.01 | -14.98 | 25.88 | -0.03 |
| gbr | 17.17 | 20.81 | 4.16 | 0.82 | 40.66 | -19.28 | 19.52 | -0.03 |
| ita | 16.76 | 20.40 | 4.33 | 0.82 | 47.34 | -18.90 | 18.97 | -0.02 |
| jpn | 9.48 | 19.43 | 2.76 | 0.49 | 41.83 | -18.98 | 17.34 | 0.03 |

- Strong local factors among G7 factors
- Virtually uncorrelated with US market
- Much lower correlation than market correlation

Correlation: G7 Country-Specific Portfolios

Country Specific Factor Correlation

| | | | | | | |
|-----|---|-------|-------|-------|-------|------|
| jpn | | | | | | 1 |
| ita | | | | 1 | 0.04 | |
| gbr | | | 1 | 0.05 | -0.04 | |
| fra | | 1 | 0.05 | -0.03 | 0.02 | |
| deu | 1 | -0.01 | -0.02 | 0.04 | 0 | |
| can | 1 | -0.03 | 0.05 | 0.07 | 0.08 | 0.06 |

Correlation of Country Market Portfolios

| | | | | | | | |
|-----|---|------|------|------|------|------|------|
| usa | | | | | | 1 | |
| jpn | | | | | | 1 | 0.34 |
| ita | | | | | 1 | 0.36 | 0.57 |
| gbr | | | | 1 | 0.65 | 0.46 | 0.69 |
| fra | | | 1 | 0.76 | 0.77 | 0.41 | 0.66 |
| deu | | 1 | 0.87 | 0.73 | 0.74 | 0.41 | 0.64 |
| can | 1 | 0.66 | 0.67 | 0.72 | 0.59 | 0.37 | 0.72 |

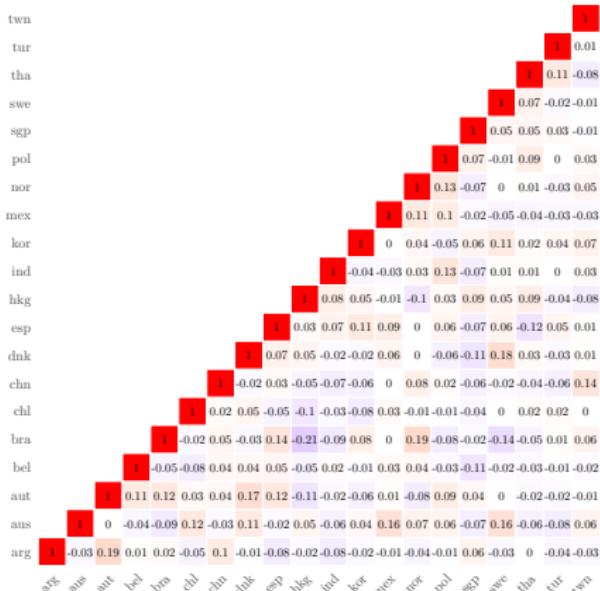
Country-Specific Portfolio: Non-G7 ($K_s = 3$)

| Country | Mean (%) | Standard Deviation (%) | t | Sharpe Ratio | Maximum Drawdown | Worst Month | Best Month | ρ_{US} |
|---------|----------|------------------------|------|--------------|------------------|-------------|------------|-------------|
| arg | 20.14 | 19.07 | 3.54 | 1.06 | 22.96 | -10.47 | 23.99 | 0.02 |
| aus | 16.66 | 20.96 | 4.57 | 0.79 | 29.42 | -16.23 | 40.68 | 0.05 |
| aut | 19.54 | 18.93 | 3.81 | 1.03 | 36.14 | -13.09 | 18.61 | -0.14 |
| bel | 27.49 | 20.93 | 7.23 | 1.31 | 33.22 | -14.78 | 24.83 | -0.08 |
| bra | 10.68 | 21.33 | 1.63 | 0.50 | 32.95 | -17.79 | 18.59 | 0.10 |
| chl | 19.78 | 21.75 | 4.65 | 0.91 | 50.88 | -14.36 | 27.19 | -0.03 |
| chn | 17.14 | 21.38 | 3.26 | 0.80 | 47.28 | -19.60 | 22.59 | -0.04 |
| dnk | 26.71 | 22.06 | 5.87 | 1.21 | 29.91 | -18.43 | 21.02 | -0.02 |
| esp | 18.79 | 20.11 | 5.39 | 0.93 | 34.48 | -16.84 | 24.38 | 0.05 |
| hkg | 21.87 | 20.40 | 6.94 | 1.07 | 32.45 | -23.01 | 26.09 | 0.10 |
| ind | 18.01 | 20.02 | 4.06 | 0.90 | 40.30 | -16.09 | 21.62 | -0.07 |
| kor | 11.39 | 19.68 | 2.62 | 0.58 | 50.19 | -15.89 | 25.05 | 0.02 |
| mex | 22.92 | 21.03 | 4.19 | 1.09 | 39.66 | -21.93 | 20.86 | 0.03 |
| nor | 18.14 | 22.34 | 3.88 | 0.81 | 60.58 | -22.76 | 25.27 | -0.01 |
| pol | 19.54 | 21.67 | 3.64 | 0.90 | 37.27 | -22.24 | 23.41 | -0.07 |
| sgp | 24.22 | 20.84 | 5.40 | 1.16 | 34.21 | -13.83 | 29.82 | 0.09 |
| swe | 20.18 | 22.10 | 4.56 | 0.91 | 40.78 | -28.07 | 37.55 | -0.03 |
| tha | 12.47 | 19.23 | 3.62 | 0.65 | 44.87 | -25.35 | 25.65 | -0.02 |
| tur | 22.67 | 20.81 | 4.71 | 1.09 | 32.10 | -18.28 | 29.24 | -0.03 |
| twn | 12.10 | 19.32 | 3.14 | 0.63 | 49.63 | -12.60 | 17.14 | 0.02 |

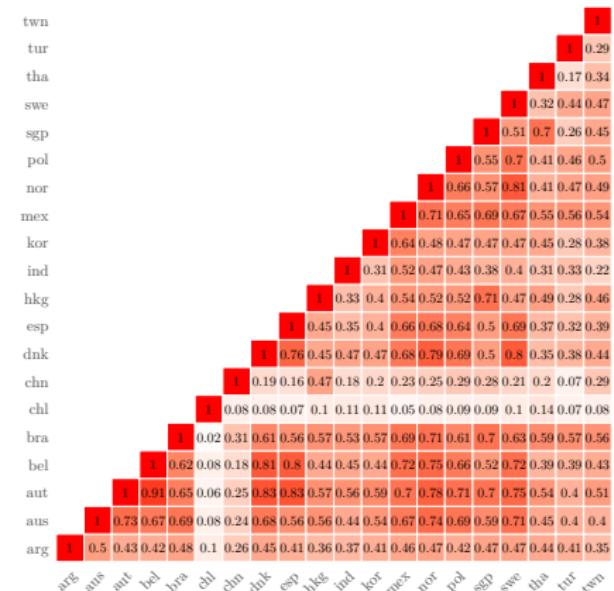
- Very low correlation with US market

Correlation: Non-G7 Country-Specific Portfolios

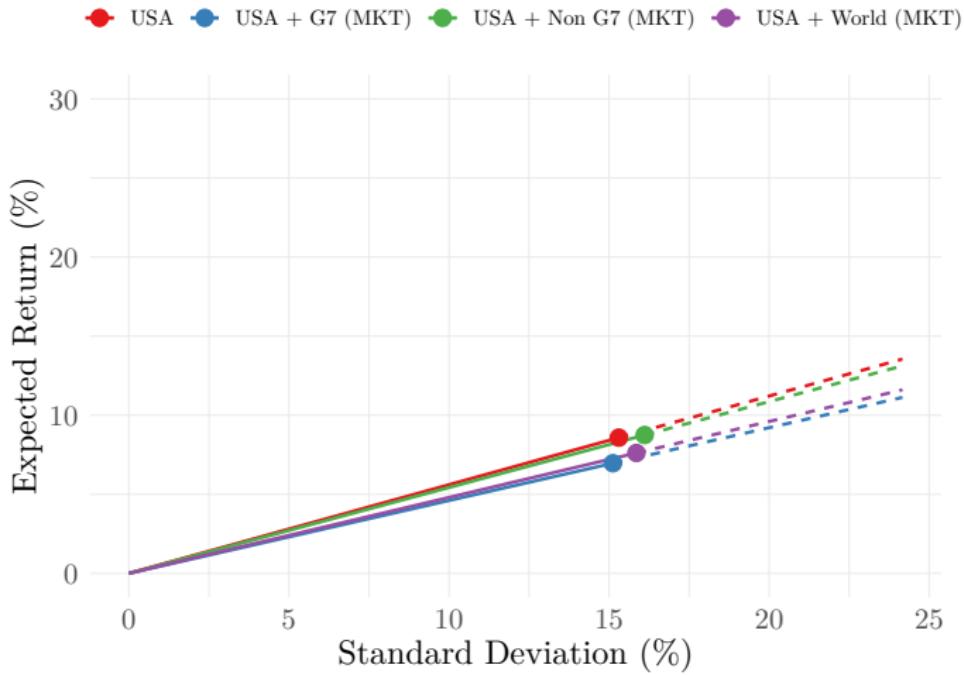
Country Specific Factor Correlation



Correlation of Country Market Portfolios

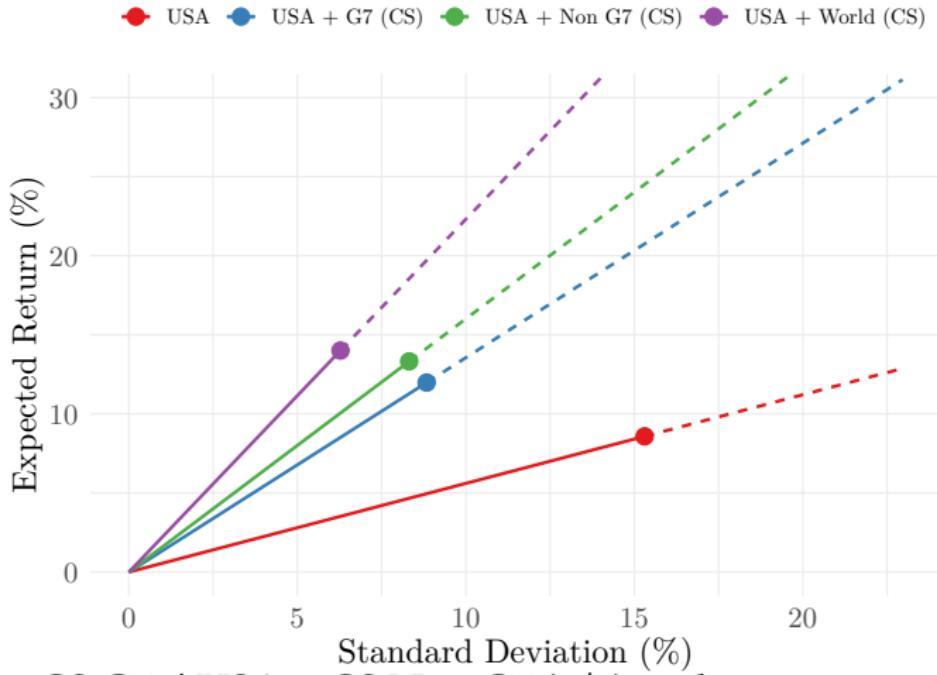


Revisiting Diversification Benefits I



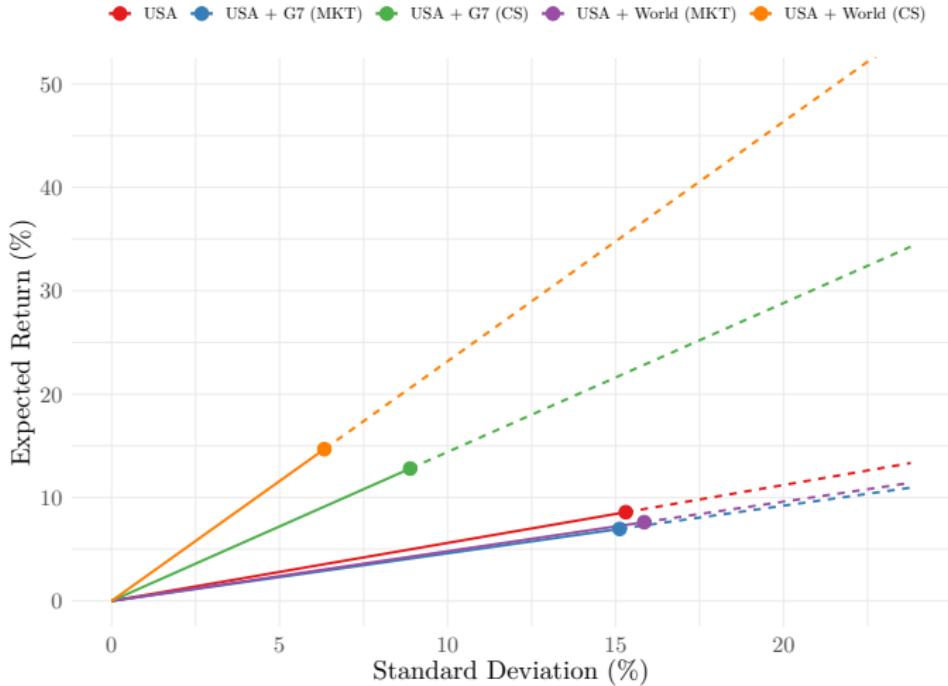
- USA + EW G7 / USA + EW Non G7 (1/2) each
- USA + World (1/3 each)

Revisiting Diversification Benefits II



- USA + CS G7 / USA + CS Non G7 (1/2) each
- USA + World (1/3 each)

Revisiting Diversification Benefits III



- More effective diversification through country specific factors

Diversification Benefits

| Portfolio | Mean (%) | Standard Deviation (%) | Sharpe Ratio |
|--------------------|----------|------------------------|--------------|
| USA | 8.58 | 15.30 | 0.56 |
| G7 (MKT) | 5.34 | 16.31 | 0.33 |
| Non G7 (MKT) | 8.91 | 18.92 | 0.47 |
| USA + G7 (MKT) | 6.96 | 15.12 | 0.46 |
| USA + Non G7 (MKT) | 8.74 | 16.11 | 0.54 |
| USA + World (MKT) | 7.61 | 15.85 | 0.48 |
| USA + G7 (CS) | 11.98 | 8.84 | 1.36 |
| USA + Non G7 (CS) | 13.32 | 8.32 | 1.60 |
| USA + World (CS) | 14.01 | 6.28 | 2.23 |

- Portfolios
 - 1/2 US + 1/2 (Non) G7 for either index or CS
 - 1/3 US + 1/3 G7 + 1/3 Non G7 for either index or CS
- Only moderate benefits from international index investing
- Strong diversification benefits from isolating country specific risk

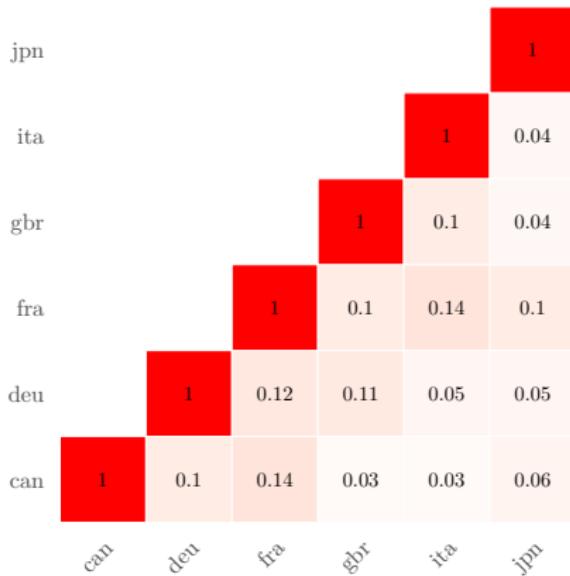
Segmentation Portfolio: G7

| Country | Mean (%) | Standard Deviation (%) | t | Sharpe Ratio | Maximum Drawdown | Worst Month | Best Month | ρ_{US} |
|---------|----------|------------------------|------|--------------|------------------|-------------|------------|-------------|
| can | 15.75 | 23.43 | 4.17 | 0.67 | 43.44 | -31.34 | 25.18 | -0.02 |
| deu | 30.25 | 22.90 | 7.85 | 1.32 | 22.48 | -16.02 | 22.23 | -0.01 |
| fra | 23.16 | 23.04 | 5.13 | 1.01 | 42.92 | -23.47 | 29.87 | 0.03 |
| gbr | 21.60 | 25.96 | 4.77 | 0.83 | 61.98 | -60.43 | 19.95 | -0.04 |
| ita | 23.94 | 22.63 | 5.42 | 1.06 | 49.48 | -17.57 | 23.35 | -0.08 |
| jpn | 26.25 | 22.45 | 5.48 | 1.17 | 54.24 | -20.03 | 27.05 | 0.03 |

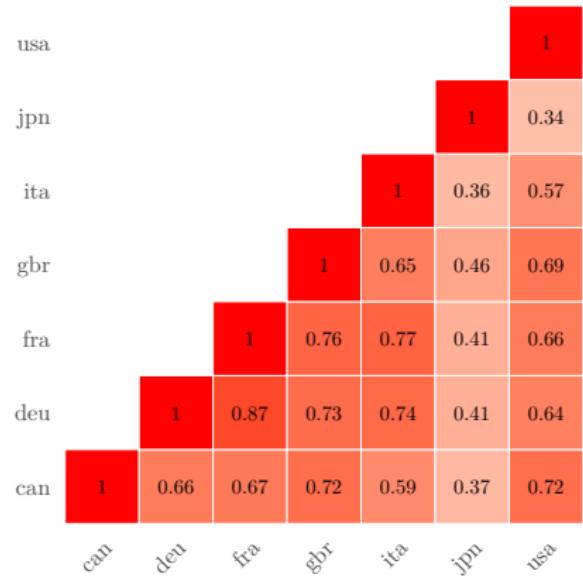
- Surprisingly large segmentation
- **Not an arbitrage**, maybe a good deal
- Structural breaks such as tariffs, capital flow suspensions, or martial law can occur even among developed countries
- “Placebo test” of NASDAQ vs. NYSE shows no evidence of segmentation
- Normally need large T to estimate $\widehat{\lambda} = \frac{1}{T} \sum_{t=1}^T f_t$, but here we estimate $\widehat{\lambda}$ from two sources and cancel f_t

Correlation: G7 Segmentation Portfolios

Segmentation Portfolio Correlation



Correlation of Country Market Portfolios



- Correlation of country specific and segmentation portfolio ≈ 0.21

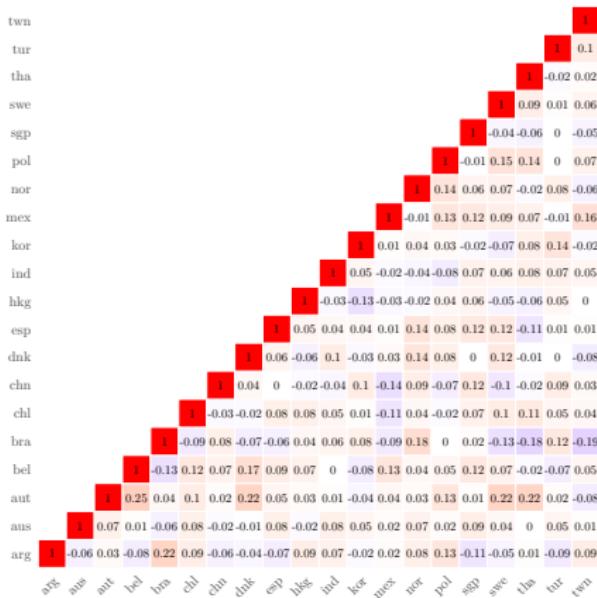
Segmentation Portfolio: Non-G7 ($K_c = 3$)

| Country | Mean (%) | Standard Deviation (%) | t | Sharpe Ratio | Maximum Drawdown | Worst Month | Best Month | ρ_{US} |
|---------|----------|------------------------|------|--------------|------------------|-------------|------------|-------------|
| arg | 16.19 | 21.20 | 2.46 | 0.76 | 42.13 | -15.57 | 34.76 | 0.05 |
| aus | 19.60 | 25.31 | 3.92 | 0.77 | 51.90 | -29.09 | 74.09 | 0.02 |
| aut | 10.68 | 21.08 | 1.76 | 0.51 | 50.42 | -19.86 | 21.48 | 0.11 |
| bel | 17.29 | 22.16 | 3.85 | 0.78 | 68.21 | -16.21 | 25.67 | 0.08 |
| bra | 10.73 | 20.69 | 1.59 | 0.52 | 37.15 | -11.12 | 16.66 | 0.15 |
| chl | 18.07 | 20.12 | 3.81 | 0.90 | 28.83 | -18.69 | 25.73 | -0.10 |
| chn | 25.03 | 23.55 | 4.51 | 1.06 | 36.17 | -18.44 | 29.84 | 0.01 |
| dnk | 28.28 | 23.87 | 5.91 | 1.18 | 28.84 | -21.08 | 26.53 | 0.07 |
| esp | 23.27 | 21.30 | 5.48 | 1.09 | 38.77 | -19.41 | 19.22 | 0.02 |
| hkg | 16.12 | 21.33 | 4.24 | 0.76 | 38.91 | -25.03 | 33.22 | 0.00 |
| ind | 19.04 | 24.94 | 3.44 | 0.76 | 47.56 | -20.75 | 27.88 | -0.02 |
| kor | 21.75 | 22.16 | 4.92 | 0.98 | 54.72 | -19.10 | 32.48 | 0.02 |
| mex | 15.47 | 21.29 | 2.92 | 0.73 | 40.42 | -13.81 | 17.29 | 0.03 |
| nor | 19.74 | 21.65 | 4.47 | 0.91 | 45.88 | -18.09 | 22.47 | -0.04 |
| pol | 21.24 | 24.48 | 3.49 | 0.87 | 32.26 | -25.90 | 59.76 | 0.03 |
| sgp | 33.13 | 21.32 | 6.93 | 1.55 | 27.67 | -18.46 | 29.27 | 0.03 |
| swe | 22.91 | 22.60 | 5.22 | 1.01 | 27.44 | -22.34 | 21.99 | -0.02 |
| tha | 21.81 | 24.41 | 4.07 | 0.89 | 48.48 | -35.35 | 30.99 | -0.03 |
| tur | 11.52 | 21.26 | 2.74 | 0.54 | 50.19 | -14.29 | 25.45 | 0.02 |
| twn | 23.47 | 20.44 | 6.21 | 1.15 | 28.68 | -18.87 | 23.68 | -0.02 |

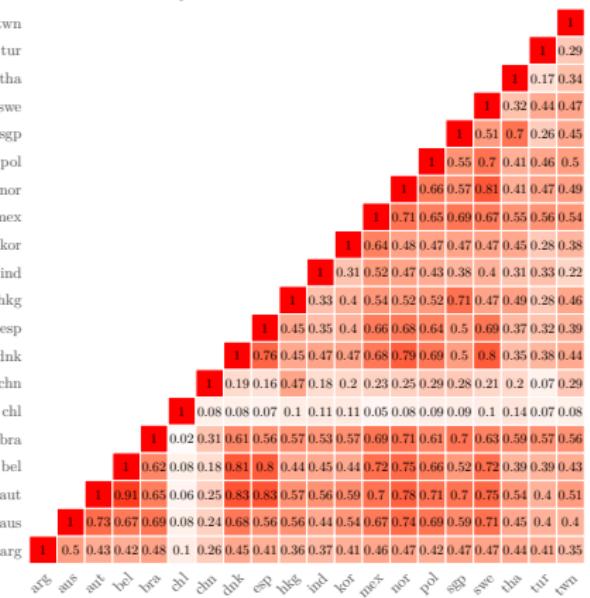
- Large differences in common factor risk

Correlation: Non-G7 Segmentation Portfolios

Segmentation Portfolio Correlation



Correlation of Country Market Portfolios



- Correlation of country specific and segmentation portfolio ≈ 0.16

Intuition: Segmentation Portfolio – G7

| Panel A: G7 Countries as Country 2 | | | | | | | | | |
|------------------------------------|----------|------------------------|------|--------------|----------------------|-----------------|----------------|-------------|-------------|
| Country | Mean (%) | Standard Deviation (%) | t | Sharpe Ratio | Maximum Drawdown (%) | Worst Month (%) | Best Month (%) | ρ_{US} | Pr[US Long] |
| can | 4.09 | 6.67 | 3.70 | 0.61 | 15.04 | -7.41 | 8.89 | -0.09 | 0.34 |
| deu | 12.75 | 11.54 | 6.50 | 1.10 | 22.50 | -13.01 | 16.16 | 0.09 | 0.76 |
| fra | 10.16 | 10.20 | 5.73 | 1.00 | 15.23 | -7.81 | 14.35 | 0.15 | 0.78 |
| gbr | 6.00 | 9.39 | 3.26 | 0.64 | 33.85 | -8.74 | 8.95 | -0.03 | 0.69 |
| ita | 8.60 | 11.72 | 3.57 | 0.73 | 34.65 | -12.77 | 15.25 | 0.22 | 0.80 |
| jpn | 6.75 | 11.07 | 3.11 | 0.61 | 28.10 | -9.31 | 10.61 | 0.20 | 0.67 |

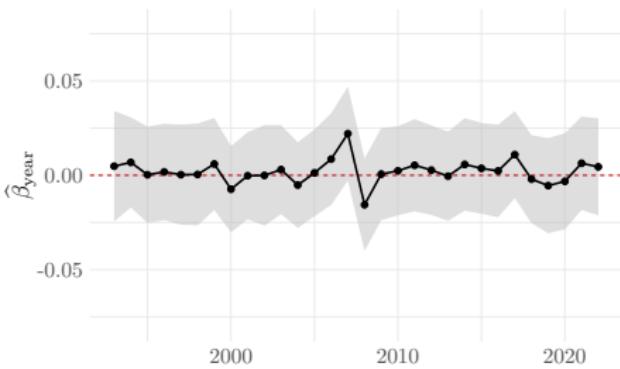
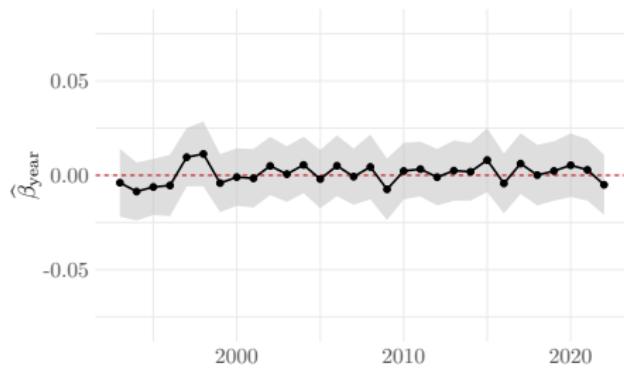
- Implementation for the single factor case
- Tends to be long the US and short the “other country”
- Single factor is likely understating the results, because Gonçalves et al. (2025) rejects a single common factor most periods (for all pairs)

Intuition: Segmentation Portfolio – Non-G7

Panel B: Non-G7 Countries as Country 2

| Country | Mean (%) | Standard Deviation (%) | t | Sharpe Ratio | Maximum Drawdown (%) | Worst Month (%) | Best Month (%) | ρ_{US} | Pr[US Long] |
|---------|----------|------------------------|------|--------------|----------------------|-----------------|----------------|-------------|-------------|
| arg | 5.99 | 13.36 | 1.52 | 0.45 | 36.68 | -12.26 | 12.04 | 0.04 | 0.64 |
| aus | 6.42 | 12.20 | 2.90 | 0.53 | 30.84 | -13.16 | 12.74 | 0.02 | 0.53 |
| aut | 8.28 | 8.26 | 4.88 | 1.00 | 10.29 | -5.16 | 7.61 | 0.08 | 0.67 |
| bel | 11.16 | 11.10 | 4.95 | 1.01 | 24.18 | -7.78 | 13.63 | 0.13 | 0.72 |
| bra | 5.56 | 9.15 | 2.59 | 0.61 | 12.38 | -7.19 | 7.21 | -0.11 | 0.54 |
| chl | 6.69 | 11.39 | 2.26 | 0.59 | 28.63 | -10.38 | 11.82 | -0.18 | 0.52 |
| chn | 9.89 | 14.65 | 3.16 | 0.67 | 29.41 | -10.25 | 13.97 | 0.15 | 0.65 |
| dnk | 5.78 | 10.34 | 3.21 | 0.56 | 18.73 | -7.82 | 9.65 | -0.14 | 0.53 |
| esp | 15.48 | 11.59 | 6.77 | 1.33 | 17.09 | -9.74 | 14.05 | 0.16 | 0.83 |
| hkg | 5.63 | 12.57 | 2.09 | 0.45 | 36.20 | -11.48 | 15.60 | 0.07 | 0.62 |
| ind | 10.53 | 12.40 | 4.13 | 0.85 | 24.08 | -10.57 | 13.35 | -0.08 | 0.59 |
| kor | 9.30 | 12.01 | 4.25 | 0.77 | 24.86 | -12.75 | 12.97 | 0.05 | 0.77 |
| mex | 6.70 | 8.88 | 3.37 | 0.75 | 17.20 | -7.45 | 6.95 | 0.05 | 0.43 |
| nor | 6.39 | 8.35 | 3.46 | 0.77 | 25.26 | -11.47 | 8.90 | -0.07 | 0.36 |
| pol | 5.39 | 9.70 | 2.79 | 0.56 | 22.77 | -7.38 | 10.16 | 0.03 | 0.65 |
| sgp | 10.61 | 10.49 | 4.97 | 1.01 | 20.43 | -7.72 | 10.52 | 0.24 | 0.76 |
| swe | 5.24 | 10.79 | 2.48 | 0.49 | 34.25 | -9.69 | 10.19 | 0.05 | 0.61 |
| tha | 9.38 | 12.40 | 3.85 | 0.76 | 39.03 | -12.02 | 10.66 | 0.20 | 0.74 |
| tur | 10.28 | 13.07 | 3.61 | 0.79 | 31.45 | -15.93 | 10.75 | -0.09 | 0.53 |
| twn | 7.87 | 11.65 | 3.60 | 0.68 | 24.52 | -9.82 | 11.84 | -0.06 | 0.51 |

Further Analysis



- Results are robust to
 - Different base currencies, GBP and JPY
 - Diversification does not break during US downturns
 - Currency risk factors (carry, dollar factor)
 - No extreme weights

Conclusion

- Develop latent cross-country factor model
- Feasible estimation for fixed T and large N
- New angle on cross-country investing with larger benefits
- International version of α -portfolio (▶ details) largely confirms the US findings of Kim, Korajczyk, and Neuhierl (2021)
- Surprisingly large differences in risk prices
- **Thank you for your comments**

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Country-Specific Portfolio: Non-G7 ($K_s = 1$)

| Country | Mean (%) | Standard Deviation (%) | t | Sharpe Ratio | Maximum Drawdown | Worst Month | Best Month | ρ_{US} |
|---------|----------|------------------------|------|--------------|------------------|-------------|------------|-------------|
| arg | 20.49 | 20.03 | 3.49 | 1.02 | 23.66 | -11.44 | 24.86 | 0.01 |
| aus | 16.68 | 20.87 | 4.73 | 0.80 | 26.83 | -14.60 | 44.65 | 0.06 |
| aut | 19.96 | 19.47 | 3.69 | 1.02 | 31.63 | -19.53 | 20.73 | -0.12 |
| bel | 27.90 | 19.45 | 7.62 | 1.43 | 25.50 | -21.40 | 18.70 | -0.06 |
| bra | 17.80 | 20.61 | 2.56 | 0.86 | 31.55 | -19.10 | 18.11 | -0.05 |
| chl | 19.99 | 21.78 | 4.41 | 0.92 | 53.23 | -19.16 | 26.02 | -0.01 |
| chn | 14.83 | 22.02 | 2.87 | 0.67 | 49.72 | -19.55 | 26.68 | -0.05 |
| dnk | 26.31 | 22.96 | 5.83 | 1.15 | 26.77 | -19.47 | 21.01 | -0.08 |
| esp | 20.29 | 22.49 | 5.69 | 0.90 | 34.16 | -18.76 | 26.48 | 0.03 |
| hkg | 19.22 | 19.37 | 5.58 | 0.99 | 42.81 | -25.11 | 24.13 | 0.08 |
| ind | 18.67 | 20.76 | 4.21 | 0.90 | 35.76 | -21.18 | 20.99 | -0.01 |
| kor | 15.10 | 18.80 | 4.04 | 0.80 | 34.64 | -16.25 | 19.79 | -0.01 |
| mex | 20.88 | 20.89 | 3.99 | 1.00 | 42.02 | -15.15 | 18.20 | 0.08 |
| nor | 19.49 | 21.67 | 4.12 | 0.90 | 53.49 | -23.85 | 19.65 | -0.02 |
| pol | 24.26 | 21.26 | 4.82 | 1.14 | 33.57 | -19.87 | 25.95 | 0.00 |
| sgp | 23.31 | 21.24 | 5.17 | 1.10 | 31.07 | -27.04 | 25.62 | 0.08 |
| swe | 22.20 | 22.87 | 5.28 | 0.97 | 58.37 | -14.87 | 34.46 | -0.01 |
| tha | 11.96 | 20.52 | 3.51 | 0.58 | 31.96 | -31.20 | 30.07 | -0.02 |
| tur | 20.29 | 19.99 | 4.24 | 1.01 | 31.52 | -13.87 | 30.88 | -0.07 |
| twn | 12.50 | 18.49 | 3.30 | 0.68 | 36.75 | -12.65 | 18.49 | 0.01 |

- Very low correlation with US market

Segmentation Portfolio: Non-G7 ($K_c = 1$)

| Country | Mean (%) | Standard Deviation (%) | t | Sharpe Ratio | Maximum Drawdown | Worst Month | Best Month | ρ_{US} |
|---------|----------|------------------------|-------|--------------|------------------|-------------|------------|-------------|
| arg | 10.85 | 23.08 | 1.67 | 0.47 | 53.32 | -14.34 | 38.04 | 0.09 |
| aus | 10.86 | 26.62 | 2.60 | 0.41 | 59.75 | -20.99 | 67.66 | -0.02 |
| aut | 13.24 | 21.41 | 2.45 | 0.62 | 40.33 | -16.89 | 23.13 | 0.18 |
| bel | 10.57 | 21.86 | 2.20 | 0.48 | 57.85 | -25.30 | 24.92 | 0.06 |
| bra | -0.81 | 23.18 | -0.11 | -0.03 | 48.88 | -19.07 | 15.65 | -0.02 |
| chl | 15.42 | 19.38 | 3.39 | 0.80 | 55.44 | -14.43 | 19.08 | 0.05 |
| chn | 4.93 | 23.93 | 1.03 | 0.21 | 55.97 | -19.93 | 41.88 | 0.06 |
| dnk | 10.75 | 20.50 | 2.40 | 0.52 | 53.77 | -12.62 | 20.02 | -0.05 |
| esp | 10.67 | 21.05 | 2.58 | 0.51 | 45.99 | -18.29 | 17.80 | 0.03 |
| hkg | 2.61 | 20.92 | 0.60 | 0.12 | 67.01 | -26.01 | 21.24 | 0.09 |
| ind | 18.14 | 21.46 | 3.88 | 0.85 | 43.11 | -20.03 | 24.82 | -0.14 |
| kor | 7.88 | 21.75 | 1.83 | 0.36 | 58.47 | -22.64 | 23.55 | -0.03 |
| mex | 4.54 | 20.06 | 0.87 | 0.23 | 52.68 | -12.30 | 16.35 | 0.02 |
| nor | 5.55 | 21.63 | 1.30 | 0.26 | 66.79 | -16.93 | 18.31 | 0.00 |
| pol | 7.77 | 19.50 | 1.98 | 0.40 | 46.79 | -25.70 | 20.13 | 0.17 |
| sgp | 16.19 | 23.07 | 3.44 | 0.70 | 48.76 | -26.98 | 35.76 | 0.19 |
| swe | 7.13 | 20.23 | 1.76 | 0.35 | 61.25 | -34.35 | 14.39 | -0.01 |
| tha | 6.89 | 21.74 | 1.73 | 0.32 | 65.94 | -38.16 | 23.34 | 0.00 |
| tur | 12.34 | 22.08 | 2.81 | 0.56 | 48.59 | -20.24 | 25.68 | 0.15 |
| twn | 8.96 | 20.54 | 2.72 | 0.44 | 35.25 | -19.23 | 17.16 | 0.00 |

- Large differences in common factor risk

Closed form Solution for Θ_g^c and Θ_g^s

- Hence, the solution $(\widehat{\Theta}_g^c, \widehat{\Theta}_g^s)$ will be determined by

$$\text{vec}([\widehat{\Theta}_g^c \ \widehat{\Theta}_g^s]) = (\mathcal{X}' \mathcal{X})^{-1} \mathcal{X}' \text{vec}(\widehat{\mathbf{R}}_g),$$

where $\mathcal{X} = ([\widehat{\mathbf{F}} \ \widehat{\mathbf{G}}] \otimes \mathbf{X}_g)$.

- [Return to Step3](#)

Multivariate Country-specific Portfolio

- Estimate the country-specific risk premia and its variance

$$\widehat{\lambda}_2^s = \frac{\widehat{\mathbf{w}}_2^{s'} \mathbf{R}_2 \mathbf{1}_T}{T}$$

$$\text{Var}(\widehat{\lambda}_2^s) = \frac{\widehat{\mathbf{w}}_2^{s'} \mathbf{R}_2 \mathbf{R}'_2 \widehat{\mathbf{w}}_2^s}{T - 1} - \widehat{\lambda}_2^s \widehat{\lambda}_2^{s'}$$

- Then, we hold

$$\left(\text{Var}(\widehat{\lambda}_2^s)^{-1} \widehat{\lambda}_2^s \right)' (\widehat{\mathbf{w}}_2^{s'} \mathbf{r}_{2,T+1})$$

- [Return to Univariate Construction](#)

Multivariate Segmentation Portfolio

- Run vector regression

$$\underbrace{\widehat{\mathbf{w}}_2^{c'} \mathbf{R}_2}_{K^c \times T} = \beta_0 + \beta_1 \left(\underbrace{\widehat{\mathbf{w}}_1^{c'} \mathbf{R}_1}_{K^c \times T} \right) + \mathbf{e}$$

- Intercept captures difference in risk premia, i.e.

$$\beta_0 \xrightarrow{P} \mathcal{S}^{c'} (\lambda_2^c - \lambda_1^c)$$

- After fitting the regression, we hold

$$\left(\text{Var}(\widehat{\beta}_0)^{-1} \widehat{\beta}_0 \right)' \left(\widehat{\mathbf{w}}_2^{c'} \mathbf{r}_{2,T+1} - \widehat{\beta}_1 \widehat{\mathbf{w}}_1^{c'} \mathbf{r}_{1,T+1} \right)$$

- [Return to Univariate Construction](#)

Alpha Portfolio

- Allow for non-zero alpha

$$R_{git} = \alpha_{ig} + \beta_{ig} f_t + \delta_{ig} g_t + e_{itg}, \quad \alpha_{ig} = \mathbf{x}'_{ig} \Theta_g^a + \gamma_{ig}^a$$

Solve constrained regression

$$\widehat{\Theta}_g^a = \arg \min_{\Theta_g^a} \| \bar{\mathbf{R}}_g - \mathbf{X}_g \Theta_g^a \|,$$

$$\text{subject to } [\widehat{\Theta}_g^c \ \widehat{\Theta}_g^s]' \mathbf{X}'_g \mathbf{X}_g \Theta_g^a = \mathbf{0}_{K_g \times 1}$$

- $\widehat{\Theta}_g^a \xrightarrow{P} \Theta_g^a$
- Construct the portfolio in an out-of-sample manner:
 $\mathbf{w}_{g,T}^a = \frac{1}{N} \mathbf{X}_g \widehat{\Theta}_{g,T}^a$

$$\mathbf{w}_{g,T}^{a'} \mathbf{r}_{g,T+1} \xrightarrow{P} \lim_{N_g} \frac{1}{N_g} \Theta_g^{a'} \mathbf{X}'_g \mathbf{X}_g \Theta_g^a \geq 0$$

▶ back

Segmentation Portfolio - Relation to Currencies

- Regression on changes in exchange rate

$$r_{i,t}^{\Delta\lambda} = \alpha_i + \beta \times \Delta(\text{Exchange Rate vs. USD})_{it} + u_{it}$$

$$\widehat{\beta} \approx -0.052, \quad t_{\widehat{\beta}} \approx -1.4, \quad R^2 < 0.01$$

- Regression on currency risk factors (Verdelhan (2018))

$$r_i^{\Delta\lambda} = \alpha + \beta_1 \times \text{Carry}_t + \beta_2 \times \text{Dollar}_t + u_t$$

$$\widehat{\beta}_1 \approx -0.043, \quad t_{\widehat{\beta}_1} \approx -1.1, \quad \widehat{\beta}_2 \approx -0.0042, \quad t_{\widehat{\beta}_2} \approx -0.08, \quad R^2 < 0.01$$

- **Ongoing work**

- No strong time trends
- NB Patton and Weller (2022) find λ -cluster even for the US