PLSC 502 – Autumn 2016 Probability Distributions, I

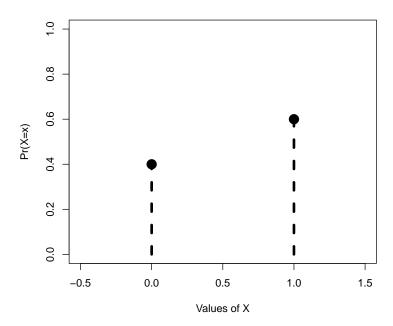
September 27, 2016

Bernoulli

$$X = 0$$
 with probability $1 - \pi$
= 1 with probability π .

$$f(x) = \begin{cases} 1 - \pi & \text{for } X = 0 \\ \pi & \text{for } X = 1 \end{cases}$$
$$= \pi^{x} (1 - \pi)^{1 - x}, x \in \{0, 1\}$$

 $X \sim \text{Bernoulli}(\pi)$



Bernoulli, continued

CDF:

$$F(x) = \sum_{x} f(x)$$

$$= \begin{cases} 1 - \pi & \text{for } X = 0 \\ 1 & \text{for } X = 1 \end{cases}$$

Expectation:

$$E(X) = \sum_{x} xf(x)$$
= (0)(1 - \pi) + (1)(\pi)
= \pi

Bernoulli, continued

Variance:

$$Var(X) = \sum_{x} [X - E(X)]^{2} f(x)$$

$$= \sum_{x} [X - \pi]^{2} f(x)$$

$$= (0 - \pi)^{2} (1 - \pi) + (1 - \pi)^{2} \pi$$

$$= \pi^{2} - \pi^{3} + \pi - 2\pi^{2} + \pi^{3}$$

$$= \pi^{2} - \pi$$

$$= \pi(1 - \pi)$$

Skewness:

Skewness
$$=rac{(1-\pi)-\pi}{\sqrt{(1-\pi)\pi}}$$

Even More Bernoulli

MGF:

$$\psi(t) = \int_{-\infty}^{\infty} \exp(tx) dF(x)$$

$$= \sum_{n=0}^{1} \exp(tn)\pi^{n} (1-\pi)^{1-n}$$

$$= \exp(0)(1-\pi) + \exp(t)\pi$$

$$= (1-\pi) + \pi \exp(t)$$

Implying:

$$\frac{\partial^k \psi(t)}{\partial^k t} = \pi \exp(t) \ \forall \ k$$

and raw moments:

$$\mathsf{E}(X^k) = \pi \ \forall \ k > 0$$

Feel the Bern...

Central moments:

$$M_1=\pi$$
,

$$M_2=\pi(1-\pi),$$

$$M_3 = \pi(1-\pi)(1-2\pi),$$

etc.

Binomial

Assume n independent binary "trials," each with identical probability of "success" π . Then the number of "successes" in n trials follows a *binomial* distribution:

$$f(x) = \binom{n}{x} \pi^{x} (1 - \pi)^{n - x}$$

where recall that

$$\binom{n}{x} \equiv \frac{n!}{x!(n-x)!}.$$

$$X \sim \text{binomial}(n, \pi)$$
.

Why "binomial"?

Binomial Theorem:

$$(a+b)^n = \sum_{k=0}^n \binom{n}{k} a^k b^{n-k}.$$

e.g.

$$(a+b)^2 = a^2 + 2ab + b^2,$$

$$(a+b)^3 = a^3 + 3a^2b + 3ab^2 + y^3,$$

$$(a+b)^4 = a^4 + 4a^3b + 6a^2b^2 + 4ab^3 + b^4,$$

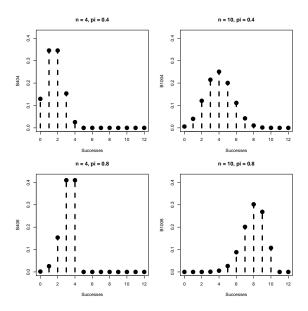
Binomial: Derivation

For n = 2:

$$Pr(X = 0)$$
 = $Pr(X_1 = 0, X_2 = 0)$
 = $Pr(X_1 = 0) \times Pr(X_2 = 0)$
 = $(1 - \pi)^2$

$$\begin{array}{lcl} \Pr(X=1) & = & \Pr(X_1=1, X_2=0 \ or \ X_1=0, X_2=1) \\ & = & \Pr(X_1=1) \times \Pr(X_2=0) + \Pr(X_1=0) \times \Pr(X_2=1) \\ & = & \pi(1-\pi) + (1-\pi)\pi \end{array}$$

$$\Pr(X = 2)$$
 = $\Pr(X_1 = 1, X_2 = 1)$
 = $\Pr(X_1 = 1) \times \Pr(X_2 = 1)$
 = π^2



More Binomial

CDF:

$$F(x) = \sum_{x}^{x} f(x)$$
$$= \sum_{j=0}^{x} {n \choose j} \pi^{j} (1-\pi)^{n-j}$$

Expectation:

$$\mathsf{E}(\mathsf{X}) = \mathsf{n}\pi,$$

More Binomial

Variance:

$$Var(X) = \sum_{x} [X - E(X)]^{2} f(x)$$

$$= \sum_{x} (X - \pi n)^{2} {n \choose x} \pi^{x} (1 - \pi)^{n-x}$$

$$= n\pi (1 - \pi).$$

Skewness:

Skewness =
$$\frac{1 - 2\pi}{\sqrt{n\pi(1 - \pi)}}$$

The Binomial...

- Is unimodal (except in certain cases),
- has median $\lceil n\pi \rceil$ or $\lceil n\pi \rceil$,
- has mode $\lceil (n+1)\pi \rceil$ or $\lfloor (n+1)\pi \rfloor$,
- has skewness that is:
 - \cdot increasing in n, and
 - · is largest when $\pi = 0.5$ for a fixed value of n.

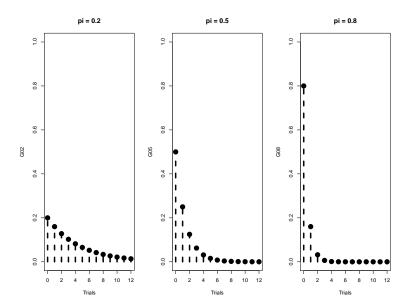
Geometric

The number of independent Bernoulli trials needed to achieve one success is a *geometric* random variable.

PDF:

$$f(x) = \pi (1-\pi)^{x-1}$$

$$X \sim \text{geometric}(\pi)$$
.



Geometric

CDF:

$$F(x) = \sum_{j=1}^{x} \pi (1-\pi)^{x-1}$$
$$= 1 - (1-\pi)^{x}$$

Expectation:

$$\mathsf{E}(X) = rac{1}{\pi}$$

Variance:

$$\mathsf{Var}(X) = \frac{1-\pi}{\pi^2}$$

Negative Binomial

The number of *failures we observe* (x) before achieving the rth success in n independent binomial trials (each with probability of success π) is distributed according to a *negative binomial* distribution.

PDF:

$$f(x) = \binom{r+x-1}{r-1} \pi^r (1-\pi)^x$$

More Negative Binomial

CDF:

$$F(x) = \sum_{j=0}^{x} {r+j-1 \choose r-1} \pi^{r} (1-\pi)^{j}$$
$$= 1 - CDF_{binomial}$$

Expected value:

$$\mathsf{E}(X) = \frac{(1-\pi)r}{\pi}$$

Even More Negative Binomial

Variance:

$$Var(X) = \frac{(1-\pi)r}{\pi^2}.$$

Skewness:

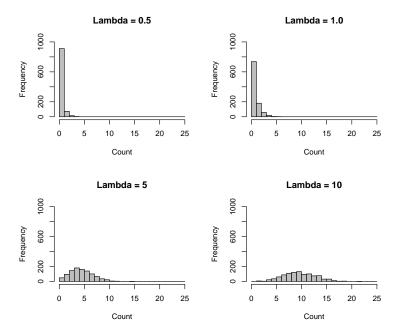
$$\mathsf{Skewness} = \frac{1+\pi}{\sqrt{\pi r}}$$

Poisson

For n independent Bernoulli trials with (sufficiently small) probability of success π and where $n\pi \equiv \lambda > 0$, the probability of observing exactly x total "successes" as the number of trials grows without limit is the *Poisson distribution*.

PDF:

$$f(x) = \lim_{n \to \infty} \left[\binom{n}{x} \left(\frac{\lambda}{n} \right)^{x} \left(1 - \frac{\lambda}{n} \right)^{n-x} \right]$$
$$= \frac{\lambda^{x} \exp(-\lambda)}{x!}.$$



More Poisson

CDF:

$$F(x) = \sum_{j=0}^{x} \frac{\lambda^{j} \exp(-\lambda)}{x!}.$$

Mean & Variance:

$$E(X) = Var(X) = \lambda$$

All higher moments are zero...

Alternative Poisson

Independent, constant-probability events occurring in time...

"arrival rate" =
$$\lambda$$

Implies:

$$Pr(\text{Event in } (t, t+h]) = \lambda$$

$$Pr(\text{No event in } (t, t+h]) = 1 - \lambda$$

$$N_{\text{Events occurring in } (t,t+h]} = \frac{\exp(-\lambda h)\lambda h^{x}}{x!}$$

If $h = 1 \forall h$, then:

$$f(x) = \frac{\exp(-\lambda)\lambda^x}{x!}$$

Multinomial

Imagine K possible distinct *outcomes* for each "trial," where each possible outcome has π_k and $\sum_{k=1}^K \pi_k = 1$.

 x_k = number of times we observe outcome k out of n trials.

Then for:

$$\mathbf{X} = \left(egin{array}{c} X_1 \ X_2 \ dots \ X_K \end{array}
ight)$$

the PDF is:

$$f(\mathbf{x}) = \frac{n!}{x_1!x_2!...x_K!} \pi_1^{x_1} \pi_2^{x_2} ... \pi_K^{x_K}$$

Multinomial, continued

Expected value:

$$\mathsf{E}(\mathsf{X}) \equiv \mathsf{E} \left(\begin{array}{c} \mathsf{X}_1 \\ \mathsf{X}_2 \\ \vdots \\ \mathsf{X}_K \end{array} \right) = n \left(\begin{array}{c} \pi_1 \\ \pi_2 \\ \vdots \\ \pi_K \end{array} \right)$$

Variance:

$$\mathsf{Var}(\mathbf{X}) \equiv \mathsf{Var} \left(egin{array}{c} X_1 \ X_2 \ dots \ X_K \end{array}
ight) = n \left[egin{array}{c} \pi_1(1-\pi_1) \ \pi_2(1-\pi_2) \ dots \ \pi_K(1-\pi_K) \end{array}
ight]$$

Covariance between X_s and X_t , $s \neq t$:

$$Cov(X_s, X_t) = -n\pi_s\pi_t$$

Schematic

