PLSC 502 – Autumn 2016 Estimation and Estimators

October 11, 2016

Random Variables, Take Two

For:

$$X_i = \mu + u_i$$

and so

$$u_i = X_i - \mu,$$

Random Variables, Take Two

That means

$$E(u) = E(X - \mu)$$

$$= E(X) - E(\mu)$$

$$= E(X) - \mu$$

$$= \mu - \mu$$

$$= 0$$

and

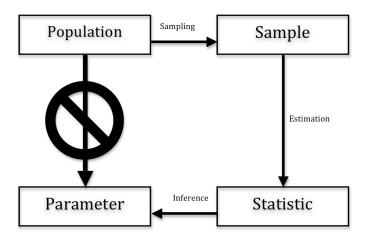
$$Var(X) = E[(X - \mu)^{2}]$$

$$= E(u^{2})$$

$$Var(u) = E[(u - E(u))^{2}]$$

$$= E[(u - 0)^{2}]$$

$$= E(u^{2}).$$



Estimation Example: \bar{X}

$$\bar{X} = \frac{1}{N} \sum_{i=1}^{N} X_i
= \frac{1}{N} \sum_{i=1}^{N} (\mu + u_i)
= \frac{1}{N} \sum_{i=1}^{N} (\mu) + \frac{1}{N} \sum_{i=1}^{N} (u_i)
= \frac{1}{N} (N\mu) + \frac{1}{N} \sum_{i=1}^{N} (u_i)
= \mu + \bar{u}$$

Properties of Estimators

Small-Sample Properties

- Hold irrespective of N
- "Small sample estimators"

Large-Sample (Asymptotic) Properties

- Hold as $N \to \infty$
- "More is better"

Unbiasedness

Means:

$$\mathsf{E}(\hat{\theta}) = \theta$$

"Bias":

$$B(\hat{\theta}) = \mathsf{E}(\hat{\theta}) - \theta$$

Example: \bar{X}

$$E(\bar{X}) = E(\mu + \bar{u})$$

$$= E(\mu) + E(\bar{u})$$

$$= \mu + 0$$

$$= \mu$$

Multiple Unbiased Estimators

For N=2:

$$Z = \lambda_1 X_1 + \lambda_2 X_2.$$

note that

$$E(Z) = E(\lambda_1 X_1 + \lambda_2 X_2)$$

$$= E(\lambda_1 X_1) + E(\lambda_2 X_2)$$

$$= \lambda_1 E(X_1) + \lambda_2 E(X_2)$$

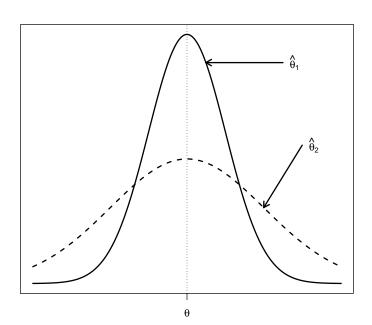
$$= \lambda_1 \mu + \lambda_2 \mu$$

$$= (\lambda_1 + \lambda_2) \mu$$

Means

$$E(Z) = \mu \iff (\lambda_1 + \lambda_2) = 1.0$$

Efficiency



Efficiency (continued)

Note:

$$Var(Z) = Var(\lambda_1 X_1 + \lambda_2 X_2)$$
$$= (\lambda_1^2 + \lambda_2^2)\sigma^2$$

and:

$$\begin{array}{rcl} \lambda_1^2 + \lambda_2^2 & = & \lambda_1^2 + (1 - \lambda_1)^2 \\ & = & \lambda_1^2 + (1 - 2\lambda_1 + \lambda_1^2) \\ & = & 2\lambda_1^2 - 2\lambda_1 + 1. \end{array}$$

Minimize:

$$\begin{array}{ccc} \frac{\partial 2\lambda_1^2-2\lambda_1+1}{\partial \lambda_1} & = & 4\lambda_1-2 \\ \\ 4\lambda_1-2 & = & 0 \\ \\ \lambda_1 & = & 0.5 \end{array}$$

Mean Squared Error

$$\mathsf{MSE}(\hat{\theta}) = \mathsf{E}[(\hat{\theta} - \theta)^2]$$
$$= \mathsf{E}[B(\hat{\theta})^2]$$
$$= \mathsf{Var}(\hat{\theta}) + [B(\hat{\theta})^2]$$

Among unbiased estimators, the efficient estimator will always have the smallest MSE (because $B(\hat{\theta}) = [B(\hat{\theta})]^2 = 0$).

Comparing Estimators via MSE

\bar{X} has:

- $B(\bar{X})=0$
- · Var(\bar{X}) = σ^2/N , so
- · $MSE(\bar{X}) = \sigma^2/N + (0)^2 = \sigma^2/N$.

$\lambda = 6$ has:

$$B(\lambda) = E(\lambda - \mu)$$

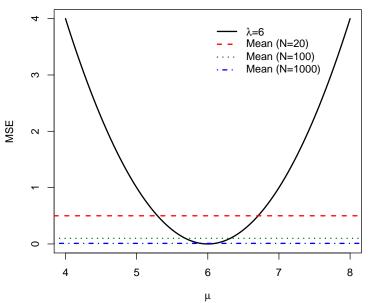
$$= E(6) - E(\mu)$$

$$= 6 - \mu,$$

$$Var(\lambda) = Var(6)$$
$$= 0$$

and so:

$$\begin{aligned} \mathsf{MSE}(\lambda) &=& \mathsf{Var}(\lambda) + [B(\lambda)]^2 \\ &=& 0 + (6 - \mu)^2 \\ &=& 36 - 12\mu + \mu^2 \end{aligned}$$



The black line is the MSE of λ , expressed as a function of the "true" population mean μ . The other colored lines are the MSEs for \bar{X} , under the assumption that $\sigma^2 = 10$ and $N = \{20, 100, 1000\}$, respectively.

Large-Sample Properties: Consistency

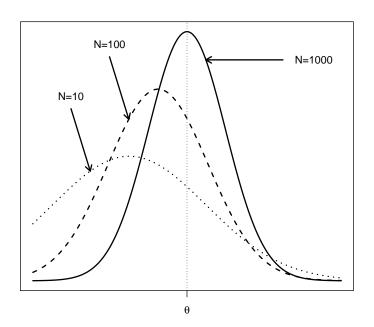
An estimator $\hat{\theta}$ is *consistent* if:

$$\underset{N \rightarrow \infty}{\lim} \Pr[|\hat{\theta} - \theta| < \epsilon] = 1.0$$

for an arbitrarily small $\epsilon > 0$

Equivalently:

$$\mathsf{E}(\hat{ heta}_{\mathsf{N}}) o heta$$
 as $\mathsf{N} o \infty$



Estimation, Generally

- Unbiased > Consistent > Biased
- Fully Efficient > Asymptotically Efficient > Inefficient
- MSE is <u>one</u> way to trade off bias vs. efficiency