# IRT Model

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Jason Qiang Guo IRT Model December 1st 1/6

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- ► The central issue of doing IRT model is to scale the distance between the latent positions of individuals. Consider the following problem:
- A legislator i has an ideal point  $\theta_i$ . If he votes for bill j, he derives utility  $u_i(\eta_j) = -(\theta_i \eta_j)^2 + \epsilon_{ij}^{[g]}$ . If he votes against the bill, his utility is  $u_i(s_j) = -(\theta_i s_j)^2 + \epsilon_{ij}^{[s]}$ , where  $s_j$  is the status quo. The terms  $\epsilon_{ij}^{[s]}$  and  $\epsilon_{ij}^{[n]}$  are i.i.d. standard normal variables (this is a random utility model). The legislator votes for bill  $j = 1, \ldots, J$  if and only if  $u_i(\eta_j) > u_i(s_j)$ .

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 $\triangleright$  The probability that a legislator *i* votes for proposition *j* is given by:

$$Pr(y_{i,j} = 1 | \theta_i) = Pr(u_i(\eta_j) > u_i(s_j)).$$

Replacing the values of  $u_i(\eta_j)$  and  $u_i(s_j)$ , we have the following expression:

$$Pr(y_{i,j} = 1 | \theta_i) = Pr(-(\theta_i - \eta_j)^2 + \epsilon'_{i,j} > -(\theta_i - s_j)^2 + \epsilon_{i,j})$$

$$= Pr((\theta_i - s_j)^2 - (\theta_i - \eta_j)^2 > \epsilon_{i,j} - \epsilon'_{i,j})$$

$$= Pr(2\theta_i(s_j - \eta_j) + (\eta_j^2 - s_j^2) > \epsilon_{i,j} - \epsilon'_{i,j})$$

$$= Pr(\theta_i d_j - b_j > \epsilon_{i,j} - \epsilon'_{i,j}), \quad \textit{where} \quad d_j = 2(s_j - \eta_j) \quad \textit{and} \quad b_j = s_j^2 - \eta_j^2.$$
So we have:  $Pr(y_{i,j} = 1 | \theta_i) = \Phi(\theta_i d_j - b_j). \quad d_j \text{ here is an item-discrimination parameter} \quad \textit{and} \quad b_i \text{ is an item-difficulty parameter}.$ 

Jason Qiang Guo IRT Model December 1st 3/6

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➤ To express the parameters of the voting model as functions of the estimated IRT parameters we use Bayes rules:

$$P(oldsymbol{ heta},b_j|y) \propto P(y|oldsymbol{ heta},d_j,b_j)P(oldsymbol{ heta})P(b)P(d) \ P(oldsymbol{ heta},d_j,b_j|y) \propto \left(\prod_{i=1}^N \prod_{i=1}^J \Phi( heta_id_j+b_j)^{y_{ij}} (1-\Phi( heta_id_j+b_j))^{1-y_{ij}}
ight)P(oldsymbol{ heta})P(oldsymbol{ heta})P(oldsymbol{$$

Jason Qiang Guo IRT Model December 1st 3 / 6

# IRT Model

The key steps in Stan are to specify the prior distributions for d, b and  $\theta$ , and also specify the model that links the choice and parameters.

```
model {
   alpha ~ normal(0, 25);
   beta ~ normal(0, 25);
   theta ~ normal(0, 1);
   for (n in 1:N)
     y[n] ~ bernoulli_logit(theta[j[n]] * beta[k[n]] - alpha[k[n]]);
}
```

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Jason Qiang GuoIRT ModelDecember 1st5 / 6

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- The critical element in the solution to this problem of course is to model how ideological positions of senators are influenced by their party ID.

 $\mbox{Jason Qiang Guo} \qquad \qquad \mbox{IRT Model} \qquad \qquad \mbox{December 1st} \qquad \mbox{5} \ / \ 6$ 

- Now suppose partisan ID plays an important role in the roll call voting (entirely reasonable assumption), we still want to scale the ideological positions of senators taking into account their party ID.
- The critical element in the solution to this problem of course is to model how ideological positions of senators are influenced by their party ID.
- One reasonable modeling assumption to make is to have a hierarchical structure for ideology:

$$\gamma_{j} \sim \textit{N}(\beta_{\textit{party}} * \textit{PartyID}_{j}, \sigma^{2}),$$
  $\theta_{j} \sim \textit{N}(\gamma_{j}, \tau^{2})$ 

Jason Qiang Guo IRT Model December 1st 5 / 6

- Now suppose partisan ID plays an important role in the roll call voting (entirely reasonable assumption), we still want to scale the ideological positions of senators taking into account their party ID.
- The critical element in the solution to this problem of course is to model how ideological positions of senators are influenced by their party ID.
- One reasonable modeling assumption to make is to have a hierarchical structure for ideology:

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  $heta_{j} \sim \textit{N}(\gamma_{j}, au^{2})$ 

- Since we have one more parameter  $\beta_{party}$  in our model, we need to give a prior for this parameter as well.
- ▶ In Stan, we just have to add this:

```
beta_party ~ normal(0, 2);
for (i in 1:J){
  theta[i] ~ normal(gamma[i], 1);
  gamma[i] ~ normal(beta_party * party[i], 1); }
```