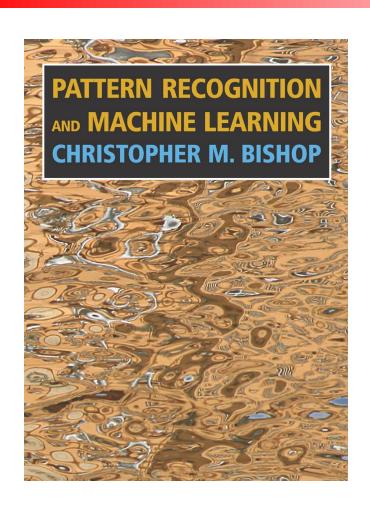
Gaussian Identities (cont.)

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Office Hour: Wednesday 4pm-5pm

NO LAB this week.

Reference



Today's class *roughly* follows Chapter 2.3-2.34

Pattern Recognition and Machine Learning

Christopher Bishop, 2006

Recap

•
$$N_{\mathbf{x}}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \coloneqq \frac{1}{(2\pi)^{\frac{d}{2}}|\boldsymbol{\Sigma}|^{\frac{1}{2}}} \cdot \exp\left[-\frac{(\mathbf{x}-\boldsymbol{\mu})^{\mathsf{T}}\boldsymbol{\Sigma}^{-1}(\mathbf{x}-\boldsymbol{\mu})}{2}\right]$$

- MVNs are multi-dimensional generalizations of univariate normal distributions, in the sense that:
- $\Sigma^{-1} = UDU^{\mathsf{T}}$
 - Eigen-decomposition, **D** is diagonal.
 - $\boldsymbol{U} \in R^{d \times d}$, $\boldsymbol{U} \boldsymbol{U}^{\top} = \boldsymbol{U}^{\top} \boldsymbol{U} = \boldsymbol{I}$
- $\mathbf{y} = \mathbf{U}^{\mathsf{T}}(\mathbf{x} \boldsymbol{\mu}).$
- $p(y) = \prod_i N_y(0, d_i), \frac{1}{d_i}$ is *i*-th diagonal-elem. of **D**.
- Use this to generate samples of MVN using uni-normal!

Recap

- Mahalanobis distance, $(x \mu)^{\top} \Sigma^{-1} (x \mu)$
 - Distance between a point x to the center of $N_x(\mu, \Sigma)$,
 - Distance between x and μ rotated by U.
 - Can be used to define the confidence region.

Moments of MVN.

- $\mathbb{E}[x] = \mu$
 - Apply the transform $z = x \mu$.

•
$$\int_{-\infty}^{0} \exp\left[-\frac{z\Sigma^{-1}z}{2}\right] z dz = -\int_{0}^{\infty} \exp\left[-\frac{z\Sigma^{-1}z}{2}\right] z dz$$

- $\mathbb{E}[xx^{\mathsf{T}}] = \mu^{\mathsf{T}}\mu + \Sigma$
 - Apply the transform $z = x \mu$.
 - Use z = Uy and $UU^{T} = I$.

Partitioned MVNs

• Given:

•
$$p(\mathbf{x}_a, \mathbf{x}_b) = N_{\mathbf{x}_a, \mathbf{x}_b} \left(\begin{bmatrix} \boldsymbol{\mu}_a \\ \boldsymbol{\mu}_b \end{bmatrix}, \begin{bmatrix} \boldsymbol{\Sigma}_{aa}, \boldsymbol{\Sigma}_{ab} \\ \boldsymbol{\Sigma}_{ba}, \boldsymbol{\Sigma}_{bb} \end{bmatrix} \right)$$

- Represent $p(x_a|x_b)$ and $p(x_a)$ using $\frac{\mu_a}{\mu_b}$ and $\frac{\Sigma_{aa}, \Sigma_{ab}}{\Sigma_{ba}, \Sigma_{bb}}$.
- Partitioned MVN formulas have huge applications in Bayesian regression, Gaussian graphical models etc.
- For simplicity, we let $\Sigma^{-1} = \Theta = \begin{bmatrix} \Theta_{aa}, \Theta_{ab} \\ \Theta_{ba}, \Theta_{bb} \end{bmatrix}$.

- You can prove by following the def. of conditional dist.
- However, observe:

$$\log N_{\boldsymbol{x}}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \begin{bmatrix} \boldsymbol{x}_{a} - \boldsymbol{\mu}_{a} \\ \boldsymbol{x}_{b} - \boldsymbol{\mu}_{b} \end{bmatrix}^{\mathsf{T}} \begin{bmatrix} \boldsymbol{\Theta}_{aa}, \boldsymbol{\Theta}_{ab} \\ \boldsymbol{\Theta}_{ba}, \boldsymbol{\Theta}_{bb} \end{bmatrix} \begin{bmatrix} \boldsymbol{x}_{a} - \boldsymbol{\mu}_{a} \\ \boldsymbol{x}_{b} - \boldsymbol{\mu}_{b} \end{bmatrix} + \text{const.}$$

- $\log N_x$ is merely a quadratic function w.r.t x + const.
- Expanding quad. term only leads to quad./linear terms.
 - w.r.t. x_a , x_b
- $\Rightarrow P(x_a|x_b)$ is an MVN (not rigorously speaking).

• If $p(t) = N_t(\mu; \Sigma)$, then $\log p(t) = -\frac{t^{\top} \Sigma^{-1} t}{2} + t^{\top} \Sigma^{-1} \mu + \text{const.}$



• If we spot terms in $-\begin{bmatrix} x_a - \mu_a \\ x_b - \mu_b \end{bmatrix}^{\mathsf{T}} \begin{bmatrix} \Theta_{aa}, \Theta_{ab} \\ \Theta_{ba}, \Theta_{bb} \end{bmatrix} \begin{bmatrix} x_a - \mu_a \\ x_b - \mu_b \end{bmatrix}$ /2 with respect to \boldsymbol{x}_a which has the same form as those in ______, we can directly identify the covariance and mean for $p(\boldsymbol{x}_a | \boldsymbol{x}_b)$.

- The quadratic term w.r.t. x_a after expansion:
- $-\mathbf{x}_a^{\mathsf{T}}\mathbf{\Theta}_{aa}\mathbf{x}_a/2 \Longrightarrow \mathsf{Cov}_{\mathbf{x}_a|\mathbf{x}_b}[\mathbf{x}_a] = [\mathbf{\Theta}_{aa}]^{-1}$.
- The *linear terms* w.r.t. x_a after expansion:
- $\bullet \ -\boldsymbol{x}_{a}^{\mathsf{T}}\boldsymbol{\Theta}_{ab}\boldsymbol{x}_{b} + \boldsymbol{x}_{a}^{\mathsf{T}}\boldsymbol{\Theta}_{ab}\boldsymbol{\mu}_{b} + \boldsymbol{x}_{a}^{\mathsf{T}}\boldsymbol{\Theta}_{aa}\boldsymbol{\mu}_{a}$
- Collect terms: $\mathbf{x}_a^{\mathsf{T}} \mathbf{\Theta}_{aa} (\boldsymbol{\mu}_a \mathbf{\Theta}_{aa}^{-1} \mathbf{\Theta}_{ab} \mathbf{x}_b + \mathbf{\Theta}_{aa}^{-1} \mathbf{\Theta}_{ab} \boldsymbol{\mu}_b)$
- Knowing $Cov_{x_a|x_b}[x_a] = [\Theta_{aa}]^{-1} \Longrightarrow$

$$\mathbb{E}_{\boldsymbol{x}_a|\boldsymbol{x}_b}[\boldsymbol{x}_a] = \boldsymbol{\mu}_a - \boldsymbol{\Theta}_{aa}^{-1}\boldsymbol{\Theta}_{ab}\boldsymbol{x}_b + \boldsymbol{\Theta}_{aa}^{-1}\boldsymbol{\Theta}_{ab}\boldsymbol{\mu}_b$$

Conditional MVN formula

•
$$p(\mathbf{x}_a|\mathbf{x}_b) = N_{\mathbf{x}_a}(\boldsymbol{\mu}_a - \boldsymbol{\Theta}_{aa}^{-1}\boldsymbol{\Theta}_{ab}\mathbf{x}_b + \boldsymbol{\Theta}_{aa}^{-1}\boldsymbol{\Theta}_{ab}\boldsymbol{\mu}_b, \boldsymbol{\Theta}_{aa}^{-1}).$$

- You can use block matrix inversion formula to represent Θ_{aa} , Θ_{ab} using Σ_{aa} , Σ_{ab} and Σ_{bb} .
- See 2.76 in PRML

• However, this formula is most easily expressed using block matrices of Θ .

Partitioned MVNs (Marginal)

- How to represent $p(x_a)$ using $\frac{\mu_a}{\mu_b}$ and $\frac{\Sigma_{aa}, \Sigma_{ab}}{\Sigma_{ba}, \Sigma_{bb}}$?
- First, we marginalize $p(x_a) = \int p(x_a, x_b) dx_b$
- Write terms in $\log p(x_a, x_b)$ w.r.t. x_b after expansion:

•
$$-\mathbf{x}_b^{\mathsf{T}}\mathbf{\Theta}_{bb}\mathbf{x}_b/2 + \mathbf{x}_b^{\mathsf{T}}(\mathbf{\Theta}_{bb}\mathbf{\mu}_b - \mathbf{\Theta}_{ba}\mathbf{x}_a + \mathbf{\Theta}_{ba}\mathbf{\mu}_a)$$

$$= -(\boldsymbol{x}^{\mathsf{T}} - \boldsymbol{\Theta}_{bb}^{-1} \boldsymbol{m}) \boldsymbol{\Theta}_{bb} (\boldsymbol{x} - \boldsymbol{\Theta}_{bb}^{-1} \boldsymbol{m}) / 2 + \boldsymbol{m}^{\mathsf{T}} \boldsymbol{\Theta}_{bb}^{-1} \boldsymbol{m} / 2,$$

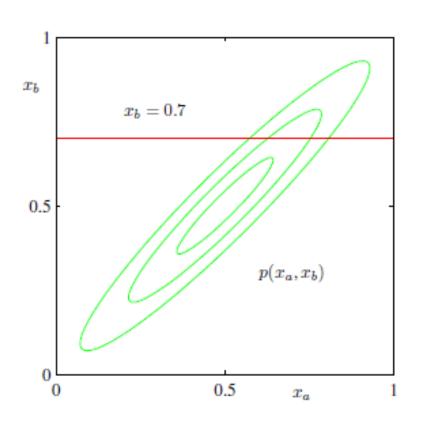
Completing the square!

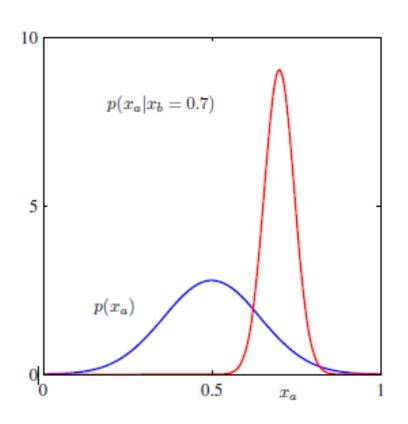
- Now we know
- $p(\mathbf{x}_a) =$ $(\dots) \exp\left(\frac{\mathbf{m}^{\mathsf{T}}\mathbf{\Theta}_{bb}^{-1}\mathbf{m}}{2}\right) \int \exp\left[-\frac{(\mathbf{x}_b^{\mathsf{T}} \mathbf{\Theta}_{bb}^{-1}\mathbf{m})\mathbf{\Theta}_{bb}(\mathbf{x}_b \mathbf{\Theta}_{bb}^{-1}\mathbf{m})}{2}\right] d\mathbf{x}_b$
- Inside integral, just a regular MVN w.r.t. $m{x}_b$ without normalizing constant, so
- $p(\mathbf{x}_a) = (\dots) \exp\left(\frac{\mathbf{m}^{\mathsf{T}}\mathbf{\Theta}_{bb}^{-1}\mathbf{m}}{2}\right) \cdot \text{const}$
- Now, let us find all terms w.r.t. x_a in above expression.

•
$$\log p(x_a) = -\frac{x_a^{\mathsf{T}}(\Theta_{aa} - \Theta_{ab}\Theta_{bb}^{-1}\Theta_{ba})x_a}{+x_a^{\mathsf{T}}(\Theta_{aa} - \Theta_{ab}\Theta_{bb}^{-1}\Theta_{ba})\mu_a + \text{const}}$$

- Using the block matrix inversion formula, $\Theta_{aa} \Theta_{ab}\Theta_{bb}^{-1}\Theta_{ba} = \Sigma_{aa}^{-1}$.
- Therefore, $p(x_a) = N_{x_a}(\mu_a, \Sigma_{aa})$
- The marginal of a joint MVN has mean and variance that is the same as the mean and variance of the partitioned MVN.

Visualization





• PRML 2.9

Gaussian Linear Model

- The prior: $p(x) = N_x(\mu, \Lambda^{-1})$
- The Likelihood: $p(y|x) = N_y(Ax + b, L^{-1})$ Linear model



- The marginal: $p(y) = N_y (A\mu + b, L^{-1} + A\Lambda^{-1}A^{\top})$
- The posterior: $p(x|y) = N_x(\Sigma\{A^\top L(y-b) + \Lambda\mu\}, \Sigma)$ where $\Sigma = (\Lambda + A^\top LA)^{-1}$

Proof: 1. Calculate the joint p(y, x), 2. Use formula we just derived to obtain marginal and conditional dist.

Read PRML, 2.3.3

Likelihood for MVN

• Given the dataset $D \coloneqq \{x_i\}_{i=1}^n$, the likelihood function of MVN density can be written as

•
$$L(\boldsymbol{\mu}, \boldsymbol{\Sigma}, D) = \sum_{i=1} \log N_{x_i}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

= $\operatorname{const} - \frac{n}{2} \log |\boldsymbol{\Sigma}| - \frac{\operatorname{tr}(\overline{X}\overline{X}^{\mathsf{T}}\boldsymbol{\Sigma}^{-1})}{2}$

- where $\overline{X} = [(x_1 \mu) \dots (x_n \mu)] \in \mathbb{R}^{d \times n}$ is the "centralized" dataset.
- tr is the trace operator.

Maximum Likelihood Estimator

- $\max_{\mu, \Sigma} L(\mu, \Sigma, D) = \max_{\Sigma} \max_{\mu} L(\mu, \Sigma, D).$
- First, solve the inner max by

•
$$\frac{\partial L(\mu, \Sigma, D)}{\partial \mu} = 0 \implies \mu_{\text{MLE}} \coloneqq \frac{1}{n} \sum_{i=1}^{n} x_i$$

- Then, plug in $\mu_{\rm MLE}$ and solve the outer max by
- $\bullet \frac{\partial L(\mu_{\text{MLE}}, \Sigma, D)}{\partial \Sigma} = 0 \Longrightarrow$
- $\Sigma_{\text{MLE}} := \frac{1}{n} \overline{X}_{\text{MLE}} \overline{X}_{\text{MLE}}^{\mathsf{T}}$,
- where $\overline{X}_{\mathrm{MLE}}$: = $[(x_1 \mu_{\mathrm{MLE}}) ... (x_n \mu_{\mathrm{MLE}})]$