

AMS 206 - Winter 2017

Quiz 1 - Solution

1. **[30 Points]** Show that the Beta family is a conjugate family to a negative binomial likelihood with unknown success probability and known number of successes, i.e., if $x | \theta \sim \text{NegBin}(k, \theta)$ so that the probability mass function is

$$p(x | \theta) = \binom{x-1}{k-1} \theta^k (1-\theta)^{x-k} \quad x = k, k+1, k+2, \dots$$

and $\theta \sim \text{Beta}(a, b)$ with density

$$p(\theta) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1} \quad \theta \in [0, 1]$$

then $\theta | x \sim \text{Beta}(a^*, b^*)$ for a pair of parameters $a^*, b^* > 0$. Find an expression for a^* and b^* .

Solution: From Bayes theorem

$$\begin{aligned} p(\theta | x) &\propto \binom{x-1}{k-1} \theta^k (1-\theta)^{x-k} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1} \\ &\propto \theta^{k+a-1} (1-\theta)^{x-k+b-1} \quad \theta \in [0, 1] \end{aligned}$$

which is proportional to beta kernel with parameters $a^* = a + k$ and $b^* = x - k + b$. Hence,

$$p(\theta | x) = \frac{\Gamma(a+b+x)}{\Gamma(a+k)\Gamma(b+x-k)} \theta^{a+k-1} (1-\theta)^{x-k+b-1} \quad \theta \in [0, 1]$$

2. **[70 points]** Let x_1, \dots, x_n be an independent sample with $x_i | \theta, s_i \sim \text{Uni}[0, s_i \theta]$, where the values s_1, \dots, s_n are known covariates. If θ is assigned a Pareto prior with density

$$p(\theta) = \begin{cases} \frac{ab^a}{\theta^{a+1}} & \theta > b \\ 0 & \text{otherwise} \end{cases}.$$

- (a) Find the posterior distribution for θ . *Hint: the prior is conjugate!*

Solution: Note that

$$\begin{aligned} p(x_1, \dots, x_n | \theta) &= \begin{cases} \left\{ \prod_{i=1}^n \frac{1}{s_i} \right\} \frac{1}{\theta^n} & x_1 < s_1 \theta, \dots, x_n < s_n \theta \\ 0 & \text{otherwise} \end{cases} \\ &\propto \begin{cases} \frac{1}{\theta^n} & \theta > \max_{1 \leq i \leq n} \left\{ \frac{x_i}{s_i} \right\} \\ 0 & \text{otherwise} \end{cases} \end{aligned}$$

Hence,

$$p(\theta | x_1, \dots, x_n) \propto \begin{cases} \frac{1}{\theta^n} \frac{1}{\theta^{a+1}} & \theta > \max_{1 \leq i \leq n} \left\{ \frac{x_i}{s_i} \right\}, \theta > b \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} \frac{1}{\theta^{n+a+1}} & \theta > \max \left\{ b, \max_{1 \leq i \leq n} \left\{ \frac{x_i}{s_i} \right\} \right\} \\ 0 & \text{otherwise} \end{cases}$$

This corresponds to the kernel of a Pareto distribution with parameters $a^* = a + n$ and $b^* = \max \left\{ b, \max_{1 \leq i \leq n} \left\{ \frac{x_i}{s_i} \right\} \right\}$. Hence

$$p(\theta | x_1, \dots, x_n) = \begin{cases} \frac{a^* (b^*)^{a^*}}{\theta^{a^*+1}} & \theta > b^* \\ 0 & \text{otherwise} \end{cases}.$$

- (b) Find a closed-form expressions for the posterior mean and $\alpha\%$ HPD credible intervals for θ .

Solution: The posterior expectation is

$$\begin{aligned} E\{\theta | x_1, \dots, x_n\} &= \int_{b^*}^{\infty} \theta \frac{a^* (b^*)^{a^*}}{\theta^{a^*+1}} d\theta \\ &= a^* (b^*)^{a^*} \frac{1}{a^* - 1} \left[-\frac{1}{\theta^{a^*-1}} \right]_{b^*}^{\infty} \\ &= \frac{a^*}{a^* - 1} (b^*)^{a^*} \frac{1}{(b^*)^{a^*-1}} = \frac{a^*}{a^* - 1} b^* \end{aligned}$$

For the posterior HPD, note that the Pareto has a decreasing density function. Therefore, the interval is of the form (b^*, u) where u satisfies

$$\int_{b^*}^u \frac{a^* (b^*)^{a^*}}{\theta^{a^*+1}} d\theta = \alpha$$

Solving the integral we have

$$\begin{aligned} \int_{b^*}^u \frac{a^* (b^*)^{a^*}}{\theta^{a^*+1}} d\theta &= (b^*)^{a^*} \left[-\frac{1}{\theta^{a^*}} \right]_{b^*}^u \\ &= (b^*)^{a^*} \left[\frac{1}{(b^*)^{a^*}} - \frac{1}{u^{a^*}} \right] = 1 - \left(\frac{b}{u} \right)^{a^*} \end{aligned}$$

Hence

$$1 - \left(\frac{b}{u} \right)^{a^*} = \alpha \quad \Rightarrow \quad u = \frac{b}{(1 - \alpha)^{1/a^*}}.$$

(c) Find the predictive distribution for a new observation x^* with associated covariate s^* ,

$$p(x^* \mid s^*, x_1, \dots, x_n, s_1, \dots, s_n, a, b)$$

Solution: Note that

$$p(x^* \mid s^*, x_1, \dots, x_n, s_1, \dots, s_n, a, b) = \int p(x^* \mid \theta, s^*, x_1, \dots, x_n, s_1, \dots, s_n, a, b) p(\theta \mid s^*, x_1, \dots, x_n, s_1, \dots, s_n, a, b) d\theta$$

In this setting

$$p(x^* \mid \theta, s^*, x_1, \dots, x_n, s_1, \dots, s_n, a, b) = p(x^* \mid \theta, s^*)$$

and

$$p(\theta \mid s^*, x_1, \dots, x_n, s_1, \dots, s_n, a, b) = p(\theta \mid x_1, \dots, x_n, s_1, \dots, s_n, a, b)$$

so

$$\begin{aligned} p(x^* \mid s^*, x_1, \dots, x_n, s_1, \dots, s_n, a, b) &= \int p(x^* \mid \theta, s^*) p(\theta \mid x_1, \dots, x_n, s_1, \dots, s_n, a, b) d\theta \\ &= \int_{\max\{b^*, x^*/s^*\}}^{\infty} \frac{1}{s^*} \frac{a^* (b^*)^{a^*}}{\theta^{a^*+1}} d\theta \\ &= \frac{a^* (b^*)^{a^*}}{s^*} \int_{\max\{b^*, x^*/s^*\}}^{\infty} \frac{d\theta}{\theta^{a^*+2}} \\ &= \frac{a^* (b^*)^{a^*}}{s^* (a^* + 1)} \left[-\frac{1}{\theta^{a^*+1}} \right]_{\max\{b^*, x^*/s^*\}}^{\infty} \\ &= \frac{1}{s^*} \frac{a^*}{(a^* + 1)} \frac{(b^*)^{a^*}}{(\max\{b^*, x^*/s^*\})^{a^*+1}} \quad 0 \leq x^* < \infty \end{aligned}$$