

Regression using Matrices

We will use the following toy data set to illustrate how regression is carried out via matrix algebra.

```
> myData = data.frame(  
  wage = c(12, 8, 16.26, 13.65, 8.5),  
  age = c(32, 33, 32, 33, 26),  
  sex = c("M", "F", "M", "M", "M")  
)
```

```
> myData
```

	wage	age	sex
1	12.00	32	M
2	8.00	33	F
3	16.26	32	M
4	13.65	33	M
5	8.50	26	M

We can now present linear regression in matrix terms. Begin with the linear regression model

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i \quad \text{where } i = 1, \dots, n$$

This implies

$$Y_1 = 1\beta_0 + \beta_1 X_1 + \epsilon_1$$

$$Y_2 = 1\beta_0 + \beta_1 X_2 + \epsilon_2$$

$$\vdots$$

$$Y_n = 1\beta_0 + \beta_1 X_n + \epsilon_n$$

We can now arrange these into appropriate vectors and matrices.

$$\mathbf{Y}_{n \times 1} = \begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix}$$

$$\mathbf{X}_{n \times 2} = \begin{bmatrix} 1 & X_1 \\ 1 & X_2 \\ \vdots & \vdots \\ 1 & X_n \end{bmatrix}$$

$$\boldsymbol{\beta}_{2 \times 1} = \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix}$$

$$\boldsymbol{\epsilon}_{n \times 1} = \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{bmatrix}$$

We can write the model using matrix terms very compactly as

$$\mathbf{Y}_{n \times 1} = \mathbf{X}_{n \times 2} \boldsymbol{\beta}_{2 \times 1} + \boldsymbol{\epsilon}_{n \times 1}$$

Dimension
of product

Next we turn to the assumptions of the errors. Since $\mathbf{E}\{\epsilon_i\} = 0$,

$$\underset{n \times 1}{\mathbf{E}\{\epsilon\}} = \underset{n \times 1}{\mathbf{0}}$$

We also make the assumption that the errors have constant variance σ^2 and that they are independent of one another $\sigma(\epsilon_i, \epsilon_j) = 0$ for $i \neq j$. The **variance-covariance matrix of the error terms** can therefore be expressed as

$$\underset{n \times n}{\sigma^2\{\epsilon\}} = \begin{bmatrix} \sigma_\epsilon^2 & 0 & \cdots & 0 \\ 0 & \sigma_\epsilon^2 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \sigma_\epsilon^2 \end{bmatrix}$$

This scalar matrix can also be expressed as

$$\underset{n \times n}{\sigma^2\{\epsilon\}} = \underset{n \times n}{\sigma^2 \mathbf{I}}$$

Thus, the **regression model** can be completely expressed in matrix terms as

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$$

where $\boldsymbol{\epsilon}$ is a vector of independent random variables with $\mathbf{E}\{\boldsymbol{\epsilon}\}=0$ and $\boldsymbol{\sigma}^2\{\boldsymbol{\epsilon}\}= \sigma^2\mathbf{I}$

Solving for β

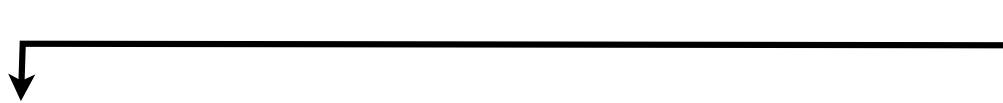
We can re-express the **error vector** as

$$\epsilon = Y - X\beta$$

The *method of least squares* minimizes the **sum of squared errors** which is expressed as

$$\epsilon' \epsilon = (Y - X\beta)'(Y - X\beta)$$

$$= \underset{1 \times n}{Y'} \underset{n \times 1}{Y} - \underset{1 \times 2}{\beta'} \underset{2 \times n}{X'} \underset{n \times 1}{Y} - \underset{1 \times n}{Y'} \underset{n \times 2}{X} \underset{2 \times 1}{\beta} + \underset{1 \times 2}{\beta'} \underset{2 \times n}{X'} \underset{n \times 2}{X} \underset{2 \times 1}{\beta}$$



Note that each term is a 1x1 matrix, which means that each term is equal to its transpose. We will re-write the third term $Y'X\beta$ as its transpose $\beta'X'Y$

$$= \underset{1 \times n}{Y'} \underset{n \times 1}{Y} - \underset{1 \times 2}{\beta'} \underset{2 \times n}{X'} \underset{n \times 1}{Y} - \underset{1 \times 2}{\beta'} \underset{2 \times n}{X'} \underset{n \times 1}{Y} + \underset{1 \times 2}{\beta'} \underset{2 \times n}{X'} \underset{n \times 2}{X} \underset{2 \times 1}{\beta}$$

Combining the two middle terms we get

$$= \underset{1 \times n}{Y'} \underset{n \times 1}{Y} - 2 \underset{1 \times 2}{\beta'} \underset{2 \times n}{X'} \underset{n \times 1}{Y} + \underset{1 \times 2}{\beta'} \underset{2 \times n}{X'} \underset{n \times 2}{X} \underset{2 \times 1}{\beta}$$

To find the values for the elements in β that minimize the equation, we differentiate with respect to β_0 and β_1 (β)

$$\frac{\partial}{\partial \beta} (\underbrace{\mathbf{Y}' \mathbf{Y}}_{1 \times n \ n \times 1} - 2 \underbrace{\beta' \mathbf{X}' \mathbf{Y}}_{1 \times 2 \ 2 \times n \ n \times 1} + \underbrace{\beta' \mathbf{X}' \mathbf{X} \beta}_{1 \times 2 \ 2 \times n \ n \times 2 \ 2 \times 1})$$

Differentiating, we get...

$$= -2 \underbrace{\mathbf{X}' \mathbf{Y}}_{2 \times n \ n \times 1} + 2 \underbrace{\mathbf{X}' \mathbf{X}}_{2 \times n \ n \times 2} \underbrace{\beta}_{2 \times 1}$$

Solving this equation for zero...

$$\begin{aligned} \underbrace{\mathbf{0}}_{2 \times 1} &= -2 \underbrace{\mathbf{X}' \mathbf{Y}}_{2 \times n \ n \times 1} + 2 \underbrace{\mathbf{X}' \mathbf{X}}_{2 \times n \ n \times 2} \underbrace{\beta}_{2 \times 1} \\ &= -\underbrace{\mathbf{X}' \mathbf{Y}}_{2 \times n \ n \times 1} + \underbrace{\mathbf{X}' \mathbf{X}}_{2 \times n \ n \times 2} \underbrace{\beta}_{2 \times 1} \end{aligned}$$

Adding $\mathbf{X}'\mathbf{Y}$ to both sides

$$\underbrace{\mathbf{X}' \mathbf{Y}}_{2 \times n \ n \times 1} = \underbrace{\mathbf{X}' \mathbf{X}}_{2 \times n \ n \times 2} \underbrace{\beta}_{2 \times 1}$$

Now we pre-multiply both sides of the equation by $(\mathbf{X}'\mathbf{X})^{-1}$

$$\underbrace{(\mathbf{X}'\mathbf{X})^{-1}}_{2 \times 2} \underbrace{\mathbf{X}' \mathbf{Y}}_{2 \times n \ n \times 1} = \underbrace{(\mathbf{X}'\mathbf{X})^{-1}}_{2 \times 2} \underbrace{\mathbf{X}' \mathbf{X}}_{2 \times n \ n \times 2} \underbrace{\beta}_{2 \times 1}$$

$$\underbrace{(\mathbf{X}'\mathbf{X})^{-1}}_{2 \times 2} \underbrace{\mathbf{X}' \mathbf{Y}}_{2 \times n \ n \times 1} = \underbrace{\mathbf{I}}_{2 \times 2} \underbrace{\beta}_{2 \times 1}$$

This means

$$\underbrace{\beta}_{2 \times 1} = \underbrace{(\mathbf{X}'\mathbf{X})^{-1}}_{2 \times 2} \underbrace{\mathbf{X}' \mathbf{Y}}_{2 \times n \ n \times 1}$$

The vector of regression coefficients can be obtained directly through manipulation of the design matrix and the vector of outcomes.

$$\underset{2 \times 1}{\beta} = (\underset{2 \times 2}{\mathbf{X}'\mathbf{X}})^{-1} \underset{2 \times n}{\mathbf{X}'} \underset{n \times 1}{\mathbf{Y}}$$

```
> Y = myData$wage
> X = matrix(c(rep(1, 5), myData$age), ncol = 2)

> b = solve(t(X) %*% X) %*% t(X) %*% Y
> b
```

[,1]

(Intercept) -3.8982759

age 0.4993678

When we estimate the β matrix from sample data we refer to it as **b**.

$$\mathbf{b} = \begin{bmatrix} -3.8983 \\ 0.4994 \end{bmatrix}$$

Fitted Values

Let the vector of **fitted values** be

$$\underset{n \times 1}{\hat{\mathbf{Y}}} = \begin{bmatrix} \hat{Y}_1 \\ \hat{Y}_2 \\ \vdots \\ \hat{Y}_n \end{bmatrix}$$

In matrix notation

$$\underset{n \times 1}{\hat{\mathbf{Y}}} = \underset{n \times 2}{\mathbf{X}} \underset{2 \times 1}{\mathbf{b}}$$

$$\hat{\mathbf{Y}}_{n \times 1} = \mathbf{X}_{n \times 2} \mathbf{b}_{2 \times 1}$$

```
> fitted_values = X %*% b  
> fitted_values
```

```
      [,1]  
1 12.081494  
2 12.580862  
3 12.081494  
4 12.580862  
5  9.085287
```

H Matrix

$$\underset{n \times 1}{\hat{\mathbf{Y}}} = \underset{n \times 2}{\mathbf{X}} \underset{2 \times 1}{\mathbf{b}}$$

Substituting the expression for \mathbf{b}

$$\underset{n \times 1}{\hat{\mathbf{Y}}} = \underset{n \times 2}{\mathbf{X}} \underset{2 \times 2}{(\mathbf{X}'\mathbf{X})^{-1}} \underset{2 \times n}{\mathbf{X}'} \underset{n \times 1}{\mathbf{Y}}$$

This is often expressed as

$$\underset{n \times 1}{\hat{\mathbf{Y}}} = \underset{n \times n}{\mathbf{H}} \underset{n \times 1}{\mathbf{Y}}$$

where

$$\underset{n \times n}{\mathbf{H}} = \underset{n \times 2}{\mathbf{X}} \underset{2 \times 2}{(\mathbf{X}'\mathbf{X})^{-1}} \underset{2 \times n}{\mathbf{X}'}$$

\mathbf{H} is a square $n \times n$ matrix that can be created completely from the design matrix and its transpose.

```
> h = X %*% solve(t(X) %*% X) %*% t(X)
```

```
> h
```

	1	2	3	4	5
1	0.21839080	0.24137931	0.21839080	0.24137931	0.08045977
2	0.24137931	0.29310345	0.24137931	0.29310345	-0.06896552
3	0.21839080	0.24137931	0.21839080	0.24137931	0.08045977
4	0.24137931	0.29310345	0.24137931	0.29310345	-0.06896552
5	0.08045977	-0.06896552	0.08045977	-0.06896552	0.97701149

The fitted values can be expressed as linear combinations of the response vector \mathbf{Y} using coefficients found in \mathbf{H} .

```
> h %*% Y
```

	[,1]
1	12.081494
2	12.580862
3	12.081494
4	12.580862
5	9.085287

Because of this, \mathbf{H} is often referred to as the *hat matrix*.

H is a symmetric matrix and also has a special property called **idempotency**:

$$\underset{n \times n}{\mathbf{H}} \underset{n \times n}{\mathbf{H}} = \underset{n \times n}{\mathbf{H}}$$

```
> h %% h
```

	1	2	3	4	5
1	0.21839080	0.24137931	0.21839080	0.24137931	0.08045977
2	0.24137931	0.29310345	0.24137931	0.29310345	-0.06896552
3	0.21839080	0.24137931	0.21839080	0.24137931	0.08045977
4	0.24137931	0.29310345	0.24137931	0.29310345	-0.06896552
5	0.08045977	-0.06896552	0.08045977	-0.06896552	0.97701149

Residuals

Let the vector of residuals be

$$\underset{n \times 1}{\mathbf{e}} = \begin{bmatrix} e_1 \\ e_2 \\ \vdots \\ e_n \end{bmatrix}$$

In matrix notation

$$\begin{aligned} \underset{n \times 1}{\mathbf{e}} &= \underset{n \times 1}{\mathbf{Y}} - \underset{n \times 2}{\mathbf{X}} \underset{2 \times 1}{\mathbf{b}} \\ &= \underset{n \times 1}{\mathbf{Y}} - \underset{n \times 1}{\hat{\mathbf{Y}}} \end{aligned}$$

Since the fitted values can be expressed as \mathbf{HY} , the residual vector can also be expressed as

$$\underset{n \times 1}{\mathbf{e}} = \underset{n \times 1}{\mathbf{Y}} - \underset{n \times n}{\mathbf{H}} \underset{n \times 1}{\mathbf{Y}}$$

The residuals can also be expressed as linear combinations of the response vector \mathbf{Y} .

Using the distributive property as it relates to matrices, we re-write this as

$$\underset{n \times 1}{\mathbf{e}} = \left(\underset{n \times n}{\mathbf{I}} - \underset{n \times n}{\mathbf{H}} \right) \underset{n \times 1}{\mathbf{Y}}$$

The matrix $(\mathbf{I} - \mathbf{H})$, like \mathbf{H} , is *symmetric* and *idempotent*.

$$\underset{n \times 1}{\mathbf{e}} = \left(\underset{n \times n}{\mathbf{I}} - \underset{n \times n}{\mathbf{H}} \right) \underset{n \times 1}{\mathbf{Y}}$$

```
> i = diag(5)

> residuals = (i - h) %*% Y
> residuals

      [,1]
1 -0.08149425
2 -4.58086207
3  4.17850575
4  1.06913793
5 -0.58528736
```

Mean Squared Error (MSE)

Recall a Mean Square (or variance estimate) is the ratio of the sum of squares to its degrees of freedom.

$$\text{MSE} = \frac{\text{SS}_{\text{error}}}{df_{\text{error}}} = \frac{\sum (e_i)^2}{n - p}$$

In matrix notation

$$\text{SS}_{\text{Error}} = \mathbf{e}'\mathbf{e}$$

```
# Compute SS for the residuals  
> SSE = t(residuals) %*% residuals  
> SSE
```

```
      [,1]  
[1,] 39.93647
```

```
# Compute df for the residuals  
> num_coef = 2  
> df_residual = nrow(myData) - num_coef  
  
> MSE = SSE / df_residual  
> MSE
```

```
      [,1]  
[1,] 13.31216
```

Standard Errors for the Coefficients

The SEs for the coefficients represent the uncertainty in the estimates.

The uncertainty is also correlated between the coefficients (unless you have a balanced design.)

We represent the uncertainty and the correlation in the estimated variance–covariance matrix of the coefficients,

$$\mathbf{V}_{\beta} \hat{\sigma}^2$$

where

$$\mathbf{V}_{\beta} = (\mathbf{X}^t \mathbf{X})^{-1}$$

and

$$\hat{\sigma}^2 = \text{MSE}$$

```
> V_b = solve(t(X) %*% X)
> V_b
```

	[,1]	[,2]
[1,]	28.1724138	-0.89655172
[2,]	-0.8965517	0.02873563

$$\mathbf{V}_{\beta} = \begin{bmatrix} 28.17 & -0.90 \\ -0.90 & 0.03 \end{bmatrix}$$

$$\hat{\sigma}^2 = 13.31$$

$$\begin{bmatrix} 28.17 & -0.90 \\ -0.90 & 0.03 \end{bmatrix} (13.31) = \begin{bmatrix} 375.0 & -11.9 \\ -11.9 & 0.04 \end{bmatrix}$$

```
> var_cov = as.numeric(MSE) * V_b
> var_cov
```

	[,1]	[,2]
[1,]	375.03555	-11.9350358
[2,]	-11.93504	0.3825332

The diagonal elements are the *variances* of the regression coefficients and the off-diagonal elements are the *covariances*.

We can also compute the estimated SE for each coefficient and the correlation between the coefficients.

$$\mathbf{V}_{\beta} = \begin{bmatrix} 28.17 & -0.90 \\ -0.90 & 0.03 \end{bmatrix}$$

$$SE_{\beta_0} = (\sqrt{\mathbf{V}_{11}})(\hat{\sigma})$$

$$SE_{\beta_1} = (\sqrt{\mathbf{V}_{22}})(\hat{\sigma})$$

$$r_{\beta_1, \beta_2} = \frac{\mathbf{V}_{12}}{\sqrt{\mathbf{V}_{11} \mathbf{V}_{22}}}$$

$$SE_{\beta_0} = (\sqrt{28.17})(3.65) = 19.37$$

$$SE_{\hat{\beta}_1} = (\sqrt{0.03})(3.65) = 0.62$$

$$r_{\hat{\beta}_1, \hat{\beta}_2} = \frac{-0.90}{\sqrt{(28.17)(0.03)}} = -0.996$$

```
# Compute SEs
> sqrt(diag(var_cov))

[1] 19.3658346  0.6184927

# Compute correlation between coefficients
> V_b[1, 2] / sqrt(prod(diag(V_b)))

[1] -0.9964441
```

There are several built-in function in R that compute these values for us in practice.

```
# Fit the linear model  
> lm.1 = lm(wage ~ 1 + age, data = myData)
```

```
# Regression summary  
> summary(lm.1)
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-3.8983	19.3658	-0.201	0.853
age	0.4994	0.6185	0.807	0.479

```
# Analysis of variance decomposition  
> anova(lm.1)
```

Analysis of Variance Table

Response: wage

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
age	1	8.678	8.678	0.6519	0.4785
Residuals	3	39.936	13.312		

```
# Obtain the regression coefficients
```

```
> coef(lm.1)
```

```
(Intercept)      age  
-3.8982759    0.4993678
```

```
# Get the fitted values
```

```
> fitted(lm.1)
```

```
      1      2      3      4      5  
12.081494 12.580862 12.081494 12.580862  9.085287
```

```
# Get the residuals
```

```
> resid(lm.1)
```

```
      1      2      3      4      5  
-0.08149425 -4.58086207  4.17850575  1.06913793 -0.58528736
```



```
# Obtain the model (X) matrix
```

```
> model.matrix(lm.1)
```

```
  (Intercept) age
1           1  32
2           1  33
3           1  32
4           1  33
5           1  26
```

```
# Obtain the variance-covariance matrix of the coefficients
```

```
> vcov(lm.1)
```

```
          (Intercept)          age
(Intercept) 375.03555 -11.9350358
age         -11.93504   0.3825332
```