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Math 669

9/2/20

### Lecture 3

DGP  $X_1, \dots, X_n$  i.i.d. with mean  $\theta$ , variance  $\sigma^2$

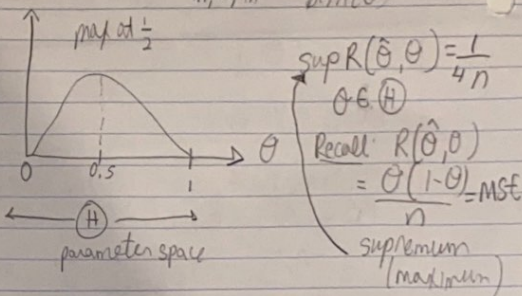
If  $\hat{\theta} = \bar{X} \Rightarrow \hat{\theta}$  is unbiased

$$SE[\hat{\theta}] = \sqrt{\text{Var}\left[\frac{1}{n}(X_1 + \dots + X_n)\right]} = \sqrt{\frac{1}{n} \sum \text{Var}[X_i]}$$

$$= \sqrt{\frac{1}{n} n \sigma^2} = \frac{\sigma}{\sqrt{n}} = \sqrt{\frac{\sigma(1-\sigma)}{n}}$$

$SE[\hat{\theta}]$

$X_1, \dots, X_n$  i.i.d.  $\text{Bern}(\theta)$



Goal #3 of inference: theory testing  
(hypothesis testing).

You have some well-specified mathematical theory about the DGP.

For example, in the iPhone survey "I think the proportion of iPhone users in the population is NOT 52.4%. I want to prove my theory to the world (using my sample).

NOTE: it is absolutely impossible to prove or disprove my theory because you cannot see the whole population (or go inside of the DGP). We must use inference which is always a guess.

Two ways to go about "proving" my theory:

- (1) I assume I'm right and wait for other people to show me data that contradicts my theory
- (2) I assume my theory is wrong. Then I advance (bring) evidence (i.e. data) to the contrary until people are convinced my theory is right

#2 is more intellectually honest and more likely to convince

A "hypothesis" is a mathematical statement about DGP e.g.  $\theta = 0.9$ ,  $\theta > 0.9$ ,  $\theta \neq 0.9$ ,  $\theta \leq 0.9$ ,  $\theta$  is in the set  $[0.89, 0.91]$ , etc

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The "alternative hypothesis" ( $H_a$ ) is the theory you want to prove. The "null hypothesis" ( $H_0$ ) is the opposite you assume in  $H_0$  for the purpose of contradicting it. Usual codes:

$$H_0: \theta \leq \theta_0, H_a: \theta > \theta_0 \quad (\text{right-tailed test})$$

↑ start here

$$H_0: \theta \geq \theta_0, H_a: \theta < \theta_0 \quad (\text{left-tailed test})$$

$$H_0: \theta = \theta_0, H_a: \theta \neq \theta_0 \quad (\text{two-tailed test})$$

How to perform this test? There are many, many options even for the same DGP.

The protocol goes as follows

- (1) you think of a "test statistic" that could measure the departure away from  $H_0$
- (2) derive the statistical estimator's distribution under  $H_0$
- (3) gauge the departure

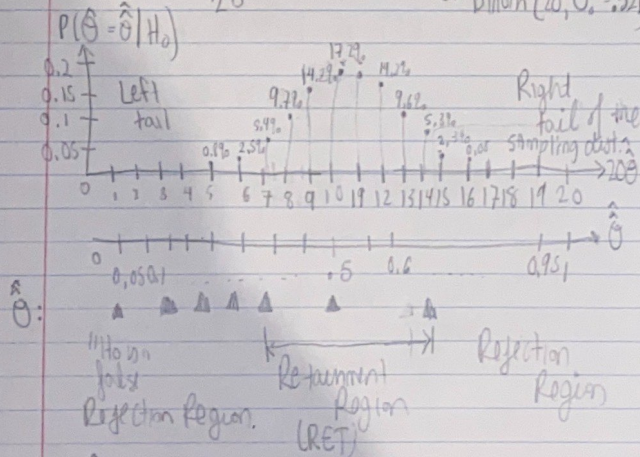
We begin with DGP: iid Bern ( $\theta$ ) and the "binomial exact test"

$$H_a: \theta \neq .524, H_0: \theta = .524 = \theta_0$$

(1) my test statistic is...  $\hat{\theta} = \bar{X}$ .  $\hat{\theta}$  is a realization from  $\hat{\theta}$

$$(2) \hat{\theta} | H_0 \sim ? \quad n = 20$$

$$\hat{\theta} = \frac{X_1 + \dots + X_{20}}{20} \Rightarrow 20\hat{\theta} | H_0 = X_1 + \dots + X_{20} \sim \text{Binom}(20, \theta_0 = .524)$$



$\hat{\theta} \in \text{RET} \Rightarrow$  Retain  $H_0$  (fail to reject  $H_0$ ).  
Not enough evidence to reject  $H_0$ . Some outliers say "accept  $H_0$ ".

$\hat{\theta} \notin \text{RET} \Rightarrow$  Reject  $H_0$  / Accept  $H_a$ . my estimate is "statistically significant".

Lets say we rejected  $H_0$  but it really was true. This is called a Type 1 error. What is the  $P(\text{Type 1 error})$  on our plot?



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$$\alpha := P(\text{Type I error}) = P(\hat{\theta} \notin \text{RET} | H_0)$$

Then in a 2-tailed test, I apportion about  $\alpha/2$  to the left tail and about  $\alpha/2$  to the right tail

$$\text{In my RET, } \alpha = P(\hat{\theta}=0 | H_0) + P(\hat{\theta}=0.3 | H_0) + P(\hat{\theta}=0.7 | H_0) + \dots + P(\hat{\theta}=1 | H_0) = 7.06\%$$

The choice of  $\alpha$  is up to you. The Scientific community's standard is 5% (sometimes 1%)

If you would like to prove your theory, you have to accept a positive probability of a Type I error

If I fail to reject  $H_0$  when  $H_a$  is true, that's a different error, a "Type II error". Failure to prove your theory

The smaller the  $\alpha$ , the larger the  $P(\text{Type II error})$ .

Truth		Decision	
		$H_0$	$H_a$
$H_0$	✓	✓	Type I error
$H_a$	Type II error	Type II error	✓
		Reject $H_0$	Reject $H_0$

As of now, we cannot calculate the  $P(\text{Type II error})$