

Lecture 60

DGP: $X_1, \dots, X_n \text{ iid } N(\theta, \sigma^2)$ standardize the estimator.

$$\hat{\theta} = \bar{X} \overset{\text{Math 241}}{\sim} N\left(\theta, \left(\frac{\sigma}{\sqrt{n}}\right)^2\right) \iff \frac{\bar{X} - \theta}{\frac{\sigma}{\sqrt{n}}} \sim N(0,1)$$

$\nearrow \text{SE}[\bar{X}]$

What if σ is unknown?

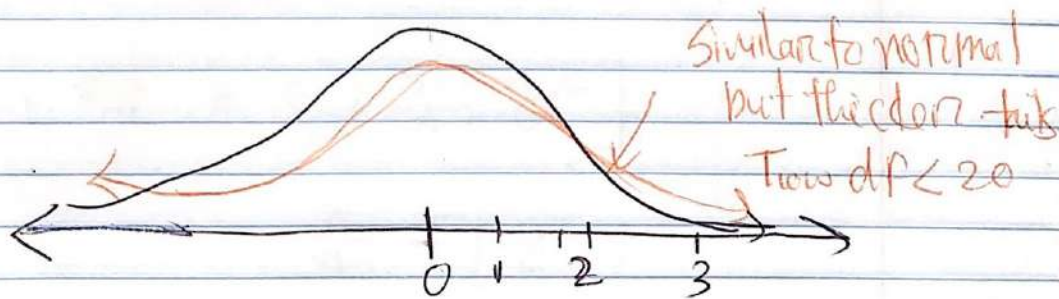
s^2 estimates $\sigma^2 \Rightarrow s = \sqrt{s^2}$ estimates σ .

In 1907 Gosset Proved:

$$\frac{\bar{X} - \theta}{\frac{s}{\sqrt{n}}} \sim T_{n-1}$$

Student's standard
T distribution with
 $n-1$ "degree of freedom"

(the parameter for the standard T
distrib)



data from $n=10$ males student heights.

$\bar{X} = 70.5, s = 2.07$

$H_a: \theta \neq 70, H_0: \theta = 70, \alpha = 5\%$

$\frac{\hat{\theta} - 70}{\frac{s}{\sqrt{n}}} = \frac{70.5 - 70}{\frac{2.07}{\sqrt{10}}} = 1.56$

$\alpha = 2.5\%$

The standardized distribution of $\hat{\theta} | H_0$

$\frac{\theta - 70}{\frac{2.07}{\sqrt{10}}} \sim T_9$

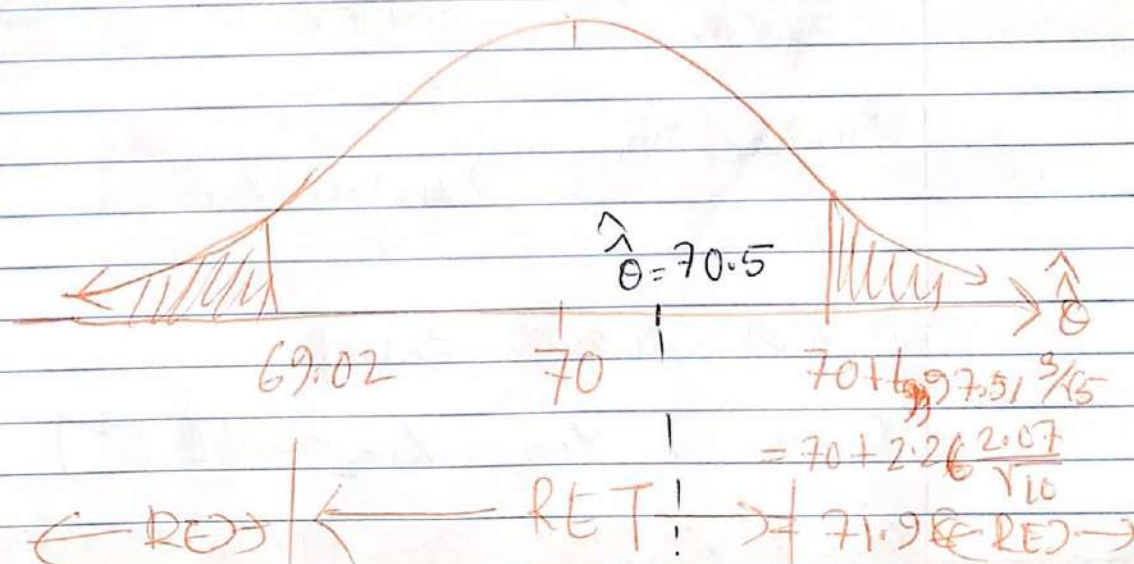
$\alpha = 2.5\%$

$t_{9, 97.5} = 2.26 > 1.96$

$= Z_{97.5\%}$

$\leftarrow REJ \rightarrow \leftarrow RET \rightarrow \leftarrow REJ \rightarrow$

$\Rightarrow REJ H_0$

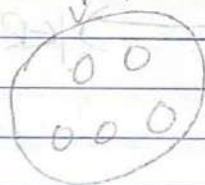
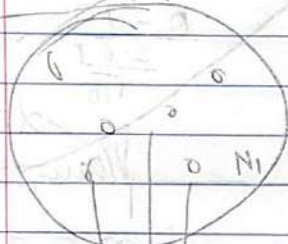


\Rightarrow Re-fair H₀

We just did our first "One sample two-sided t test" (of a mean).

new topic!

Population 1

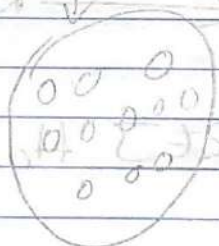
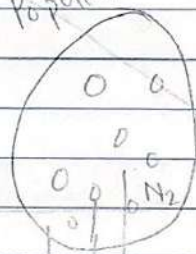


Sample 1

Size n_1

$x_{11}, x_{12}, \dots, x_{1n_1}$

Population 2



Sample 2

Size n_2

$x_{21}, x_{22}, \dots, x_{2n_2}$

$N_1 \approx \infty, N_2 \approx \infty \Rightarrow \text{iid}$

Assume, $x_{11}, \dots, x_{1n_1}, x_{21}, \dots, x_{2n_2} \text{ iid } N(\theta, \sigma^2)$

σ_1^2, σ_2^2 known but θ_1, θ_2 are unknown.

independent of

$$X_{211} \dots X_{2n_2} \text{ iid } N(\theta_2, \sigma_2^2)$$

There are 3 types of tests that are usually done:-

(I) $H_a: \theta_1 \neq \theta_2 \Rightarrow H_0: \theta_1 = \theta_2$ equivalently.

$$H_a: \theta_1 - \theta_2 \neq 0 \Rightarrow H_0: \theta_1 - \theta_2 = 0$$

(II) $H_a: \theta_1 < \theta_2 \Rightarrow H_0: \theta_1 \geq \theta_2$ equivalently.

$$H_a: \theta_1 - \theta_2 < 0 \Rightarrow H_0: \theta_1 - \theta_2 \geq 0$$

(III) $H_a: \theta_1 > \theta_2 \Rightarrow H_0: \theta_1 \leq \theta_2$ equivalently.

$$H_a: \theta_1 - \theta_2 > 0 \Rightarrow H_0: \theta_1 - \theta_2 \leq 0$$

What is a test statistic? (for $\theta_1 - \theta_2$)

$\hat{\theta}_1 - \hat{\theta}_2$ Point estimate.

The estimator that produces these estimates

is simply \bar{X}

$$\hat{\theta}_1 \sim N(\theta_1, \frac{\sigma_1^2}{n_1}) \text{ ind. } \hat{\theta}_2 \sim N(\theta_2, \frac{\sigma_2^2}{n_2})$$

Math 241 (difference of 2 normal)

$$\hat{\theta}_1 - \hat{\theta}_2 = \bar{X}_1 - \bar{X}_2 \overset{\text{exactly}}{\sim} N(\theta_1 - \theta_2, \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2})$$

$$\Rightarrow SE[\hat{\theta}_1 - \hat{\theta}_2] = \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$$

$$\frac{(\hat{\theta}_1 - \hat{\theta}_2) - (\theta_1 - \theta_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} \sim N(0, 1)$$

under H_0 (all three), $\theta_1 - \theta_2 = 0$

$$\hat{\theta}_1 - \hat{\theta}_2 \sim N(0, \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2})$$

Standardized $\rightarrow \frac{\hat{\theta}_1 - \hat{\theta}_2}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$

Let's test if male mean height is different than female mean height.

$$\vec{X}_2 = \langle 60, 59, 64, 64, 64, 63 \rangle$$

$$n_2 = 6 \quad \bar{X}_2 = 62.3 \quad n_1 = 10, \bar{X}_1 = 70.5$$

1.91 \rightarrow From male

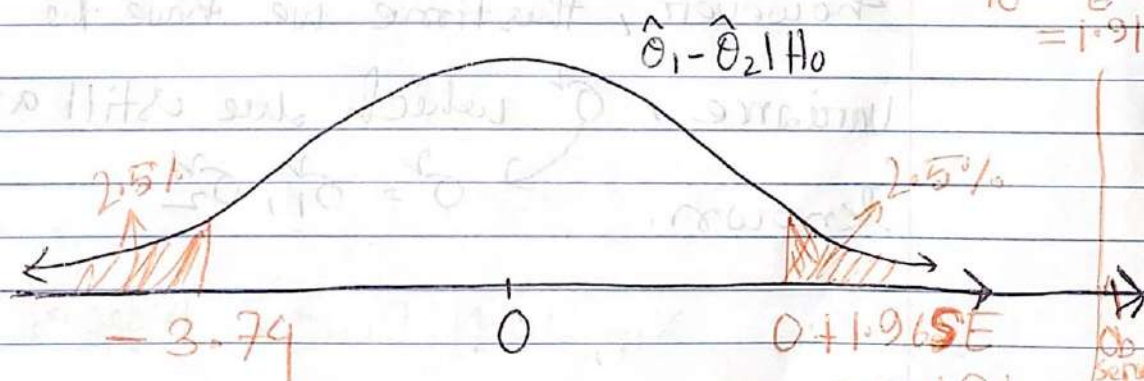
$$\hat{\theta}_1 - \hat{\theta}_2 = 70.5 - 62.3 = 8.2$$

We assumed we knew the variances.

So the variance for the men was assumed to be 4^2 and now the Variance for the Women is assumed to be 3.5^2 . so $\sigma_1^2 = 4^2$ $\sigma_2^2 = 3.5^2$
 $\alpha = 5\%$

We can do our 2 sample 2 sided test

$$SE = \sqrt{\frac{4^2}{10} + \frac{3.5^2}{6}} = 1.91$$



$$= 1.96 \times 1.91 = 3.74$$

Observed estimate

Observed estimate

$$\hat{\theta}_1 - \hat{\theta}_2 \notin \text{RET} \Rightarrow \text{Reject } H_0$$

$$P\text{Val} = 2P(\hat{\theta}_1 - \hat{\theta}_2 > 8.2)$$

$$= 2P\left(\frac{\hat{\theta}_1 - \hat{\theta}_2}{SE} > \frac{8.2}{1.91}\right)$$

$$= 2P(Z > 4.29)$$

$$= 1.8 \times 10^{-5} < \alpha$$

\Rightarrow Reject H_0

New * let's sample from two populations again

however, this time we have the same variance, σ^2 which we still assume known.

$$\sigma^2 = \sigma_1^2, \sigma_2^2$$

$X_{11}, \dots, X_{1n_1} \text{ i.i.d. } N(\theta_1, \sigma^2)$ indep of X_{21}, \dots, X_{2n_2}

$X_{2n_2} \text{ i.i.d. } N(\theta_2, \sigma^2)$

Under H_0 ,

$$\hat{\theta}_1 - \hat{\theta}_2 \sim N\left(0, \sigma^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)\right)$$

also, $\frac{\hat{\theta}_1 - \hat{\theta}_2}{\sigma \left(\frac{1}{n_1} + \frac{1}{n_2}\right)} \sim N(0, 1)$

2 Sample 2 Sided Z test of equal Variance

The test can be run again, you can probably

assume $\sigma = 3.75$.

Same as above but σ^2 unknown. How

can we estimate the standard error?

S_1^2, S_2^2 are the sample variances in

both samples 1 and 2.

$$S_1^2 = \frac{1}{n_1 - 1} \sum_{i=1}^{n_1} (X_{1,i} - \bar{X}_1)^2$$

$$S_2^2 = \frac{1}{n_2 - 1} \sum_{i=1}^{n_2} (X_{2,i} - \bar{X}_2)^2$$

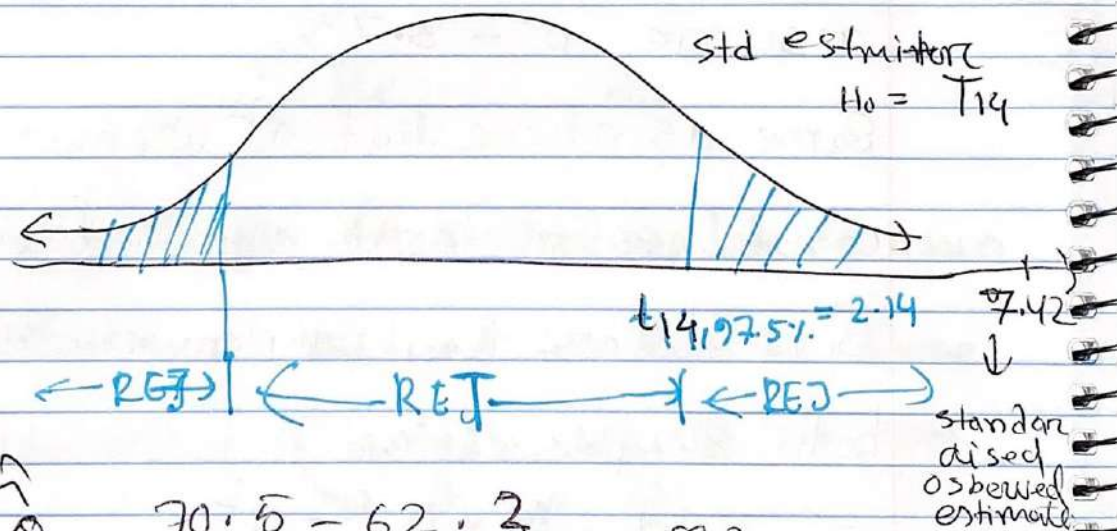
$$S_{\text{pooled}}^2 = \frac{(n_1-1)S_1^2 + (n_2-1)S_2^2}{n_1 + n_2 - 2}$$

weighted average.

You can prove that

$$\frac{\hat{\theta}_1 - \hat{\theta}_2}{S_{\text{pooled}} \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim T_{n_1 + n_2 - 2}$$

this allows you to do the "2-sample
+ test of equal variance"



$$\hat{\theta} = \frac{70.5 - 62.3}{S_{\text{pooled}} \sqrt{\frac{1}{10} + \frac{1}{6}}} = \frac{8.2}{2.14 \times 0.51} = 7.42$$

$$S_{\text{pooled}}^2 = \frac{9 \times 2.07^2 + 5 \times 2.25^2}{14}$$

$$= 4.56$$