## Problem class 1<sup>a</sup>

Nuisance parameters, the Normal model, and the Normal linear regression with unknown variance

Lecturer: Georgios Karagiannis

georgios.karagiannis@durham.ac.uk

<sup>a</sup>Author: Georgios P. Karagiannis.

## **Nuisance parameters**

<concept

**Definition 1.** Assume observable quantities  $y=(y_1,...,y_n)$ . Assume that the sampling distribution is  $dF(y|\theta)$  labeled by an unknown parameter  $\theta \in \Theta$ . Let  $\theta=(\phi,\lambda)^{\top}$  with  $\phi \in \Phi$  and  $\lambda \in \Lambda$ . Assume You are interested in learning parameter  $\phi \in \Phi$ , and You are not interested in learning the unknown parameter  $\lambda \in \Lambda$ ; but both  $\phi,\lambda$  are parts of the statistical model parameterisation. The unknown quantity  $\lambda \in \Lambda$  is called <u>nuisance parameter</u>. We an call  $\phi \in \Phi$  parameter of interest.

Note 2. In Bayesian Stats, learning (or quantifying uncertainty about) parameter of interest  $\phi$  under the presence of a nuisance parameter  $\lambda \in \Lambda$  is performed according to the Bayesian paradigm as usual: You specify a prior  $d\Pi(\phi,\lambda)$  with PDF/PMF  $\pi(\phi,\lambda) = \pi(\phi|\lambda)\pi(\lambda)$  on the joint space of ALL Your unknown parameters  $\theta = (\phi,\lambda)^{\top}$ ; you compute the joint posterior distribution  $d\Pi(\theta|y)$  of  $\theta = (\phi,\lambda)^{\top}$  via the Bayesian theorem. Reasonably, Your posterior degree of believe about the parameter of interest  $\phi$  given the data  $y = (1_1,...,y_n)$  is given through the marginal posterior distribution  $d\Pi(\phi|y)$ .

Note 3. To summarize; Specify the Bayesian model as:

<sum-up

$$\begin{cases} y | \overbrace{\phi, \lambda}^{=\theta} \sim \mathrm{d}F(y| \overbrace{\phi, \lambda}^{=\theta}) &, \text{ the statistical model} \\ (\underbrace{\phi, \lambda}_{=\theta}) \sim \mathrm{d}\Pi(\underbrace{\phi, \lambda}_{=\theta}) &, \text{ the prior model} \end{cases}$$

The joint posterior of  $\theta$  given y is  $d\Pi(\theta|y) = d\Pi(\lambda|y,\phi)d\Pi(\phi|y)$  is with PDF/PMF

$$\pi(\overbrace{\phi,\lambda}|y) = \frac{f(y|\overbrace{\phi,\lambda})\pi(\overbrace{\phi,\lambda})}{f(y)} = \underbrace{\frac{f(y|\phi,\lambda)\pi(\lambda|\phi)}{f(y|\phi)}}_{=\pi(\lambda|y,\phi)} \underbrace{\frac{f(y|\phi)\pi(\phi)}{f(y)}}_{=\pi(\phi|y)} = \pi(\lambda|y,\phi)\pi(\phi|y)$$

The (marginal) likelihood  $f(y|\phi)$  of y given  $\phi$  is

$$f(y|\phi) = \underbrace{\int_{\Lambda} f(y|\phi,\lambda) \mathrm{d}\Pi(\lambda|\phi)}_{= \mathbb{E}_{\Pi(\lambda|\phi)}(f(y|\phi,\lambda)|\phi)} = \begin{cases} \int_{\Lambda} f(y|\phi,\lambda) \pi(\lambda|\phi) \mathrm{d}\lambda & \text{, if } \lambda \text{ cont} \\ \\ \sum_{\forall \lambda \in \Lambda} f(y|\phi,\lambda) \pi(\lambda|\phi) & \text{, if } \lambda \text{ discr} \end{cases}$$

The PDF/PMF  $\pi(\phi|y)$  of marginal posterior  $d\Pi(\phi|y)$  of  $\phi$  is

$$\pi(\phi|y) = \underbrace{\int_{\Lambda} \pi(\phi, \lambda|y) \mathrm{d}\lambda}_{=\mathrm{E}_{\Pi(\lambda|y)}(\pi(\phi|y, \lambda))} \qquad \text{or equivalently} \qquad \pi(\phi|y) = \frac{f(y|\phi)\pi(\phi)}{f(y)}$$

The predictive distribution dG(z|y) of the next outcome  $z=(y_{n+1},...y_{n+m})$  given y has pdf/pmf

$$g(z|y) = \int f(y|\overbrace{\phi,\lambda}) \mathrm{d}\Pi(\overbrace{\phi,\lambda}|y)$$

and the marginal likelihood f(y) is

$$f(y) = \int f(y|\overbrace{\phi,\lambda}) \pi(\overbrace{\phi,\lambda}) \mathrm{d}\phi \mathrm{d}\lambda$$

## Practice in challenging problems

**Exercise 4.**  $(\star\star)$ (Nuisance parameters are involved)

<-story

Assume observable quantities  $y=(y_1,...,y_n)$  forming the available data set of size n. Assume that the observations are drawn i.i.d. from a sampling distribution which is judged to be in the Normal parametric family of distributions  $N(\mu, \sigma^2)$  with unknown mean  $\mu$  and variance  $\sigma^2$ . We are interested in learning  $\mu$  and the next outcome  $z=y_{n+1}$ . We do not care about  $\sigma^2$ .

Assume You specify a Bayesian model

<-set-up

$$\begin{cases} y_i|\mu,\sigma^2 \sim \mathrm{N}(\mu,\sigma^2), \text{ for all } i=1,...,n \\ \mu|\sigma^2 \sim \mathrm{N}(\mu_0,\sigma^2\frac{1}{\tau_0}) \\ \sigma^2 \sim \mathrm{IG}(a_0,k_0) \end{cases}, \text{ prior}$$

1. Show that

$$\sum_{i=1}^{n} (y_i - \theta)^2 = n(\bar{y} - \theta)^2 + ns^2,$$

where  $s^2 = \frac{1}{2} \sum_{i=1}^n (y_i - \bar{y})^2$ .

2. Show that the joint posterior distribution  $\Pi(\mu, \sigma^2|y)$  is such as

$$\mu|y, \sigma^2 \sim N(\mu_n, \sigma^2 \frac{1}{\tau_n})$$
  
 $\sigma^2|y \sim IG(a_n, k_n)$ 

with

$$\mu_n = \frac{n\bar{y} + \tau_0 \mu_0}{n + \tau_0};$$
  $\tau_n = n + \tau_0;$   $a_n = a_0 + n$ 

$$k_n = k_0 + \frac{1}{2}ns_n^2 + \frac{1}{2}\frac{\tau_0 n(\mu_0 - \bar{y})^2}{n + \tau_0}$$

**Hint:** It is

$$-\frac{1}{2}\frac{(\mu-\mu_1)^2}{v_1} - \frac{1}{2}\frac{(\mu-\mu_2)^2}{v_2}\dots - \frac{1}{2}\frac{(\mu-\mu_n)^2}{v_n} = -\frac{1}{2}\frac{(\mu-\hat{\mu})^2}{\hat{v}} + C$$

where

$$\hat{v} = \left(\sum_{i=1}^{n} \frac{1}{v_i}\right)^{-1}; \quad \hat{\mu} = \hat{v}\left(\sum_{i=1}^{n} \frac{\mu_i}{v_i}\right); \quad C = \frac{1}{2} \frac{\hat{\mu}^2}{\hat{v}} - \frac{1}{2} \sum_{i=1}^{n} \frac{\mu_i^2}{v_i}$$

3. Show that the marginal posterior distribution  $\Pi(\mu|y)$  is such as

$$\mu|y \sim \mathrm{T}_1\left(\mu_n, \frac{k_n}{a_n} \frac{1}{\tau_n}, 2a_n\right)$$

**Hint-1:** If  $x \sim IG(a, b)$ , y = cx, then  $y \sim IG(a, cb)$ .

Hint-2: The definition of Student T is considered as known

4. Show that the predictive distribution  $\Pi(z|y)$  is Student T such as

$$z|y \sim T_1\left(\mu_n, \frac{k_n}{a_n}(\frac{1}{\tau_n} + 1), 2a_n\right)$$

Hint-1: Consider that

$$N(x|\mu_1, \sigma_1^2) N(x|\mu_2, \sigma_2^2) = N(x|m, v^2) N(\mu_1|\mu_2, \sigma_1^2 + \sigma_2^2)$$

where

$$v^2 = \left(\frac{1}{\sigma_1^2} + \frac{1}{\sigma_2^2}\right)^{-1}; \quad m = v^2 \left(\frac{\mu_1}{\sigma_1^2} + \frac{\mu_2}{\sigma_2^2}\right)$$

Hint-2: The definition of Student T is considered as known

## **General practice**

**Exercise 5.**  $(\star\star)$ Consider the Bayesian model

$$\begin{cases} x_i | \theta & \stackrel{\text{iid}}{\sim} \operatorname{Ex}(\theta), \ \forall i = 1, ..., n \\ \theta & \sim \operatorname{Ga}(a, b) \end{cases}$$

**Hint-1:** The PDF of  $x \sim \mathrm{G}(a,b)$  is  $\mathrm{Ga}(x|a,b) = \frac{b^a}{\Gamma(a)} x^{a-1} \exp(-bx) 1_{(0,+\infty)}(x)$ 

**Hint-2:** The PDF of  $x \sim \text{Ex}(\theta)$  is  $\text{Ex}(x|\theta) = \text{Ga}(x|1,\theta)$ 

- 1. Show that the parametric model is member of the Exponential family, and the sufficient statistic for a sample of observables  $x = (x_1, ..., x_n)$ .
- 2. Show that the posterior distribution  $\theta$  given x is Gamma and compute its parameters.
- 3. Show that the predictive distribution G(z|x) of a future z given  $x = (x_1, ..., x_n)$ , has PDF

$$g(z|x) = \frac{a^*(b^*)^{a^*}}{(z+b^*)^{a^*+1}} \mathbf{1}(x \ge 0)$$

Exercise 6. (\*\*)Consider the Bayesian model

$$\begin{cases} x_i | \theta & \stackrel{\text{IID}}{\sim} \text{Mu}_k(\theta) \\ \theta & \sim \text{Di}_k(a) \end{cases}$$

where  $\theta \in \Theta$ , with  $\Theta = \{\theta \in (0,1)^k | \sum_{j=1}^k \theta_j = 1\}$  and  $\mathcal{X}_k = \{x \in \{0,...,n\}^k | \sum_{j=1}^k x_j = 1\}.$ 

**Hint-1:** Mu<sub>k</sub> denotes the Multinomial probability distribution with PMF

$$\mathbf{Mu}_{k}(x|\theta) = \begin{cases} \prod_{j=1}^{k} \theta_{j}^{x_{j}} & \text{, if } x \in \mathcal{X}_{k} \\ 0 & \text{, otherwise} \end{cases}$$
 (1)

**Hint-2:**  $Di_k(a)$  denotes the Dirichlet distribution with PDF

$$\mathrm{Di}_k(\theta|a) = \begin{cases} \frac{\Gamma(\sum_{j=1}^k a_j)}{\prod_{j=1}^k \Gamma(a_j)} \prod_{j=1}^k \theta_j^{a_j-1} & \text{, if } \theta \in \Theta \\ 0 & \text{, otherwise} \end{cases}$$

- 1. Show that the parametric model (1) is a member of the k-1 exponential family.
- 2. Compute the likelihood  $f(x_{1:n}|\theta)$ , and find the sufficient statistic  $t_n := t_n(x_{1:n})$ .
- 3. Compute the posterior distribution. State the name of the distribution, and expresses its parameters with respect to the observations and the hyper-parameters of the prior. Justify your answer.
- 4. Compute the probability mass function of the predictive distribution for a future observation  $y = x_{n+1}$  in closed form.

**Hint** 
$$\Gamma(x) = (x-1)\Gamma(x-1)$$
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