



# Metamodeling under Inequality Constraints

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## 2 Proposal

- Constrained Kriging: Maatouk and Bay (2016)
- Markov Chain Monte Carlo (MCMC) for constrained Kriging
- Hyperparameters Estimation: Constrained Maximum Likelihood (CML)

## 3 Conclusions and Future Works

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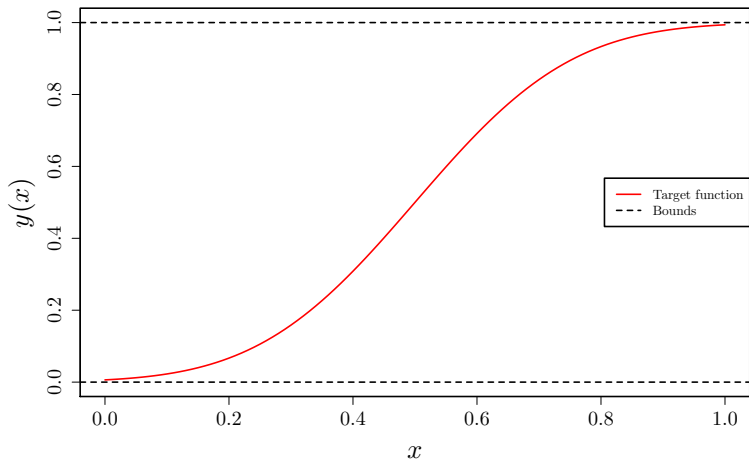
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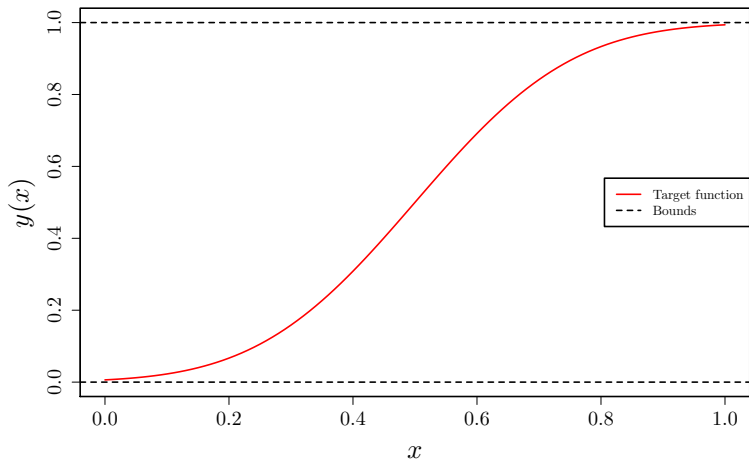
# Motivation

## Toy example.



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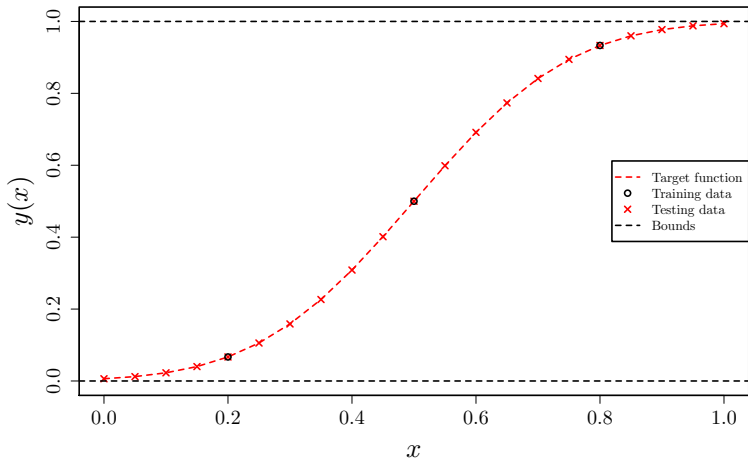
## Toy example.



⇒ The target function is bounded and monotonic!!

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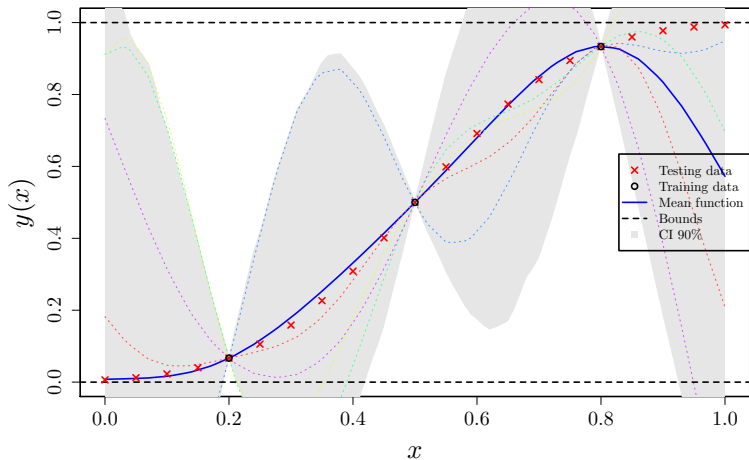
## Toy example.



Data

# Motivation

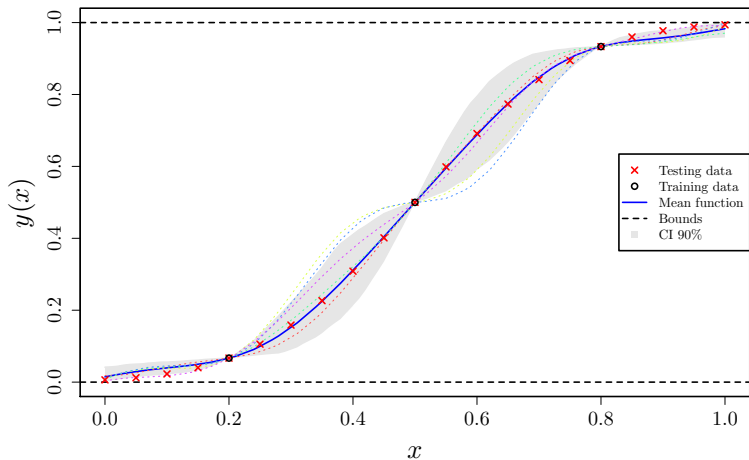
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Standard Kriging

# Motivation

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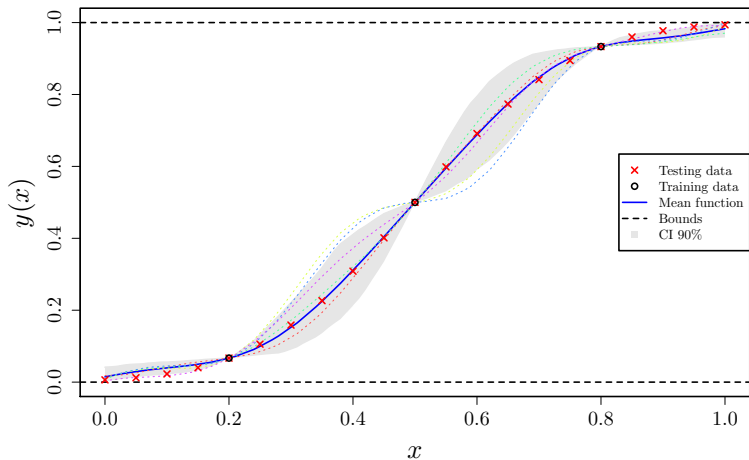


Constrained Kriging



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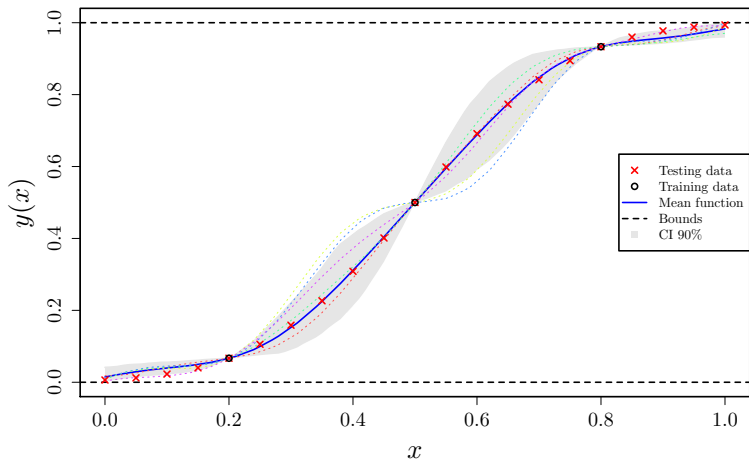
**Toy example.**  $y(x) = \frac{1}{2} \left[ 1 + \operatorname{erf} \left\{ \frac{x-0.5}{0.2\sqrt{2}} \right\} \right]$ .



Constrained Kriging

# Motivation

**Toy example.**  $y(x) = \Phi\left(\frac{x-0.5}{0.2}\right)$  (Gaussian CDF).



Constrained Kriging

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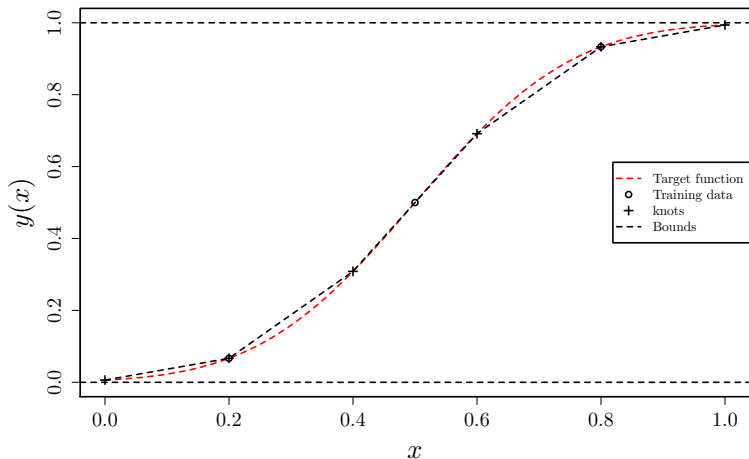
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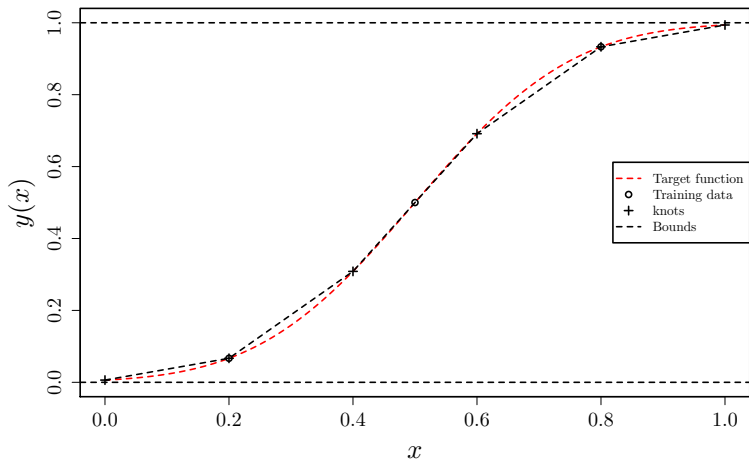
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**Figure:** toy example  $y(x) = \Phi(\frac{x-0.5}{0.2})$  (Gaussian CDF).



# Constrained Kriging: finite approximation

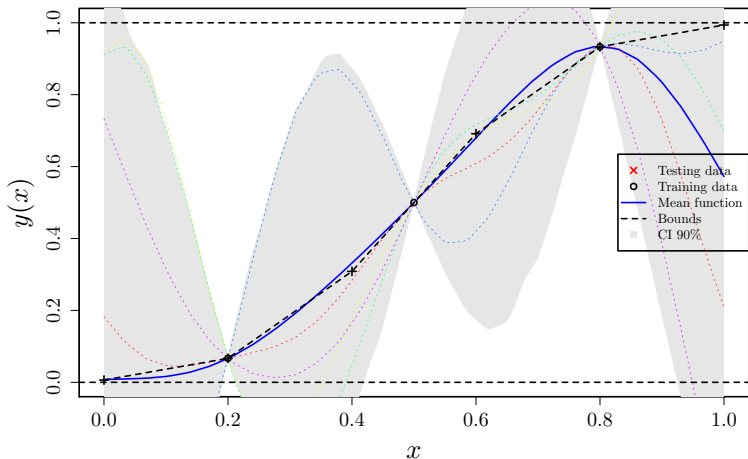
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⇒ The finite approx. is also bounded and monotonic!!

# Constrained Kriging: finite approximation

**Figure:** toy example  $y(x) = \Phi(\frac{x-0.5}{0.2})$  (Gaussian CDF).



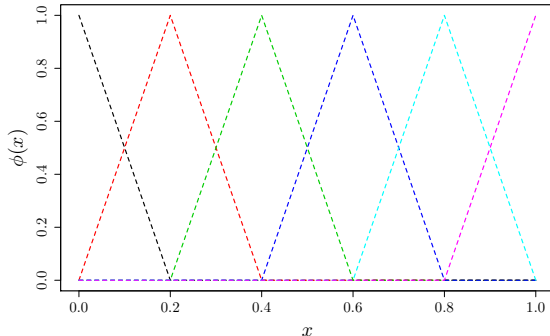
⇒ We can assume that the knots' images are random variables.

# Constrained Kriging: Maatouk and Bay (2016)

Let the finite-one-dimensional GP-based approximation be defined as

$$y(x) \approx \sum_{j=0}^{m-1} \xi_j \phi_j(x), \quad x \in [0, 1], \quad \text{subject to } \boldsymbol{\xi} \in \mathcal{C}, \quad (1)$$

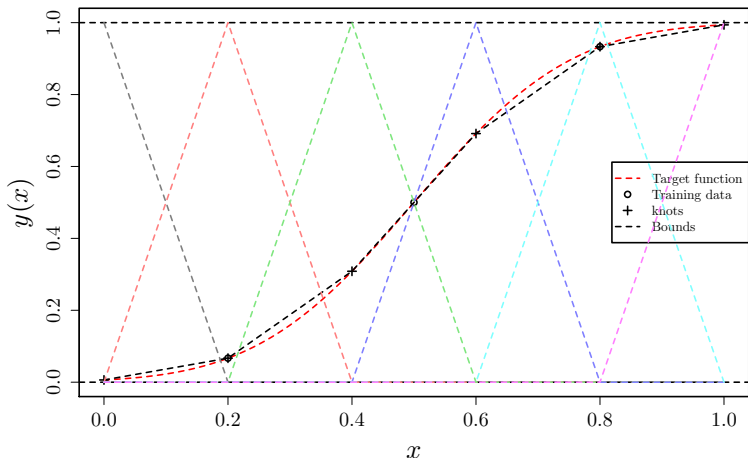
where  $\boldsymbol{\xi} = [\xi_0, \xi_1, \dots, \xi_{m-1}]^\top$  (knots' images) is a zero-mean Gaussian vector with covariance matrix  $\boldsymbol{\Gamma} \in \mathbb{R}^{m \times m}$ , and  $\phi_j : [0, 1] \rightarrow \mathbb{R}$  are hat functions.





# Constrained Kriging: finite approximation

**Figure:** toy example  $y(x) = \Phi(\frac{x-0.5}{0.2})$  (Gaussian CDF).



# Constrained Kriging: Maatouk and Bay (2016)

◆ We are interested on convex sets of the form  $\mathcal{C} : \mathbf{l} \leq \mathbf{\Lambda} \boldsymbol{\xi} \leq \mathbf{u}$ , where  $\mathbf{\Lambda} \in \mathbb{R}^{p \times m}$ , and  $\mathbf{l}$  and  $\mathbf{u}$  represent the lower and the upper bounds, respectively. Then, we have

$$\mathbf{y} \approx \mathbf{\Phi} \boldsymbol{\xi}, \quad \boldsymbol{\xi} \sim \mathcal{N}(\mathbf{0}, \mathbf{\Gamma}), \quad \text{subject to} \quad \mathbf{l} \leq \mathbf{\Lambda} \boldsymbol{\xi} \leq \mathbf{u}. \quad (2)$$

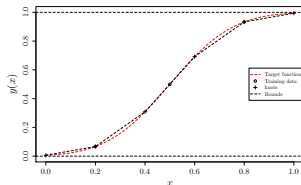
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## Boundedness constraint

$$\underbrace{\begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}}_{\mathbf{l}_b} \leq \underbrace{\begin{bmatrix} 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 1 \end{bmatrix}}_{\mathbf{\Lambda}_b} \underbrace{\begin{bmatrix} \xi_1 \\ \xi_2 \\ \xi_3 \\ \xi_4 \\ \xi_5 \end{bmatrix}}_{\boldsymbol{\xi}} \leq \underbrace{\begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \\ 1 \end{bmatrix}}_{\mathbf{u}_b}$$



The finite approximation is also bounded...

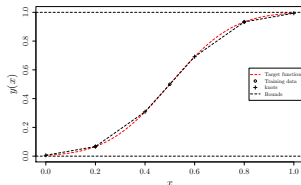
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## Monotonicity constraint

$$\underbrace{\begin{bmatrix} 0 \\ 0 \\ 0 \\ 0 \\ 0 \end{bmatrix}}_{\mathbf{l}_m} \leq \underbrace{\begin{bmatrix} 1 & 0 & 0 & 0 & 0 \\ -1 & 1 & 0 & 0 & 0 \\ 0 & -1 & 1 & 0 & 0 \\ 0 & 0 & -1 & 1 & 0 \\ 0 & 0 & 0 & -1 & 1 \end{bmatrix}}_{\mathbf{\Lambda}_m} \underbrace{\begin{bmatrix} \xi_1 \\ \xi_2 \\ \xi_3 \\ \xi_4 \\ \xi_5 \end{bmatrix}}_{\boldsymbol{\xi}} < \underbrace{\begin{bmatrix} \infty \\ \infty \\ \infty \\ \infty \\ \infty \end{bmatrix}}_{\mathbf{u}_m}$$



The finite approximation is also bounded and monotonic.

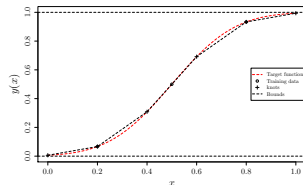
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## Boundedness + Monotonicity

$$\mathbf{l} = \begin{bmatrix} l_b \\ l_m \end{bmatrix}, \quad \mathbf{\Lambda} = \begin{bmatrix} \Lambda_b \\ \Lambda_m \end{bmatrix}, \quad \mathbf{u} = \begin{bmatrix} u_b \\ u_m \end{bmatrix}.$$



The finite approximation is also bounded and monotonic.

# Constrained Kriging: Maatouk and Bay (2016)

Finally, we obtain

$$\xi_C | \mathbf{y} \sim \mathcal{TN}(\Lambda \mu_{\xi | \mathbf{y}}, \Lambda \Sigma_{\xi | \mathbf{y}} \Lambda^\top, \mathbf{l}, \mathbf{u}),$$

where  $\xi_C : \mathbf{l} \leq \Lambda \xi_C \leq \mathbf{u}$ , and

$$\mu_{\xi | \mathbf{y}} = \Gamma \Phi^\top (\Phi \Gamma \Phi^\top)^{-1} \mathbf{y},$$

$$\Sigma_{\xi | \mathbf{y}} = \Gamma - \Gamma \Phi^\top (\Phi \Gamma \Phi^\top)^{-1} \Phi \Gamma.$$

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$$\begin{aligned} \mu_{\xi | \mathbf{y}} &= \mathbf{\Gamma} \mathbf{\Phi}^\top (\mathbf{\Phi} \mathbf{\Gamma} \mathbf{\Phi}^\top)^{-1} \mathbf{y}, \\ \Sigma_{\xi | \mathbf{y}} &= \mathbf{\Gamma} - \mathbf{\Gamma} \mathbf{\Phi}^\top (\mathbf{\Phi} \mathbf{\Gamma} \mathbf{\Phi}^\top)^{-1} \mathbf{\Phi} \mathbf{\Gamma}. \end{aligned}$$

$\Rightarrow$  Maatouk and Bay (2016) proposed an MC approach based on rejection sampling known as *Rejection Sampling from the Mode* (RSM).

# Constrained Kriging: Maatouk and Bay (2016)

Maatouk and Bay (2016) proposed an MC approach based on rejection sampling known as *Rejection Sampling from the Mode* (RSM).

## ✓ Advantages

- It is an exact approach for sampling (uncorrelated samples).
- It is easy to implement.



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## ✗ Disadvantages

- However, simulating from a truncated multivariate Gaussian is required: for higher dimensions the acceptance rate is smaller!

# Constrained Kriging: Maatouk and Bay (2016)

## ♦ Aim

- To simulate from truncated high dimensional Gaussians!

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⇒ Let's call some MCMC techniques! (Bishop, 2007; Murphy, 2012)

- Gibbs sampling.
- Metropolis-Hastings (MH) algorithm.
- Hamiltonian Monte Carlo (HMC).

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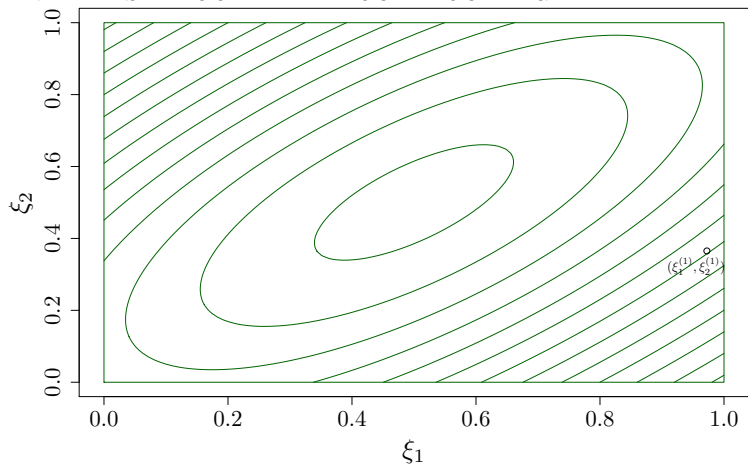
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# MCMC: Gibbs Sampling (Gibbs)

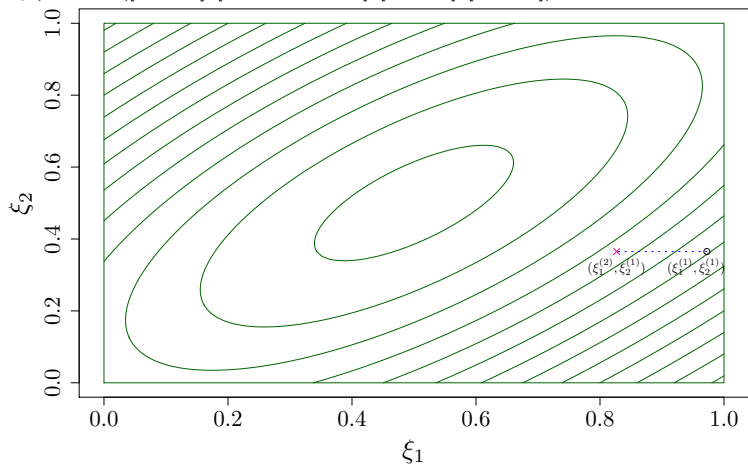
**Figure:** Solid red lines represent the contour lines of the bivariate Gaussian  $(\xi_1, \xi_2) \sim \mathcal{TN}([0.5 \ 0.5], [1.0 \ 0.7; 0.7 \ 1.0], [0.0 \ 0.0], [1.0 \ 1.0])$ .



Iteration 1

# MCMC: Gibbs Sampling (Gibbs)

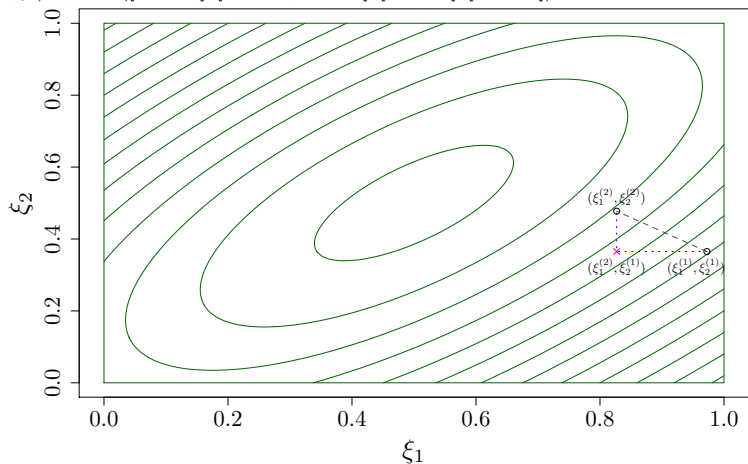
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Iteration 2: simulating the 1st coordinate using  $p(\xi_1^{(2)} | \xi_2^{(1)})$

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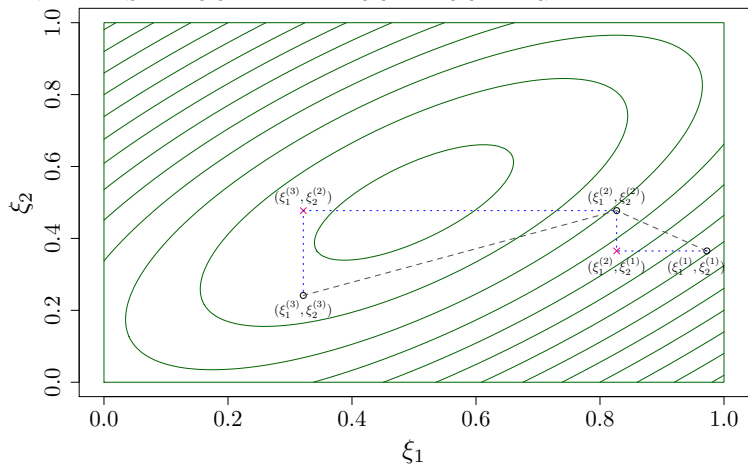
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Iteration 2: simulating the 2nd coordinate using  $p(\xi_2^{(2)} | \xi_1^{(2)})$

# MCMC: Gibbs Sampling (Gibbs)

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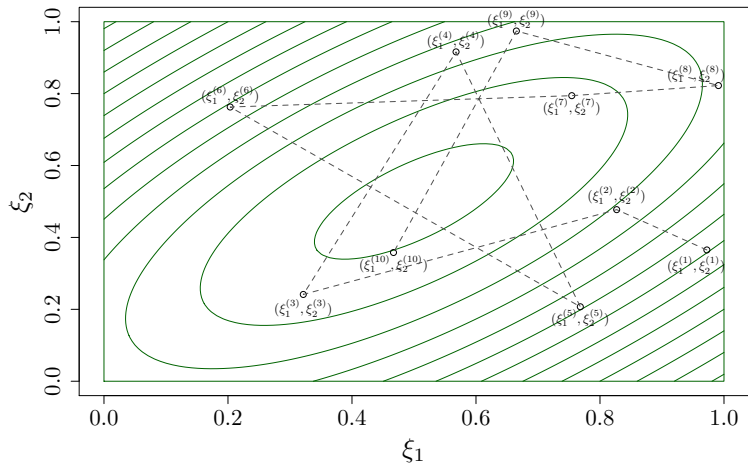


Iteration 3



# MCMC: Gibbs Sampling (Gibbs)

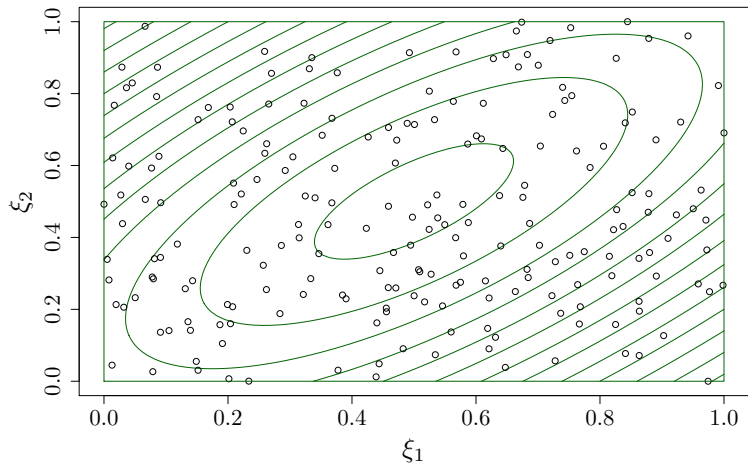
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Iteration 10

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Iteration 200

# MCMC: Gibbs Sampling (Gibbs)

## ✓ Advantages

- Unlike RSM, there is no rejection step.
- Simulating from a truncated multivariate Gaussian is reduced to a sequential sampling from truncated univariate Gaussians.
- There are efficient implementations in R (e.g `tmvtnorm` package).

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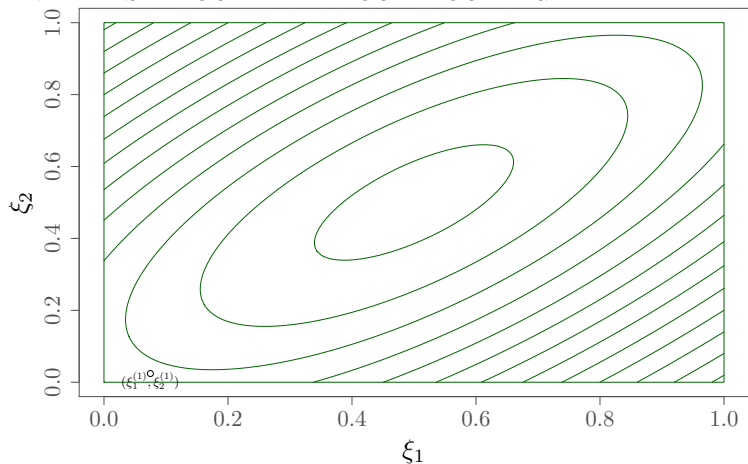
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## ✗ Disadvantages

- It is necessary to discard intermediate simulations to obtain less correlated samples (*thinning* effect).

# MCMC: Metropolis-Hastings (MH)

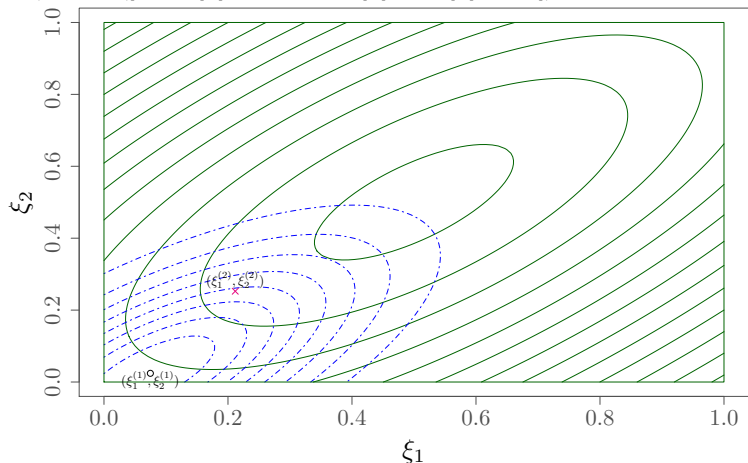
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Iteration 1

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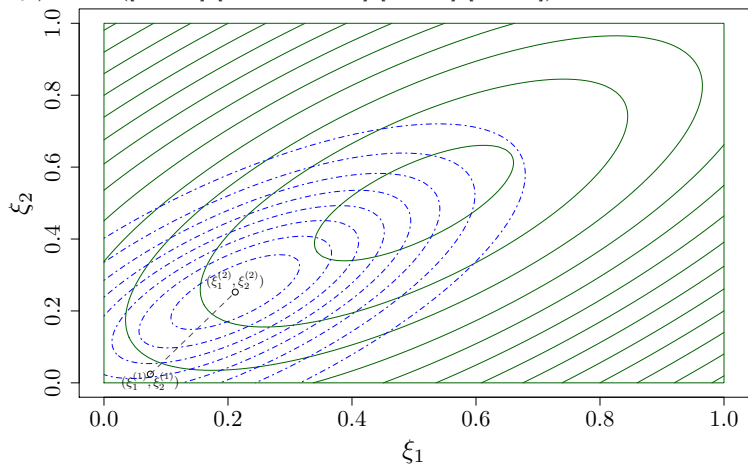
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Iteration 2: simulating from  $q(\boldsymbol{\xi}^{(2)}|\boldsymbol{\xi}^{(1)}) = \mathcal{N}(\boldsymbol{\xi}^{(2)}|\boldsymbol{\xi}^{(1)}, \eta\boldsymbol{\Sigma})$

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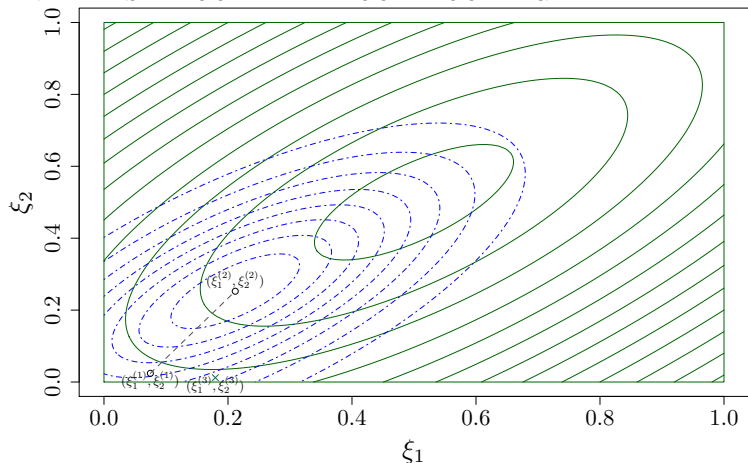
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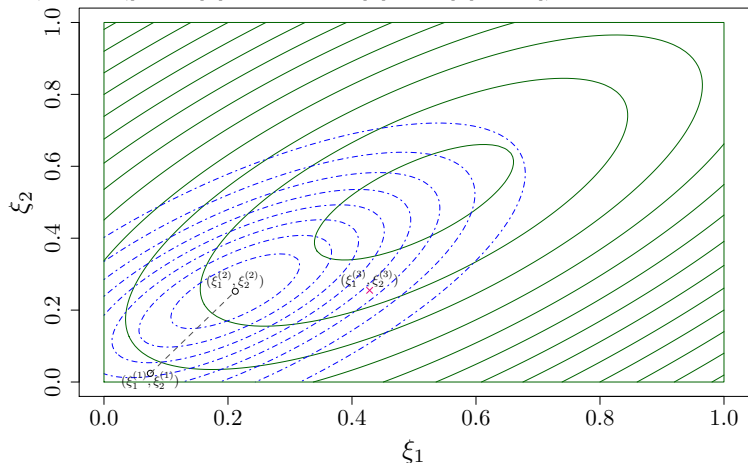


Iteration 3: simulating from  $q(\boldsymbol{\xi}^{(3)}|\boldsymbol{\xi}^{(2)}) = \mathcal{N}(\boldsymbol{\xi}^{(3)}|\boldsymbol{\xi}^{(2)}, \eta\boldsymbol{\Sigma})$



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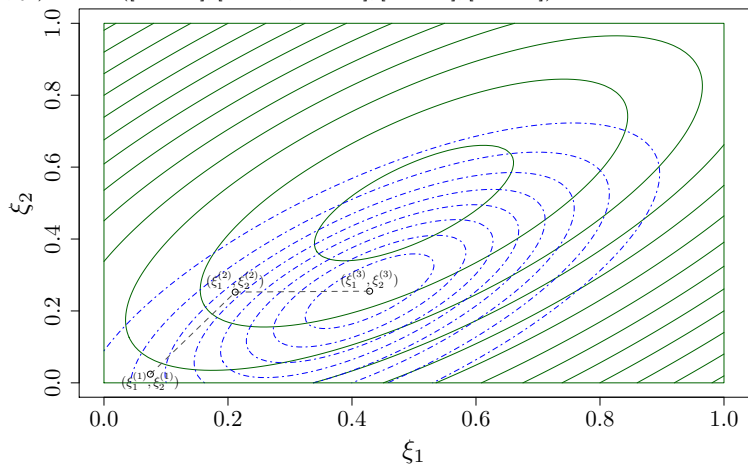
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Iteration 3: resampling from  $q(\xi^{(3)}|\xi^{(2)}) = \mathcal{N}(\xi^{(3)}|\xi^{(2)}, \eta\Sigma)$

# MCMC: Metropolis-Hastings (MH)

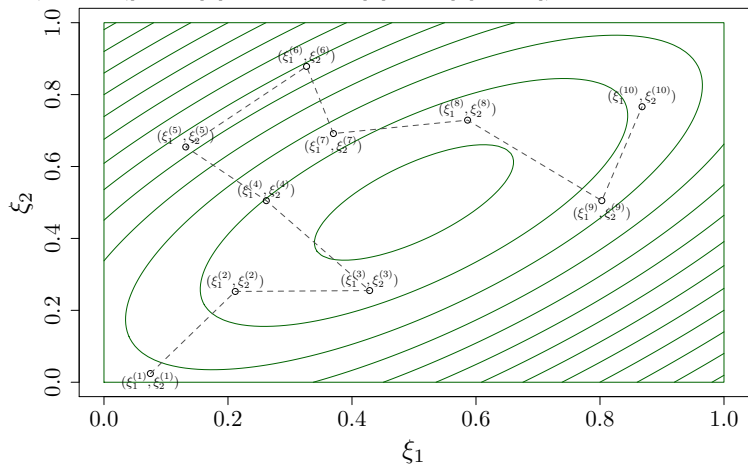
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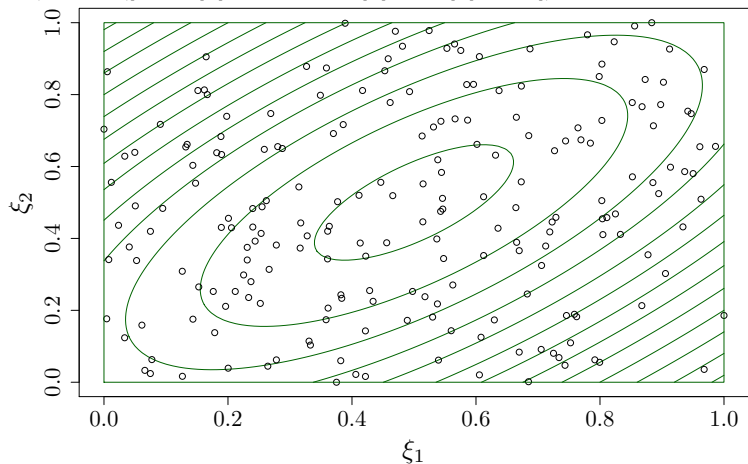
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## ✗ Disadvantages

- The scale factor  $\eta$  has to be tuned.
- Like RSM, simulating from a truncated multivariate Gaussian is required: higher dimensions smaller acceptance rate!

# MCMC: Hamiltonian Monte Carlo (HMC)

- Duane et al. (1987) introduced an efficient hybrid approach using the properties of the Hamiltonian dynamics.
- According to physical systems, HMC is employed to simulate

$$p(\mathbf{x}) \propto \exp \left\{ -\frac{U(\mathbf{x})}{T} \right\}, \quad (3)$$

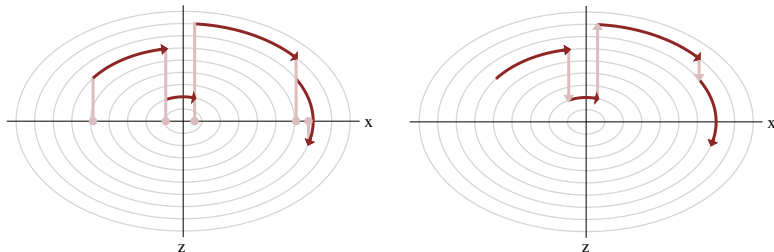
where  $U(\mathbf{x})$  is the *energy* of the state  $\mathbf{x}$ , and  $T$  is the temperature.

- For the Gaussian case,  $p(\mathbf{x}) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ ,

$$U(\mathbf{x}) = \frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu}) \quad \text{and} \quad T = 1.$$

# MCMC: Hamiltonian Monte Carlo (HMC)

**Figure:** HMC in 1-dimensional example. Figures from (Betancourt, 2017).



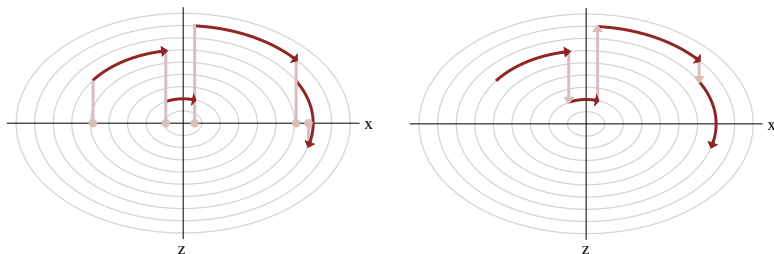
$$H = U(\mathbf{x}) + K(\mathbf{z}) = \frac{1}{2} \mathbf{x}^\top \boldsymbol{\Sigma}^{-1} \mathbf{x} + \frac{1}{2} \mathbf{z}^\top \mathbf{M}^{-1} \mathbf{z},$$

where  $\mathbf{z}$  is a vector with the auxiliary *momentum* variables.



# MCMC: Hamiltonian Monte Carlo (HMC)

**Figure:** HMC in 1-dimensional example. Figures from (Betancourt, 2017).



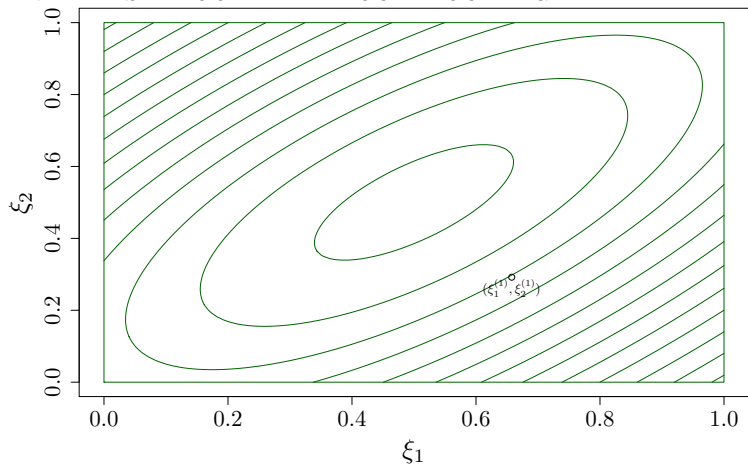
$$H = U(\mathbf{x}) + K(\mathbf{z}) = \frac{1}{2} \mathbf{x}^\top \boldsymbol{\Sigma}^{-1} \mathbf{x} + \frac{1}{2} \mathbf{z}^\top \mathbf{M}^{-1} \mathbf{z},$$

where  $\mathbf{z}$  is a vector with the auxiliary *momentum* variables.

Toy example on-line: [\[url\]](#)

# MCMC: Hamiltonian Monte Carlo (HMC)

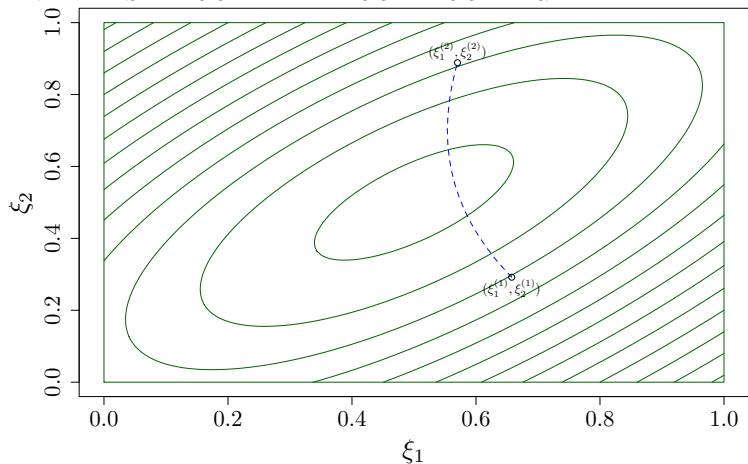
**Figure:** Solid red lines represent the contour lines of the bivariate Gaussian  $(\xi_1, \xi_2) \sim \mathcal{TN}([0.5 \ 0.5], [1.0 \ 0.7; 0.7 \ 1.0], [0.0 \ 0.0], [1.0 \ 1.0])$ .



Iteration 1

# MCMC: Hamiltonian Monte Carlo (HMC)

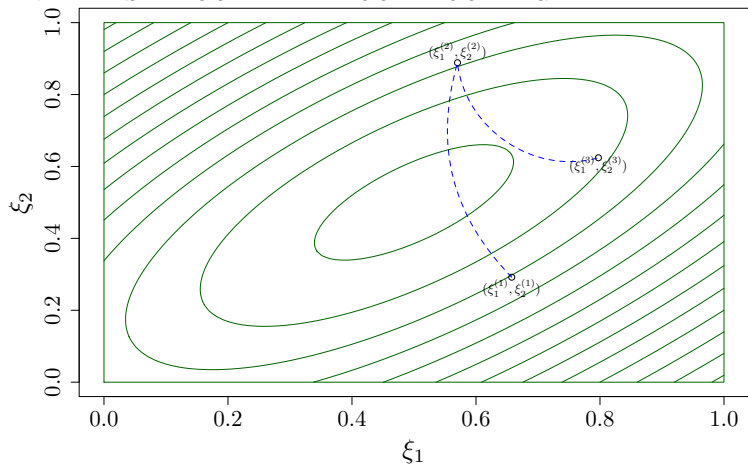
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Iteration 2

# MCMC: Hamiltonian Monte Carlo (HMC)

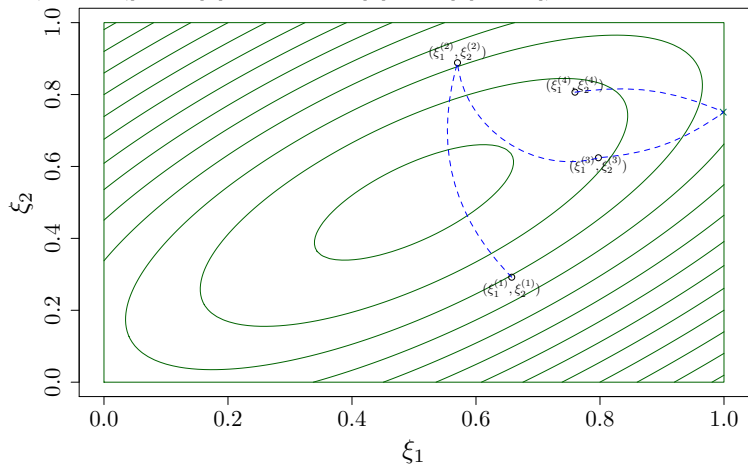
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Iteration 3

# MCMC: Hamiltonian Monte Carlo (HMC)

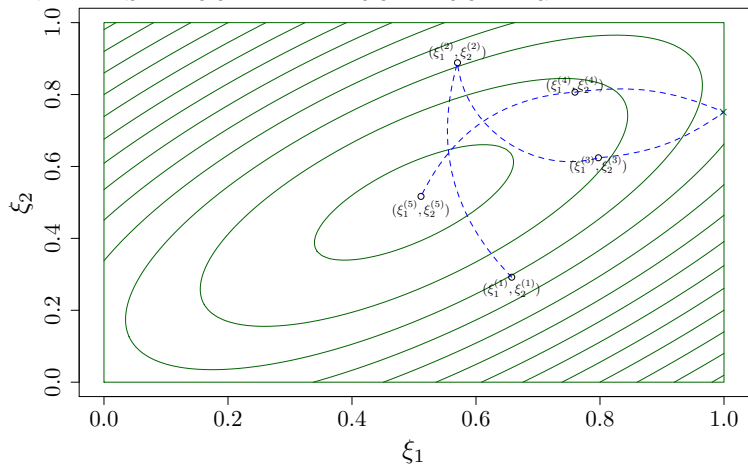
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Iteration 4

# MCMC: Hamiltonian Monte Carlo (HMC)

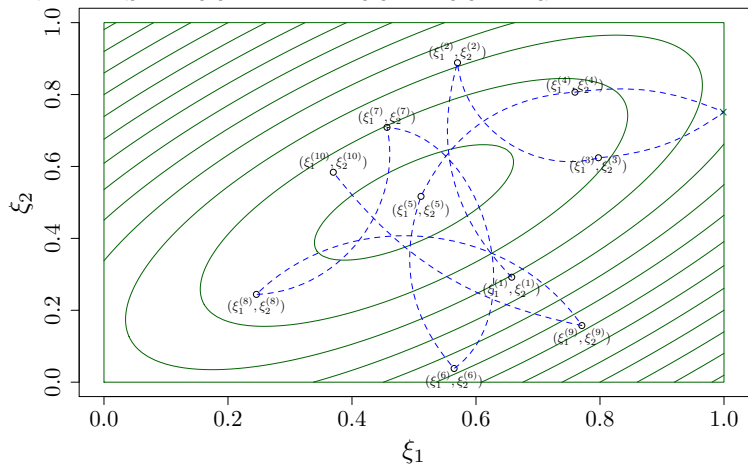
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Iteration 5

# MCMC: Hamiltonian Monte Carlo (HMC)

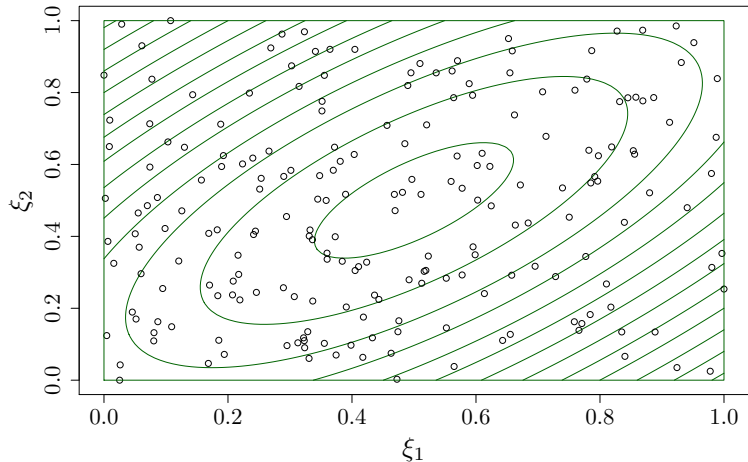
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Iteration 10

# MCMC: Hamiltonian Monte Carlo (HMC)

**Figure:** Solid red lines represent the contour lines of the bivariate Gaussian  $(\xi_1, \xi_2) \sim \mathcal{TN}([0.5 \ 0.5], [1.0 \ 0.7; 0.7 \ 1.0], [0.0 \ 0.0], [1.0 \ 1.0])$ .



Iteration 200



# MCMC: Hamiltonian Monte Carlo (HMC)

## ✓ Advantages

- It is an efficient approach to simulate less correlated samples.
- There is an exact HMC for sampling from truncated Gaussians with linear inequality constraints (Pakman and Paninski, 2014).
- The approach from (Pakman and Paninski, 2014) is already implemented in R: `tmg` package.
- It is the trending topic in the state-of-the-art.

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- It is the trending topic in the state-of-the-art.

## ✗ Disadvantages

- Some parameters have to be tuned (e.g. travel time between samples).
- Sometimes it requires some expensive computations (e.g. computing hitting instants and reflection velocities) (Pakman and Paninski, 2014).

# MCMC: Hamiltonian Monte Carlo (HMC)

**Table:** Comparison between proposed MCMC techniques w.r.t. the RSM approach

	RSM	Gibbs	MH	HMC
Exact method	✓	✗	✗	✗
Non parametric	✓	✓	✗	✗✓
Acceptance rate	✗	✓✓	✗-✓	✓
Speed	✗	✓	✗-✓	✓
Uncorrelated samples	-	✗	✗-✓	✓
Package	-	tmvtnorm	-	tmg

RSM: Rejection Sampling from the Mode

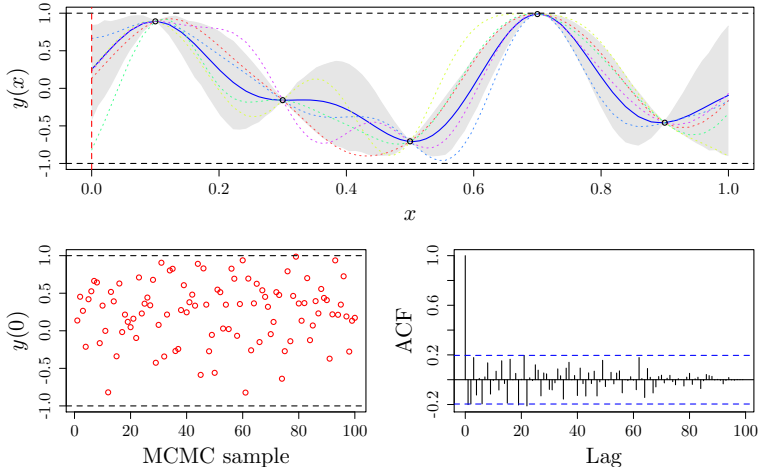
MH: Metropolis-Hastings

HMC: Hamiltonian Monte Carlo

# Partial results: simulating from the model

## Toy example 1

Performance of the proposed MCMC approaches for constrained Kriging w.r.t. the one obtained using RSM.

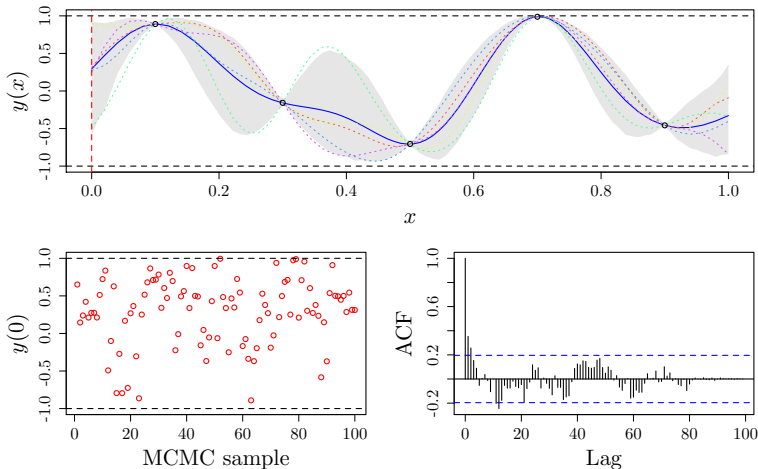


Rejection Sampling from the Mode (RSM) - Gold Standard

# Partial results: simulating from the model

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Performance of the proposed MCMC approaches for constrained Kriging w.r.t. the one obtained using RSM.

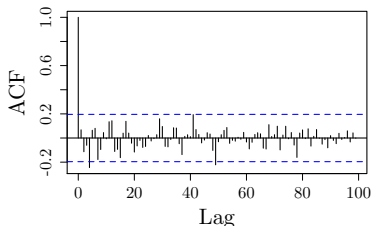
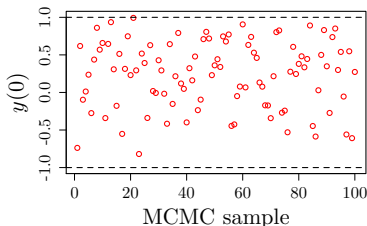
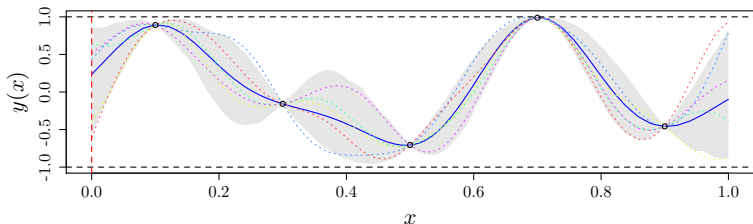


Gibbs sampling with thinning = 100

# Partial results: simulating from the model

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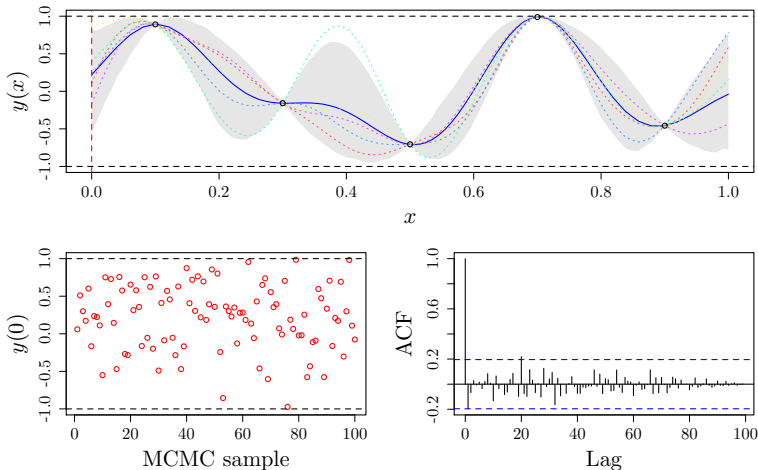


Random walk Metropolis algorithm with  $\eta = 1$

# Partial results: simulating from the model

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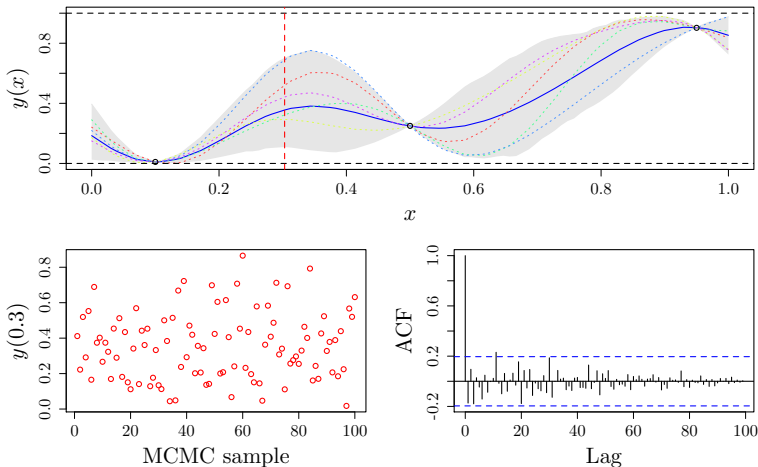


Hamiltonian Monte Carlo (HMC)

# Partial results: simulating from the model

## Toy example 2

Performance of the proposed model for Kriging under different types of constraints using HMC. Target function:  $y = x^2$ .



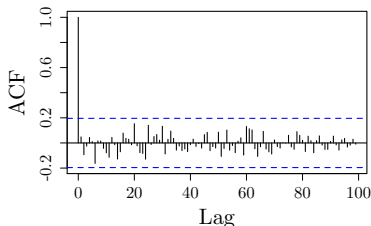
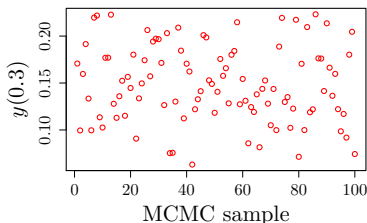
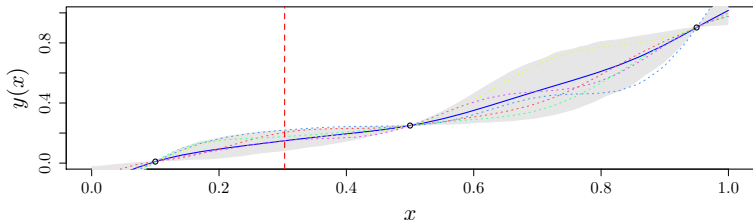
Boundedness



# Partial results: simulating from the model

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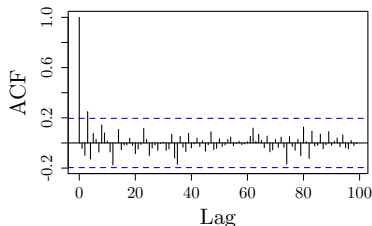
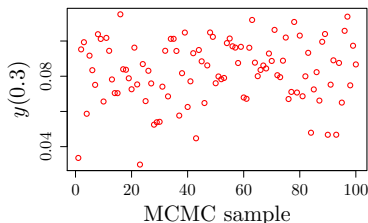
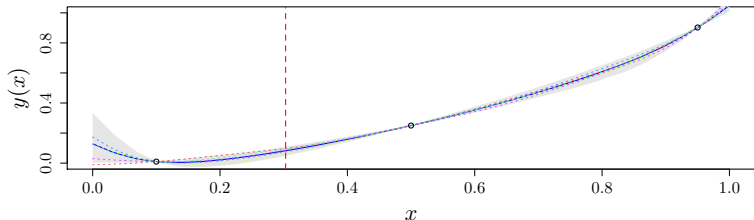


Monotonicity

# Partial results: simulating from the model

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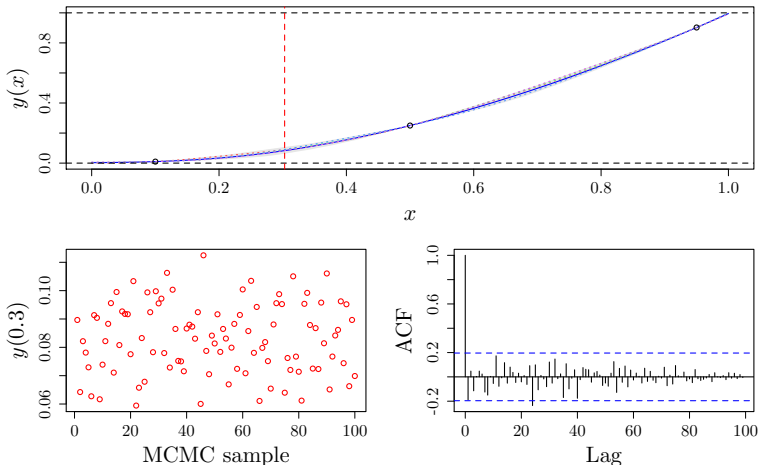


Convexity

# Partial results: simulating from the model

## Toy example 2

Performance of the proposed model for Kriging under different types of constraints using HMC. Target function:  $y = x^2$ .



Boundedness + Monotonicity + Convexity

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- Hyperparameters Estimation: Constrained Maximum Likelihood (CML)

## 3 Conclusions and Future Works

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# Proposal: Maximum Likelihood (ML)

Commonly, the hyperparameters  $\Theta$  of a Gaussian process (GP) are estimated by minimizing  $\mathcal{L} = -\log p(\mathbf{y}|\Theta)$  known as the *negative log likelihood* (NLL). Let  $\mathbf{y}|\Theta \sim \mathcal{GP}(\mathbf{0}, \Sigma_{\Theta})$ , the NLL follows

$$\mathcal{L} = -\log p(\mathbf{y}|\Theta) = \frac{d}{2} \log(2\pi) + \frac{1}{2} \log |\Sigma_{\Theta}| + \frac{1}{2} \mathbf{y}^{\top} \Sigma_{\Theta}^{-1} \mathbf{y}, \quad (4)$$

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where the covariance  $\Sigma_{\Theta}$  depends on the hyperparameters  $\Theta$ .

$\Rightarrow$  Equation (4) can be optimized using standard gradient algorithms.

$\Rightarrow$  It does not take into account the inequalities  $\mathcal{C} : l \leq \Lambda \xi \leq u$ .

# Proposal: Maximum Likelihood (ML)

In order to take into account a prior the fact that  $\xi_c : l \leq \Lambda \xi \leq u$ , we propose a constrained version of the NLL given by

$$\begin{aligned}
 \mathcal{L}^* &= -\log p(\mathbf{y}|\xi_c, \Theta) \\
 &= -\log \frac{p(\mathbf{y}|\Theta)p(\xi_c|\mathbf{y}, \Theta)}{p(\xi_c|\Theta)} \\
 &= -\log p(\mathbf{y}|\Theta) - \log p(\xi_c|\mathbf{y}, \Theta) + \log p(\xi_c|\Theta).
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$\Rightarrow$  Good news!! Green terms can be properly estimated using the theory behind Gaussian orthant probabilities (e.g. `mvtnorm` package (Genz, 1992); `TruncatedNormal` package, (Botev, 2017)).

# Partial results: estimating the hyperparameters

## 1-dimensional example

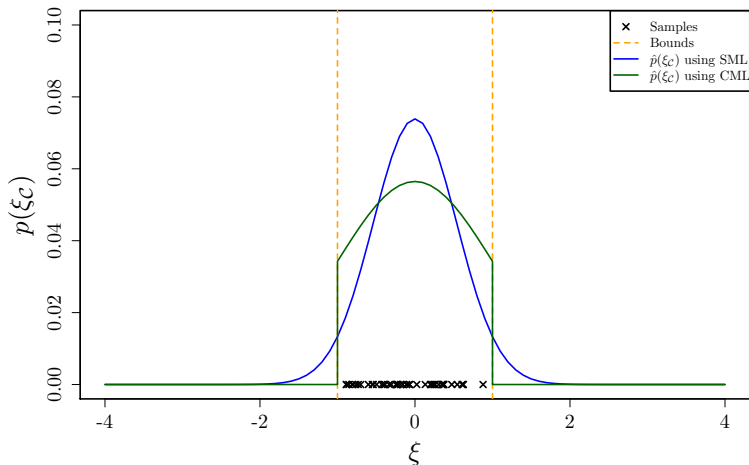
In this example, we estimate  $\hat{\sigma}$  given different nb of samples.

- Simulations:  $\xi_{C_i} \sim \mathcal{TN}(0, \sigma^2 = 1^2, -1, 1)$  for  $i = 1, \dots, n$ .
- Target function:  $y_i = \xi_{C_i}$ .
- We estimate  $\hat{\sigma}$  by maximizing each type of likelihood
  - Standard ML (SML):  $\hat{\sigma} = \arg \max_{\sigma} \log p(y|\sigma)$ .
  - Constrained ML (CML):  $\hat{\sigma} = \arg \max_{\sigma} \log p(y|\xi_C, \sigma)$ .
- We repeat the experiment 100 times.

# Partial results: estimating the hyperparameters

## 1-dimensional example

Estimating  $\hat{\sigma}$  given different nb of samples.

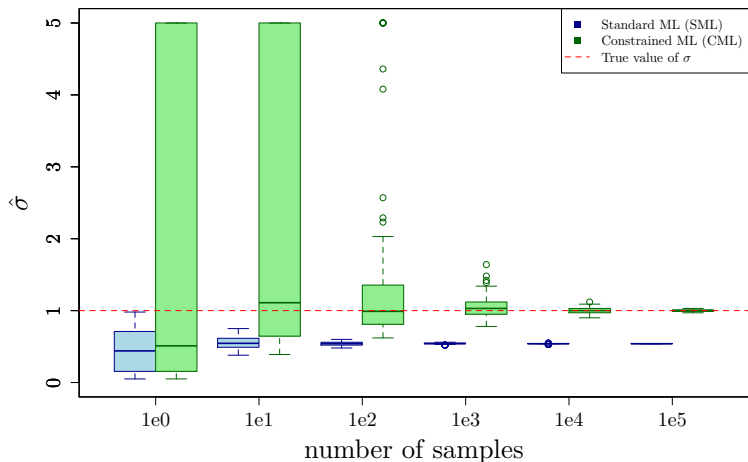


■ Standard ML (SML) ■ Constrained ML (CML)

# Partial results: estimating the hyperparameters

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# Proposal: estimating the hyperparameters

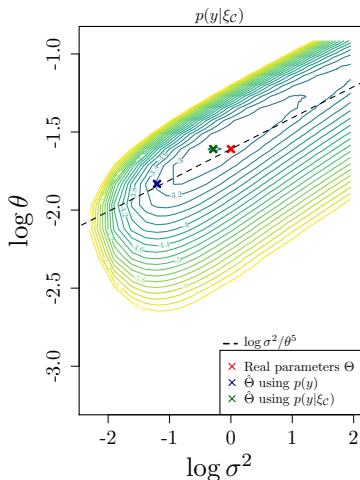
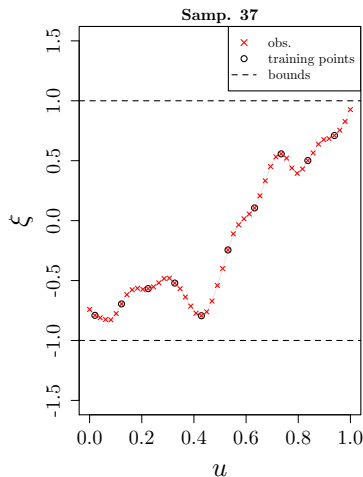
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We estimate  $\hat{\Theta} = (\hat{\sigma}^2, \hat{\theta})$  using different GP simulations.

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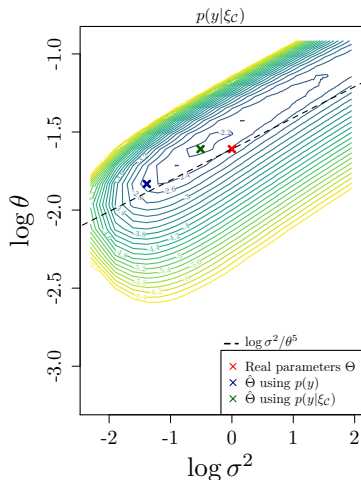
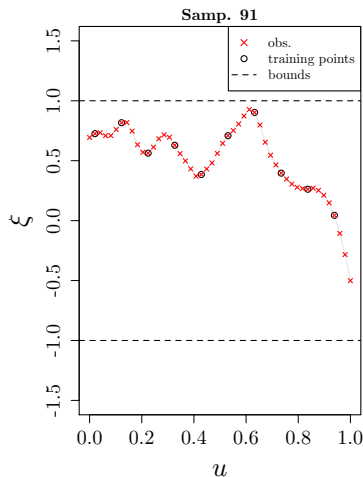


× Standard ML (SML) × Constrained ML (CML)

# Proposal: estimating the hyperparameters

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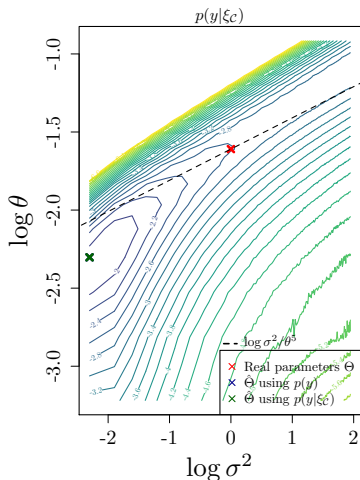
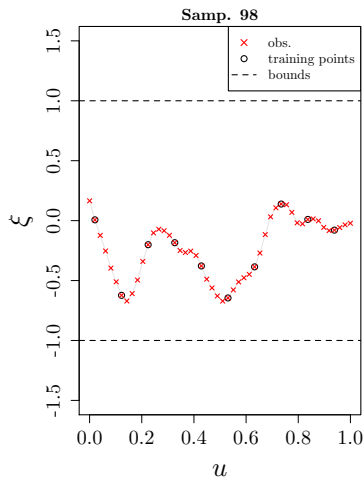


$\times$  Standard ML (SML)  $\times$  Constrained ML (CML)

# Proposal: estimating the hyperparameters

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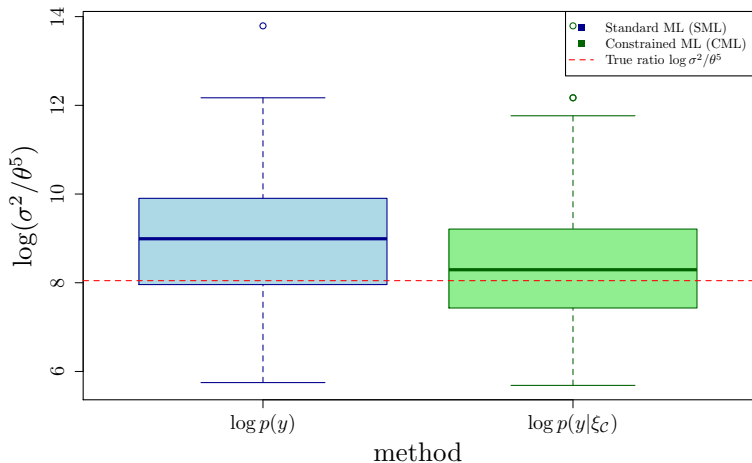
× Standard ML (SML) × Constrained ML (CML)



# Proposal: estimating the hyperparameters

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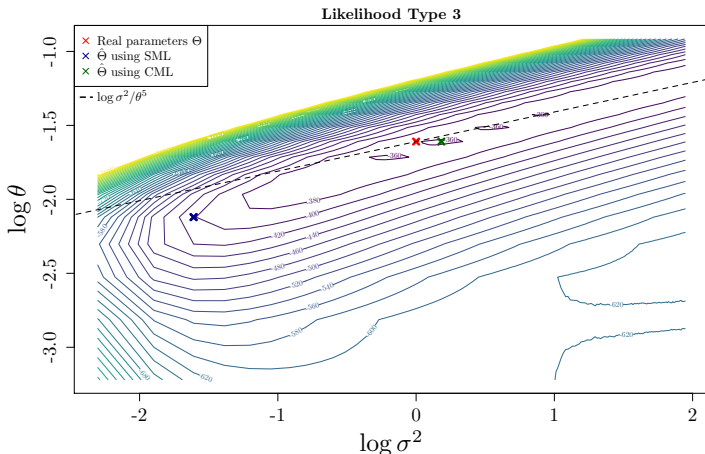


■ Real value ■ Standard ML (SML) ■ Constrained ML (CML)

# Proposal: estimating the hyperparameters

## Multi-dimensional example.

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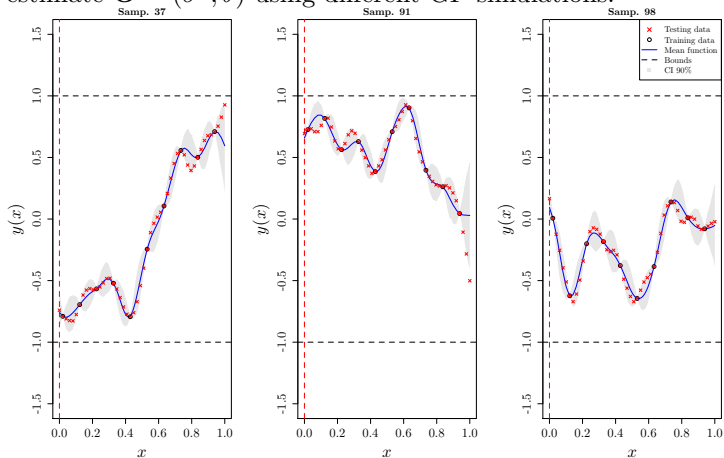


× Real value × Standard ML (SML) × Constrained ML (CML)

# Proposal: estimating the hyperparameters

## Multi-dimensional example.

We estimate  $\hat{\Theta} = (\hat{\sigma}^2, \hat{\theta})$  using different GP simulations.



SMSE: (left) 0.00778823 (centre) 0.08241287 (right) 0.0542675

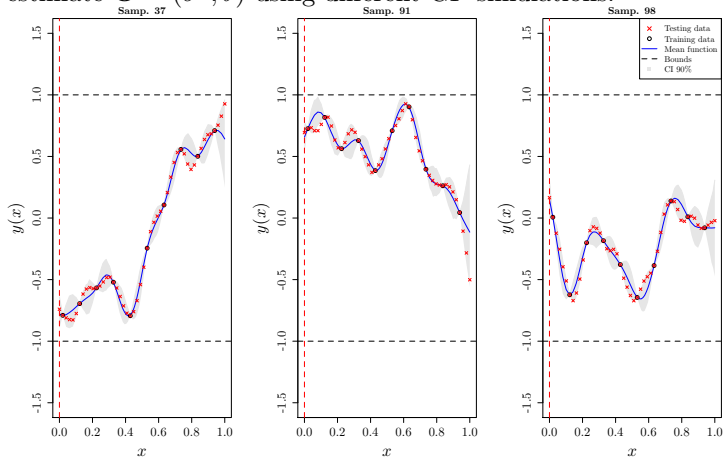
MSLL: (left) -3.935441 (centre) -3.184300 (right) -2.929120

Unconstrained ML.

# Proposal: estimating the hyperparameters

## Multi-dimensional example.

We estimate  $\hat{\Theta} = (\hat{\sigma}^2, \hat{\theta})$  using different GP simulations.



SMSE: (left) 0.0106797 (centre) 0.07928962 (right) 0.0542675

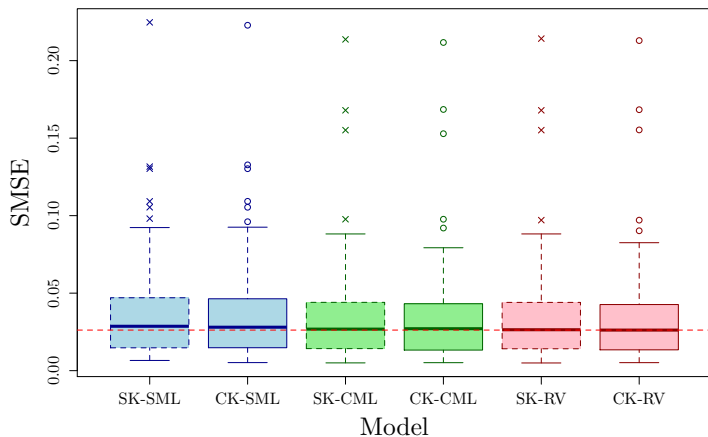
MSLL: (left) -3.960415 (centre) -3.218316 (right) -2.929120

Constrained ML.

# Proposal: estimating the hyperparameters

## Multi-dimensional example.

We estimate  $\hat{\Theta} = (\hat{\sigma}^2, \hat{\theta})$  using different GP simulations.

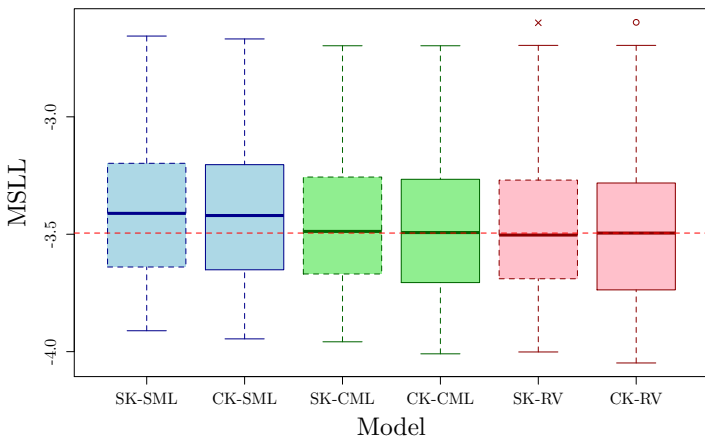


■ Standard ML (SML) 
 ■ Constrained ML (CML) 
 ■ Real  $\Theta$   
× Standard Kriging (SK) 
 ○ Constrained Kriging (CK)

# Proposal: estimating the hyperparameters

## Multi-dimensional example.

We estimate  $\hat{\Theta} = (\hat{\sigma}^2, \hat{\theta})$  using different GP simulations.



■ Standard ML (SML) 
 ■ Constrained ML (CML) 
 ■ Real  $\Theta$   
 × Standard Kriging (SK) 
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# Conclusions and Future Works

## Conclusions

- We further investigated the approach proposed in (Maatouk and Bay, 2016): now it works for any linear set of inequality constraints.
- We implemented several simulation methods based on MCMC approaches.
- We investigated a proper likelihood for hyperparameters estimation.
- We implemented the codes in R (they are almost an R package).



# Conclusions and Future Works

## Future works

- To implement a gradient-based method to estimate automatically the hyperparameters of the model.
- To investigate theoretical estimation properties from the constrained likelihood.
- To evaluate the proposed approach with real problems.
- To build the R package.
- To extend this approach when the input space is multi-dimensional, i.e.  $\mathbf{x} \in [0\ 1]^d$ .

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