

Replication Exercise 1 - Field Experiments

Marina Merlo - n. USP: 7197987

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Question 1 - What hypotheses are GGL testing? Where did they get these hypotheses from?

They're testing whether social pressure increases one's propensity to vote, getting this hypothesis from the social psychology and voting behavior literature. They test different levels of social enforcement on voting, from the single reminder of the civic duty of voting up to a higher pressure closer to the "shaming tactics".

Question 2 - What are the treatment and control conditions? What is the outcome variable?

The treatment conditions is to receive a mail before the election with incentives to vote; the control was to receive nothing. The outcome variable was the voting rate.

Question 3 - What is the unit of analysis? What is the level at which treatment is applied?

The unit of analysis is each individual, but the treatment is applied per household.

Question 4 - Did randomization work? Let's reproduce Table 1 of GGL to conduct balance tests between the treatment and control groups on pre-treatment covariates. Note that GGL evaluate balance at the household level so you first need to aggregate the individual data to the household level by finding the household mean on each of the variables we want to assess balance for. Then calculate the mean across household separately for the control and treatment groups. What do we learn from the results?

Table 1: Replication of table 1

treatment	hh_size	g2002	g2000	p2004	p2002	p2000	sex	age	n
Civic Duty	1.91	0.84	0.87	0.42	0.41	0.27	0.5	51.18	20001
Control	1.91	0.83	0.87	0.42	0.41	0.26	0.5	51.31	99999
Hawthorne	1.91	0.84	0.87	0.42	0.41	0.26	0.5	51.20	20002
Neighbors	1.91	0.84	0.87	0.42	0.41	0.26	0.5	51.34	20000
Self	1.91	0.84	0.86	0.42	0.41	0.26	0.5	51.24	20000

Question 5 - GGL don't bother to do a t-test for the difference-in-means, but let's do it ourselves. Conduct a t-test for the difference in mean household age between the Control and 'Neighbors' conditions. Interpret the result.

```
##
## Welch Two Sample t-test
##
## data: data$age[data$treatment == "Control"] and data$age[data$treatment == "Neighbors"]
## t = -0.48289, df = 54214, p-value = 0.6292
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -0.1992680 0.1204886
```

```
## sample estimates:
## mean of x mean of y
## 49.81355 49.85294
```

There's no significant difference between the mean age between control and neighbors - that means control and treatment are balanced on age and this variable won't be a source of bias on our estimates.

Question 6 - Now let's look at the results of the experiment. Perform a simple difference-in-means-test for voter turnout between the Control and 'Neighbors' groups in the individual data. Interpret the result.

```
##
## Welch Two Sample t-test
##
## data: data$voted[data$treatment == "Control"] and data$voted[data$treatment == "Neighbors"]
## t = -30.207, df = 52613, p-value < 2.2e-16
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -0.08658577 -0.07603405
## sample estimates:
## mean of x mean of y
## 0.2966383 0.3779482
```

While the voting average for the control group was 29,66%, the "Neighbors" treatment group voted, on average, 37,79%. This difference is statistically significant at 0.05, meaning there's a positive effect of the treatment (social pressure as knowing the neighbors voting) on the outcome (voting on 2006 Primary Election)

Question 7 - Now run an OLS regression to understand the effect of each treatment on voter turnout, to replicate column (a) of Table 3. (If you prefer you can use treatment as a factor variable, not a series of dummies like the authors use. For this question do not adjust the standard errors). Interpret the results in terms of the change in probability of voting in the election.

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.2966	0.001061	279.5	0
treatmentCivic Duty	0.0179	0.0026	6.884	5.849e-12
treatmentHawthorne	0.02574	0.002601	9.896	4.37e-23
treatmentSelf	0.04851	0.0026	18.66	1.219e-77
treatmentNeighbors	0.08131	0.002601	31.26	2.94e-214

Table 3: Fitting linear model: voted ~ treatment

Observations	Residual Std. Error	R^2	Adjusted R^2
344084	0.4641	0.003394	0.003383

The results show a progressive increase in the voting turnout along the four groups of treatments and their social pressure effect, confirming the hypothesis. For the highest social pressure treatment (Neighbors), for

example, there's a 8.13% increase in the voting turnout compared to the control group.

Question 9 - Repeat your regression but this time with standard errors clustered to the household level. What difference does this make? Why do we do this?

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.2966	0.00131	226.5	0
treatmentCivic Duty	0.0179	0.003237	5.53	3.2e-08
treatmentHawthorne	0.02574	0.003258	7.9	2.804e-15
treatmentSelf	0.04851	0.0033	14.7	6.584e-49
treatmentNeighbors	0.08131	0.00337	24.13	1.522e-128

Table 5: Fitting linear model: $\text{voted} \sim \text{treatment}$ Now the standard errors are slightly higher for all the estimates - for example, the standard error for the Neighbors treatment coefficient was 0.0026 and now it's 0.0033. We do this so we can take unobserved intra-household correlation into account on our estimation.

Observations	Residual Std. Error	R^2	Adjusted R^2
344084	0.4641	0.003394	0.003383

Question 10 - Next, we want to add block-level fixed effects to our model to reproduce column (b) of Table 3. How do the results change? How does this change the comparisons we are making in the regression?

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.3102	0.001208	256.8	0
treat.neighbor2	0.08207	0.003373	24.33	1.471e-130

Table 7: Fitting linear model: $\text{voted} \sim \text{treat.neighbor2}$

Observations	Residual Std. Error	R^2	Adjusted R^2
229444	0.4616	0.004318	0.004314

The voting turnout on control group and the effect from the Neighbor treatment now it's a little bit higher when compared to the regression without the block-level fixed effects. [Not sure how to interpret this change, however]

Question 11 - Add covariates (g2000,g2002,p2000,p2002,p2004) to your model from Q10 to reproduce column 3 of Table 3. How do the results change when we add covariates?

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	0.09466	0.002386	39.67	0
treat.neighbor2	0.0817	0.00324	25.22	3.8e-140

	Estimate	Std. Error	t value	Pr(> t)
g2000	-0.004778	0.002935	-1.628	0.1035
g2002	0.09883	0.002654	37.24	1.164e-302
p2000	0.099	0.00262	37.78	2.014e-311
p2002	0.1343	0.002322	57.83	0
p2004	0.1551	0.00227	68.32	0

Table 9: Fitting linear model: $\text{voted} \sim \text{treat.neighbor2} + \text{g2000} + \text{g2002} + \text{p2000} + \text{p2002} + \text{p2004}$

Observations	Residual Std. Error	R^2	Adjusted R^2
229444	0.4447	0.07582	0.0758

The results don't change much, with slight reduction for the Neighbor estimate. The estimate goes from 0.08207 (without covariates) to 0.08170 (with covariates). This means the covariates we added have so little effect on the voting turnout that they don't change the effect of the treatment.

Question 12 - In place of an OLS regression, use a logit regression model to run the same model as in Q8. How would you interpret the results?

```
## How to cite this model in Zelig:
##   R Core Team. 2007.
##   logit: Logistic Regression for Dichotomous Dependent Variables
##   in Christine Choirat, Christopher Gandrud, James Honaker, Kosuke Imai, Gary King, and Olivia Lau,
##   "Zelig: Everyone's Statistical Software," http://zeligproject.org/
```

Table 10: Fitting generalized (binomial/logit) linear model: $\text{voted} \sim \text{treatment}$

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-0.8634	0.005006	-172.5	0
treatmentCivic Duty	0.08437	0.0121	6.972	3.118e-12
treatmentHawthorne	0.1205	0.01204	10.01	1.39e-23
treatmentSelf	0.2229	0.01187	18.79	9.797e-79
treatmentNeighbors	0.3651	0.01168	31.26	1.644e-214

The estimates of the logit regression are the log of the odds (probability of success over the probability of failure), so we need to exponentiate the estimates to get the odds ratio. For example, the coefficient for Neighbors treatment is the log of odds ratio between the Neighbor treated group and the Control group. Exponentiating it:

```
exp(0.3651)
```

```
## [1] 1.440658
```

Or else, in Zelig, we can just add "odds_ratio = TRUE" in the summary to the conversion is automatic:

```
summary(reg4.logit,odds_ratios = TRUE)
```

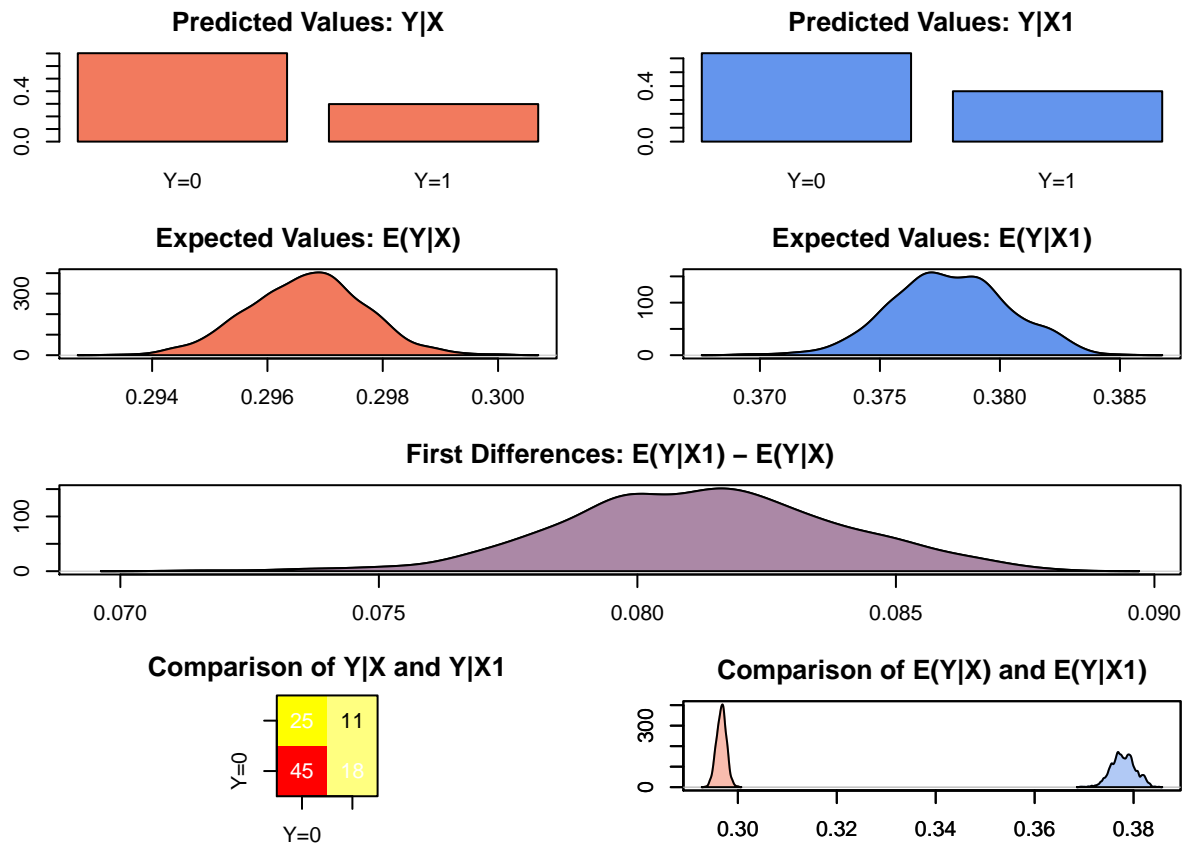
```
## Model:
##
## Call:
## stats::glm(formula = voted ~ treatment, family = binomial("logit"),
##   data = as.data.frame(.))
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -0.9744  -0.8691  -0.8389   1.4586   1.5590
##
## Coefficients:
##              Estimate (OR) Std. Error (OR)  z value Pr(>|z|)
## (Intercept)      0.421744      0.002111 -172.459  < 2e-16 ***
## treatmentCivic Duty  1.088029      0.013166   6.972 3.12e-12 ***
## treatmentHawthorne  1.128035      0.013578  10.009  < 2e-16 ***
## treatmentSelf       1.249742      0.014831  18.786  < 2e-16 ***
## treatmentNeighbors  1.440646      0.016826  31.260  < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 429238  on 344083  degrees of freedom
## Residual deviance: 428090  on 344079  degrees of freedom
## AIC: 428100
##
## Number of Fisher Scoring iterations: 4
```

So, the odds of voting for Neighbors-treated group is 44% higher than the odds for the control group.

Question 13 - Predict the first difference (the change) of the probability of voting when moving from the ‘Control’ to ‘Neighbors’ treatment category using your logit model from Q12.

```
##
## sim x :
## -----
## ev
##      mean      sd      50%      2.5%      97.5%
## [1,] 0.2966815 0.0009972789 0.2967058 0.2946763 0.2985833
## pv
##      0      1
## [1,] 0.702 0.298
##
## sim x1 :
## -----
## ev
##      mean      sd      50%      2.5%      97.5%
## [1,] 0.3779259 0.002429126 0.3778734 0.3733004 0.3825576
## pv
```

```
##          0          1
## [1,] 0.637 0.363
## fd
##          mean          sd          50%          2.5%          97.5%
## [1,] 0.08124433 0.002610724 0.08127775 0.07628583 0.08634921
```



There's a change of 0.0812 for the expected value for voting when moving from control to treatment (but I'm not sure yet how to actually interpret this mean in a logit)

Question 14 - How does the data processing the authors conduct on p.36-37 (under 'Study Population') affect your interpretation of the conclusions?

The missing ZIP codes or vote history might have a different voting behavior than those who were included, i.e., the missingness wasn't completely at random. The exclusion of blocks with high number of apartments also can affect the results given people who live in these blocks can have a different voting pattern, or different propensity for social pressure than single-family homes

Question 15 - How generalizable are the findings of this study to other elections? To the same set of elections in 2010? To neighbouring Indiana in the same year? To elections in Brazil?

It's plausible to believe the same effects will hold for other north-american elections, but maybe with smaller impact. For more recent elections, the social pressure measured by the mail experiment might be less effective when interacting with other sources of social norms and expectations as those in the social media.

The results could also change on neighborhoods with higher urban density, where each individual can go by almost anonymous and will experience less social pressure. On the other hand, less urban neighborhoods or countryside cities where everyone know each other the social pressure effect can be even higher.

For Brazil, due to the compulsory voting and where the voter turnout hasn't been a problem, the parameters of the voting calculus presented by the authors (p. 35) is probably different. Given the cost of not voting is higher than voting (if you decide not to vote, you have to justify your absence or pay a fine later; otherwise, you'll lose your passport and other rights), these incentives might have a marginal effect on a already high voter turnout.