

# **An algorithm for inference of a class of Markovian Regime-Switching Models**

Angus Lewis with Nigel Bean & Giang Nguyen

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# MRS models

- Theory developed by Hamilton:
  - *Rational expectations econometric analysis of changes in regimes: An investigation of the term structure of interest rates*, 1988
- Motivating idea – Different economic regimes determine behaviour of financial markets
- Widely applied since (10,000+ citations):
  - Weather modelling
  - Speech recognition
- Generalises HMMs
- Extended to hidden semi-Markov models/semi-Markovian regime-switching models
- We consider a new type of MRS specification

# MRS models

Typically specified as

$$X_t = \alpha^{(i)} + \phi_1^{(i)} X_{t-1} + \cdots + \phi_p^{(i)} X_{t-p} + \sigma^{(i)} \varepsilon_t \text{ when } R_t = i$$

where

$$\{R_t\}_{t \in \mathbb{N}}$$

is a Markov chain on state space  $\mathcal{S} = \{1, 2, \dots, M\}$ , and

$$\{\varepsilon_t\} \sim \text{i.i.d. } N(0, 1).$$

# Problems & Solutions

Two problems

- Likelihood evaluation
- Likelihood maximisation

Three algorithms

- Forward – likelihood evaluation
- Backward – enables the EM algorithm & quantities of interest
- Expectation Maximisation (EM) – likelihood maximisation

# Evaluating likelihoods (naively)

- The regime process,  $\{R_t\}$ , is hidden
- To evaluate the likelihood.

Let  $\mathbf{x}_{0:t} = (x_0, \dots, x_t)$ ,

$$\begin{aligned} f_{\theta}(\mathbf{x}_{0:T}) &= \sum_{\mathbf{R} \in \mathcal{R}} f_{\theta}(\mathbf{x}_{0:T}, R_0, \dots, R_T) \\ &= \sum_{\mathbf{R} \in \mathcal{R}} f_{\theta}(\mathbf{x}_{0:T} \mid R_0, \dots, R_T) f_{\theta}(R_0, \dots, R_T), \end{aligned}$$

where

- $\mathbf{R} = (R_0, \dots, R_T)$  is a hidden regime sequence
- $T$  is the number of observations
- Naive complexity is  $\mathcal{O}(M^T)$ 
  - $M$  is the number of regimes

# Evaluating likelihoods

## HMMs and dependent regime models

### The forward algorithm<sup>1</sup>

For  $t = 1, \dots, T$

$$f_{\theta}(\mathbf{x}_{0:t}, R_t = j) = \overbrace{f_{\theta}(x_t \mid R_t = j, \mathbf{x}_{0:t-1})}^{\text{e.g. AR(p) densities}} \sum_{i \in \mathcal{S}} p_{ij} \underbrace{f_{\theta}(R_{t-1} = i, \mathbf{x}_{0:t-1})}_{\text{value from last iteration}}$$

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<sup>1</sup>A computationally stable *normalised* algorithm can be derived from this.

# Evaluating likelihoods

## HMMs and dependent regime models

### The forward algorithm<sup>1</sup>

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Then

$$L(\theta) = \sum_{j \in \mathcal{S}} f_{\theta}(\mathbf{x}_{0:T}, R_T = j)$$

- $\mathcal{O}(M^2 T)$  complexity

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<sup>1</sup>A computationally stable *normalised* algorithm can be derived from this.

# Forward algorithm – More than just likelihoods

HMMs and dependent regime models

$$\overbrace{P_{\theta}(R_t = i \mid \mathbf{x}_{0:t})}^{\text{filtered probabilities}} = \frac{f_{\theta}(R_t = i, \mathbf{x}_{0:t})}{\sum_{j \in \mathcal{S}} f_{\theta}(R_t = j, \mathbf{x}_{0:t})}$$

Similarly,

$$\overbrace{P_{\theta}(R_t = i \mid \mathbf{x}_{0:t-1})}^{\text{prediction probabilities}}$$



# Smoothed probabilities

## HMMs and dependent regime models

### Backward algorithm<sup>2</sup>

For  $t = T - 1, \dots, 0$

$$\underbrace{P_{\theta}(R_t = i \mid \mathbf{x}_{0:T})}_{\text{smoothed probabilities}} = \underbrace{P_{\theta}(R_t = i \mid \mathbf{x}_{0:t})}_{\text{filtered probabilities}} \sum_{j \in \mathcal{S}} p_{ij} \underbrace{\frac{P_{\theta}(R_{t+1} = j \mid \mathbf{x}_{0:T})}{P_{\theta}(R_{t+1} = j \mid \mathbf{x}_{0:t})}}_{\text{prediction probabilities}}$$

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<sup>2</sup>Kim, 1994, *Dynamic linear models with Markov-switching*

# Maximising likelihoods

## HMMs and dependent regime models

### EM algorithm

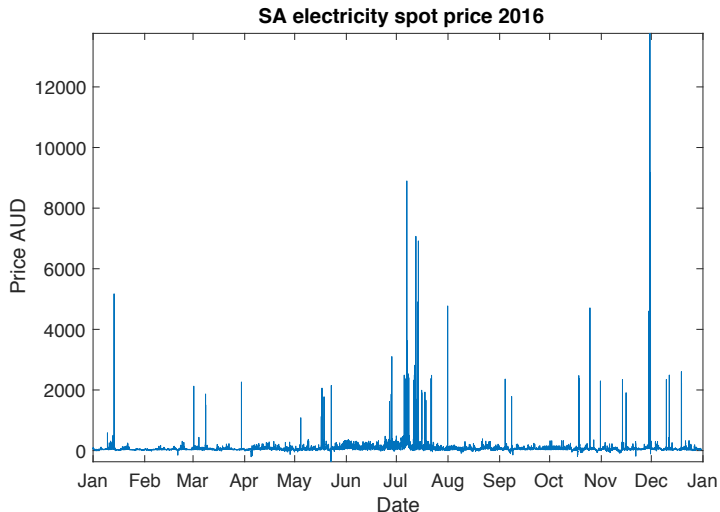
#### E. Construct

$$\begin{aligned} Q(\theta, \theta_n) &= \mathbb{E}[\log f_{\theta}(\mathbf{x}_{0:T}, \mathbf{R}) \mid \mathbf{x}_{0:T}, \theta_n] \\ &= \sum_{t=1}^T \sum_{i \in \mathcal{S}} \overbrace{P_{\theta_n}(R_t = i \mid \mathbf{x}_{0:T})}^{\text{smoothed probabilities}} \log f_{\theta}(x_t \mid R_t = i, \mathbf{x}_{0:t-1}) \\ &\quad + \sum_{i,j \in \mathcal{S}} \sum_{t=1}^T P_{\theta_n}(R_t = j, R_{t-1} = i \mid \mathbf{x}_{0:T}) \log(p_{ij}) \end{aligned}$$

**M.** Set  $\theta_{n+1} = \arg \max_{\theta} Q(\theta, \theta_n)$

Works well when M-step is analytic.

# Our motivation – Electricity price modelling



# Independent regime MRS models

Consider the MRS model

$$X_t = \begin{cases} B_t^1, & R_t = 1 \\ \vdots & \\ B_t^k, & R_t = k \\ S_t^{k+1} & R_t = k + 1 \\ \vdots & \\ S_t^M & R_t = M \end{cases}$$

where

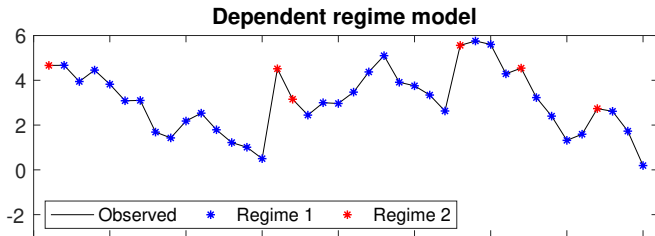
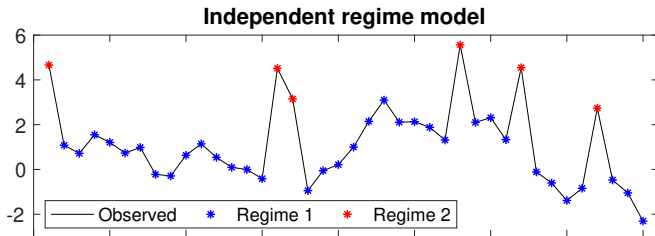
$$B_t^i = \alpha_i + \phi_i B_{t-1}^i + \sigma_i \varepsilon_t^i$$

are AR(1) and

$$S_t^j \sim \text{i.i.d. } N(\mu_j, \sigma_j^2).$$

# MRS models of electricity prices

- *Independent regime* MRS models allow 'jumpy' behaviour



# Independent vs dependent regimes

## Independent regime model

- Multiple time series evolving in parallel
- Which process is observed is determined by the regime

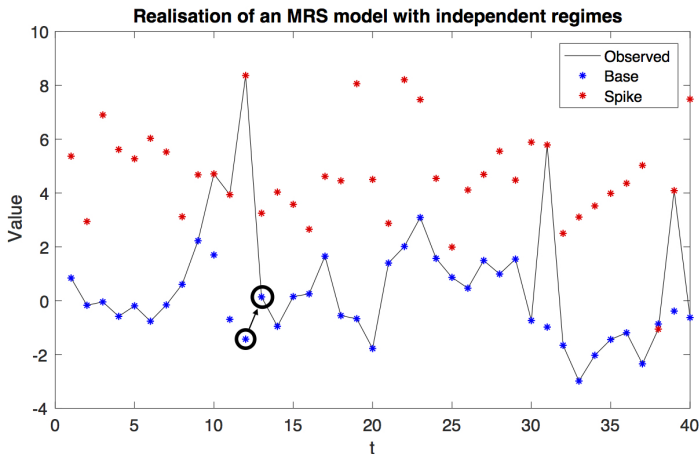
$$X_t = \begin{cases} B_t = \alpha + \phi B_{t-1} + \sigma \varepsilon_t, & R_t = 1 \\ \text{i.i.d. } N(\mu, \sigma_2^2), & R_t = 2 \end{cases}$$

## Dependent regime model

- A single time series
- Dynamics (parameters) determined by the regime

$$X_t = \begin{cases} \alpha + \phi X_{t-1} + \sigma \varepsilon_t, & R_t = 1 \\ \text{i.i.d. } N(\mu, \sigma_2^2), & R_t = 2 \end{cases}$$

# Independent regimes – Example



- $B_t$  depends on  $B_{t-1}$ , but it might be unobserved

# Evaluating likelihoods – Independent regime models

This idea motivates an approximate algorithm, the EM-like algorithm

## The forward algorithm – Hamilton

For  $t = 1, \dots, T$

$$f_{\theta}(\mathbf{x}_{0:t}, R_t = j) = \overbrace{f_{\theta}(\mathbf{x}_t \mid R_t = j, \mathbf{x}_{0:t-1})}^{\text{naive } \mathcal{O}(M^t) \text{ calculation}} \sum_{i \in \mathcal{S}} p_{ij} f_{\theta}(R_{t-1} = i, \mathbf{x}_{0:t-1})$$

Then

$$L(\theta) = \sum_{j \in \mathcal{S}} f_{\theta}(\mathbf{x}_{0:T}, R_T = j)$$



# Maximising likelihoods – Independent regime models

## An approximate algorithm – EM-like

- At time  $t$  suppose  $R_t$  and  $B_{t-1}^{R_t}$  are known
- Then density of  $X_t$  would be known

$$x_t \mid R_t = i, B_{t-1}^i \sim N(\alpha_i + \phi_i B_{t-1}^i, \sigma_i^2).$$

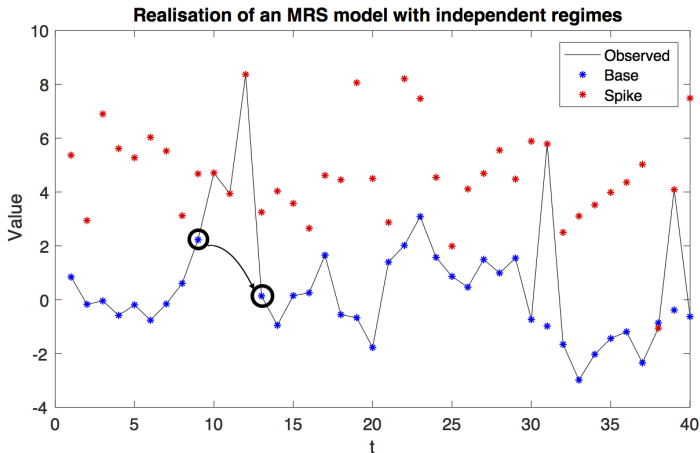
- The EM-like algorithm replaces  $B_{t-1}^i$  with approximate values

$$\begin{aligned} \tilde{B}_{t-1}^i &= P_\theta(R_{t-1} = i \mid \mathbf{x}_{0:t-1})x_t \\ &\quad + P_\theta(R_{t-1} \neq i \mid \mathbf{x}_{0:t-1})(\alpha_i + \phi_i \tilde{B}_{t-2}^i) \end{aligned}$$

calculated iteratively.

Can be shown to perform poorly.

# Independent regimes – Example



- Can integrate unobserved prices away
- $B_t$  depends on a random lagged observation

## Evaluating likelihoods – Independent regime models

Alternatively, to write down the density of  $x_t$  we need to know

- $R_t$ ,
- Time since the last visit to  $R_t$ .

# Evaluating likelihoods – Independent regime models

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- $R_t$ ,
- Time since the last visit to  $R_t$ .

Idea: augment  $\{R_t\}$  with time-since-last-visit counters

- Let  $N_{t,i} = \ell$  denote the event that the last visit to state  $i$  before time  $t$  was  $\ell$  transitions ago.
- Define an augmented process

$$\{\mathbf{H}_t\} = \{(\mathbf{N}_t, R_t)\} = \{(N_{t,1}, \dots, N_{t,k}, R_t)\}$$

- Similar ideas for hidden semi-Markov models
  - Counters for time since last transition

## Facts about $\mathbf{H}_t = (\mathbf{N}_t, R_t)$

- $\{\mathbf{H}_t\}$  is a Markov chain.
- Transition probabilities

$$\begin{aligned} &P_{\theta}(\mathbf{H}_{t+1} = (\mathbf{N}_{t+1}, j) \mid \mathbf{H}_t = (\mathbf{N}_t, i)) \\ &= \begin{cases} p_{ij} & \text{for } i \in \{k+1, \dots, M\}, j \in \mathcal{S}, \mathbf{N}_{t+1} = \mathbf{N}_t + \mathbf{1}, \\ p_{ij} & \text{for } i \in \{1, \dots, k\}, j \in \mathcal{S}, \mathbf{N}_{t+1} = \mathbf{N}_t^{(-i)} + \mathbf{1}, \\ 0 & \text{otherwise,} \end{cases} \end{aligned}$$

where  $\mathbf{N}_t^{(-i)} = (N_{t,1}, \dots, N_{t,i-1}, 0, N_{t,i+1}, \dots, N_{t,k})$ .

# Evaluating likelihoods

## Independent regime models

### A new forward algorithm<sup>3</sup>

For  $t = 1, \dots, T$

$$\begin{aligned} f_{\theta}(\mathbf{x}_{0:t}, \mathbf{H}_t = (\mathbf{N}_t, j)) \\ = f_{\theta}(x_t \mid \mathbf{H}_t = (\mathbf{N}_t, j), \mathbf{x}_{0:t-1}) f_{\theta}(\mathbf{H}_t = (\mathbf{N}_t, j), \mathbf{x}_{0:t-1}) \end{aligned}$$

where

$$\begin{aligned} f_{\theta}(\mathbf{H}_t = (\mathbf{N}_t, j), \mathbf{x}_{0:t-1}) \\ = \begin{cases} \sum_{i \in \mathcal{S}} p_{ij} f_{\theta}(\mathbf{H}_{t-1} = (\mathbf{N}_t - \mathbf{1}, i), \mathbf{x}_{0:t-1}), & N_{t,r} > 1 \forall r \\ \sum_{m=1}^t p_{ij} f_{\theta}(\mathbf{H}_{t-1} = (\mathbf{N}_t - \mathbf{1} + m\mathbf{e}_r, r), \mathbf{x}_{0:t-1}), & N_{t,r} = 1 \end{cases} \end{aligned}$$

<sup>3</sup>A computationally stable normalised algorithm can also be derived

# Evaluating likelihoods

Independent regime models

A new forward algorithm – continued

$$L(\theta) = \sum_{j \in \mathcal{S}} \sum_{\mathbf{N}_T} f_{\theta}(\mathbf{x}_{0:T}, \mathbf{H}_T = (\mathbf{N}_T, j))$$

## Forward algorithm – More than just likelihoods

$$\overbrace{P_{\theta}(\mathbf{H}_t = (\mathbf{N}_t, i) \mid \mathbf{x}_{0:t})}^{\text{filtered probabilities}} = \frac{f_{\theta}(\mathbf{H}_t = (\mathbf{N}_t, i), \mathbf{x}_{0:t})}{\sum_{j \in S} \sum_{\mathbf{N}_t} f_{\theta}(\mathbf{H}_t = (\mathbf{N}_t, j), \mathbf{x}_{0:t})}$$

Similarly,

$$\overbrace{P_{\theta}(\mathbf{H}_{t+1} = (\mathbf{N}_{t+1}, i) \mid \mathbf{x}_{0:t})}^{\text{prediction probabilities}}$$



# Smoothed probabilities

## Independent regime models

### Backward algorithm

For  $t = T - 1, \dots, 0$

$$\begin{aligned} & \overbrace{P_{\theta}(\mathbf{H}_t = (\mathbf{N}_t, i) \mid \mathbf{x}_{0:T})}^{\text{smoothed probabilities}} \\ &= \overbrace{P_{\theta}(\mathbf{H}_t = (\mathbf{N}_t, i) \mid \mathbf{x}_{0:t})}^{\text{filtered probabilities}} \sum_{j \in \mathcal{S}} p_{ij} \underbrace{\frac{P_{\theta}(\mathbf{H}_{t+1} = (\mathbf{N}_{t+1}, j) \mid \mathbf{x}_{0:T})}{P_{\theta}(\mathbf{H}_{t+1} = (\mathbf{N}_{t+1}, j) \mid \mathbf{x}_{0:t})}}_{\text{prediction probabilities}} \end{aligned}$$

where  $\mathbf{N}_{t+1} = (N_{t,1} + 1, \dots, N_{t,i-1} + 1, 1, N_{t,i+1} + 1, \dots, N_{t,k} + 1)$   
for  $i \in \{1, 2, \dots, k\}$  and  $\mathbf{N}_{t+1} = \mathbf{N}_t + \mathbf{1}$  for  $i \in \{k + 1, \dots, M\}$ .

# Maximising likelihoods

## HMMs and dependent regime models

### EM algorithm

Iterate

**E.** Construct

$$\begin{aligned} Q(\theta, \theta_n) &= \sum_{t=1}^T \sum_{i \in \mathcal{S}} \sum_{\mathbf{N}_t} \overbrace{P_{\theta_n}(\mathbf{H}_t = (\mathbf{N}_t, i) \mid \mathbf{x}_{0:T})}^{\text{smoothed probabilities}} \log f_{\theta}(\mathbf{x}_t \mid R_t = i, \mathbf{N}_t, \mathbf{x}_{0:t-1}) \\ &\quad + \sum_{i,j \in \mathcal{S}} \log(p_{ij}) \sum_{t=1}^T P_{\theta_n}(R_t = j, R_{t-1} = i \mid \mathbf{x}_{0:T}) \end{aligned}$$

**M.** Set  $\theta_{n+1} = \arg \max_{\theta} Q(\theta, \theta_n)$

# Complexity

- Complexity is  $\mathcal{O}(M^2 T^{k+1} k^k)$
- Compare to naive complexity  $\mathcal{O}(M^T)$ 
  - $M$  – the number of regimes
  - $T$  – the number of observations
  - $k$  – the number of AR(1) processes/counters
- Number of accessible states of  $\{\mathbf{N}_t\}$  by time  $t$  is

$$\sum_{m=0}^{\min(t,k)} \binom{t}{m} \binom{k}{m} m! \approx \mathcal{O}(t^k k^k)$$

given

$$\mathbf{H}_0 = \begin{cases} (\mathbf{1}, i) & \text{with probability } \pi_i \\ (\mathbf{N}_0, j) & \text{with probability 0 otherwise,} \end{cases}$$

for some distribution  $\boldsymbol{\pi} = (\pi_1, \dots, \pi_M)$ .

# Truncation

- $\mathcal{O}(M^2 T^{k+1} k^k)$  may be large
- Truncate

$$N_{t,i} \in \{1, 2, \dots, D-1, D\}, \text{ where } D \ll T$$

and allow self-transitions  $D \rightarrow D$

- Complexity is now  $\mathcal{O}(M^2 D^k T k^k)$
- Compare to the complexity of the naive approach with the same assumptions:  $\mathcal{O}(M^D T)$
- We set

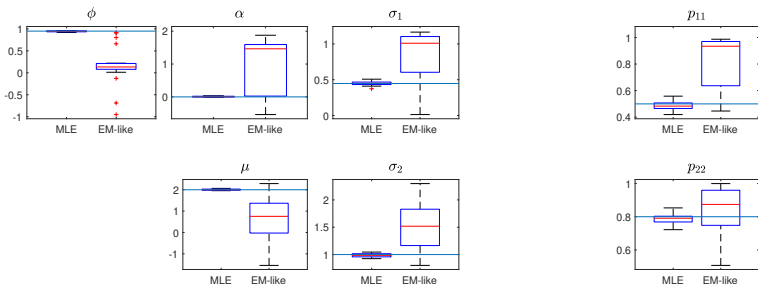
$$f(x_t \mid R_t = i, N_{t,i} = D, \mathbf{x}_{0:t-1}) = f(x_t \mid R_t = i, N_{t,i} = \infty)$$

i.e. the stationary density for Regime  $i$

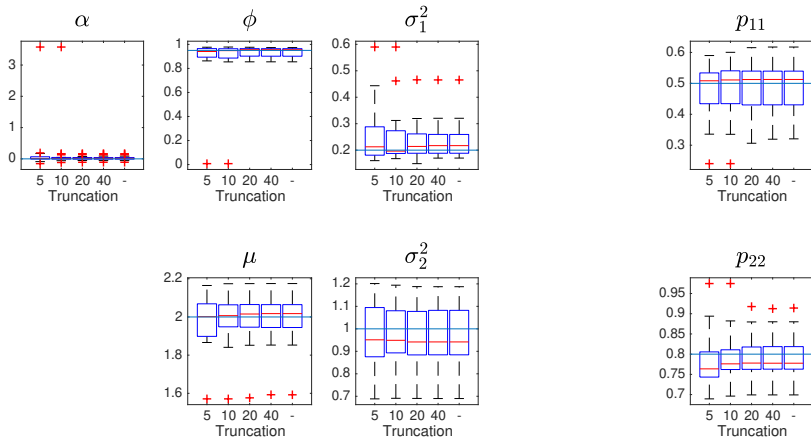
# Comparison with existing method

$$X_t = \begin{cases} B_t, & \text{if } R_t = 1, \\ S_t, & \text{if } R_t = 2, \end{cases}$$

- $B_t = 0.95B_{t-1} + 0.2\varepsilon_t$ ,
- $S_t \sim N(2, 1)$
- $p_{11} = 0.5, p_{22} = 0.8$



# An example – Truncation



# Summary

- Developed novel algorithms for independent regime MRS models
  - Forward, backward, EM
  - Efficient approximations to all of the above
  - Important concept was the augmented hidden chain  $\{\mathbf{H}_t\}$
- We believe these ideas can be extended to more general processes
  - more general AR(p) processes
  - other Markovian processes?
- *Estimation of Markovian-Regime-Switching models with independent regimes*, Lewis, Nguyen, Bean, To be submitted

# An example – Consistency

