

Firms that trade

1 Stylized facts and the discrete choice approach

Looking at plant-level¹ data (typically censuses or surveys that cover only manufacturing) from several countries and time periods, a few big things jump out.

1. About 20 percent of plants have non-zero exports. [This is a statement about direct exporting. Firms that export through intermediaries—like wholesalers—would not be picked up in this data.]
2. Exporting is a persistent state: Conditional on exporting this year, there is about a 90 percent chance of exporting next year.
3. Exporters are bigger, pay higher wages, and are more productive.

These facts motivate a discrete choice approach. Suppose that there are start-up costs to exporting: relabeling goods, upgrading quality, making business connections, finding distributors, doing market research. Let $\pi_x(s_{it}, p_t)$ be the profit plant i earns if it exports, where s_{it} are plant level variables and p_t are aggregate variables. Let F^0 be a fixed cost of entry. The net present value of a firm contemplating export entry is

$$V_{it}^x = E_t \sum_{n=t}^{\infty} m_n(p_n) [\pi_x(s_{nt}, p_n) - \pi_d(s_{nt}, p_n)] - F_t^0, \quad (1)$$

where m_n is the (possibly stochastic) discount factor, and π_d is the profit from only selling to the domestic market. The export entry decision is straightforward: Enter the export market if $V_{it}^x > 0$, do not enter the export market otherwise.

Suppose firms are heterogeneous in the profits they can earn in the export market. The entry cost induces selection on firm type, so that only the firms profitable enough will enter. With a large enough entry cost, many firms will not export, giving us fact #1. If higher profitability is associated with larger plants, or more productive plants, we also have fact #3. We can generate fact #2 if we introduce a fixed cost that the exporter must pay in each period to continue operating in the export market, but the sequence notation gets clunky. It is clearer in the recursive problem below.

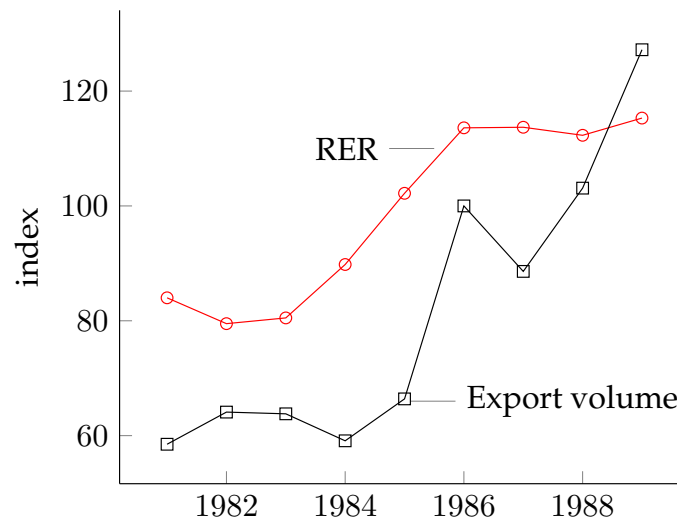
2 Roberts and Tybout (1997)

The data covers Colombian manufacturing firms with more than 10 employees. The data cover 1981–1989. Table 1 highlights the dynamics of this period

1. 1981–1983: small appreciation of RER, export participation falls
2. 1983–1989: large devaluation (30 percent), increase in volume of exports and number of exporters, but number of new exporters is pretty small. Why?

¹Most people use the word *plant* to describe a physical production location, and the word *firm* to describe an organizational unit, which may be made up of several plants. Most models feature one-plant firms, and I tend to use the two terms interchangeably. Particularly when dealing with data, we should be careful about how we use these two terms.

Figure 1: The real exchange rate and export volumes in Colombia.



3. Evidence from surveys (pg. 550) about obstacles to exporting.

This is related to a question in the literature that concerns the loose relationship between trade flows and real exchange rates, which we can see in this episode (figure 1). Until this paper, most evidence has been based on aggregate or sectoral models and data, so the sunk cost hypothesis remained an untested hypothesis.

Table 2 introduces the transition rates. A firm can be either an exporter, x , or a non-exporter, d . The transition rate from t to $t + 1$ is

$$\rho_t^{jk} = \frac{N_{t+1}^k}{N_t^j} \quad (2)$$

which is the number of plants of type k in period $t + 1$ divided by the number of plants of type j in period t . In table 2, for 1982–1983, we have $\rho^{dd} = 0.97$ and $\rho^{xx} = 0.83$: Exporting (or not exporting) is a persistent state. Export entry rates are about 3 percent and exit rates are about 13 percent.

This persistence could come from two sources:

1. Sunk costs of entry
2. Persistent firm heterogeneity

Identifying these sources is empirically difficult, and one of the major goals of the paper. A second tension arises from the “churning” of plants into and out of exporting. 60 percent of plants that exported at some point changed status more than once. (Data issues?) Here the possibilities are

1. Small entry costs
2. A benefit to having exported recently

So the main questions asked here are **Are sunk costs important for explaining firm export behavior?** and **How quickly does export experience decay in easing reentry into the export market?**

2.1 Model

This is a model of a plant's decision problem. Assume that exporting profits are separable from domestic profits. (Constant returns to scale in production will generate this.) Plant i 's gross profit from exporting is $\pi_i(p_t, s_{it})$, where p_t is a vector of exogenous variables (exchange rates, demand shifters, etc.) and s_{it} is a vector of state variables (capital stock, etc.).

If a plant has never exported it must pay F_{it}^0 to enter, and if the plant last exported in year $t-j$ ($j \geq 2$), pays a reentry cost of F_{it}^j . Note: these costs vary across time and firm.

Let Y_{it} be equal to 1 if plant i exports in t and 0 otherwise. The exporting history of plant of age J_i is a vector of zeros and ones,

$$Y_{it}^{(-)} = \{Y_{i,t-j} | j = 0, \dots, J_i\}. \quad (3)$$

The profit in period t of the plant is

$$R_{it}(Y_{it}^{(-)}) = Y_{it} \left[\pi_{it}(p_t, s_{it}) - F_{it}^0 (1 - Y_{i,t-1}) - \sum_{j=2}^{J_i} (F_{it}^j - F_{it}^0) \tilde{Y}_{i,t-j} \right] - X_i Y_{i,t-1} (1 - Y_{it}), \quad (4)$$

where π^* is the maximized value of period profits, X_i is the scrap value of the export operation, and

$$\tilde{Y}_{i,t-j} = \Pi_{k=1}^{j-1} (1 - Y_{i,t-k}) \quad (5)$$

is equal to 1 if the plant last exported in $t-j$, and 0 otherwise. Notice that if you were exporting last period, this is equal to 0.

Firms maximize discounted profits by choosing

$$Y_{it}^{(+)} = \{Y_{i,t+j} | j = 0, \dots\}. \quad (6)$$

The maximized value of the firm is

$$V_{it}(\Omega_{it}) = \max_{Y_{it}^{(+)}} E_t \sum_{j=t}^{\infty} \delta^{j-t} R_{ij} | \Omega_{it}, \quad (7)$$

where Ω is the plant's information set. This is a hard problem in sequence form. Recursively we can write it as

$$V_{it}(\Omega_{it}) = \max_{Y_{it}} \left[R_{it}(Y_{it}^{(-)}) + \delta E_t V_{i,t+1}(\Omega_{i,t+1} | Y_{it}^{(-)}) \right]. \quad (8)$$

Rather than choose the entire future sequence of exporting decisions, choose only today's. This is now a dynamic programming problem, although it is history dependent: All of the previous exporting decisions matter, because they determine the size of the reentry cost. A firm will choose to export when

$$\begin{aligned} & \pi_{it}(p_t, s_{it}) + \delta E_t [V_{i,t+1}(\Omega_{i,t+1} | Y_{it} = 1) - V_{i,t+1}(\Omega_{i,t+1} | Y_{it} = 0)] \\ & \geq F_i^0 - (F_i^0 + X_i) Y_{i,t-1} + \sum_{j=2}^{J_i} (F_{it}^j - F_{it}^0) \tilde{Y}_{i,t-j}, \end{aligned} \quad (9)$$

which just says that the value of exporting must be greater than the cost. Special cases:

- No sunk costs, but only a fixed cost that must be paid in each period that the firm exports. Export whenever $\pi_{it} > F_{it}$. The future is irrelevant.
- No entry costs at all. Export whenever $\pi_{it} > 0$.

In the full model

1. If a firm has never exported before, it needs LHS of (9) to be greater than F^0 .
2. If a firm exported last period, it needs LHS of (9) to be greater than $-X_i$ to stay in the market. Notice that this means that current profits can actually be negative. This creates hysteresis (and persistence).
3. If a firm last exported j periods ago, it needs LHS of (9) to be greater than F^j . This is “reentry” into the export market.

Notice that 1 and 2 imply that the entry and exit thresholds are not identical. Firm value has to be greater than F^0 to enter, but only greater than X to stay in the market. For firm values that are between these two costs, the behavior of a firm depends on its history: If the firm is already in the export market, it will stay there, but if it is not already in the market it will not enter. Two otherwise identical firms will be behaving differently.

In this framework we can ask, does exporting history help explain exporting status?

1. Cost show up as coefficients on the indicator variables on the right hand side. Notice that costs influence the left hand side, too, but in a nonlinear way, so they are not identified.
2. Realizations of p_t and s_{it} influence decisions both by the effect on current profits and on the expected future profits.

2.2 Empirical model

Define the LHS of (9) as

$$\pi_{it}^* = \pi_{it}(p_t, s_{it}) + \delta E_t [V_{i,t+1}(\Omega_{i,t+1}|Y_{it} = 1) - V_{i,t+1}(\Omega_{i,t+1}|Y_{it} = 0)], \quad (10)$$

which is a latent variable. The discrete choice model is

$$Y_{it} = \begin{cases} 1 & \text{if } \pi_{it}^* - F_i^0 + (F_i^0 + X_i) Y_{i,t-1} + \sum_{j=2}^{J_i} (F_{it}^0 - F_{it}^j) \tilde{Y}_{i,t-j} \geq 0 \\ 0 & \text{otherwise.} \end{cases} \quad (11)$$

How should we estimate this? Two possibilities: (1) Assume functional forms for π and the processes over s and p . (2) Estimate a reduced form equation for $\pi^* - F^0$.

The advantage of (1) over (2) is that we could recover of all the structural parameters of the profit function and use it to perform counterfactuals. The difficulty with (1) is that you must make strong assumptions about functional forms, and, since fixed costs depend on how long you have been in the market, the decisions are conditioned on a ninth-order Markov Chain, which would

be difficult to compute. The advantage of (2) over (1) is that it is easy, and still reveals information on sunk costs. The difficulty is that we don't get structural parameters. We are going with (2) in this paper. See Das, Roberts, and Tybout (2007) and Ruhl and Willis (2017) for examples of (1).

Implementation

Specify the reduced-form model of export profits as

$$\pi_{it}^* - F_i^0 = \mu_t + \beta Z_{it} + \epsilon_{it} \quad (12)$$

The μ_t are time specific industry- or macro-level variables that affect macro variability, but common to all plants (time fixed effect). These are things like exchange rates, trade policy, and other things that would be in p_t .

Z_{it} is a vector of observable plant level characteristics: a constant, industry dummies, the ownership structure of the plant, some geography dummies, the terms of trade, wage rate, capital stock and plant age.

Restrict fixed costs and scrap values to be the same across plants: Drop the i on F and X . Define $\gamma^j = F^0 - F^j$ and $\gamma^0 = F^0 + X$ so that

$$Y_{it} = \begin{cases} 1 & \text{if } \mu_t + \beta Z_{it} + \gamma^0 Y_{i,t-1} + \sum_{j=2}^{J_i} \gamma^j \tilde{Y}_{i,t-j} + \epsilon_{it} \geq 0 \\ 0 & \text{otherwise.} \end{cases} \quad (13)$$

Once we've estimated this, we can

1. Test for the importance of sunk costs by testing if the γ are jointly zero.
2. See how fast export experience decays by observing the magnitudes of the γ^j .
3. Interact the Y with the macro or plant level variables to see if sunk costs vary across time and plants.
4. Can look at terms in Z to see how plant entry and exit behavior changes with changes in other variables.

Econometric issues

1. Some unobserved variables, like managerial quality, will induce serial correlation in the error term. If not accounted for, the persistence in the error term will show up as "evidence of sunk costs."

Solution: Let ϵ be the sum of a permanent, plant-specific component and a 1st order AR component

$$\epsilon_{it} = \alpha_i + \omega_{it} \quad (14)$$

with $\omega_{it} = \rho \omega_{i,t-1} + \eta_{it}$. Assume α_i is i.i.d. normal across plants and η_{it} is i.i.d. normal across plants and time, ensuring that $\text{cov}(Z_{it}, \epsilon_{it}) = \text{cov}(\alpha_i, \omega_{it}) = 0$. Can't identify the epsilon variance, so normalize it to 1. Notice the α is permanent: It is the unobserved "quality" of the plant that can give rise to the persistence in export behavior. The horse race is between α and the γ s.

The serial correlation in the error term is now a function of α_i and ρ .

2. What about pre sample years? Since histories mattering is the null hypothesis, we care about what has happened before the sample. Cannot take starting point as exogenous, since $Y_{i,T-J}$ is a function of α_i and realizations of ω_{it} which are correlated with ϵ_{it} .

In the paper, they have a 9 year window, 1981-1989 and use a 3 lag model. So what about 1981-1983? How do we determine the “initial conditions?”

Solution: (Heckman, 1981) let Y_{it} when $t \leq J$ be a function of lagged explanatory variables (make an auxiliary model $\pi - F^0$ for the years that don't have lags),

$$\pi_{it}^* - F^0 = \lambda Z_{it}^P + \epsilon_{it}^P, \quad (15)$$

where Z_{it}^P is the distributed lag (i.e., many lags) in pre-sample realizations of the exogenous variables, plus 2-year lags of wages, capital stock and export price. Note the P superscripts — these are different variables.

So the pre-sample export decision are described by

$$Y_{it} = \begin{cases} 1 & \text{if } 0 \leq \lambda Z_{it}^P + \epsilon_{it}^P \\ 0 & \text{otherwise,} \end{cases} \quad (16)$$

where $\epsilon_{it}^P = \alpha_i^P + \omega_{it}^P$ with $\omega_{it}^P = \rho \omega_{i,t-1}^P + \eta_{it}^P$. Notice the lack of lagged export decisions in this problem.

A critical part of this procedure is to allow for α_i and α_i^P to be correlated. This allows for serial correlation in ϵ_i^P and ϵ_i . This allows the model to attribute export status persistence to plant-level idiosyncratic characteristics that otherwise would get loaded onto the fixed costs.

Estimation

Estimate (13) for $t > J$ and (16) for $t \leq J$. Estimate: $\beta, \lambda, \mu, \text{var}(\alpha), \rho, \text{var}(\alpha^P), \rho^P, \text{cor}(\alpha, \alpha^P)$ by simulated method of moments.

1. Choose a parameter vector
2. Draw error terms for each time and plant, and use them with the exogenous data to get decisions over Y_{it}
3. Repeat 1 and 2 a lot of times and get probabilities of the trajectory Y_{it}
4. Use method of moment weights to get a metric of fit
5. Vary parameters until you have minimized the metric

Results

Table 3. Restricted $\alpha = \alpha^P$ and restrict F to not vary over time. Models without these restrictions didn't do any better. Model 2 in column iii is their favorite.

1. Jointly, the γ are not zero, but only the first lag is significant. **Exporting history matters, but dies off quickly.**
2. Older, bigger, corporate plants have a higher probability of exporting. **Big could imply productive or IRS. Age probably correlated with productivity.**
3. Geography matters: being near the coast increases the probability of exporting.
4. Prices are insignificant, but time dummies are probably soaking this up.
5. $\text{var}(\alpha_i)$ is 0.67, so **two-thirds of the variation is due to firm heterogeneity.**
6. Once we control for everything, ρ is not significant. Serially correlated idiosyncratic “shocks” are not the driver of persistence.

Table 4. Aggregate the predicted Y paths and compute transition rates. Pretty good.

Table 5. Compute estimated probability of exporting for different plant types. $\alpha = 0$ is the mean.

1. For below mean α export history doesn’t matter. These are bad plants, who are not profitable in the export market, so they don’t ever really export.
2. For average plants — especially with above average observables — the probability of exporting is increasing in 1) how long ago you last exported, 2) how unobservably good you are relative to the mean (read down a column), 3) how observably good you are relative to the mean (read across a row by group)
3. History dependence decays quickly. (Look at $\alpha = 1$ and 75th percentile group.)

What next?

The big takeaway here is that both unobserved, permanent heterogeneity in export profitability and sunk entry costs are important in describing the export behavior of firms. This result tees up a model in which firms with heterogeneous productivity make export decisions in the face of fixed entry costs. This is the Melitz (2003) model.

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