

Variational Inference for Dirichlet Process Mixtures and Beyond

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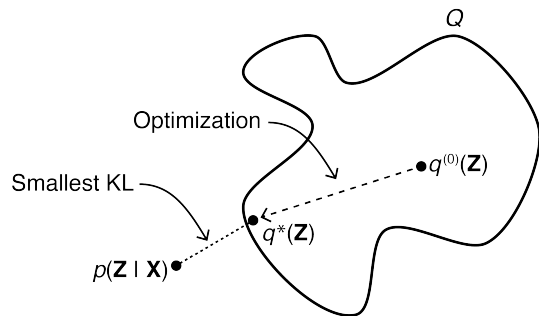


Figure 1: Visualization of VI optimization

Minimization of the Kullback-Leibler (KL) divergence:

$$q^*(z) = \operatorname{argmin}_{q(z) \in \mathcal{Q}} \operatorname{KL}(q(z) \| p(z|x))$$

We choose \mathcal{Q} to be flexible enough to capture a density close to $p(z|x)$, but simple enough for efficient optimization. [2]

Since we cannot compute the KL, we optimize an alternative objective that is equivalent to the KL up to an added constant,

$$\text{ELBO}(q) = \mathbb{E}_q[\log p(z, x)] - \mathbb{E}_q[\log q(z)]$$

This function is called the evidence lower bound (ELBO). Maximizing the ELBO is equivalent to minimizing the KL divergence.

Finally, we approximate the posterior with the optimized member of the family $q^*(.)$.

Mean-field

We work in a mean-field variational framework, which means that latent variables \mathbf{z} are mutually independent and each governed by a distinct factor in the variational density.

$$q(\mathbf{z}) = \prod_{j=1}^m q_j(z_j)$$

Where m is the dimension of the latent variables.

Comparing variational inference and MCMC

MCMC

- Produce exact solution asymptotically
- Computationally intensive
- Suited for:
 1. small datasets
 2. when more precise samples are needed

VARIATIONAL INFERENCE

- Does not provide an exact solution
- Fast convergence
- Suited for:
 1. large datasets
 2. when we want to quickly explore many models.

Algorithm: Mixture of Gaussians

Bayesian Multivariate Gaussian Mixture Model using VI

Hierarchical model:

$$x_i \mid c_i, \mu \stackrel{\text{ind}}{\sim} \mathcal{N}_d \left(c_i^T \mu, \mathbb{I}_d \right) \quad i = 1, \dots, n$$

$$\mu_k \stackrel{\text{iid}}{\sim} \mathcal{N}_d \left(0, \sigma^2 \mathbb{I}_d \right) \quad k = 1, \dots, K$$

$$c_i \stackrel{\text{iid}}{\sim} \text{Cat} \left(\frac{1}{K}, \dots, \frac{1}{K} \right) \quad i = 1, \dots, n$$

- i. K is the number of clusters μ
- ii. c_i is the cluster assignment, it indicates which latent cluster x_i comes from
- iii. n is the size of the sample
- iv. d is the data dimension
- v. σ^2 is a fixed hyperparameter

Variational densities

We want to approximate the posterior of μ_k , which is still a Gaussian, so we consider the following variational densities q :

$$q(\mu_k) \stackrel{\text{iid}}{\sim} \mathcal{N}_d(m_k, s_k^2 \mathbb{I}_d) \quad k = 1, \dots, K$$

While to approximate the cluster allocation parameters we have:

$$P(c_i = j) = \varphi_{ij} \quad k = 1, \dots, K; \quad i = 1, \dots, n$$

To define our variational densities we have to compute the following parameters:

- m : K vectors of means
- s^2 : vector of variances
- φ : $N \times K$ matrix of cluster assignments probabilities

Updates

Variational update for the cluster assignment c_i :

$$\varphi_{ik} \propto \exp[\mathbb{E}[\mu_k; m_k, s_k^2]x_i - \mathbb{E}[\mu_k^2; m_k, s_k^2]/2]$$

Updates

Variational update for the cluster assignment c_i :

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Variational density of the k th mixture component expressed in terms of the variational mean and variance:

$$m_k = \frac{\sum_i \varphi_{ik} x_i}{\frac{1}{\sigma^2} + \sum_i \varphi_{ik}} \qquad s_k^2 = \frac{1}{\frac{1}{\sigma^2} + \sum_i \varphi_{ik}}$$

Optimization

Since VI can be considered an optimization problem, here we have the function to be optimized. In particular, the goal is to maximize the ELBO.

$$\begin{aligned} ELBO(m, s^2, \varphi) = & \sum_{k=1}^K \mathbb{E}[\log p(\mu_k); m_k, s_k^2] + \\ & \sum_{i=1}^n (\mathbb{E}[\log p(c_i); \varphi_i] + \mathbb{E}[\log p(x_i|c_i, \mu); \varphi_i, \hat{m}, s^2]) + \\ & - \sum_{i=1}^n \mathbb{E}[\log q(c_i; \varphi_i)] - \sum_{k=1}^K \mathbb{E}[\log q(\mu_k; m_k, s_k^2)] \end{aligned}$$

Each expectation can be computed in closed form. [3]

Results

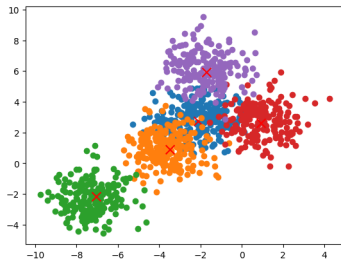


Figure 2: Our generated data

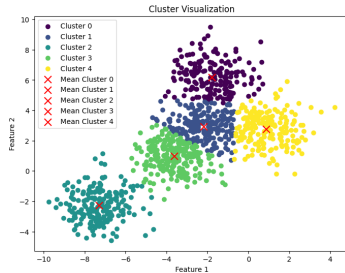


Figure 3: VI clustering

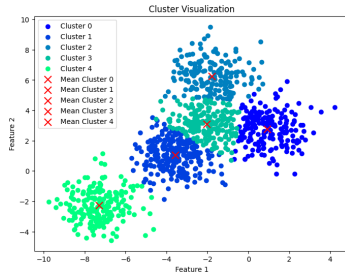


Figure 4: MCMC clustering

Comparison with MCMC (K=5, d=2)

TIME	N = 100	N = 1000	N = 10000
VI	508 ms \pm 6.78 ms	2.31 s \pm 304 ms	1min 16s \pm 778 ms
MCMC	19 s \pm 2 s	1min 47 s \pm 7 s	21 min \pm 1 min

CLUSTER ERROR	N = 100	N = 1000	N = 10000
VI	0.292078797	0.240227441	0.195773545
MCMC	0.326697788	0.246452642	0.197457232

DISTR. ERROR	N = 100	N = 1000	N = 10000
VI	33.17399	34.65825	64.832405
MCMC	0.24917	4.21288	33.59156

What's next?

Our goal

- We will focus on implementing an algorithm for Dirichlet process mixture models using variational inference. The algorithm is described in the paper from Blei and Jordan (2006). [1]
- Moreover we'll compare the results of the algorithm implemented with VI with the results obtained from MCMC implementation.

To do so we'll move to a nonparametric setting.

The Dirichlet process is a stochastic process used in Bayesian nonparametric. It is a distribution over distributions, which means that each draw from a Dirichlet process is itself a distribution.

Distributions drawn from a Dirichlet process are discrete, but cannot be described using a finite number of parameters, thus the classification as a nonparametric model.

Hierarchical model:

$$G \mid \alpha, G_0 \sim DP(\alpha, G_0)$$

$$\eta_n \mid G \stackrel{\text{ind}}{\sim} G$$

$$n = 1, \dots, N$$

$$X_n \mid \eta_n \stackrel{\text{ind}}{\sim} p(x_n, \eta_n)$$

$$n = 1, \dots, N$$

- i. α is a positive scaling parameter
- ii. G_0 is a non-atomic probability distribution
- iii. η_n follows an Pòlya urn distribution
- iv. N is the total number of drawn η_n

VI adaptation

We consider this factorized family of variational distributions for meanfield variational inference:

$$q(v, \hat{\eta}, z) = \prod_{t=1}^{T-1} q_{\gamma_t}(v_t) \prod_{t=1}^T q_{\tau_t}(\hat{\eta}_t) \prod_{n=1}^N q_{\phi_n}(z_n)$$

Definition of the ELBO, the function to maximize:

$$\begin{aligned} ELBO(V, \hat{\eta}, Z) = & \mathbb{E}_q[\log p(V \mid \alpha)] + \mathbb{E}_q[\log p(\hat{\eta} \mid \lambda)] + \\ & \sum_{n=1}^N (\mathbb{E}_q[\log p(Z_n \mid V)] + \mathbb{E}_q[\log p(x_n \mid Z_n)]) - \mathbb{E}_q[\log q(V, \hat{\eta}, Z)] \end{aligned}$$

Mean-field coordinate ascent algorithm

We can define the updates to be implemented:

$$\gamma_{t,1} = 1 + \sum_{n=1}^N \phi_{n,t}$$

$$\gamma_{t,2} = \alpha + \sum_{n=1}^N \sum_{j=t+1}^T \phi_{n,j} \phi_{n,t}$$

$$\tau_{t,1} = \lambda_1 + \sum_{n=1}^N \phi_{n,t} x_n$$

$$\tau_{t,2} = \lambda_2 + \sum_{n=1}^N \phi_{n,t}$$

$$\phi_{n,t} \propto \exp(S_t)$$

Where $t \in \{1, \dots, T\}$ and $n \in \{1, \dots, N\}$, where

$$S_t = \mathbb{E}_q[\log V_t] + \sum_{i=1}^{t-1} \mathbb{E}_q[\log(1 - V_t)] + \mathbb{E}_q[\hat{\eta}_t]^T X_n - \mathbb{E}_q[a(\hat{\eta})]$$

Iterating these updates optimizes the ELBO with respect to the variational parameters. 17/18

References

- [1] David M. Blei and Michael I. Jordan. **“Variational inference for Dirichlet process mixtures”**. In: *Bayesian Analysis* 1.1 (2006), pp. 121–143. DOI: 10.1214/06-BA104. URL: <https://doi.org/10.1214/06-BA104>.
- [2] David M. Blei, Alp Kucukelbir, and Jon D. McAuliffe. **“Variational Inference: A Review for Statisticians”**. In: *Journal of the American Statistical Association* 112.518 (Apr. 2017), pp. 859–877. DOI: 10.1080/01621459.2017.1285773. URL: <https://doi.org/10.1080%2F01621459.2017.1285773>.
- [3] Haoliang. **“Variational Inference in Bayesian Multivariate Gaussian Mixture Model”**. In: *Bayesian Analysis* (2020). DOI: Hao-2020.