

# Causal Disparity Analysis

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# Takeaway messages

- ① Causal disparity analysis (CDA) has a similar structure to causal mediation analysis.
- ② But these two are different species.
- ③ Causal disparity analysis is a hot topic, feel free to consider writing your final paper on it.

# What questions do we ask in CDA?

- How much, and in what ways, do racial differences in medical care contribute to racial disparities in health?  
race  $\rightarrow$  medical care  $\rightarrow$  health
- In which ways does occupational sorting contribute to the gender wage gap?  
gender  $\rightarrow$  occupation  $\rightarrow$  wage
- What are the roles of educational attainment in the relationship between social class origin and destination?  
origin  $\rightarrow$  education  $\rightarrow$  destination
- Does racial differences in incarceration explain the racial disparity in income?  
race  $\rightarrow$  incarceration  $\rightarrow$  income

# The common structure of these questions

- Three temporally ordered variables  
group  $\rightarrow$  treatment  $\rightarrow$  outcome
- Goal of CDA:  
Quantify the mechanisms by which a treatment variable causally explains a descriptive group disparity
- Why do we care about *descriptive* group disparity?
- Why do we care about *causal* explanation of group disparities?

## Wait, isn't this a mediation class?

- In causal mediation analysis (CMA), we also have a three variable structure  
treatment  $\rightarrow$  mediator  $\rightarrow$  outcome
- But, CMA explains causal effect of treatment, not *descriptive* group disparity, e.g.,

$$E(Y^a - Y^{a*}) = E(Y^{aM^a} - Y^{aM^{a*}}) + E(Y^{aM^{a*}} - Y^{a^*M^{a*}}).$$

Randomized interventional analogues:

$$E(Y^{aG^a} - Y^{a^*G^{a*}}) = E(Y^{aG^a} - Y^{aG^{a*}}) + E(Y^{aG^{a*}} - Y^{a^*G^{a*}}).$$

# How can we do CDA (causal disparity analysis)?

- 1 Kitagawa-Blinder-Oaxaca decomposition (Kitagawa, 1955; Blinder, 1973; Oaxaca, 1973)
- 2 Random equalization decomposition (VanderWeele and Robinson, 2014; Jackson and VanderWeele, 2018; Jackson, 2022; Lundberg, 2022)
- 3 Self-promotion: Yu and Elwert (2023)

Group:  $G \in \{a, b\}$ ; treatment:  $D$ ; outcome:  $Y$ ; baseline covariates:  $X$

Group-specific linear regression:

$$Y = \alpha_g + \beta_g D + \gamma_g^T X + \epsilon.$$

KBO decomposition:

$$\begin{aligned} E_a(Y) - E_b(Y) = & \underbrace{\alpha_a - \alpha_b}_{\text{intercept}} + \underbrace{\beta_b[E_a(D) - E_b(D)]}_{\text{characteristic}} + \underbrace{E_a(D)[\beta_a - \beta_b]}_{\text{slope}} \\ & + \gamma_b^T [E_a(X) - E_b(X)] + E_a(X)^T [\gamma_a - \gamma_b]. \end{aligned}$$

Problem?

# Random equalization decomposition

Two versions: Jackson's and Lundberg's

$E_g(Y^{R(D|g')})$ : post-intervention mean outcome for group  $g$  after each member of group  $g$  is counterfactually given a treatment value randomly drawn from group  $g'$ .

The Jackson version:

$$E_a(Y) - E_b(Y) = \underbrace{E_b(Y^{R(D|a)}) - E_b(Y)}_{\text{change in disparity}} + \underbrace{E_a(Y) - E_b(Y^{R(D|a)})}_{\text{remaining disparity}}.$$

This corresponds to Propositions 1 and 4 in Jackson and VanderWeele (2018). Let's verify this.



# Random equalization decomposition

$E_g(Y^{R(D)})$ : post-intervention mean outcome for group  $g$  after each member of group  $g$  is counterfactually given a treatment value randomly drawn from the whole population.

The Lundberg version:

$$E_a(Y) - E_b(Y) = \underbrace{E_a(Y) - E_b(Y) - \left[ E_a(Y^{R(D)}) - E_b(Y^{R(D)}) \right]}_{\text{change in disparity}} + \underbrace{E_a(Y^{R(D)}) - E_b(Y^{R(D)})}_{\text{remaining disparity}}$$

Recall equation (3) in Lundberg (2022), p.15

This corresponds to Intervention c in Figure 6 in Lundberg (2022).

What's good about random equalization decompositions?

Do they have any problems, though?

# Problem with random equalization decomposition

We want a decomposition that can capture the contribution of differential prevalence of  $D$  between the two groups.

Does change in disparity work? Not really...

Null criterion: a valid measure of the contribution of differential prevalence should be zero when there is no group difference in treatment prevalence, i.e.,  $E_a(D) = E_b(D)$ .

Change in disparity in Jackson's version  $E_b(\tau)[E_a(D) - E_b(D)] - \text{Cov}_b(D, \tau)$ .

(the self-equalization paradox)

Change in disparity in Lundberg's version

$E(\tau)[E_a(D) - E_b(D)] - [p_a - p_b][E_a(D) - E_b(D)][E_a(\tau) - E_b(\tau)] + \text{Cov}_a(D, \tau) - \text{Cov}_b(D, \tau)$ .

# A new decomposition

Notation:  $Y^0$ : baseline potential outcome;  $\tau = Y^1 - Y^0$ : treatment effect.  $D \in \{0, 1\}$

$$\begin{aligned} & E_a(Y) - E_b(Y) \\ &= \underbrace{E_a(Y^0) - E_b(Y^0)}_{\text{baseline}} + \underbrace{E_b(\tau)[E_a(D) - E_b(D)]}_{\text{prevalence}} \\ &\quad + \underbrace{E_a(D)[E_a(\tau) - E_b(\tau)]}_{\text{effect}} + \underbrace{\text{Cov}_a(D, \tau) - \text{Cov}_b(D, \tau)}_{\text{selection}}. \end{aligned}$$

Example:

- baseline: racial disparity in outcome unrelated to college attendance
- prevalence: differential college attendance rates
- effect: differential average returns to college attendance
- selection: differential selection into college based on individual-level returns from college

# A new decomposition

Nice things about this:

- ① Nonparametric causal decomposition of group disparities
- ② Four-way
- ③ Novel selection component (related to effectiveness, fairness)
- ④ Satisfies the null criterion
- ⑤ Has interventional interpretation (prescriptive)
- ⑥ Nonparametric estimation with theoretically guaranteed good properties

# Interventional and graphical representations

Two-step intervention: within-group randomization first, between-group equalization second

$$\begin{aligned} E_a(Y) - E_b(Y) - \left[ E_a \left( Y^{R(D|a)} \right) - E_b \left( Y^{R(D|b)} \right) \right] &= \text{selection} \\ E_b \left( Y^{R(D|a)} \right) - E_b \left( Y^{R(D|b)} \right) &= \text{prevalence} \\ E_a \left( Y^{R(D|a)} \right) - E_b \left( Y^{R(D|a)} \right) &= \text{baseline} + \text{effect} \end{aligned}$$

Why we can avoid violating the null criterion: we separate equalization from randomization

Graph on board

# Conditional decomposition

Natural question arising from the interventional/prescriptive interpretation: what if we are not satisfied with the marginal interventions?

Conditional decomposition comes to our rescue

# Conditional decomposition

$$\begin{aligned} E_a(Y) - E_b(Y) = & \underbrace{E_a(Y^0) - E_b(Y^0)}_{\text{baseline}} \\ & + \underbrace{\int [E_a(D | q) - E_b(D | q)] E_b(\tau | q) f_b(q) dq}_{\text{conditional prevalence}} \\ & + \underbrace{\int [E_a(\tau | q) - E_b(\tau | q)] E_a(D | q) f_a(q) dq}_{\text{conditional effect}} \\ & + \underbrace{E_a[\text{Cov}_a(D, \tau | Q)] - E_b[\text{Cov}_b(D, \tau | Q)]}_{\text{conditional selection}} \\ & + \underbrace{\int E_a(D | q) E_b(\tau | q) [f_a(q) - f_b(q)] dq}_{Q \text{ distribution}}. \end{aligned}$$

# Conditional decomposition

Interventional interpretation: within-group conditional randomization first, between-group conditional equalization second

$$\begin{aligned} E_a(Y) - E_b(Y) - \left[ E_a \left( Y^{R(D|a,Q)} \right) - E_b \left( Y^{R(D|b,Q)} \right) \right] &= \text{conditional selection} \\ E_b \left( Y^{R(D|a,Q)} \right) - E_b \left( Y^{R(D|b,Q)} \right) &= \text{conditional prevalence} \\ E_a \left( Y^{R(D|a,Q)} \right) - E_b \left( Y^{R(D|a,Q)} \right) &= \text{baseline} + \text{conditional effect} \\ &\quad + Q \text{ distribution.} \end{aligned}$$

Example:  $Q$ : prior achievement

Why/when would we want a conditional equalization instead of a marginal equalization?



# Conditional decomposition

Proposition 2 in Jackson and VanderWeele (2018):

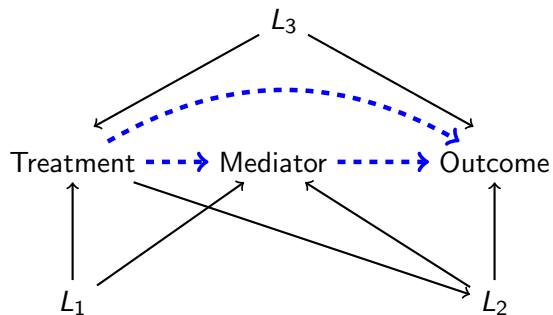
$$E_a(Y) - E_b(Y) = \underbrace{E_b\left(Y^{R(D|a,Q)}\right) - E_b(Y)}_{\text{change in disparity}} + \underbrace{E_a(Y) - E_b\left(Y^{R(D|a,Q)}\right)}_{\text{remaining disparity}}.$$

Similar for Lundberg's version (Intervention d in Figure 6)

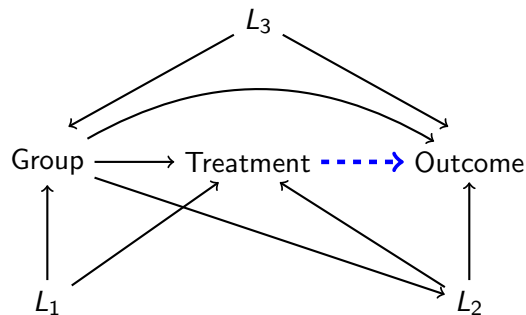
Violate a conditional null criterion

# Now let's talk about identification

Mediation? Random equalization decompositions? Yu and Elwert?



(a) Causal mediation



(b) Group disparity

## A little bit of estimation

Regression, weighting, or a mixture of the two (can be strengthened by data-adaptive techniques like machine learning)

Generally, regression estimators are derived by plugging in parametric functions.

For example, we may assume  $E_g(Y \mid D, X) = \alpha_g + \beta_g D + \gamma_g X$ .

$$\begin{aligned} & E_b \left( Y^{R(D|a)} \right) \\ &= E_b \left( Y^0 \right) [1 - E_a(D)] + E_b \left( Y^1 \right) E_a(D) \\ &= E_b [E_b(Y \mid D = 0, X)] [1 - E_a(D)] + E_b [E_b(Y \mid D = 1, X)] E_a(D) \\ &= [\alpha_b + \gamma_b E_b(X)] [1 - E_a(D)] + [\alpha_b + \beta_b + \gamma_b E_b(X)] E_a(D) \\ &= \alpha_b + \gamma_b E_b(X) + \beta_b E_a(D). \end{aligned}$$

$$\text{Prevalence component: } E_b \left( Y^{R(D|a)} \right) - E_b \left( Y^{R(D|b)} \right) = \beta_b [E_a(D) - E_b(D)].$$

# Estimation

Weighting estimators are generally derived by using tilting functions to “tilt” the factual distribution to an interventional distribution. For  $E_b(Y^{R(D|a)})$ , we can use the weight

$$\frac{f_a(D)}{f_b(D | X)}.$$

For  $E_b(Y^{R(D|a,Q)})$ , we can use the weight

$$\frac{f_a(D | Q)}{f_b(D | X, Q)}.$$

Yu and Elwert (2023) develop a mixture of regression and weighting that is implemented using ML. It has every nice property you can ask for: flexible, efficient, robust.

# Summary

- ① Causal disparity analysis is related to but different from causal mediation analysis.
- ② There are various approaches to do CDA. I think ours is the winner...
- ③ In R, you can use the “cdgd” package.