

PLSC 504 - Fall 2017

Cure Models

October 10, 2017

Standard models (e.g.):

$$h(T_i|\mathbf{X}_i, \beta) = \frac{f(T_i|\mathbf{X}_i, \beta)}{S(T_i|\mathbf{X}_i, \beta)}$$

assume:

$$\int_0^{\infty} f(t) dt = 1 \quad \forall i.$$

All observations will (eventually) experience the event of interest.

Mixture Cure Model

Assume (unobserved):

$$Y_i = \begin{cases} 1 & \text{for observations that will eventually fail,} \\ 0 & \text{for those that will not.} \end{cases}$$

For observations with $Y = 1$:

$$\begin{aligned} f(T_i | \mathbf{X}_i, \beta, Y_i = 1) &= g(T | \mathbf{X}_i, \beta) \\ F(T_i | \mathbf{X}_i, \beta, Y_i = 1) &= G(T | \mathbf{X}_i, \beta) \end{aligned}$$

For observations with $Y = 0$, $f(T)$ and $F(T)$ are undefined.

Mixture Cure Model (continued)

Define:

$$\Pr(Y_i = 1) = \delta_i.$$

Overall survival is then just:

$$S_i(T) = (1 - \delta_i) + \delta_i[1 - G_i(t)]$$

Mixture Cure Model: Likelihood

Then for $C_i = 1$:

$$\begin{aligned} L_i | C_i = 1 &= \Pr(Y_i = 1) \Pr(T_i = t | Y_i = 1, \mathbf{X}_i, \beta) \\ &= \delta_i g(T_i | \mathbf{X}_i, \beta) \end{aligned}$$

For $C_i = 0$:

$$\begin{aligned} L_i | C_i = 0 &= \Pr(Y_i = 0) + \Pr(Y_i = 1) \Pr(T_i > t_i | Y_i = 1, \mathbf{X}_i, \beta) \\ &= (1 - \delta_i) + \delta_i [1 - G(T_i | \mathbf{X}_i, \beta)] \end{aligned}$$

Mixture Cure Model: Likelihood

Implies:

$$\mathbf{L} = \prod_{i=1}^N [\delta_i g(T_i | \mathbf{X}_i, \beta)]^{C_i} \{(1 - \delta_i) + \delta_i [1 - G(T_i | \mathbf{X}_i, \beta)]\}^{(1 - C_i)}$$

and:

$$\begin{aligned} \ln \mathbf{L} = & \sum_{i=1}^N C_i \{ \ln(\delta_i) + \ln [g(T_i | \mathbf{X}_i, \beta)] \} \\ & + (1 - C_i) \ln \{ (1 - \delta_i) + \delta_i [1 - G(T_i | \mathbf{X}_i, \beta)] \} \end{aligned}$$

Mixture Cure Model: Specification

Typically:

$$\delta_i = \frac{\exp(\mathbf{Z}_i\gamma)}{1 + \exp(\mathbf{Z}_i\gamma)}$$

or:

$$\delta_i = \Phi(\mathbf{Z}_i\gamma).$$

Identified even if $\mathbf{Z} \equiv \mathbf{X}$.

Non-Mixture Cure Model (e.g. Sposto 2002)

N_i = number of pre-cancerous cell clusters, with:

$$N_i \sim \text{Poisson}(\lambda).$$

$\text{Pr}(\text{Cure})$ is:

$$\pi_i = \text{Pr}(N_i = 0).$$

Time to cancer onset for cluster j of observation i is:

$$Z_{ij} \sim F(t), \quad j = \{1, 2, \dots, N_i\}.$$

Non-Mixture Cure Model (continued)

Survival to first onset:

$$S(t) = \pi^{F(t)}$$

with hazard function:

$$h(t) = -\ln(\pi)f(t)$$

which reflects the fact that $\int_0^\infty h(t)dt = -\ln(\pi)$.

Non-Mixture Cure Model (continued)

Rewritten $S(t)$:

$$S(t) = \exp[\ln(\pi)F(t)].$$

Assuming:

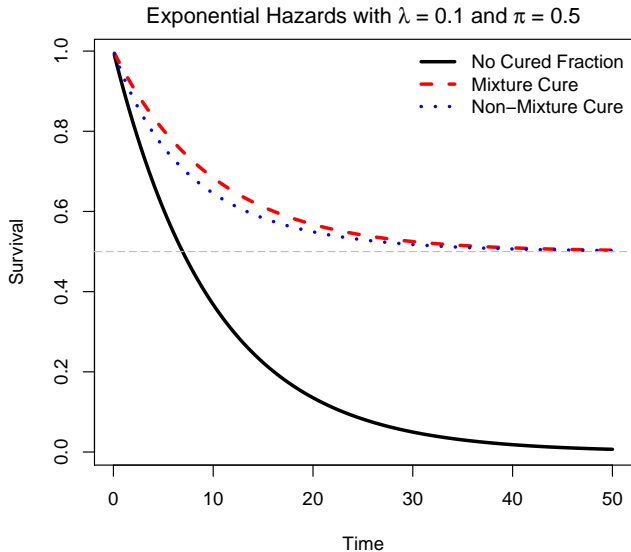
$$\pi_i = \exp[-\exp(\mathbf{X}_i\beta)]$$

we get:

$$S(t) = \exp\{[-\exp(\mathbf{X}_i\beta)]F(t)\}.$$

which is the Cox.

Mixture vs. Non-Mixture Models



Discrete-Time Cure Models

- Parametric / Cox \longrightarrow Poisson
- Mixture Cure Model \longrightarrow Zero-Inflated Poisson
- Non-Mixture Cure Model \longrightarrow “Hurdle” Poisson

R

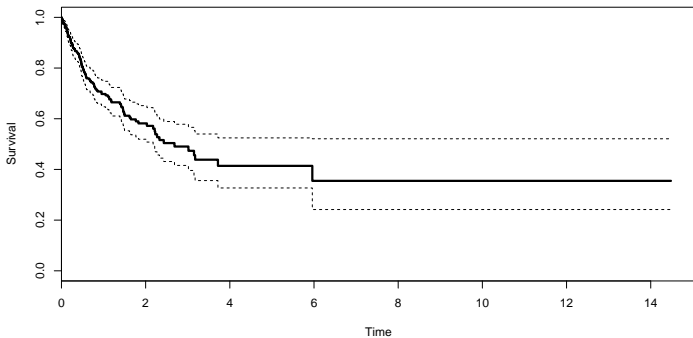
- `smcure` (semiparametric mixture models via EM)
- `semicure` (same; old)
- `nltm` (various; see Tsodikov 2003)
- CR, NPHMC (power analysis for cure models)

Stata

- `strsmix` and `strsnmix` (general parametric mixture & non-mixture cure models)
- `cureregr` (an old version)
- `lncure` (log-normal cure model)
- `spsurv` (discrete-time cure model)
- `zip` / `zinb` (discrete-time kludge)

A Simulated Example

```
> set.seed=7222009  
> X<-rnorm(500)  
> Z<-rbinom(500,1,0.5)  
> T<-rweibull(500,shape=1.2,scale=1/(exp(0.5+1*X)))  
> C<-rbinom(500,1,(0.4-0.3*Z))  
> S<-Surv(T,C)
```



Cox Models

```
> coxph(S~X)
```

```
Call:
```

```
coxph(formula = S ~ X)
```

	coef	exp(coef)	se(coef)	z	p
X	1.05	2.85	0.124	8.44	0

```
Likelihood ratio test=77.7 on 1 df, p=0 n= 500, number of events= 130
```

```
> coxph(S~X+Z)
```

```
Call:
```

```
coxph(formula = S ~ X + Z)
```

	coef	exp(coef)	se(coef)	z	p
X	1.08	2.956	0.122	8.9	0.0e+00
Z	-1.59	0.204	0.230	-6.9	5.4e-12

```
Likelihood ratio test=140 on 2 df, p=0 n= 500, number of events= 130
```

Cure Model

```
> cure.fit<-smcure(S~X,cureform=~Z,data=data.cure,model="ph")
```

Program is running..be patient... done.

Call:

```
smcure(formula = S ~ X, cureform = ~Z, data = data.cure, model = "ph")
```

Cure probability model:

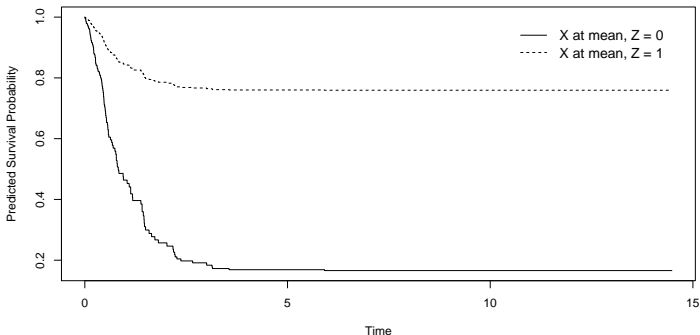
	Estimate	Std.Error	Z value	Pr(> Z)
(Intercept)	1.6	0.39	4.1	3.4e-05
Z	-2.8	0.41	-6.7	2.5e-11

Failure time distribution model:

	Estimate	Std.Error	Z value	Pr(> Z)
X	1.1	0.14	8.1	6.7e-16

An Interesting Plot

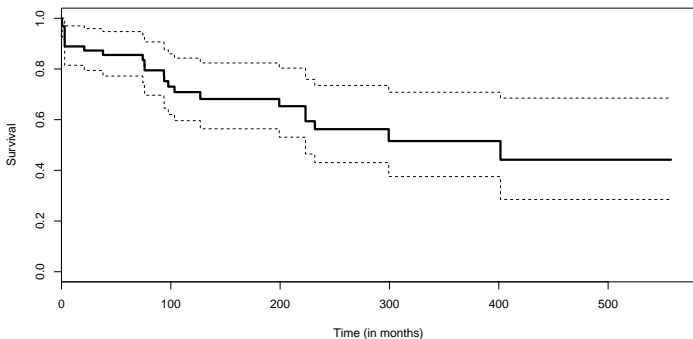
```
> cure.hat<-predictsmcure(cure.fit,c(rep(mean(X),times=2)),  
                           c(0,1),model="ph")  
  
> cure.pic<-plotpredictsmcure(cure.hat,type="S",model="ph")
```



An Example: Ceasefire Durability

Data are a subset used in Fortna (2004) (full data are [here](#)).

- $N = 63$
- Non-time-varying



Ceasefires: Cox Model

```
> CF.cox<-coxph(CF.S~tie+imposed+lndeaths+contig+onedem+twodem,  
                data=CF,method="efron")
```

```
> CF.cox
```

Call:

```
coxph(formula = CF.S ~ tie + imposed + lndeaths + contig + onedem +  
      twodem, data = CF, method = "efron")
```

	coef	exp(coef)	se(coef)	z	p
tie	1.845	6.327	0.557	3.314	0.00092
imposed	0.210	1.233	0.594	0.353	0.72000
lndeaths	-0.135	0.874	0.193	-0.699	0.48000
contigyes	2.898	18.143	0.948	3.058	0.00220
onedem	3.423	30.648	1.144	2.991	0.00280
twodem	-0.723	0.485	1.209	-0.598	0.55000

Likelihood ratio test=36.8 on 6 df, p=0.00000197 n= 63, number of events= 23

(hours of fiddling...)

A Typical Result

```
> CF.cure1.fit<-smcure(CF.S~tie+lndeaths+imposed,  
                      cureform=~contig,data=CF,model="ph",  
                      link="logit",emmax=500)
```

Program is running..be patient... done.

Call:

```
smcure(formula = CF.S ~ tie + lndeaths + imposed, cureform = ~contig,  
      data = CF, model = "ph", link = "logit", emmax = 500)
```

Cure probability model:

	Estimate	Std.Error	Z value	Pr(> Z)
(Intercept)	-3.4	12.4	-0.27	0.79
contig	2.1	7.4	0.28	0.78

Failure time distribution model:

	Estimate	Std.Error	Z value	Pr(> Z)
tie	2.05	4.06	0.50	0.61
lndeaths	-0.37	0.34	-1.10	0.27
imposed	0.97	2.40	0.41	0.68

There were 50 or more warnings (use warnings() to see the first 50)

From Svulik (2008)

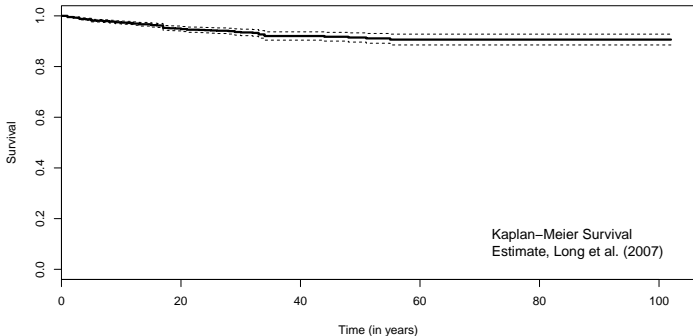
Consolidation status model^b

<i>GDP per capita</i>	2.121*** (0.586)	—	2.045*** (0.555)	2.121*** (0.586)
<i>GDP growth</i>	-0.014 (0.227)	—	-0.048 (0.246)	-0.014 (0.227)
<i>Military (vs. Not independent)</i>	-4.061** (1.895)	—	-3.985** (1.857)	-4.061** (1.895)
<i>Civilian (vs. Not independent)</i>	-0.421 (1.097)	—	-0.549 (1.067)	-0.421 (1.097)
<i>Monarchy (vs. Not independent)</i>	-20.158 (2888.609)	—	-15.844 (680.185)	-13.965 (891.870)
<i>Parliamentary (vs. Mixed)</i>	2.231 (2.230)	—	2.290 (2.326)	2.231 (2.230)
<i>Presidential (vs. Mixed)</i>	-8.310** (3.958)	—	-8.186** (4.035)	-8.310** (3.958)
<i>Intercept</i>	-6.144** (2.646)	—	-5.920** (2.644)	-6.145** (2.647)

Another Example: Peace Duration

Long, Nordstrom and Baek (2007 *JOP*)

- Peace duration among allies
- Time-varying dyadic data, 1816-2001 ($NT = 57,819$)



Cox Model (replicating LNB)

```
> LNB.cox<-coxph(LNB.S~relcap+major+jdem+border+wartime+s_wt_glo+
  medarb+noagg+arbcom+organ+milinst+cluster(dyad),
  data=LNB,method="breslow")

> LNB.cox
Call:
coxph(formula = LNB.S ~ relcap + major + jdem + border + wartime +
  s_wt_glo + medarb + noagg + arbcom + organ + milinst + cluster(dyad),
  data = LNB, method = "breslow")
```

	coef	exp(coef)	se(coef)	robust se	z	p
relcap	-1.431	0.239	0.614	0.683	-2.096	0.036000
major	1.137	3.118	0.241	0.280	4.064	0.000048
jdem	-0.987	0.373	0.367	0.380	-2.600	0.009300
border	1.931	6.897	0.190	0.206	9.378	0.000000
wartime	-0.359	0.699	0.367	0.467	-0.768	0.440000
s_wt_glo	-0.284	0.752	0.332	0.355	-0.802	0.420000
medarb	-0.367	0.693	0.285	0.306	-1.202	0.230000
noagg	-0.463	0.630	0.126	0.152	-3.051	0.002300
arbcom	1.306	3.690	0.325	0.316	4.133	0.000036
organ	0.353	1.423	0.280	0.285	1.236	0.220000
milinst	-0.373	0.689	0.187	0.177	-2.101	0.036000

Cure Models

(hours of fiddling...)

```
> LNB.cure<-smcure(LNB.altS~relcap+major+jdem+border+wartime+s_wt_glo+  
  medarb+noagg+arbcom+organ+milinst,  
  cureform=~border,model="ph",data=LNB)
```

Program is running..be patient...

Cure Models (Stata Remix)

```
. stset count1, id(episode) f(buofmzmid==1)
. gen h0=0
. strsmix major jdem border wartime, bhazard(h0) distribution(weibull) link(logistic) k1
> (relcap major jdem border wartime s_wt_glo medarb noagg arbcom organ milinst)
```

					Number of obs	=	57819
					Wald chi2(4)	=	36.82
Log likelihood = -793.21263					Prob > chi2	=	0.0000

	_t	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	

pi							
	major	-7.921296	3.764002	-2.10	0.035	-15.2986	-.5439877
	jdem	-.6177566	.7656096	-0.81	0.420	-2.118324	.8828107
	border	-1.943181	.3786093	-5.13	0.000	-2.685241	-1.20112
	wartime	2.583909	1.051959	2.46	0.014	.5221065	4.645711
	_cons	2.659179	.3980719	6.68	0.000	1.878972	3.439385

ln_lambda							
	relcap	-1.408332	.7129111	-1.98	0.048	-2.805613	-.0110523
	major	-1.232928	.395653	-3.12	0.002	-2.008394	-.4574626
	jdem	-1.69796	.4596442	-3.69	0.000	-2.598846	-.7970736
	border	1.224114	.2622007	4.67	0.000	.7102103	1.738018
	wartime	.42086	.4072876	1.03	0.301	-.377409	1.219129
	s_wt_glo	-.274703	.3579769	-0.77	0.443	-.9763249	.4269188
	medarb	-.8221547	.3503126	-2.35	0.019	-1.508755	-.1355545
	noagg	-.68365	.1465971	-4.66	0.000	-.970975	-.3963251
	arbcom	1.667284	.4562532	3.65	0.000	.7730438	2.561524
	organ	.9298395	.3595899	2.59	0.010	.2250563	1.634623
	milinst	-.4428979	.2251323	-1.97	0.049	-.8841491	-.0016468
	_cons	-2.060399	.7260061	-2.84	0.005	-3.483344	-.6374528

ln_gamma							
	_cons	.0969349	.0733007	1.32	0.186	-.0467319	.2406018

Cure models...

- ...Powerful
- ...Intuitive
- ...Temperamental
- ...Ask a lot of your data