

# Chapter 13.3 Negative Binomial Sampling

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Chapter 13 Case Studies

# Introduction

- ▶ We presented evidence that the observed counts of “can” from a group of Federalist Papers of Alexander Hamilton were overdispersed.
- ▶ There was more variability in the counts than predicted by the Poisson sampling model.
- ▶ Want to find an alternative sampling density for the counts that is able to accommodate this additional variation.

# Negative Binomial

- ▶ Recall  $y_i$  represents the number of “can”s in the  $i$ -th Federalist Papers.
- ▶ Conditional on parameters  $\alpha$  and  $\beta$ , one assigns  $y_i$  the Negative Binomial density defined as

$$f(Y_i = y_i \mid \alpha, \beta) = \frac{\Gamma(y_i + \alpha)}{\Gamma(\alpha)} p_i^\alpha (1 - p_i)^{y_i},$$

where

$$p_i = \frac{\beta}{\beta + n_i/1000}.$$

# Negative Binomial Generalizes Poisson

- ▶ Mean count is given by  $E(y_i) = \mu_i$  where

$$\mu_i = \frac{n_i}{1000} \frac{\alpha}{\beta}.$$

- ▶ The ratio  $\alpha/\beta$  is playing the same role as  $\lambda$  – one can regard  $\alpha/\beta$  as the true rate of the particular word per 1000 words.
- ▶ The variance of the count  $y_j$  is given by

$$\text{Var}(y_i) = \mu_i \left( 1 + \frac{n_i}{1000\beta} \right)$$

- ▶ With the extra multiplicative term  $\left( 1 + \frac{n_i}{1000\beta} \right)$ , Negative Binomial is able to accommodate the additional variability in the counts.

# Posterior Analysis

- ▶ Counts  $y_1, \dots, y_N$  are independent Negative Binomial with parameters  $\alpha$  and  $\beta$
- ▶ Likelihood function is equal to

$$L(\alpha, \beta) = \prod_{i=1}^N f(y_i \mid \alpha, \beta).$$

- ▶ Assume  $\alpha$  and  $\beta$  are independent and assign to each  $\alpha$  and  $\beta$  a Gamma density with parameters 0.001 and 0.001. Then the posterior density is given by

$$\pi(\alpha, \beta \mid y_1, \dots, y_N) \propto L(\alpha, \beta) \pi(\alpha, \beta)$$

where  $\pi(\alpha, \beta)$  is the product of Gamma densities.

## R Work

- ▶ Using JAGS, the Negative Binomial density is represented by the JAGS function `dnegbin()` with parameters `p[i]` and `alpha`.
- ▶ One first defines `p[i]` in terms of the parameter `beta` and the sample size `n[i]`, and then expresses the Negative Binomial density in terms of `p[i]` and `alpha`.

```
modelString = "  
model{  
  for(i in 1:N){  
    p[i] <- beta / (beta + n[i] / 1000)  
    y[i] ~ dnegbin(p[i], alpha)  
  }  
  mu <- alpha / beta  
  alpha ~ dgamma(.001, .001)  
  beta ~ dgamma(.001, .001)  
}
```

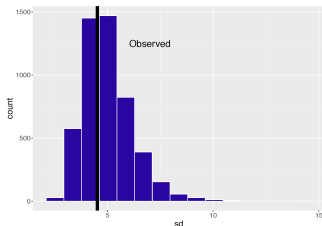
# Posterior Predictive Checking

- ▶ Can the Negative Binomial density accommodate the extra variability in the word counts?
- ▶ One can check this statement by a posterior predictive check implemented in the R function `one_rep()`.
- ▶ Start with a simulated value  $(\alpha^*, \beta^*)$  from the posterior distribution. Then we simulated a replicated dataset  $y_1^R, \dots, y_N^R$  where  $y_i^R$  has a Negative Binomial distribution with parameters  $\alpha^*$  and  $\beta^*/(\beta^* + n_i/1000)$ . Then we compute the standard deviation of the  $\{y_i^R\}$ .

```
one_rep <- function(i){  
  p <- post$beta[i] / (post$beta[i] + n / 1000)  
  sd(rnbinom(length(y), size = post$alpha[i], prob = p))  
}
```

# Repeat Process

- ▶ By repeating this algorithm for 5000 iterations, store 5000 draws of the standard deviation of samples from the predictive distribution.
- ▶ Figure displays a histogram of the standard deviations from the predictive samples and the observed standard deviation of the counts is shown as a vertical line.
- ▶ Predictions with a Negative Binomial sampling model appear consistent with the spread in the observed word counts.





# Inference

- ▶ One performs inferences about the mean use of the word “can” in Hamilton essays measured by the parameter  $\mu = \alpha/\beta$ .
- ▶ Figure displays MCMC diagnostic plots for the parameter  $\mu$ . A 90% posterior interval estimate for the rate of “can” is (2.20, 3.29)

