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MATH 369

9/9/20

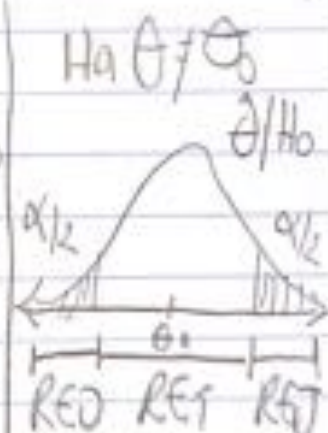
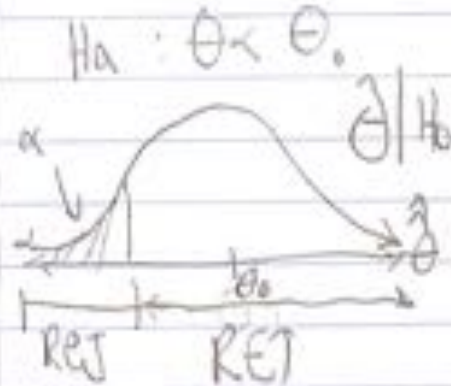
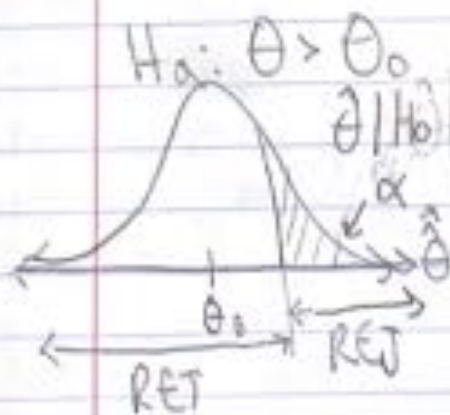
Lecture 4

I don't think I'll give exam q's on this:
"Level of a test" α is defined as $P(\text{Type 1 error})$
 $\geq \alpha$
"size of a test" is exactly $P(\text{Type 1 error})$

In our ex., the level was 5% but the size was 7.06%. Since $\alpha = 5\%$ was unattainable

If $\hat{\theta} | H_0$ is continuous, then level = size = α .
If it's discrete, some sizes won't be attainable.

If I want a level of $\alpha = 5\%$ and the size is lower, then I'm "cheating" (we'll see next class)



what we did in the previous lecture (2) was called a "binomial exact test of one proportion". Downsides: (1) you need a binomial pmf calculator and it's a lot of work to get the rejection region (2) not all sizes are attainable. This is the recommended test.

Let X_1, X_2, \dots, X_n iid some distribution with mean μ and variance σ^2 . The central limit theorem (CLT) shows that:

$$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \xrightarrow{d} N(0, 1)$$

"convergence in distribution". It means as n gets large, the CDF of the LHS looks more and more like the CDF of the RHS

* $\Rightarrow \bar{X} \overset{\text{approx. distr.}}{\sim} N\left(\mu, \frac{\sigma^2}{n}\right)$

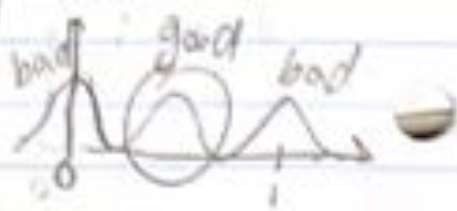
and $T = X_1 + \dots + X_n \sim N(n\mu, n\sigma^2)$

If X_1, \dots, X_n iid Bern(θ) and n is "big" then:

$$\hat{\theta} = \bar{X} \sim N\left(\theta, \frac{\theta(1-\theta)}{n}\right)$$

this is a pretty good approximation if θ is not too close to 0 or 1

How to perform an "approximate test"? There are many, many options even for the same DGP. The protocol



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goes as follows:

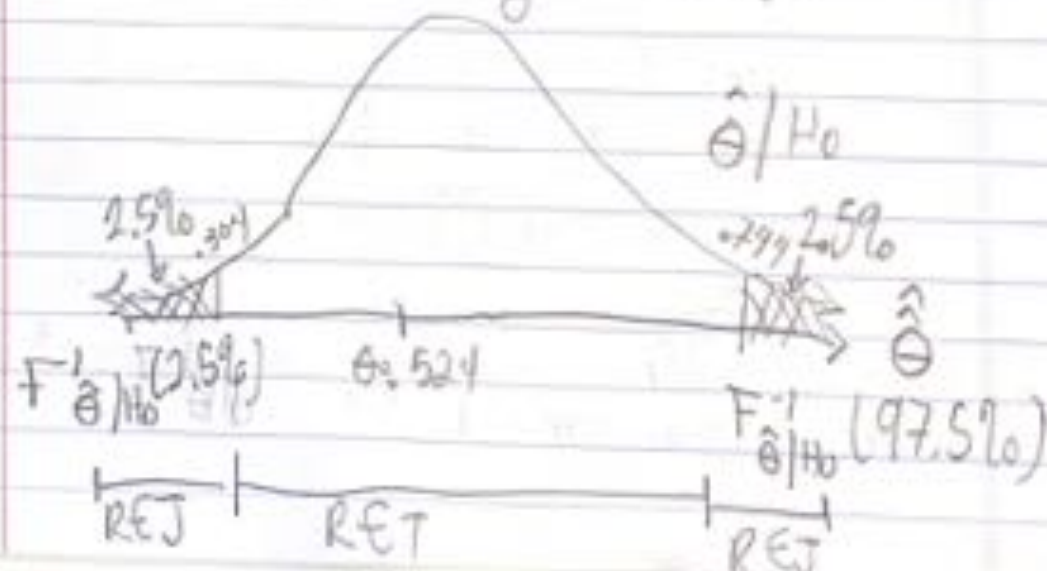
- (1) you think of a "test statistic" that could measure the departure away from H_0
- (2) Derive the statistical estimator's **approx** distribution under H_0 , $\hat{\theta}|H_0$
- (3) Gauge the departure of $\hat{\theta}$ from the bulk of the distribution $\hat{\theta}|H_0$ at level α .

$H_0: \theta = .524, H_a: \theta \neq .524, n=20, \hat{\theta} = 0.6$
(same as last class)

$$\hat{\theta}|H_0 \sim N\left(.524, \frac{.524(1-.524)}{20}\right)$$

$$= N(.524, \underbrace{.112}_{\sigma^2})$$

Set $\alpha = 5\% \Rightarrow \frac{\alpha}{2} = 2.5\%$



$$F_{\hat{\theta}|H_0}^{-1}(2.5\%)$$

(4)

$$P(\hat{\theta}|H_0 \leq \hat{\theta}) = 2.5\% \text{ solve for } \hat{\theta}$$

$$\Rightarrow P\left(\frac{\hat{\theta}|H_0 - .524}{.112} \leq \frac{\hat{\theta} - .524}{.112}\right) = 2.5\%$$

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 $N(0,1)$

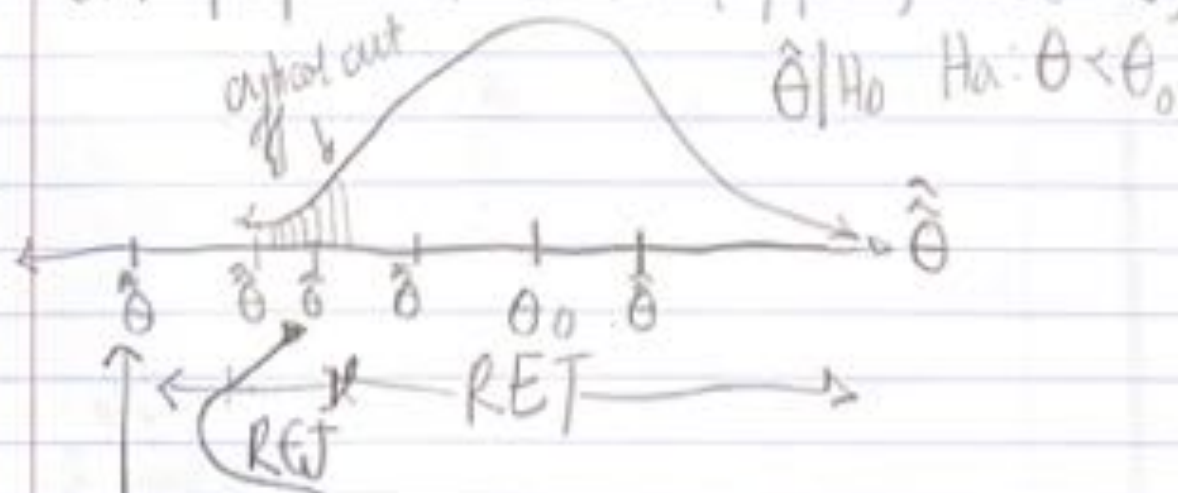
$$\Rightarrow P\left(Z \leq \frac{\hat{\theta} - .524}{.112}\right) = 2.5\%$$

Look at table or use computer

$$\Rightarrow \frac{\hat{\theta} - .524}{.112} \approx 1.96 \approx 2$$

$$\Rightarrow \hat{\theta} = .304$$

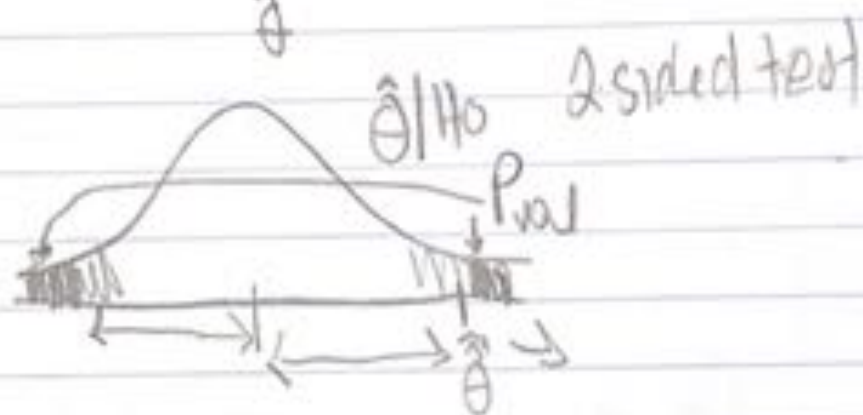
One proportion z-test (approximate test).



the estimate should imply a "stronger" rejection, that this estimate

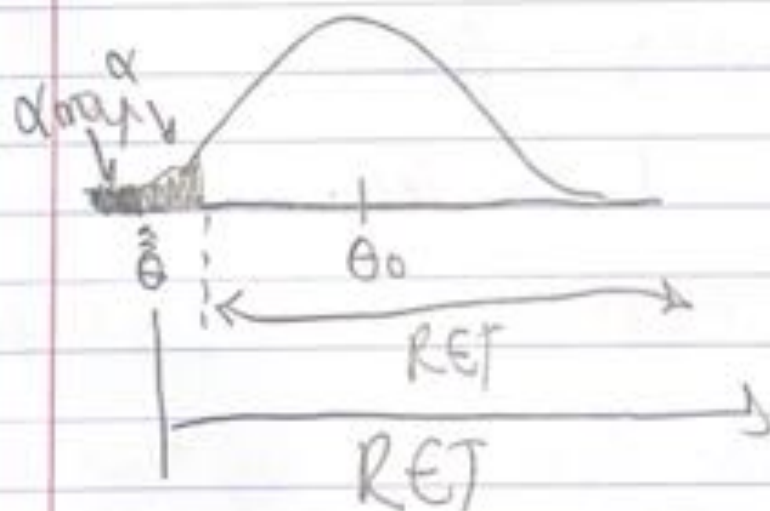
To measure the "strength" of a rejection (or "weakness" of a retainment), Fisher (5) introduced the "p-value" also called the level of statistical significance as:

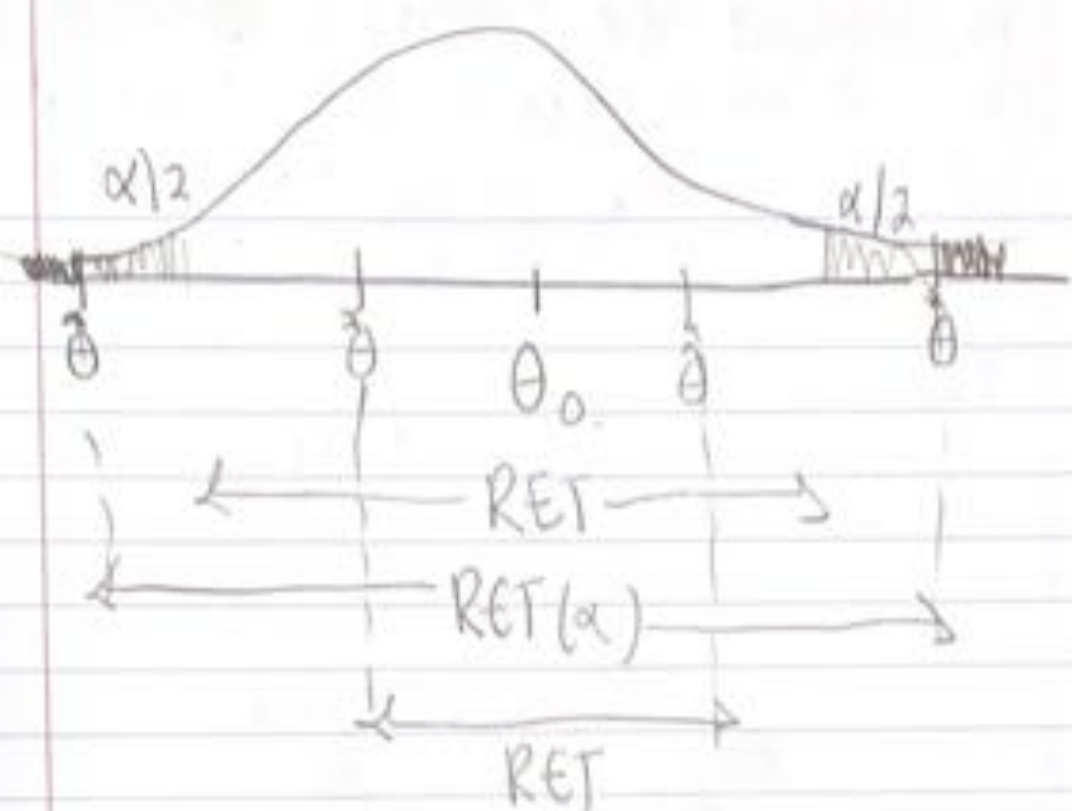
$P_{val} := P(\text{estimate is more extreme than one observed} | H_0)$



Real definition

$$P_{val} = \max \{ \alpha : \hat{\theta} \in RET(\alpha) \}$$





If H_0 is retained $\Rightarrow p\text{-val} \geq \alpha$

If H_0 is rejected $\Rightarrow p\text{-val} < \alpha$

Type II errors & POWER

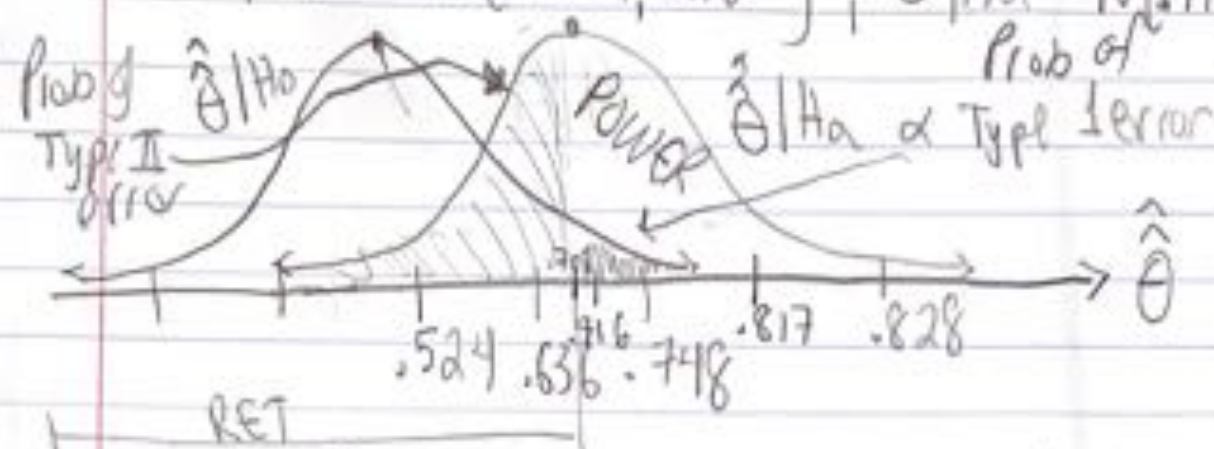
DGP $X_1, \dots, X_n \stackrel{iid}{\sim} \text{Bern}(\theta)$

$H_0: \theta = .524 = \theta_0$ but $H_a: \theta = .716 = \theta_a$

This is a non-standard setup since both H_0 and H_a are "point hypotheses". This makes the outcome weird: either you retain $\theta = .524$ or you accept $\theta = .716$. But ignore this for now

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$$\hat{\theta}|H_0 \sim N(.524, .112^2), \hat{\theta}|H_a \sim N(.716, .101^2)$$



At $\alpha = 5\%$, the z value is 1.645 which means the Rejection Region ends at $\hat{\theta} = .524 + 1.645(.112) = .708$

$$P(\text{Type II error}) = P(\hat{\theta}|H_a \in \text{RET})$$

$$= P\left(\frac{\hat{\theta}|H_a - .716}{.101} \leq \frac{.708 - .716}{.101}\right)$$

$$= P(Z \leq -.079) \approx 47\%$$

Errors

Decision \ Truth	RET	RET
	Ho	Type I
Ha	Type II	

POWER = $P(\text{Rejecting } H_0 | H_a)$

$= 1 - P(\text{Retaining } H_0 | H_a)$

$= 1 - P(\text{Type II error})$

Power $\approx 53\%$ pretty bad

Power is the probability of proving your theory is true!! You want Power to be large i.e. near 100%