#### Linear Regression Models P8111

3/8 5:00

Lecture 13

Hammer LL 109

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#### Today's Lecture

- Model selection vs. model checking
- Continue with model checking (regression diagnostics)

### Model selection vs. model checking

In a model of the form

$$y|x = f(x) + \epsilon \qquad \epsilon \sim (0, \sigma^* I)$$

model selection focuses on how you construct  $f(\cdot)$ ; model checking asks whether the  $\epsilon$  match the assumed form.

### Model checking

#### Two major areas of concern:

- Global lack of fit, or general breakdown of model Linearity" ×p is 'Right" assumptions

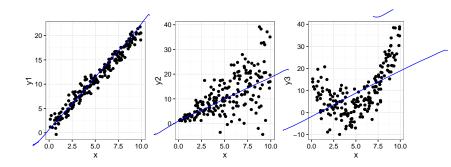
▶ Unbiased, uncorrelated errors 
$$E(\epsilon|x) = E(\epsilon) = 0$$
▶ Constant variance  $Var(y|x) = Var(\epsilon|x) = \sigma^2$ 
▶ Independent errors
▶ Normality of errors

Effect of influential points and outliers

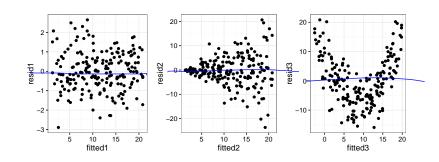
### Model checking

- Global lack of fit, or general breakdown of model assumptions
  - Residual analysis QQ plots, residual plots against fitted values and predictors Adjusted variable plots
- Effect of influential points and outliers
  - ► Measure of leverage, influence, outlying-ness

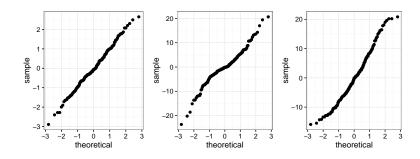
## Some data plots



# Some residual plots



### Checking Normality assumption



#### Non-constant variance

$$\mathcal{E} \sim \mathcal{N}(0, \mathcal{E})$$
 What to do ... 
$$\mathcal{I} \sim \mathcal{N}(x\beta, \mathcal{E})$$

- Nothing; just use least squares and bootstrap
- Use weighted LS, GLS (later)
- Use a variance stabilizing transformation

### Variance-stabilizing transformation



Suppose y is strictly positive,  $\mu = E(y|x)$ ,  $Var(y|x) = \sigma^2 g(\mu)$ 

- Replace y with  $y^* = T(y)$  such that  $Var(y^*|x)$  is approximately constant
- Delta method says  $Var(T(y)) = (T'(\mu))^2 \sigma^2 g(\mu)$

#### Variance-stabilizing transformation

To get constant variance, we want

So the transformation necessary to stabilize the variance really depends on the variance function itself, e.g.  $g(\cdot)$ 

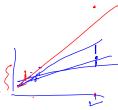
# Variance-stabilizing transformation examples

■ Example 1: If  $Var(y|x) = \sigma^2 \mu^2$ , i.e. if  $g(y) = y^2$ , T(y) = ?

$$T(y) = \int \frac{dy}{\sqrt{g(y)}} dy = \int \frac{dy}{\sqrt{y}} dy = ln(y)$$

■ Example 2: If  $Var(y|x) = \sigma^2 \underline{\mu}$ , i.e. if g(y) = y, T(y) = ?

## Isolated points



Points can be isolated in three ways

- Leverage point outlier in *x*
- Outlier outlier in y|x
- Influential point a point that largely affects *β* 
  - Deletion influence;  $|\hat{\boldsymbol{\beta}} \hat{\boldsymbol{\beta}}_{(-i)}|$
  - ► Basically, a high-leverage outlier

Leverage is measured by the hat matrix, outlying-ness by the residual

## Quantifying leverage

We measure leverage (the "distance" of  $x_i$  from the distribution of x) using

$$h_{ii} = \mathbf{x}_i^T (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{x}_i$$

 $h_{ii} = x_i^T (X^T X)^{-1} x_i$  where  $h_{ii}$  is the  $(i, i)^{th}$  entry of the hat matrix.

$$\begin{array}{c} \times (x^{T}x)^{-1}x^{T} \\ \hline \\ \left(x^{T}x^{T}\right) \\ \end{array}$$

### Leverage

Some notes about the hat matrix

$$\sum_{i} h_{ii} \stackrel{\text{def}}{=} tr(\mathbf{H}) = (p+1)$$

$$tr(\underbrace{x(x^{T}x)^{T}(x^{T})}_{x}) = tr(\underbrace{x(x^{T}x)^{T}}_{x})$$

(Note – the trace of the hat matrix generalizes to non-parametric methods, where you don't have a specific number of parameters to count. This is a useful measure of "model size" or "effective degrees of freedom" in these cases.)

$$5y = 3$$

# Leverage

Some notes about the hat matrix

$$\hat{y}_i = \sum_j \underline{h}_{ij} y_j$$

$$\hat{y}_i = \hat{h}_{ij} y_j$$

$$\sum_{i} h_{ij} = \sum_{j} h_{ij} = 1$$

$$\underbrace{h_{ii}}_{ii} \approx ($$

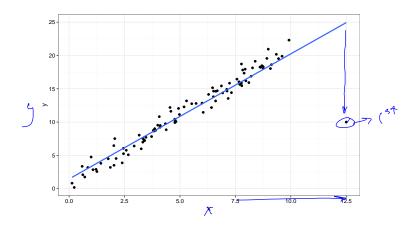
$$h_{ij} \approx 0$$

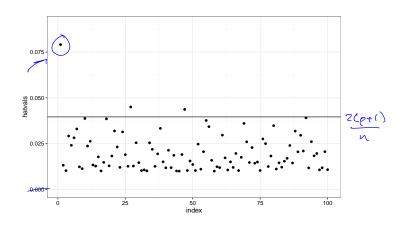
These mean that  $h_{ii}$  is the weight given to  $y_i$  in determining  $\hat{y}_i$ 

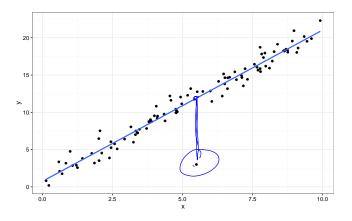
## Leverage

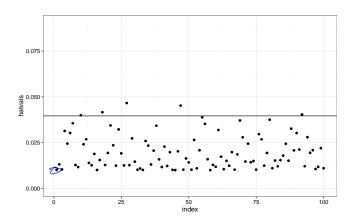
What counts as "big" leverage?

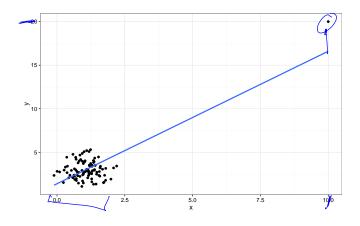
- Average leverage is (p+1)/n
- Typical rules of thumb are 2(p+1)/n or 3(p+1)/n
- Leverage plots can be useful as well

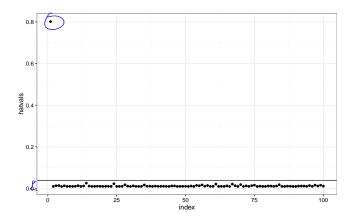












#### **Outliers**

- When we refer to "outliers" we typically mean "points that don't have the same mean structure as the rest of the data"
  - Residuals give an idea of "outlying-ness", but we need to standardize somehow
- Remember (from last lecture)  $Var(\hat{\epsilon}_i) = \sigma^2(1 h_{ii}) \dots$

#### Outliers

The standardized residual is given by

$$\hat{\epsilon}_i^* = \frac{\underline{\hat{\epsilon}_i}}{\sqrt{Var(\hat{\epsilon}_i)}} = \frac{\hat{\epsilon}_i}{\hat{\sigma}\sqrt{(1 - h_{ii})}}$$

The Studentized residual is given by

$$t_{i} = \frac{\hat{\epsilon}_{i}}{\hat{\sigma}_{(-i)}\sqrt{(1-h_{ii})}} = \hat{\epsilon}_{i}^{*} \left(\frac{n-(p+1)}{n-(p+1)-\hat{\epsilon}_{i}^{*2}}\right)^{1/2}$$

Studentized residuals follow a  $t_{n-(p+1)-1}$  distribution.

#### Influence

Specifically, deletion influence

$$|\hat{\beta} - \hat{\beta}_{(-i)}| = (\hat{\beta} - \hat{\beta}_{(-i)})^{\mathsf{T}} (\hat{\beta} - \hat{\beta}_{(-i)})^{\mathsf{T}}$$

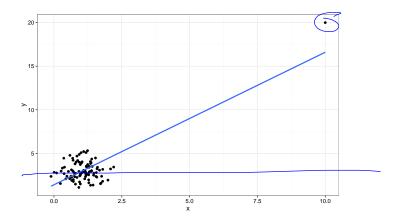
Cook's distance is

$$D_{i} = \frac{(\hat{\beta} - \hat{\beta}_{(-i)})^{T} (X^{T}X)(\hat{\beta} - \hat{\beta}_{(-i)})}{(p+1)\hat{\sigma}^{2}}$$

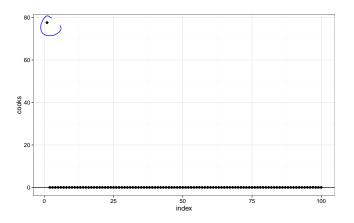
$$= \frac{(\hat{y} - \hat{y}_{(-i)})^{T} (\hat{y} - \hat{y}_{(-i)})}{(p+1)\hat{\sigma}^{2}}$$

$$= \frac{1}{p+1} \hat{\epsilon}_{i}^{2} \frac{h_{ii}}{1 - h_{ii}}$$

# Cook's distance plot



## Cook's distance plot



#### Handy R functions

#### Suppose you fit a linear model in R;

- hatvalues gives the diagonal elements of the hat matrix  $h_{ii}$  (leverages)
- rstandard gives the standardized residuals
- rstudent gives the studentized residuals
- cooks.distance gives the Cook's distances

### Today's big ideas

■ Model checking

■ Suggested reading: Faraway Ch 7